

9.3 – Compound Distributions

S is the sum of a discrete random number, N , of random loss variables, X :

$$S = X_1 + X_2 + \dots X_N$$

Back Matter: Convolution

$$(f * g)(t) \stackrel{\text{def}}{=} \int_0^x f(t)g(x-t)dt$$

The Compound Model

$$\left. \begin{aligned} F_S &= \sum_n p_n F_X^{*n}(x) \\ f_S &= \sum_n p_n f_X^{*n}(x) \end{aligned} \right\} \begin{aligned} F_X^{*n} &= \int_0^x F_X^{*(n-1)}(x-t)f_X(t)dt, n=2,3\dots \\ f_X^{*n} &= \int_0^x f_X^{*(n-1)}(x-t)f_X(t)dt, n=2,3\dots \end{aligned}$$

Note that this is the specific case where the X 's are i.i.d.

This is computed recursively, since n is a random variable, so the aggregate severity distribution must be computed by brute force summing of all possible values of n .

$P(S=0)$ has a discrete, possibly non-zero, probability at $x=0$. N.B.: $P(S=0) \neq f_S(0)$.

Severity Distributions Closed Under Convolution

A severity distribution is closed under convolution if adding i.i.d members of a family produces another family of that family. Formally:

$$X_i \sim f_X(x; a), S = \sum_{i=1}^N X_i$$

$$f_S = \sum_{i=1}^{\infty} p_n f_X^{*n}(x; a) = \sum_{i=1}^{\infty} p_n f_X(x; na)$$

Properties of \mathcal{S}

MGF & PGF

$$M_S(t) = M_N[\lg M_X(t)]$$

$$G_S(z) = G_N[G_X(t)]$$

Using PGF_S can be a more convenient way to evaluate $f_S(s)$:

$$f_S(S=0) = G_S(0) = G_N[G_X(0)]$$

$$f_S(S=1) = G'_S(0) = \{G'_N[G_X(0)]\}G'_X(0)$$

⋮

$$f_S(S=k) = \frac{G_S^{(k)}(0)}{k!}$$

Moments

$$E(S) = E(N)E(X)$$

$$\text{Var}(S) = E(N)\text{Var}(X) + E(X)^2 \text{Var}(N)$$

9.3b – Net stop-loss Premium

The deductible is applied to aggregate losses, instead of individual losses:

$$E[(S-d)_+] = \int_d^{\infty} 1 - F_S(x)dx$$

$$= \int_d^{\infty} (x-d)f_S(x)dx$$

Linear Interpolation

In case there is an interval with no aggregate probability, interpolation may be used.

Suppose $P(a < S < b) = 0$. Then, for $a \leq d \leq b$,

$$E[(S-d)_+] = \frac{b-d}{b-a} E[(S-a)_+] + \frac{d-a}{b-a} E[(S-b)_+]$$

9.4 – Analytic Shortcuts for Aggregate-Severity

1. Negative Binomial-Exponential

Frequency \sim NB, Severity \sim exponential

$$F_S(x) = 1 - \left[\sum_{n=1}^r \binom{r}{n} \left(\frac{\beta}{1+\beta} \right)^n \left(\frac{1}{1+\beta} \right)^{r-n} \right] \left[\sum_{j=0}^{n-1} \left[x\theta^{-1}(1-\beta)^{-1} \right]^j e^{-x\theta^{-1}(1+\beta)^{-1}} / j! \right]$$

If $r=1$, this is a *geometric compound distribution*:

$$f_S = \frac{\beta}{\theta(1+\beta)^2} \exp\left[-\frac{x}{\theta(1+\beta)}\right], F_S = 1 - \frac{\beta}{1+\beta} \exp\left[-\frac{x}{\theta(1+\beta)}\right], x \geq 0$$

2. Exponential Severity

F_S is a gamma-mixture.

