Introduction to plm

Yves Croissant & Giovanni Millo February 29, 2008

1 Introduction

The aim of package plm is to provide an easy way to estimate panel models. Some panel models may be estimated with package nlme (non-linear mixed effect models), but not in an intuitive way for an econometrician. plm provides methods to read panel data, to estimate a wide range of models and to make some tests. This library is loaded using:

> library(plm)

This document illustrates the features of plm, using data available in package

> library(Ecdat)

These data are used in Baltagi (2001).

2 Model estimation

 ${\tt plm}$ provides four functions for estimation :

- plm: estimation of the basic panel models, *i.e.* within, between and random effect models. Models are estimated using the lm function to transformed data,
- ullet pvcm: estimation of models with variable coefficients,
- pgmm : estimation of general method of moments models,
- pggls: estimation of general feasible generalized least squares models.

All these functions share the same 4 first arguments:

- formula: the symbolic description of the model to be estimated,
- data: a data.frame,
- effect: the kind of effects to include in the model, *i.e.* individual effects, time effects or both,
- model: the kind of model to be estimated, most of the time a model with fixed effects or a model with random effects,

- indexes: the indexes.
- NULL (the default value), it is then assumed that the first two columns contain the individual and the time index,
- a character string, which should be the name of the individual index,
- a character vector of length two containing the names of the individual and the time index,
- an integer which is the number of individuals (only in case of balanced panel with observations sorted by individual.

The plm.data function is then called, which returns a data.frame with the two first columns containing the individual and the time indexes.

The results of this four functions are stored in an object which class has the same name of the function. They all inherit from class panelmodel. A panelmodel object contains: coefficients, residuals, fitted.values, vcov, df.residual and call.

Functions that extract these elements and to print the object are provided.

2.1 Estimation of the basic models with plm

There are two ways to use plm: the first one is to estimate a list of models (the default behavior), the second to estimate just one model. In the first case, the estimated models are:

- the fixed effects model (within),
- the pooling model (pooling),
- the between model (between),
- the error components model (random).

The basic use of plm is to indicate the model formula, the data.frame and the name of the model to be estimated 1 :

```
> data("Produc", package = "Ecdat")
> zzwith <- plm(log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
+ data = Produc)</pre>
```

A particular model to be estimated may also be indicated by filling the model argument of plm.

```
> zzra <- plm(log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
+ data = Produc, model = "random")
> print(zzra)
```

 $^{^1\}mathrm{The}$ following example is from Baltagi (2001), pp. 25–28.

```
Model Formula: log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp
Coefficients:
(intercept)
           log(pcap)
                        log(pc)
                                  log(emp)
                                                unemp
 2.1354110
            0.0044386
                       0.3105484
                                 0.7296705 -0.0061725
  summary and print.summary methods are provided.
> summary(zzwith)
Oneway (individual) effect Within Model
plm(formula = log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
   data = Produc)
Balanced Panel: n=48, T=17, N=816
Residuals :
   Min. 1st Qu.
                 Median 3rd Qu.
-0.12000 -0.02370 -0.00204 0.01810 0.17500
Coefficients :
           Estimate Std. Error t-value Pr(>|t|)
log(pcap) -0.02614965 0.02900158 -0.9017
                                       0.3672
         log(pc)
        log(emp)
         unemp
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Total Sum of Squares: 18.941
Residual Sum of Squares: 1.1112
Multiple R-Squared: 0.94134
F-statistic: 3064.81 on 764 and 4 DF, p-value: 2.1339e-07
> summary(zzra)
Oneway (individual) effect Random Effect Model (Swamy-Arora's transformation)
plm(formula = log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
   data = Produc, model = "random")
Balanced Panel: n=48, T=17, N=816
Effects:
                       std.dev share
                 var
idiosyncratic 0.0014544 0.0381371 0.1754
           0.0068377 0.0826905 0.8246
individual
```

theta: 0.88884

```
Residuals :
   Min. 1st Qu.
                 Median 3rd Qu.
-0.10700 -0.02460 -0.00237 0.02170 0.20000
Coefficients:
             Estimate Std. Error t-value Pr(>|t|)
(intercept) 2.13541100 0.13346149 16.0002 < 2.2e-16 ***
           0.00443859 0.02341732 0.1895
log(pcap)
                                          0.8497
           log(pc)
log(emp)
           unemp
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' '1
Total Sum of Squares: 29.209
Residual Sum of Squares: 1.1879
Multiple R-Squared: 0.95933
F-statistic: 4782.77 on 811 and 4 DF, p-value: 8.7623e-08
  For a random model, the summary method gives information about the vari-
ance of the components of the errors.
  plm objects can be updated using the update method:
> zzwithmod <- update(zzwith, . ~ . - unemp - log(emp) + emp)
> summary(zzwithmod)
Oneway (individual) effect Within Model
Call:
plm(formula = log(gsp) ~ log(pcap) + log(pc) + emp, data = Produc)
Balanced Panel: n=48, T=17, N=816
Residuals :
    Min.
          1st Qu.
                     Median
                             3rd Qu.
                                         Max.
-0.194000 -0.037400 0.000373 0.035700 0.274000
Coefficients:
          Estimate Std. Error t-value Pr(>|t|)
log(pcap) 1.7888e-01 4.0690e-02 4.3961 1.102e-05 ***
log(pc)
         6.9975e-01 2.9154e-02 24.0019 < 2.2e-16 ***
         3.7909e-05 8.7824e-06 4.3165 1.585e-05 ***
emp
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Total Sum of Squares: 18.941
Residual Sum of Squares: 2.7948
Multiple R-Squared: 0.85245
```

F-statistic: 1473.23 on 765 and 3 DF, p-value: 2.4449e-05

Fixed effects may be extracted easily from a plm object using fixef:

> fixef(zzwithmod)[1:10]

```
ALABAMA
                ARIZONA
                           ARKANSAS
                                     CALIFORNIA
                                                    COLORADO CONNECTICUT
-0.15044698 -0.01596112 -0.13449962
                                     0.29699815
                                                  0.13601482
                                                              0.38383408
  DELAWARE
                FLORIDA
                            GEORGIA
                                           IDAHO
-0.11862549
             0.23429687
                         0.12381708 -0.22199517
```

The fixef function returns an object of class fixef. A summary method is provided, which prints the effects (in deviation from the overall intercept), their standard errors and the test of equality to the overall intercept.

