

* Introduction to fundamentals :

Standard data transformation for time series is lagging the data then taking successive differences.

* Mean : $\mu_{x_t} = \mathbb{E}(x_t)$

$$\mu_{w_t} = \mathbb{E}(w_t) = 0 \quad \text{where } w_t \sim N(0, 1) \text{ white noise}$$

$$\mu_{v_t} = \mathbb{E}(v_t) = \frac{1}{3} (\mathbb{E}(w_{t+1}) + \mathbb{E}(w_t) + \mathbb{E}(w_{t-1})) = 0 \quad \text{where}$$

v_t is the moving average series (smoothing series)

$$\mu_{x_t} = \mathbb{E}(x_t) + \mathbb{E}(\sigma_t + \sum_{j=1}^t w_j) = \sigma_t + \sum_{j=1}^t \mathbb{E}(w_j)$$

where x_t is a random walk with a drift

$$\mu_{x_t} = \mathbb{E}(\cos(2\pi \frac{t+1s}{50}) + w_t) = \cos(2\pi \frac{t+1s}{50}) + 0$$

The mean at x_t is a cosine wave

* Autocovariance : $\gamma_n(s, t) = \text{Cov}(n_s, n_t)$
 $= \mathbb{E}((n_s - \mu_s)(n_t - \mu_t))$

If $\gamma_n(s, t) = 0$ then n_t and n_s are not linearly dependent

$$\gamma_w(s, t) = \begin{cases} \sigma^2, & s=t \\ 0 & s \neq t \end{cases} \quad \text{as } w_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

$$\text{Cov}\left(\sum_{j=1}^m a_j X_j, \sum_{k=1}^r b_k Y_k\right) = \sum_{j=1}^m \sum_{k=1}^r a_j b_k \text{Cov}(X_j, Y_k)$$

* Autocorrelation Function (ACF) : $\rho(s, t) = \frac{\gamma(t, s)}{\sqrt{\gamma(t, t) \gamma(s, s)}}$

- * Stationarity : 1) μ_t constant and independent of t
 2) $\gamma(s, t)$ depends on times t and s only

Ranaleem walk is not stationary as $\gamma(s, t) = \min(s, t)$
 σ^2 depends on time.

When referring to a stationary time series process we can use $\mu = \mu_t$ as it is independent of t and not only depends on t and s so suppose $s = h + t$ where h is the time shift or lag.

$$\gamma(t, t+h) = \text{Cov}(n_t, n_{t+h}) = \text{Cov}(n_0, n_h) = \sigma(0, h)$$

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)}, \quad \sigma(h) = \text{Cov}(n_{t+h}, n_t)$$

White noise and moving average are stationary

- * Trend Stationarity : $n_t = \beta t + \gamma_t$, γ_t stationary

$$1) \mu_{n_t} = E(n_t) = \beta t + \mu_\gamma$$

$$\begin{aligned}
 2) \sigma_n(h) &= \text{Cov}(n_{t+h}, n_t) \\
 &= \mathbb{E}((n_{t+h} - \mu_{n_{t+h}})(n_t - \mu_n)) \\
 &= \mathbb{E}((y_{t+h} - \mu_y)(y_t - \mu_y)) \\
 &= \sigma_y(h)
 \end{aligned}$$

$$3) \sigma(0) = \mathbb{E}((n_t - \mu)^2) = \text{Var}(n_t)$$

$$4) \sigma(h) = \sigma(-h)$$

* Wold Decomposition : Any stationary time series n_t can be written as linear comb. (filter) of white noise terms

$$n_t = \sum_{j \geq 0} \Psi_j w_{t-j} + \mu$$

$\sum_{j \geq 0} \Psi_j^2 < +\infty$ & $\Psi_0 = 1$, we call them linear processes.

* Jointly Stationary : If both n_t & y_t are stationary and their cross-covariance function is at lag h

$$\begin{aligned}
 \sigma_{ny}(h) &= \text{Cov}(n_{t+h}, y_t) \\
 &= \mathbb{E}((n_{t+h} - \mu_n)(y_t - \mu_y))
 \end{aligned}$$

* Ordinary Least Squares for time series :

$$* n_t = \beta_0 + \sum_{j=1}^{q-1} \beta_j z_{t-j}$$

$$* S = \sum_{i=1}^n w_i^2 = \sum_{i=1}^n (n_i - \beta_0 - \sum_{j=1}^{q-1} \beta_j z_{t+j})^2$$

We solve for $\frac{\partial S}{\partial \beta_i} = 0 \quad \forall i=0, \dots, q-1$ to get the coefs.