The MGF of the sum of N exponentially distributed r.v.s with same mean θ is gamma:

$$M_{X_1+X_2+\dots+X_N}(t) = (1-\theta t)^{-n}$$

$$\Rightarrow F_X^{*n}(x) = \Gamma(n; x/\theta) = 1 - \sum_{j=0}^{n-1} \frac{(x/\theta)^j}{j!} e^{-x/\theta}, n=1,2,3\dots$$

(Discrete form of lower incomplete Gamma function)

Distribution of aggregate claims:

$$f_S = \sum_{n=1}^{\infty} p_n \left\{ \frac{1}{\theta} e^{-x/\theta} \frac{(x/\theta)^{n-1}}{(n-1)!} \right\}, F_S = p_0 + \sum_{n=1}^{\infty} p_n \Gamma(n; x/\theta)$$

Which is a weighted gamma-mixture.

3. Compound Poisson Aggregate Severity

$$X_i \sim Po(\lambda_i) \Rightarrow S \sim Po\left(\lambda = \sum_{i=1}^n \lambda_i\right)$$

$$\Rightarrow F_S = \sum_{i=1}^n \frac{\lambda_i}{\lambda} F_{X_i}(x)$$

Note that $P(N=n) = \frac{\lambda_i}{\lambda}$

9.6 – Computing Aggregate Distribution

1. Recursion for (a, b, 0), (a, b, 1) Class

(a, b, 1) N Distribution

Severity distribution is discrete and defined on $1, 2, \dots, m$ (m can be infinite)

$$f_S = \frac{[p_1 - (a - b)p_0]f_X(x) + \sum_{y=1}^{\min(x,m)} \left[a + \frac{by}{x} \right] f_X(y)f_S(x-y)}{1 - af_X(0)}$$

(a, b, 0) IN Distribution

$$f_S = \frac{\sum_{x=1}^s \left[a + \frac{bx}{s} \right] f_X(y)f_S(x-y)}{1 - af_X(0)}, x \in \mathbb{Z}_+$$

Both cases start with calculating the value of $P(S=0) = P_N(f_X(0))$.

2. Continuous Severity

(a, b, 1) & (a, b, 0) Frequency Distribution

$$f_S = p_1 f_X(x) + \int_0^x \left[a + \frac{by}{x} \right] f_X(y)f_S(x-y)dy$$

3. Discretization of Continuous Distributions

3.1 Mass Dispersal

For an arbitrary interval of span h , index $j = 0, 1, 2, \dots$

$$f_0 = P\left(X < \frac{h}{2}\right) = F_X\left(\frac{h}{2} - 0\right)$$

$$f_j = P\left(jh - \frac{h}{2} \leq X < jh + \frac{h}{2}\right) = F_X\left(jh + \frac{h}{2} - 0\right) - F_X\left(jh - \frac{h}{2} - 0\right)$$

3.2 Method of Moment (matching)

The idea is to construct an arithmetic distribution function that matches the first p moments.

1. Construct an arbitrary interval of length ph , denoted $[x_k, x_k + ph)$.
2. Compute point masses m_0, m_1, \dots, m_p at points $x_k, x_k + h, \dots, x_k + ph$
3. Lagrange interpolation:

$$\sum_{j=0}^p (x_k + jh)^r m_j^k = \int_{x_k-0}^{x_k+ph-0} x^r f_X(x) dx, r = 0, 1, 2, \dots, p$$

$$m_j^k = \int_{x_k-0}^{x_k+ph-0} \prod_{i \neq j} \frac{x - x_k - ih}{(j-i)h} f_X(x) dx, j = i \dots p$$

9.7 – Effect of Individual Modifiers on Aggregate Severity

Per-Payment and Per-Loss

Relationships

$$F_{Y^L}(y) = (1 - v) + vF_{Y^P}(y), y \geq 0$$

$$M_{Y^L}(t) = (1 - v) + vM_{Y^P}(t)$$

$$\text{Where : } v = P(X > d)$$

Changes to Frequency

$$G_{N^P}(z) = G_{N^L}(1 - v + vz)$$

Moments

It is more convenient to use:

$$E[S] = E(N^L)E(Y^L)$$

$$Var(S) = E[N]Var(Y^L) + E[Y^L]^2 Var(N)$$

9.11 – Individual Risk Models

This is the compound model with one additional complication: severities are compound Bernoulli:

Parametric Approximation

That is, just assuming a normal, gamma, or lognormal distribution.