

**Donner Party Survival: A Reassessment With Nonlinear Regression With R**

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### **Abstract**

The survivorship data for the 1846-1847 Donner Party is used to demonstrate the use of restricted cubic splines (rcs), Generalized Additive Models (GAMs), and survivorship curves for the analyses of survival, a binary response variable. The Donner Party data are challenging because there were only 87 members of the traveling party. Of those 87 travelers, 8 died of causes other than weather and starvation, leaving only 79 travelers for the key analyses. Despite the small sample size, the rcs and GAM analyses reveal striking curvilinear patterns in survivorship, patterns not previously noted by Donner scholars. Survivorship was nearly perfect for females aged 5 to 36, whereas males had lower survival decreasing linearly with age. Of the 79 snow-trapped travelers, no females and only one male past age 40 survived. Only 1 female (Age 25) in the range of 4 to 36 died. Survival was perfect for two families of size 6 and 9, but much lower for larger and smaller groups. Kaplan-Meier survivorship curves revealed the poor survival of males and especially employees, like teamsters. Males and employees died earlier and at much higher rates than females and family members.

### **Key words**

Cox survivorship, Generalized Additive Models, *k*-fold cross-validation, Kaplan-Meier, restricted cubic splines

## 1. INTRODUCTION

The Donner Party expedition is an American tragedy. The Donner-Reed wagon train with 87 travelers followed the recommendation of Lansford Hastings to take an untried shortcut to California across the Sierra Nevada Mountains and were trapped by a huge 28 October 1846 snowstorm. Only 47 members of the party survived until the final rescue party on 22 April 1847. Most Donner Party members, survivors and non-survivors, resorted to cannibalism.

Grayson (1990, 1994, 1997, 2018) analyzed the demographics of Donner Party survival, concluding “...survivorship within the party was mediated almost entirely by three factors: age, sex and the size of the kin group with which each member traveled” (Grayson 1990, p. 223). Grayson’s later papers and 2018 provide more support for the importance of those variables. Rarick (2008) also added a fourth important pattern, the low rates of survival of the teamsters and servants who worked for the families. Note, that I follow Grayson (2018) in using Sex rather than Gender to analyze the differential mortality among Donner Party travelers, acknowledging that gender-based differences in behavior could very well have played a role in survival.

Ramsey & Schafer (2013) analyzed a subset of the data for individuals aged greater than or equal to 15 and used the Donner Party data in their textbook *Statistical Sleuth* to introduce binary logistic regression. Their final model of Donner Party survival was an additive binary generalized linear logistic regression model using Age and Sex as explanatory variables. They briefly discuss an interaction model, but they did not analyze the survivorship of the 42 individuals deleted from their analysis nor did they consider Family Group Size or the survivorship of teamsters and

servants. The use of those deleted travelers will allow me to present curvilinear regression and survivorship analyses, the keys to understanding Donner Party survival.

I'll analyze the effects of Sex, Age, and Family Group Size using restricted cubic spline regressions within a binomial Generalized Linear Model (Harrell 2015) and Generalized Additive Models (GAMs, Wood 2017, 2019). I'll then analyze survivorship with Cox Proportional Hazards models, Kaplan-Meier survivorship curves, and Fisher's exact hypergeometric test.

For the statistics instructor, these data provide an interesting case study for the introduction of binary logistic regression, nonlinear regression using restricted cubic splines, generalized additive models (GAMs) and survival analysis. The following analyses are in the spirit of Harrell (2015) and Andrews (2021) whose textbooks introduce binary logistic regression with both restricted cubic splines and GAMs.

## **2. ANALYSIS OF DATA**

All analyses were performed with R (R Core Team 2023) with the following R packages: *caret* for GAM *k*-fold cross-validation, *mgcv* (Wood 2017, 2019) for GAMs, *rms* (Harrell 2015) for Generalized linear models and restricted cubic splines, *survival* & *survminer* for survivorship analyses, and the *tidyverse* (Wickham & Grolemund 2017) which includes *dplyr* and *ggplot2*.