> summary(fixef(zzwithmod))[1:10,]

```
Estimate Std. Error
                             t-value
                                     Pr(>|t|)
ALABAMA
         ARIZONA
         -0.01596112 0.2180845 -0.07318777 0.94165670
ARKANSAS
         -0.13449962 0.2071487 -0.64929021 0.51615081
CALIFORNIA
         0.29699815  0.2526566  1.17550143  0.23979417
COLORADO
          CONNECTICUT 0.38383408 0.2222083 1.72736143 0.08410277
DELAWARE
         -0.11862549
                  0.1950720 -0.60811143 0.54311357
FLORIDA
          0.23429687
                  0.2339542
                           1.00146486 0.31660212
GEORGIA
          0.12381708 0.2261564
                           0.54748435 0.58404602
IDAHO
```

2.2 More advanced use of plm

2.2.1 Options for the random effect model

The random effect model is obtained as a linear estimation on quasi-differentiated data. The parameter of this transformation is obtained using preliminary estimations. Four estimators of this parameter are available, depending on the value of the argument random.method:

- swar: from SWAMY and ARORA (1972), the default value,
- walhus: from Wallace and Hussain (1969),
- amemiya: from AMEMIYIA (1971),
- nerlove : from Nerlove (1971).

For exemple, to use the amemiya estimator:

```
> zzra <- plm(log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
+ data = Produc, model = "random", random.method = "amemiya")</pre>
```

2.2.2 Choosing the effects

The default behavior of plm is to introduce individual effects. Using the effect argument, one may also introduce :

- time effects (effect="time"),
- individual and time effects (effect="twoways").

For example, to estimate a two-ways effect model for the Grunfeld data:

```
> data("Grunfeld", package = "Ecdat")
> z <- plm(inv ~ value + capital, data = Grunfeld, model = "random",
      effect = "twoways", random.method = "amemiya")
> summary(z)
Twoways effects Random Effect Model (Amemiya's transformation)
Call:
plm(formula = inv ~ value + capital, data = Grunfeld, effect = "twoways",
   model = "random", random.method = "amemiya")
Balanced Panel: n=10, T=20, N=200
Effects:
                   var std.dev share
idiosyncratic 2644.135 51.421 0.2359
individual
             8294.716 91.075 0.7400
              270.529 16.448 0.0241
theta : 0.87475 (id) 0.29695 (time) 0.29595 (total)
Residuals :
  Min. 1st Qu. Median 3rd Qu.
                                  Max.
-176.00 -18.00
                  3.02
                         18.00
                                233.00
Coefficients:
             Estimate Std. Error t-value Pr(>|t|)
(intercept) -64.351811 31.183651 -2.0636 0.03905 *
                       0.011028 10.1192 < 2e-16 ***
value
             0.111593
capital
             0.324625
                        0.018850 17.2214 < 2e-16 ***
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
Total Sum of Squares: 2038000
Residual Sum of Squares: 514120
Multiple R-Squared: 0.74774
F-statistic: 291.965 on 197 and 2 DF, p-value: 0.0034191
```

In the "effects" section of the result is printed now the variance of the three elements of the error term and the three parameters used in the transformation.

The two–ways effect model is for the moment only available for balanced panels.

2.2.3 Hausman-Taylor's model

HAUSMAN-TAYLOR's model may be estimated with plm by equating the model argument to "ht" and filling the second argument instruments with a formula indicating the variables used as instruments.

```
> data("Wages", package = "Ecdat")
> Wages <- plm.data(Wages, 595)
> form <- lwage ~ wks + south + smsa + married + exp + I(exp^2) +
      bluecol + ind + union + sex + black + ed | sex + black +
      bluecol + south + smsa + ind
> ht <- plm(form, data = Wages, model = "ht")
> summary(ht)
Oneway (individual) effect Hausman-Taylor Model
plm(formula = lwage ~ wks + south + smsa + married + exp + I(exp^2) +
    bluecol + ind + union + sex + black + ed, data = Wages, model = "ht",
    instruments = ~sex + black + bluecol + south + smsa + ind)
T.V. exo : bluecolyes, southyes, smsayes, ind
T.V. endo : wks, marriedyes, exp, I(exp^2), unionyes
T.I. exo : sexmale, blackyes
T.I. endo : ed
Balanced Panel: n=595, T=7, N=4165
Effects:
                   var std.dev share
idiosyncratic 0.023044 0.151803 0.0253
             0.886993 0.941803 0.9747
individual
theta: 0.93919
Residuals :
   Min. 1st Qu.
                   Median 3rd Qu.
                                       Max.
-1.92000 -0.07070 0.00657 0.07970 2.03000
Coefficients:
              Estimate Std. Error t-value Pr(>|t|)
(intercept) 2.7818e+00 3.0765e-01 9.0422 < 2.2e-16 ***
            8.3740e-04 5.9973e-04 1.3963
wks
                                            0.16263
            7.4398e-03 3.1955e-02 0.2328
                                             0.81590
southyes
smsayes
           -4.1833e-02 1.8958e-02 -2.2066
                                             0.02734 *
marriedyes -2.9851e-02 1.8980e-02 -1.5728
                                             0.11578
            1.1313e-01 2.4710e-03 45.7851 < 2.2e-16 ***
I(exp^2)
           -4.1886e-04 5.4598e-05 -7.6718 1.696e-14 ***
bluecolyes -2.0705e-02 1.3781e-02 -1.5024 0.13299
            1.3604e-02 1.5237e-02 0.8928
                                            0.37196
ind
            3.2771e-02 1.4908e-02 2.1982
                                            0.02794 *
unionyes
            1.3092e-01 1.2666e-01 1.0337
```