$$* SSE = \sum_{i=1}^n (n_i - \hat{n}_i)^2$$

$$* E(MS_{Res}) = E\left(\frac{SSE}{n-q}\right) = \sigma^2$$

$$* t = \frac{\hat{\beta}_j - \beta_j}{SE(\beta_j)}, \quad \text{in OLS } SE(\beta_j) = \sqrt{(X^T X)^{-1}_{jj}} MS_{Res}$$

$$* F = \frac{MS_{Reg}}{MS_{Res}}$$

$$* \text{Coefficient of determination: } R^2 = \frac{SS_{Reg}}{SST}$$

$$* \text{Akaike's Information Criterion: } AIC = \log\left(\frac{SSE_k}{n}\right) + \frac{k+2}{n}$$

$$* \text{Bayesian Information Criterion: } BIC = \log\left(\frac{SSE_k}{n}\right) + \frac{k \log n}{n}$$

* AIC is better with small samples while BIC is superior in larger samples.

* Suppose the data is not stationary:

1) Detrending: $n_t = r_t + y_t, \quad r_t = \beta_0 + \beta_1 t$

$$\hat{y}_t = \hat{n}_t - \hat{r}_t$$

2) Differencing: $\nabla n_t = n_t - n_{t-1}$ removes linear trend
 $\nabla^2 n_t = n_t - 2n_{t-1} + n_{t-2}$ removes quadratic trend

Differencering af ørcler d : $\nabla^d = (1 - B)^d$

$$\begin{aligned}\nabla^3 u_t &= (1 - B)^3 u_t \\&= (1 - 3B + 3B^2 - B^3) u_t \\&= u_t - 3u_{t-1} + 3u_{t-2} - u_{t-3}\end{aligned}$$

3) Heteroscedasticity : $y_t = \log(u_t)$ transformativen

4) Box-Cox : $y_t = \begin{cases} \frac{u_t^\lambda - 1}{\lambda}, & \lambda \neq 0 \\ \log(u_t), & \lambda = 0 \end{cases}$

* Classical Structural Modelling :

We decompose $u_t = T_t + S_t + N_t + C_t$
↑ ↑ ↑ ↓
trend seasonal noise cyclic

* ARIMA Modelling

* Autoregressive model AR(p) :

$u_t = \alpha + \phi_1 u_{t-1} + \phi_2 u_{t-2} + \dots + \phi_p u_{t-p} + w_t$
where u_t is stationary and w_t is white noise

* The models $u_t = \pm u_{t-1} + w_t$ are not AR as they are a random walk which is not stationary

If u_t is an AR model then $|\phi| < 1$ $u_t = \sum_{j=0}^p \phi_j w_{t-j}$

- * Causal time series : $n_t = \mu + \sum_{j \geq 0} \phi_j w_{t-j}$
- * Moving Average Model at order q :

$$n_t = w_t + \theta_1 w_{t-1} + \theta_2 w_{t-2} + \dots + \theta_q w_{t-q}$$
 always stationary
- * Invertible time series : $w_t = \sum_{j \geq 0} \pi_j n_{t-j}$; $\sum_{j \geq 0} \pi_j^2 < +\infty$
- * ARMA(p,q) : $n_t = \alpha + \phi_1 n_{t-1} + \dots + \phi_p n_{t-p} + w_t + \theta_1 w_{t-1} + \dots + \theta_q w_{t-q}$
 where n_t is causal and invertible

$$n_t = \beta_0 + \beta_1 n_{t-1} + \dots + \beta_p n_{t-p} + \epsilon_t \text{ where}$$

$$\epsilon_t = w_t + \theta_1 w_{t-1} + \dots + \theta_q w_{t-q}$$

ϕ is restricted to obtain causality
 θ is restricted to obtain invertibility

- * Model redundancy :

$$n_t = 0.3 n_{t-1} + 0.4 n_{t-2} + w_t + 0.5 w_{t-1} \text{ looks like ARMA(2,1)}$$

$$(n_t - 0.3 n_{t-1} - 0.4 n_{t-2}) = w_t + 0.5 w_{t-1}$$

$$(1 - 0.3B - 0.4B^2) n_t = (1 + 0.5B) w_t$$

$$(1 + 0.5B)(1 - 0.8B) n_t = (1 + 0.5B) w_t$$

$$(1 - 0.8B) n_t = w_t$$

Hence, $n_t = 0.8 n_{t-1} + w_t$ which is AR(1)

We can check this on R by observing the roots of the polynomial.

* Partial Autocorrelation Function (PACF) :

$$\phi_{hh} = \rho(u_h - \hat{u}_h, u_0 - \hat{u}_0), \quad h \geq 2$$

$$\phi_{11} = \rho(u_1, u_0)$$

where \hat{u}_h is the regression of u_h on $\{u_1, \dots, u_{h-1}\}$
and \hat{u}_0 is the regression of u_0 on $\{u_1, \dots, u_{h-1}\}$