The code and data are available from a github site (links provided). The demographic data were taken from Grayson (2018) Table 2.1. Age, occupation, cause of death, and date of death were

obtained from Stewart (1960), Rarick (2008), Brown (2009), and Grayson (2018). When there were conflicting dates or ages, Grayson's (2018) dates based on more recent primary sources were used.

I'll test combinations of four covariates—Age, Sex, Family Group Size and Teamster & Servant status— on pared data ( $n=79$ ) from which the 8 travelers who died or crossed the Sierra Nevada mountains before the first major snowfall on 28 October 1846 have been deleted. The primary statistical method is binomial logistic regression using Harrell's (2015) Generalized Linear Model function (`rms::Glm`) or the base R `glm` function. Family Group Size differs from Family Size based on surnames. For example, Elizabeth Donner traveled with two children from a previous marriage: Solomon Hook (14) and William Hook (12); the Fosdicks traveled with the Graves family, and the Fosters traveled with the Murphy family, and so on. Grayson (2018) provides Rarek (2008 pp. x-xi) and Grayson (2018) provide similar classifications of families, based largely on Stewart (1960). Age and Family Group Size will be tested as restricted cubic splines using Harrell's (2015) `rms` packages' `rcs` function and as GAMs using Wood's (2019) `mgcv` package. As recommended by Harrell (2021), the Akaike Information Content (AIC) was used to find the appropriate number of knots for the restricted cubic spline regression. That  $k$  was usually the lowest AIC, hence the highest likelihood after penalization for the number of parameters. Preserving degrees of freedom and reducing overfitting was a high priority, so a lower number of knots, with 3 being the minimum allowed, was used if its AIC was within 4 units of the minimum AIC found with a higher number of knots. The  $AIC < 4$  threshold was chosen because Burnham and Anderson (2004 p. 271) and Bolker (2015 p. 210) regard AICs less than 2 apart as roughly equivalent, AICs 4-7 apart as clearly distinguishable, and models with

AICs more than 10 apart as definitely different. This 4-unit AIC threshold serves as an additional penalty in selecting the number of knots for the rcs models above that already incorporated in the AIC calculation. In two cases of rcs regression, the AIC minimization criterion produced fits that appeared to be based on too many knots. For those cases,  $k$ -fold cross validation using the log loss (binomial deviance) for minimization was used to find a lower rcs knot size.

There are three different parameters designated  $k$  in the analyses reported here and in the curvilinear regression literature. First, the number of knots in rcs regression is  $k$ . Second,  $k$  is the basis function parameter controlling the smoothing of GAMs, which Wood (2017) sometimes refers to as knots. Finally, there is  $k$ -fold cross validation using the caret package, which is used to find the appropriate GAM basis function (smoothing parameter) or rcs knot number that produces the lowest log loss (binomial deviance) statistic. For  $k$ -fold cross validation, ten folds were used, so that a tenth of the cases were removed for each of 10 analyses, and the average log loss when predicting the removed data was used to judge the fit of the model for the GAM  $k$  parameter or rcs knot size. Once the GAM  $k$  or rcs knot size or sizes were determined, the full data ( $n = 79$ ) was fit with the appropriate GAM basis function.

Tests of null hypotheses use Wald Chi-square and Z tests and Wilks' drop-in-deviance tests, described in Ramsey & Schafer (2013).

Cox proportional hazards models and Kaplan-Meier survivorship curves were generated using the survival and survminer packages. R coding was assisted by Open AI's GPT-4.