sexmale

0.30129

```
blackyes -2.8575e-01 1.5570e-01 -1.8352 0.06647 .
ed 1.3794e-01 2.1248e-02 6.4919 8.474e-11 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Total Sum of Squares: 243.04
Residual Sum of Squares: 95.947
Multiple R-Squared: 0.60522
F-statistic: 489.524 on 4151 and 13 DF, p-value: 3.3651e-16
```

2.2.4 Instrumental variables estimation

One or all of the models may be estimated using instrumental variables. The instruments are specified whether as a one side formula in the argument ${\tt instruments}$, or at the end of the formula after a | sign. The following four commands are similar:

We illustrate instrumental variables estimation with the Crime data². The prbarr and polpc variables are assumed to be endogenous and there are two external instruments taxpc and mix:

```
> data("Crime", package = "Ecdat")
> form <- log(crmrte) ~ log(prbarr) + log(polpc) + log(prbconv) +</pre>
      log(prbpris) + log(avgsen) + log(density) + log(wcon) + log(wtuc) +
      log(wtrd) + log(wfir) + log(wser) + log(wmfg) + log(wfed) +
      log(wsta) + log(wloc) + log(pctymle) + log(pctmin) + region +
      smsa + year
> inst <- ~. - log(prbarr) - log(polpc) + log(taxpc) + log(mix)</pre>
> cr <- plm(form, data = Crime, model = "random", instruments = inst,
      pvar = TRUE)
> form2 <- log(crmrte) ~ log(prbarr) + log(polpc) + log(prbconv) +</pre>
      log(prbpris) + log(avgsen) + log(density) + log(wcon) + log(wtuc) +
      log(wtrd) + log(wfir) + log(wser) + log(wmfg) + log(wfed) +
      log(wsta) + log(wloc) + log(pctymle) + log(pctmin) + region +
      smsa + year | . - log(prbarr) - log(polpc) + log(taxpc) +
      log(mix)
> cr1 <- plm(form, data = Crime, model = "random", instruments = inst,
      pvar = TRUE)
> cr2 <- plm(form2, data = Crime, model = "random", pvar = TRUE)
> summary(cr2)
Oneway (individual) effect Random Effect Model (Swamy-Arora's transformation)
Instrumental variable estimation (Balestra-Varadharajan-Krishnakumar's transformation)
Call:
plm(formula = log(crmrte) ~ log(prbarr) + log(polpc) + log(prbconv) +
    log(prbpris) + log(avgsen) + log(density) + log(wcon) + log(wtuc) +
    log(wtrd) + log(wfir) + log(wser) + log(wmfg) + log(wfed) +
    log(wsta) + log(wloc) + log(pctymle) + log(pctmin) + region +
    smsa + year, data = Crime, model = "random", pvar = TRUE,
```

²See Baltagi (2001), pp.119–120.

```
instruments = ~. - log(prbarr) - log(polpc) + log(taxpc) +
      log(mix))
Instrumental Variables:
~log(prbconv) + log(prbpris) + log(avgsen) + log(density) + log(wcon) + log(wtuc) +
   log(wtrd) + log(wfir) + log(wser) + log(wmfg) + log(wfed) + log(wsta) + log(wloc) +
   log(pctymle) + log(pctmin) + region + smsa + year + log(taxpc) + log(mix)
Balanced Panel: n=90, T=7, N=630
Effects:
               var std.dev share
idiosyncratic 0.022269 0.149228 0.326
individual
           0.046036 0.214561 0.674
theta: 0.74576
Residuals :
  Min. 1st Qu. Median 3rd Qu.
                            Max.
-5.0200 -0.4760 0.0273 0.5260 3.1900
Coefficients:
             Estimate Std. Error t-value Pr(>|t|)
(intercept)
           -0.4538241 1.7029840 -0.2665 0.789864
           -0.4141200 0.2210540 -1.8734 0.061015 .
log(prbarr)
           0.5049285 0.2277811 2.2167 0.026642 *
log(polpc)
log(prbconv) -0.3432383 0.1324679 -2.5911 0.009567 **
           log(prbpris)
           log(avgsen)
           log(density)
           -0.0042963 0.0414225 -0.1037 0.917392
log(wcon)
log(wtuc)
           0.0444572 0.0215449 2.0635 0.039068 *
           -0.0085626 0.0419822 -0.2040 0.838387
log(wtrd)
           -0.0040302 0.0294565 -0.1368 0.891175
log(wfir)
            0.0105604 0.0215822 0.4893 0.624620
log(wser)
           -0.2017917 0.0839423 -2.4039 0.016220 *
log(wmfg)
log(wfed)
           -0.2134634 0.2151074 -0.9924 0.321023
           -0.0601083 0.1203146 -0.4996 0.617362
log(wsta)
           log(wloc)
           log(pctymle)
           0.1948760 0.0459409 4.2419 2.217e-05 ***
log(pctmin)
           -0.2281780 0.1010317 -2.2585 0.023916 *
regionwest
regioncentral -0.1987675  0.0607510 -3.2718  0.001068 **
           smsayes
            0.0132140 0.0299923 0.4406 0.659518
year82
           year83
           year84
year85
           -0.0977398  0.0511685  -1.9102  0.056113 .
           -0.0719390 0.0605821 -1.1875 0.235045
year86
           -0.0396520 0.0758537 -0.5227 0.601153
year87
```

Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' '1

```
Total Sum of Squares: 1354.7
Residual Sum of Squares: 557.64
Multiple R-Squared: 0.58836
```

F-statistic: 33.1494 on 603 and 26 DF, p-value: 7.3608e-16

The instrumental variables estimator used may be indicated with the ${\tt inst.method}$ argument:

- bvk, from Balestra and Varadharajan (1987), the default value,
- baltagi, from BALTAGI (1981).

2.2.5 Unbalanced panel

plm enables the estimation of unbalanced panel data, with a few restrictions (twoways effects models are not supported and the only transformation for random effects models is swar).

The following example is based on the Hedonic data³:

```
> data("Hedonic", package = "Ecdat")
> form <- mv ~ crim + zn + indus + chas + nox + rm + age + dis +
      rad + tax + ptratio + blacks + lstat
> ba <- plm(form, model = "random", data = Hedonic, index = "townid")
> summary(ba)
Oneway (individual) effect Random Effect Model (Swamy-Arora's transformation)
Call:
plm(formula = mv ~ crim + zn + indus + chas + nox + rm + age +
    dis + rad + tax + ptratio + blacks + lstat, data = Hedonic,
    model = "random", index = "townid")
Unbalanced Panel: n=92, T=1-30, N=506
Effects:
                   var std.dev share
idiosyncratic 0.016965 0.130249 0.502
individual
             0.016832 0.129738 0.498
theta :
  Min. 1st Qu. Median
                           Mean 3rd Qu.
                                           Max.
```

Residuals :

```
Min. 1st Qu. Median Mean 3rd Qu. Max. -0.641000 -0.066100 -0.000519 -0.001990 0.069800 0.527000
```

0.2915 0.5904 0.6655 0.6499 0.7447 0.8197

Coefficients :

 $^{^3 \}mathrm{See}$ Baltagi (2001), p. 174.

```
-7.2338e-03 1.0346e-03
                                   -6.9921 2.707e-12 ***
crim
zn
            3.9575e-05 6.8778e-04
                                     0.0575 0.9541153
indus
            2.0794e-03
                        4.3403e-03
                                     0.4791 0.6318706
chasyes
           -1.0591e-02 2.8960e-02 -0.3657 0.7145720
           -5.8630e-03 1.2455e-03 -4.7074 2.509e-06 ***
nox
            9.1773e-03 1.1792e-03
                                    7.7828 7.095e-15 ***
rm
           -9.2715e-04 4.6468e-04 -1.9952 0.0460159 *
age
           -1.3288e-01 4.5683e-02 -2.9088 0.0036279 **
dis
            9.6863e-02
                        2.8350e-02
                                    3.4168 0.0006337 ***
rad
                                    -1.9824 0.0474298 *
                        1.8902e-04
tax
           -3.7472e-04
           -2.9723e-02
                        9.7538e-03
                                    -3.0473 0.0023089 **
ptratio
blacks
            5.7506e-01
                        1.0103e-01
                                     5.6920 1.256e-08 ***
           -2.8514e-01 2.3855e-02 -11.9533 < 2.2e-16 ***
lstat
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' '1
Total Sum of Squares: 893.08
Residual Sum of Squares: 8.6843
Multiple R-Squared: 0.99028
F-statistic: 3854.18 on 492 and 13 DF, p-value: < 2.22e-16
```

2.3 Variable coefficients model

The pvcm function enables the estimation of variable coefficients models. Time or individual effects are introduced if effect is fixed to "time" or "individual" (the default value).

Coefficients are assumed to be fixed if model="within" and random if model="random". In the first case, a different model is estimated for each individual (or time period). In the second case, the SWAMY (1970) model is estimated. It is a generalized least squares model which use the result of the previous model.

With the Grunfeld data, we get:

> summary(znp)

```
> znp <- pvcm(inv ~ value + capital, data = Grunfeld, model = "within")</pre>
> znp
Model Formula: inv ~ value + capital
Coefficients:
   (Intercept)
                    value
                            capital
    -149.78245 0.1192808 0.3714448
1
2
     -49.19832 0.1748560 0.3896419
3
      -9.95631 0.0265512 0.1516939
4
      -6.18996 0.0779478 0.3157182
5
      22.70712 0.1623777 0.0031017
6
      -8.68554 0.1314548 0.0853743
7
      -4.49953 0.0875272 0.1237814
8
      -0.50939 0.0528941 0.0924065
9
      -7.72284 0.0753879 0.0821036
       0.16152 0.0045734 0.4373692
10
```

```
Oneway (individual) effect No-pooling model
Call:
pvcm(formula = inv ~ value + capital, data = Grunfeld, model = "within")
Balanced Panel: n=10, T=20, N=200
Residuals:
     Min.
             1st Qu.
                         Median
                                     Mean
                                             3rd Qu.
                                                           Max.
-1.845e+02 -7.118e+00 -3.926e-01 3.438e-16 5.703e+00 1.440e+02
Coefficients:
  (Intercept)
                       value
                                        capital
Min. :-149.782 Min. :0.004573 Min. :0.003102
1st Qu.: -9.639 1st Qu.:0.058518 1st Qu.:0.087132
Median: -6.956 Median: 0.082738 Median: 0.137738
Mean : -21.368
                  Mean :0.091285 Mean :0.205264
3rd Qu.: -1.507
                   3rd Qu.:0.128411
                                     3rd Qu.:0.357513
Max. : 22.707
                  Max.
                          :0.174856 Max. :0.437369
Total Sum of Squares: 9359900
Residual Sum of Squares: 324730
Multiple R-Squared: 0.96531
> form <- inv ~ value + capital
> sw <- plm(form, data = Grunfeld, model = "random")
> summary(sw)
Oneway (individual) effect Random Effect Model (Swamy-Arora's transformation)
plm(formula = inv ~ value + capital, data = Grunfeld, model = "random")
Balanced Panel: n=10, T=20, N=200
Effects:
                  var std.dev share
                       52.768 0.282
idiosyncratic 2784.458
individual
            7089.800
                        84.201 0.718
theta: 0.86122
Residuals :
  Min. 1st Qu. Median 3rd Qu.
                                  Max.
-178.00 -19.70
                  4.69
                         19.50 253.00
Coefficients:
             Estimate Std. Error t-value Pr(>|t|)
(intercept) -57.834415 28.898935 -2.0013 0.04536 *
             0.109781
value
                        0.010493 10.4627 < 2e-16 ***
capital
             0.308113
                      0.017180 17.9339 < 2e-16 ***
```

--Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Total Sum of Squares: 2381400
Residual Sum of Squares: 548900
Multiple R-Squared: 0.7695
F-statistic: 328.837 on 197 and 2 DF, p-value: 0.0030364

2.4 General method of moments estimator

The general method of moments is provided by the pgmm function. It's main argument is a dynformula which describe the variables of the model and the lag structure.

The effect argument is either NULL, "individual" (the default), or "twoways". In the first case, the model is estimated in levels. In the second case, the model is estimated in first differences to get rid of the individuals effects. In the last case, the model is estimated in first differences and time dummies are included.

In a gmm estimation, there are "normal" instruments and "gmm" instruments. gmm instruments are indicated with the gmm.inst argument (a one side formula) and the lags by with the lag.gmm argument. By default, all the variables of the model that are not used as gmm instruments are used as normal instruments, with the same lag structure.

The complete list of instruments can also be specified with the argument instruments which should be a one side formula (or dynformula).

The model argument specifies whether a one-step or a two-steps model is required ("onestep" or "twosteps").

The following example is from Arellano (2003). Employment in different firms is explained by past values of employment and wages (two lags). All available lags are used up to t-2.

```
> data("Snmesp", package = "plm")
> z \leftarrow pgmm(dynformula(n ~w, lag = list(c(1, 2), c(1, 2))), effect = "twoways",
      model = "twosteps", Snmesp, gmm.inst = "n + w, lag.gmm = c(2, m)
          99), transformation = c("d")
> summary(z)
Twoways effects Two steps model
pgmm(formula = n - lag(n, 1) + lag(n, 2) + lag(w, 1) + lag(w, 1)
    2), data = Snmesp, effect = "twoways", model = "twosteps",
    gmm.inst = \tilde{n} + w, lag.gmm = c(2, 99), transformation = c("d"))
Balanced Panel: n=738, T=8, N=5904
Number of Observations Used:
Residuals
      Min.
              1st Qu.
                           Median
                                         Mean
                                                 3rd Qu.
                                                                Max.
-1.5390000 -0.0511100 0.0010240 0.0001746 0.0549800 1.2780000
```

```
Coefficients
           Estimate Std. Error z-value Pr(>|z|)
lag(n, 1) 0.8415278 0.0883895 9.5207 < 2e-16 ***
lag(n, 2) -0.0031454 0.0290445 -0.1083 0.91376
lag(w, 1) 0.0779827 0.0836384 0.9324 0.35114
lag(w, 2) -0.0525764 0.0249418 -2.1080 0.03503 *
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
Sargan Test: chisq(36) = 36.91417 (p.value=0.42648)
Autocorrelation test (1): normal = -6.709587 (p.value=9.7588e-12)
Autocorrelation test (2): normal = 0.1986467 (p.value=0.42127)
Wald test for coefficients: chisq(4) = 234.7444 (p.value=< 2.22e-16)
Wald test for time dummies: chisq(5) = 44.47645 (p.value=1.8536e-08)
  In the following example, a pure auto-regressive model is estimated.
> z \leftarrow pgmm(dynformula(n ~ 1, lag = list(c(1, 2))), effect = "twoways",
     model = "twosteps", Snmesp, gmm.inst = ~n, lag.gmm = c(2,
          99), transformation = c("d")
> summary(z)
Twoways effects Two steps model
pgmm(formula = n ~ lag(n, 1) + lag(n, 2), data = Snmesp, effect = "twoways",
   model = "twosteps", gmm.inst = ~n, lag.gmm = c(2, 99), transformation = c("d"))
Balanced Panel: n=738, T=8, N=5904
Number of Observations Used: 3690
Residuals
              1st Qu.
                          Median
                                       Mean
                                               3rd Qu.
-1.4530000 -0.0499300 -0.0002421 0.0000663 0.0520800 1.2020000
Coefficients
         Estimate Std. Error z-value Pr(>|z|)
lag(n, 1) 0.747547
                    0.088270 8.4688 < 2e-16 ***
                    0.021952 1.7165 0.08607 .
lag(n, 2) 0.037680
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Sargan Test: chisq(18) = 14.40912 (p.value=0.70206)
Autocorrelation test (1): normal = -5.947032 (p.value=1.3652e-09)
Autocorrelation test (2): normal = 0.2629247 (p.value=0.39630)
Wald test for coefficients: chisq(2) = 105.3612 (p.value=< 2.22e-16)
Wald test for time dummies: chisq(5) = 59.15637 (p.value=1.8156e-11)
```

2.5 General FGLS models

General FGLS estimators are based on a two-step estimation process: first an OLS model is estimated, then its residuals are used to estimate an error covariance matrix for use in a feasible-GLS analysis. Formally, the structure of the error covariance matrix is $V = I_N \otimes \Omega$, with symmetry being the only requisite for Ω : $\Omega(ij) = \Omega(ji)$ (see Wooldridge (2002), 10.4.3 and 10.5.5).