### 3. RESULTS

#### 3.1. Effects of Age and Sex analyzed with a restricted cubic spline regression

There are 79 Donner Party travelers analyzed: the 87 original members of the Party minus five who died before the major snowfall (Halloran, Hardcoop, Pike, Snyder, Wolfinger) and the three (William McCutcheon, James Reed & Walter Herron) who crossed the Sierras before the 28 October 1846 storm that stopped the Donner Party's advance. The effects of Age and Sex and their interaction were analyzed with a binomial logistic regression with a restricted cubic spline regression for Age with three knots. The age variable was fit with a 3-knot restricted cubic spline with AIC=95.408, which was within 4 AIC units of the 4-knot AIC=92.358. The results are shown in Table 1 and Figure 1.

The most striking pattern in Figure 1 is the strong curvilinear pattern in female survivorship. From ages 4 to 36, only one female died, Eleanor Eddy aged 25. There was a tremendous difference between female and male survival, with only one male in his 20's surviving: William Eddy aged 28, the husband of Eleanor Eddy. Only 2 of 8 of the 79 Donner Party travelers aged 40 or above survived. Two others of the original 87 were also over 40: James Reed aged 45 went over the Sierras before the 28 October snowfall, and Hardcoop aged 60 died on 8 October.

**Table 1.** Wald & Wilks statistics and effect sizes for the rcs(Age, 3 knot) \* Sex binary generalized linear model.

**1.1 Formulae:** Null Model: Status ~ 1

Model 1: Status ~ rcs(Age, 3) \* Sex

Model 2: Status ~ rcs(Age, 3) + Sex

**1.2 Wald Statistics from glm**

| <b>Coefficients</b>     | <b>Estimate</b> | <b>Std. Error</b> | <b>z value</b> | <b>Pr(&gt; z )</b> |
|-------------------------|-----------------|-------------------|----------------|--------------------|
| (Intercept)             | -0.8355         | 0.8105            | -1.031         | 0.30260            |
| rcs(Age, 3)Age          | 0.3912          | 0.1435            | 2.727          | 0.00640            |
| rcs(Age, 3)Age'         | -0.9380         | 0.3359            | -2.792         | 0.00523            |
| SexMale                 | 0.8534          | 1.0756            | 0.793          | 0.42754            |
| rcs(Age, 3)Age:SexMale  | -0.4029         | 0.1577            | -2.554         | 0.01064            |
| rcs(Age, 3)Age':SexMale | 0.8973          | 0.3588            | 2.501          | 0.01239            |

**1.3 Wilks (Drop-in-Deviance) Tests** The interaction model had considerably more explanatory

value than the null model, and the interaction model had more explanatory value than the additive model.

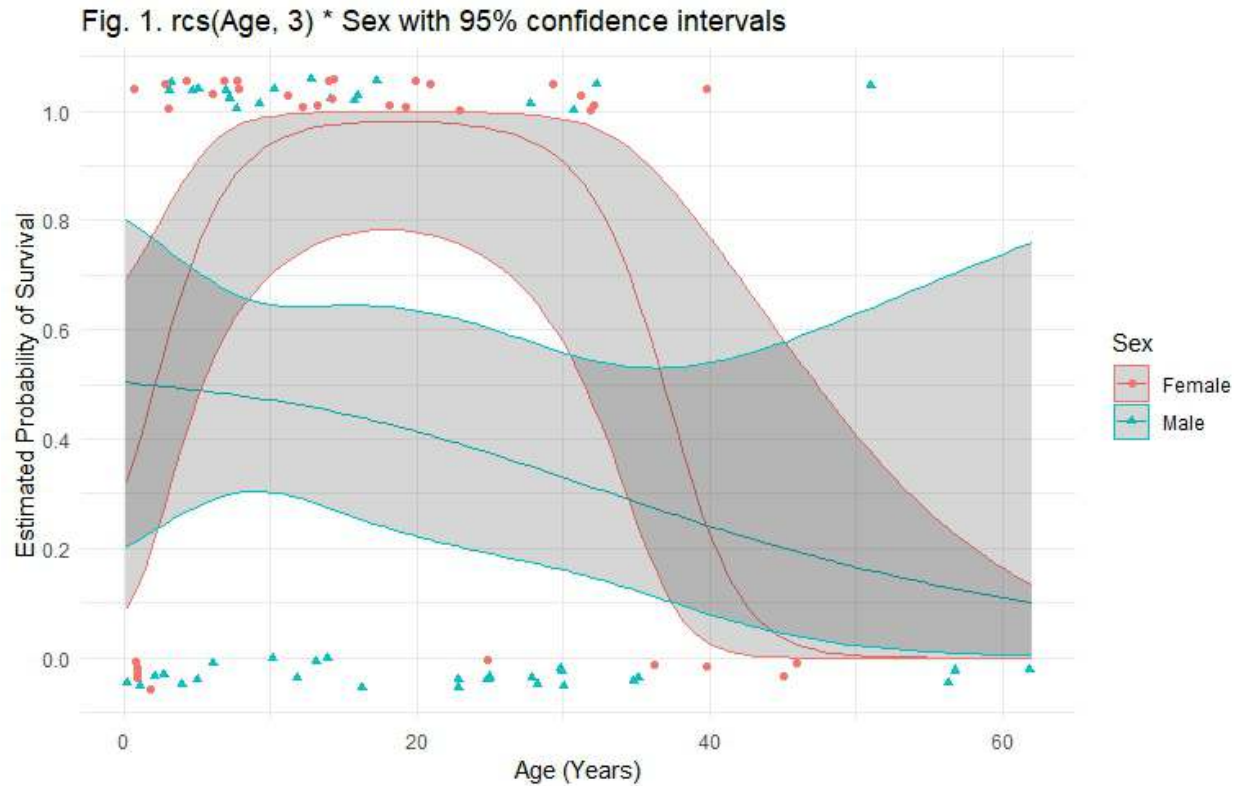
|                          | <b>Deviance</b> | <b>Chi-square</b> | <b>df</b> | <b>P</b>          |
|--------------------------|-----------------|-------------------|-----------|-------------------|
| Null Model:              | 109.201         | 78                |           |                   |
| Residual Model 1:        | 83.408          | 73                |           |                   |
| <b>Model 1 vs. Null:</b> | <b>25.793</b>   | <b>5</b>          |           | <b>&lt;0.0001</b> |
| Residual Model 2:        | 92.651          | 75                |           |                   |
| <b>Model 1 vs. 2:</b>    | <b>9.244</b>    | <b>2</b>          |           | <b>&lt;0.01</b>   |

**1.4 Effect sizes from Harrel's rms::Glm**

A 14-y old female's odds of survival were 45 times higher ( $\exp(3.814)$ ) than a 14-y old male (95% CI: 3.5 to 590 times). A 24-y old male's survival odds were only 27% of a 14 year old ( $\exp(10 \cdot -0.59288) \cdot 100$ ), but the 95% Confidence Intervals (CIs) are huge:  $<1e-6$  to 88,000). The male curve and confidence intervals in Figure 1 at ages 14 and 24 shows why the effect CIs are so broad.

| <b>Factor</b>                | <b>Effect</b> | <b>S.E.</b> | <b>Lower 0.95</b> | <b>Upper 0.95</b> |
|------------------------------|---------------|-------------|-------------------|-------------------|
| Age                          | -0.59288      | 0.63738     | -1.8632           | 0.67743           |
| Sex - Female:Male            | 3.81400       | 1.28540     | 1.2523            | 6.37570           |
| Adjusted to: Age=14 Sex=Male |               |             |                   |                   |