This framework allows the error covariance structure inside every group (if effect="individual") of observations to be fully unrestricted and is therefore robust against any type of intragroup heteroskedasticity and serial correlation. This structure, by converse, is assumed identical across groups and thus ggls is inefficient under groupwise heteroskedasticity. Cross-sectional correlation is excluded a priory.

Moreover, the number of variance parameters to be estimated with NT data points is T(T+1)/2, which makes these estimators particularly suited for situations where N >> T, as e.g. in labour or household income surveys, while problematic for "long" panels.

In a pooled time series context (effect="time"), symmetrically, this estimator is able to account for arbitrary cross-sectional correlation, provided that the latter is time-invariant (see Greene (2003) 13.9.1-2, p.321-2). In this case serial correlation has to be assumed away and the estimator is consistent with respect to the time dimension, keeping N fixed.

The function pggls estimates general FGLS models, with either fixed of "random" effects⁴.

The "random effect" general FGLS is estimated by

```
> zz <- pggls(log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
      data = Produc, model = "random")
> summary(zz)
Oneway (individual) effect Random effects model
Call:
pggls(formula = log(gsp) ~ log(pcap) + log(pc) + log(emp) + unemp,
    data = Produc, model = "random")
Balanced Panel: n=48, T=17, N=816
Residuals
     Min.
            1st Qu.
                       Median
                                    Mean
                                           3rd Qu.
                                                        Max.
-0.255700 -0.070200 -0.014120 -0.008909
                                         0.039120
                                                    0.455500
Coefficients
               Estimate
                         Std. Error z-value Pr(>|z|)
(intercept)
             2.26388494
                         0.10077679 22.4643 < 2.2e-16 ***
                         0.02004106 5.2725 1.346e-07 ***
log(pcap)
             0.10566584
log(pc)
             0.21643137
                         0.01539471 14.0588 < 2.2e-16 ***
```

0.01863632 38.2553 < 2.2e-16 ***

log(emp)

0.71293894

⁴The "random effect" is better termed "general FGLS" model, as in fact it does not have a proper random effects structure, but we keep this terminology for consistency with plm.

The fixed effects pggls (see WOOLDRIDGE (2002, p.276)) is based on estimation of a within model in the first step; the rest follows as above. It is estimated by

The pggls function is similar to plm in many respects (e.g., Hausman tests may be carried out on pggls objects much the same way they are done on plm ones). An exception is that the estimate of the group covariance matrix of errors (zz\$sigma, 17x17 matrix, not shown) is reported in the model objects instead of the usual estimated variances of the two error components.

3 Tests

3.1 Tests of poolability

pooltest tests the hypothesis that the same coefficients apply to each individual. It is a standard F test, based on the comparison of a model obtained for the full sample and a model based on the estimation of an equation for each individual. The main argument of pooltest is a plms or a plm object. The second argument is a pvcm object obtained with model=within. If the first argument is a plms object, a third argument effect should be fixed to FALSE if the intercepts are assumed to be identical (the default value) or TRUE if not⁵.

```
> form <- inv ~ value + capital
> znp <- pvcm(form, data = Grunfeld, model = "within")
> zplm <- plm(form, data = Grunfeld, model = "within")
> pooltest(zplm, znp)

        F statistic

data: inv ~ value + capital
F = 5.7805, df1 = 18, df2 = 170, p-value = 1.219e-10
alternative hypothesis: unstability

> z <- plm(form, data = Grunfeld, effect = "time")
> znpt <- pvcm(form, data = Grunfeld, effect = "time", model = "within")
> pooltest(z, znpt)

        F statistic

data: inv ~ value + capital
F = 1.5495, df1 = 38, df2 = 140, p-value = 0.03553
alternative hypothesis: unstability
```

⁵The following examples are from Baltagi (2001), pp. 57–58.

3.2 Tests for individual and time effects

3.2.1 Lagrange multiplier tests

plmtest implements tests of individual or/and time effects based on the results of the pooling model. It's main argument is a plm object (the result of a pooling model) or a plms object.

Two additional arguments can be added to indicate the kind of test to be computed. The argument type is whether :

- bp : Breusch-Pagan (1980), the default value,
- honda: HONDA (1985),
- kw: King and Wu (1997).

The effects tested are indicated with the ${\tt effect}$ argument :

- individual for individual effects (the default value),
- time for time effects,
- twoways for individuals and time effects.