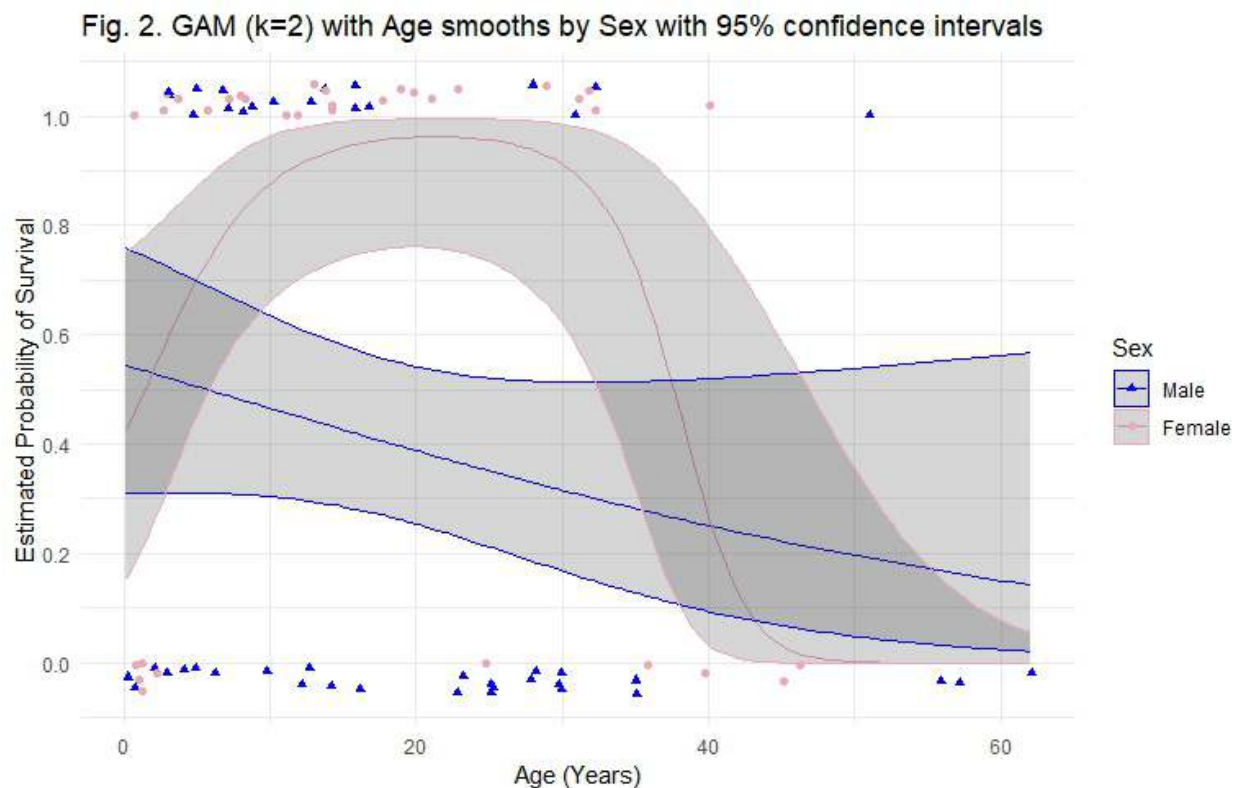
**Figure 1.** Display of the restricted cubic spline regression of Survival versus the interaction of Age (3 knots) and Sex. Between the ages of 4 and 36 only one Female died (Eleanor Eddy, Age 25). Only two of 79 Donner travelers survived past age 40: Margaret Breen (40) and Patrick Breen (51).

### 3.2 Effects of Age and Sex analyzed with a Generalized Additive Model

The effects of Age and Sex on survivorship were analyzed using a GAM. A  $k$ -fold cross-validation determined the best basis function or smoothing parameter  $k$  for the GAM. With  $k = 2$ , the Root Mean Square Error (RMES) was minimized. The results of the GAM analysis of Age, Sex and their interaction are shown in Table 2 and Figure 2.

**Table 2.** Statistics for the binomial GAM of Survival versus Age\* Sex with a logit link with the GAM basis function  $k = 2$ , determined by a  $k$ -fold cross validation. The effective degrees of freedom or edf indicates a quadratic pattern (Wood 2017, p. 83). UBRE is the unbiased risk estimator (Wood, 2017, p. 255).

|  | Estimate | Std. Error | z value | Pr(> z ) |         |
|--|----------|------------|---------|----------|---------|
| (Intercept)  | 9.193    | 3.127      | 2.94    | 0.00328  |         |
| <b>Approximate significance of smooth terms:</b>             |          |            |         |          |         |
|  |          | edf        | Ref.df  | Chi.sq   | p-value |
| s(Age):as.numeric(Sex == "Male")                             |          | 2.0        | 2.00    | 11.264   | 0.00358 |
| s(Age):as.numeric(Sex == "Female")                           |          | 1.9        | 1.99    | 7.495    | 0.02286 |
| R-sq.(adj) = 0.229    Deviance explained = 21.9%             |          |            |         |          |         |
| UBRE = 0.20352    Scale est. = 1                      n = 79 |          |            |         |          |         |



**Figure 2.** Display of the GAM of Sex and Age, with the model basis function  $k$  determined to be 2 by a  $k$ -fold cross-validation analysis. The curvilinear relationship is similar to that shown in Figure 1 using a restricted cubic spline regression.

### 3.3 Effects of Family Group Size analyzed with restricted cubic spline regression

The rcs regression revealed that Family Group Size was strongly related to Survival as shown by Wilks drop in deviance tests (Table 3), but the relationship was not linear as shown in Figure 3.

The AIC for knots 3 through 7 were 103.10, 100.96, **100.414**, 98.8629 and 96.8564 making 5 knots the appropriate choice based on an AIC threshold difference of 4 AIC units.

**Table 3.** Wald & Wilks statistics and effect sizes for the restricted cubic spline regression of Survival versus Family Group Size. Despite the somewhat large Wald p-values, the Wilks Drop-in-deviance chi square test and Figure 3 clearly show that the model is explaining more than chance variability in Donner Party Survivorship.

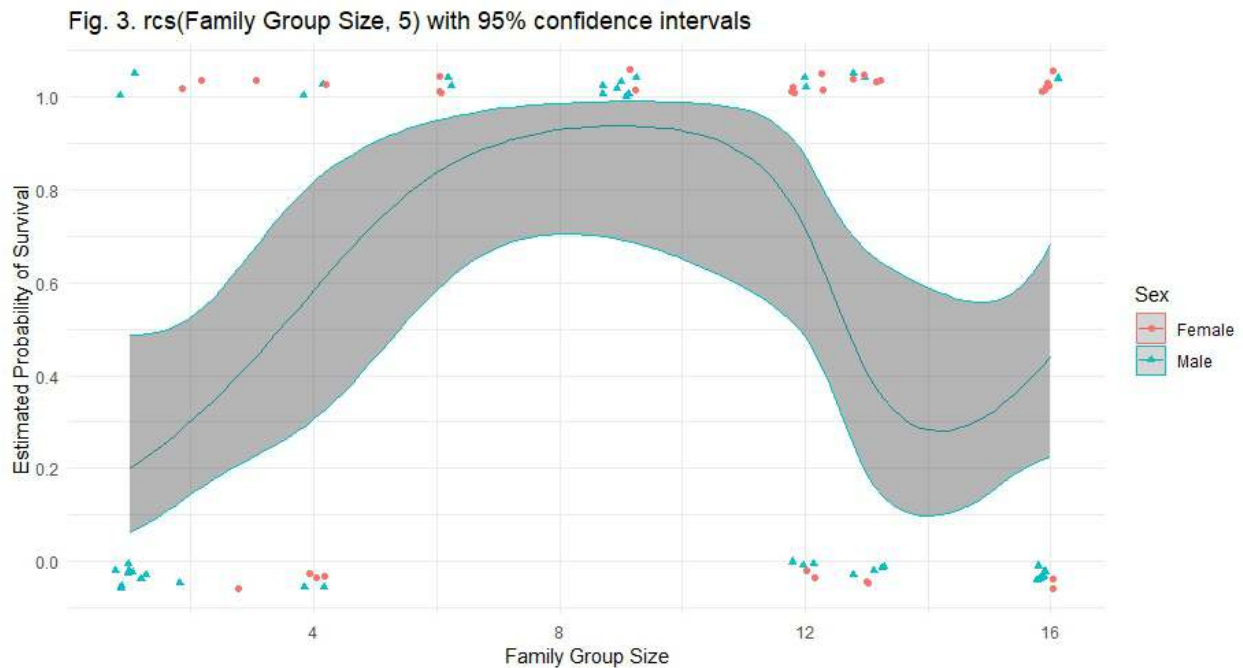
**3.1 Formulae:** Null Model: Status ~ 1  
Model 3: Status ~ rcs(Family\_Group\_Size, 5)