Some examples of the use of plmtest are shown below⁶:

```
> library(Ecdat)
> g <- plm(inv ~ value + capital, data = Grunfeld, model = "pooling")
> plmtest(g)
        Lagrange Multiplier Test - (Honda)
data: inv ~ value + capital
normal = 28.2518, p-value < 2.2e-16
alternative hypothesis: significant effects
> plmtest(g, effect = "time")
        Lagrange Multiplier Test - time effects (Honda)
data: inv ~ value + capital
normal = -2.5404, p-value = 0.002768
alternative hypothesis: significant effects
> plmtest(g, type = "honda")
        Lagrange Multiplier Test - (Honda)
data: inv ~ value + capital
normal = 28.2518, p-value < 2.2e-16
alternative hypothesis: significant effects
> plmtest(g, type = "ghm", effect = "twoways")
  <sup>6</sup>See Baltagi (2001), p. 65.
```

```
Lagrange Multiplier Test - two-ways effects (Gourieroux, Holly and
        Monfort)
data: inv ~ value + capital
chisq = 798.1615, df = 2, p-value < 2.2e-16
alternative hypothesis: significant effects
> plmtest(g, type = "kw", effect = "twoways")
        Lagrange Multiplier Test - two-ways effects (King and Wu)
data: inv ~ value + capital
normal = 21.8322, df = 2, p-value < 2.2e-16
alternative hypothesis: significant effects
3.2.2 F tests
pFtest computes F tests of effects based on the comparison of the within and
the pooling models. Its arguments are whether a plms object or two plm objects
(the results of a pooling and a within model). Some examples of the use of
pFtest are shown below<sup>7</sup>:
> library(Ecdat)
> gp <- plm(inv ~ value + capital, data = Grunfeld, model = "pooling")
> gw <- plm(inv ~ value + capital, data = Grunfeld, model = "within")
> gt <- plm(inv ~ value + capital, data = Grunfeld, model = "within",
      effect = "time")
> gd <- plm(inv ~ value + capital, data = Grunfeld, model = "within",
      effect = "twoways")
> pFtest(gw, gp)
        F test for effects
data: inv ~ value + capital
F = 49.1766, df1 = 9, df2 = 188, p-value < 2.2e-16
alternative hypothesis: significant effects
> pFtest(gt, gp)
        F test for effects
data: inv ~ value + capital
F = 0.2345, df1 = 19, df2 = 178, p-value = 0.9997
alternative hypothesis: significant effects
> pFtest(gd, gw)
        F test for effects
data: inv ~ value + capital
F = 1.4032, df1 = 19, df2 = 169, p-value = 0.1309
alternative hypothesis: significant effects
```

⁷See Baltagi (2001), p. 65.

3.3 Hausman's test

phtest computes the Hausman's test which is based on the comparison of two models. It's main argument may be:

- a plms object. In this case, the two models used in the test are the within and the random models (the most usual case with panel data),
- two plm objects.

Some examples of the use of phtest are shown below 8:

```
> gw <- plm(inv ~ value + capital, data = Grunfeld, model = "within")
> gr <- plm(inv ~ value + capital, data = Grunfeld, model = "random")
> phtest(gw, gr)
```

Hausman Test

```
data: inv ~ value + capital
chisq = 2.3304, df = 2, p-value = 0.3119
alternative hypothesis: one model is inconsistent
```

3.4 Robust covariance matrix estimation

Robust estimators of the covariance matrix of coefficients are provided, mostly for use in Wald-type tests. pvcovHC estimates three "flavours" of White (1980, 1984)'s heteroskedasticity-consistent covariance matrix (known as the *sandwich* estimator). Interestingly, in the context of panel data the most general version also proves consistent vs. serial correlation.

All types assume no correlation between errors of different groups while allowing for heteroskedasticity across groups, so that the full covariance matrix of errors is $V = I_n \otimes \Omega_i$; i = 1, ..., n. As for the *intragroup* error covariance matrix of every single group of observations, "white1" allows for general heteroskedasticity but no serial correlation, i.e

$$\Omega_i = \begin{bmatrix}
\sigma_{i1}^2 & \dots & \dots & 0 \\
0 & \sigma_{i2}^2 & & \vdots \\
\vdots & & \ddots & 0 \\
0 & & & \sigma_{iT}^2
\end{bmatrix}$$
(1)

while "white2" is "white1" restricted to a common variance inside every group, estimated as $\sigma_i^2 = \sum_{t=1}^T e_{it}^2/T$, so that $\Omega_i = I_T \otimes \sigma_i^2$ (see Greene (2003), 13.7.1-2 and Wooldridge (2003), 10.7.2); "arellano" (see ibid. and the original ref. Arellano (1987)) allows a fully general structure w.r.t. heteroskedasticity and serial correlation:

⁸See Baltagi (2001), p. 71.

$$\Omega_{i} = \begin{bmatrix}
\sigma_{i1}^{2} & \sigma_{i1,i2} & \dots & & \sigma_{i1,iT} \\
\sigma_{i2,i1} & \sigma_{i2}^{2} & & & \vdots \\
\vdots & & \ddots & & \vdots \\
\vdots & & & \sigma_{iT-1}^{2} & \sigma_{iT-1,iT} \\
\sigma_{iT,i1} & \dots & \dots & \sigma_{iT,iT-1} & \sigma_{iT}^{2}
\end{bmatrix}$$
(2)

The latter is, as already observed, consistent w.r.t. timewise correlation of the errors, but on the converse, unlike the White 1 and 2 methods, it relies on large N asymptotics with small T.

The errors may be weighted according to the schemes proposed by MacKinnon and White (1985) and Cribari-Neto (2004) to improve small-sample performance.