### 3.2 Wald Statistics from glm.

| Coefficients:                                 | Estimate | Std. Error | z value | Pr(> z ) |
|---|----------|------------|---------|----------|
| (Intercept)                                   | -1.9266  | 0.9855     | -1.955  | 0.051    |
| rcs(Family_Group_Size, 5)Family_Group_Size    | 0.5424   | 0.3839     | 1.413   | 0.158    |
| rcs(Family_Group_Size, 5)Family_Group_Size'   | 0.7826   | 2.4304     | 0.322   | 0.747    |
| rcs(Family_Group_Size, 5)Family_Group_Size''  | -3.6399  | 4.7616     | -0.764  | 0.445    |
| rcs(Family_Group_Size, 5)Family_Group_Size''' | 63.8362  | 36.1297    | 1.767   | 0.077    |

**3.3 Wilks (Drop-in-Deviance) Test** The 5-knot Family Group Size rcs model had considerably more explanatory value than the null model.

|                    | Chi-square | df |            |
|--------------------|------------|----|------------|
| Null deviance:     | 109.201    | 78 |            |
| Residual deviance: | 90.414     | 74 |            |
| Drop in deviance:  | 18.787     | 4  | p < 0.0001 |



**Figure 3.** Effect of Family Group Size on Survival, modeled with a restricted cubic spline with 5 knots. In Family Group Sizes of six (the Reed Family) and nine (the Breen family), everyone survived. Lone travelers (Family Group Size 1) had the lowest survival.

### 3.4 Effects of Family Group Size analyzed with a Generalized Additive Model

A  $k$ -fold cross-validation determined that the GAM basis function  $k=5$  minimized the RMSE.

Similar to the rcs analysis, the GAM revealed that Family Group Size was strongly related to Survival as shown by Wald tests (Table 4), but the relationship was not linear (Figure 4). As with rcs regression (Figure 3), lone travelers in Family Group Size 1 had the lowest survival.