Main use of pvcovHC is together with testing functions from lmtest and car packages. These typically allow passing the vcov parameter to be either a matrix or a function (see Zeileis 2004). If one is happy with the defaults, it is easiest to pass the function itself:

```
> library(lmtest)
> data("Airline", package = "Ecdat")
> form <- log(cost) ~ log(output) + log(pf) + lf</pre>
> z <- plm(form, data = Airline, model = "within")
> coeftest(z, pvcovHC)
t test of coefficients:
             Estimate Std. Error t value Pr(>|t|)
log(output) 0.919285
                        0.029498 31.1640 < 2.2e-16 ***
                        0.017362 24.0457 < 2.2e-16 ***
log(pf)
             0.417492
lf
            -1.070396
                        0.384669 -2.7826  0.006707 **
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
   else one may do the covariance computation inside the call to coeftest,
thus passing on a matrix:
> coeftest(z, pvcovHC(z, method = "white2", type = "HC3"))
t test of coefficients:
             Estimate Std. Error t value Pr(>|t|)
log(output)
            0.919285
                        0.029021 31.6769 < 2.2e-16 ***
             0.417492
                        0.014301 29.1928 < 2.2e-16 ***
log(pf)
lf
            -1.070396
                        0.211686 -5.0565 2.605e-06 ***
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' '1
```

For some tests, e.g. for multiple model comparisons by waldtest, one should always provide a function⁹. In this case, optional parameters are provided as shown below (see also Zeileis, 2004, p.12):

⁹Joint zero-restriction testing still allows providing the vcov of the unrestricted model as a matrix, see the documentation of package lmtest

```
> waldtest(z, update(z, . ~ . - log(pf) - lf), vcov = function(x) pvcovHC(x,
      method = "white2", type = "HC3"))
Wald test
Model 1: log(cost) ~ log(output) + log(pf) + lf
Model 2: log(cost) ~ log(output)
 Res.Df Df Chisq Pr(>Chisq)
2
      83 -2 858.92 < 2.2e-16 ***
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
  linear.hypothesis from package car may be used to test for linear restric-
tions:
> library(car)
> linear.hypothesis(zz, "2*log(pc)=log(emp)", vcov = pvcovHC)
Linear hypothesis test
Hypothesis:
2 \log(pc) - \log(emp) = 0
Model 1: log(gsp) \sim log(pcap) + log(pc) + log(emp) + unemp
Model 2: restricted model
Note: Coefficient covariance matrix supplied.
 Res.Df Df Chisq Pr(>Chisq)
     811
1
2
     812 -1 2.2928
                        0.1300
```

4 References

Amemiyia, T. (1971), The estimation of the variances in a variance–components model, *International Economic Review*, 12, pp.1–13.

Arellano M. (1987), Computing robust standard errors for within group estimators, Oxford bulletin of Economics and Statistics, 49, 431–434.

Arellano, M. (2003), Panel data econometrics, Oxford University Press.

Arellano M. and S. Bond (1991), Some tests of specification for panel data: monte carlo evidence and an application to employment equations, *Review of Economic Studies*, 58, pp.277–297.

Balestra, P. and J. Varadharajan–Krishnakumar (1987), Full information estimations of a system of simultaneous equations with error components structure, *Econometric Theory*, 3, pp.223–246.

Baltagi, B.H. (1981), Simultaneous equations with error components, *Journal* of econometrics, 17, pp.21–49.

Baltagi, B.H. (2001) Econometric Analysis of Panel Data. John Wiley and sons. ltd.

Breusch, T.S. and A.R. Pagan (1980), The Lagrange multiplier test and its applications to model specification in econometrics, *Review of Economic Studies*, 47, pp.239–253.

Cribari-Neto F. (2004), Asymptotic inference under heteroskedasticity of unknown form. Computational Statistics & Data Analysis 45, 215–233.

Gourieroux, C., A. Holly and A. Monfort (1982), Likelihood ratio test, Wald test, and Kuhn–Tucker test in linear models with inequality constraints on the regression parameters, *Econometrica*, 50, pp.63–80.

Greene W. H. (2003), Econometric Analysis, 5th ed. Prentice Hall.

Hausman, G. (1978), Specification tests in econometrics, *Econometrica*, 46, pp.1251–1271.

Hausman, J.A. and W.E. Taylor (1981), Panel data and unobservable individual effects, *Econometrica*, 49, pp.1377–1398.

Honda, Y. (1985), Testing the error components model with non–normal disturbances, *Review of Economic Studies*, 52, pp.681–690.

King, M.L. and P.X. Wu (1997), Locally optimal one–sided tests for multiparameter hypotheses, *Econometric Reviews*, 33, pp.523–529.

MacKinnon J. G., White H. (1985), Some heteroskedasticity-consistent covariance matrix estimators with improved finite sample properties. *Journal of Econometrics* **29**, 305–325.

Nerlove, M. (1971), Further evidence on the estimation of dynamic economic relations from a time–series of cross–sections, *Econometrica*, 39, pp.359–382.

Swamy, P.A.V.B. (1970), Efficient inference in a random coefficient regression model, *Econometrica*, 38, pp.311-323.

Swamy, P.A.V.B. and S.S. Arora (1972), The exact finite sample properties of the estimators of coefficients in the error components regression models, *Econometrica*, 40, pp.261–275.

Wallace, T.D. and A. Hussain (1969), The use of error components models in combining cross section with time series data, *Econometrica*, 37(1), pp.55–72.

White H. (1980), Asymptotic Theory for Econometricians, Ch. 6, Academic Press, Orlando (FL).

White H. (1984), A heteroskedasticity-consistent covariance matrix and a direct test for heteroskedasticity. *Econometrica* 48, 817–838.

Wooldridge J. M. (2003), Econometric Analysis of Cross Section and Panel Data, MIT Press

Zeileis A. (2004), Econometric Computing with HC and HAC Covariance Matrix Estimators. *Journal of Statistical Software*, **11**(10), 1–17.

URL http://http://www.jstatsoft.org/v11/i10/.