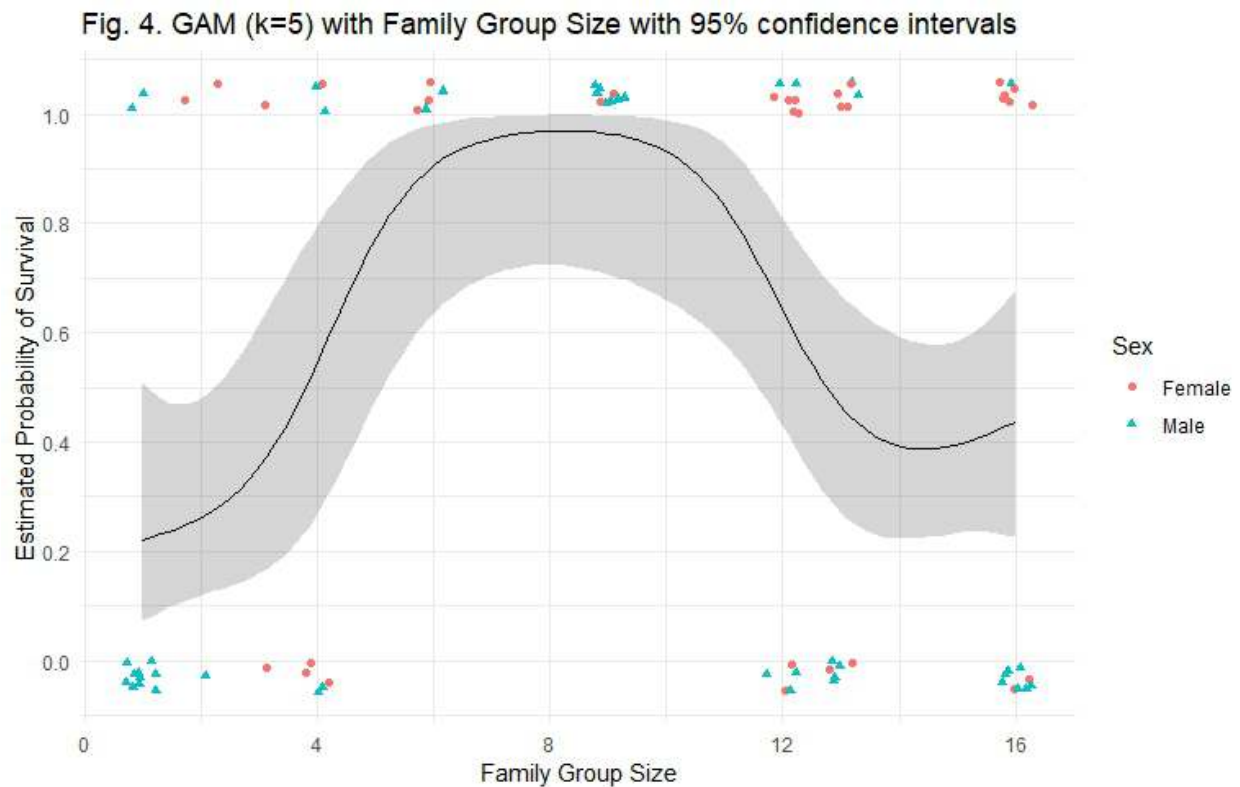
**Table 4.** Statistics for the binary logistic GAM model of Family Groups Size with the basis function (GAM smoothing parameter)  $k=5$ .

|             | <b>Estimate</b> | <b>Std. Error</b> | <b>z value</b> | <b>Pr(&gt; z )</b> |
|-------------|-----------------|-------------------|----------------|--------------------|
| (Intercept) | 0.3082          | 0.2941            | 1.048          | 0.295              |

Approximate significance of smooth terms:

|                      | <b>edf</b> | <b>Ref.df</b> | <b>Chi.sq</b> | <b>p-value</b> |
|----------------------|------------|---------------|---------------|----------------|
| s(Family_Group_Size) | 3.766      | 3.957         | 11.61         | 0.0195         |

R-sq.(adj) = 0.173    Deviance explained = 19.2%  
 UBRE = 0.23719    Scale est. = 1    n = 79



**Figure 4.** Effect of Family Group Size on Survival, modeled with a GAM with  $k=5$ , chosen using  $k$ -fold cross validation. Most of the lone travelers in Family Group Size 1 were employees.

### 3.5 Simultaneous Analysis of Sex, Age, and Family Group Size With Restricted Cubic Splines

One way to analyze the joint effects of Sex, Age, and Family Group Size is to model Survival as curved surfaces resulting from the action of all three variables. In this model, an AIC analysis indicated that 6 knots should be used for Age and 6 knots for Family Group Size. The 6,6 model had the lowest AIC value by far (56.328), but the 3-d curved surface plots were full of spikes due to single observations. A *k*-fold cross validation analysis found that Age with 3 knots and Family Group Size with 5 knots was appropriate even though the AIC was much higher at 78.369. Table 5 displays the Wald Chi-square tests and Figure 5 shows a three-dimensional view of survivorship with the *k* values determined by *k*-fold cross validation.

**Table 5. Wald statistics for the rcs(Age, 3) \*Sex + rcs(Family Group Size, 5)**

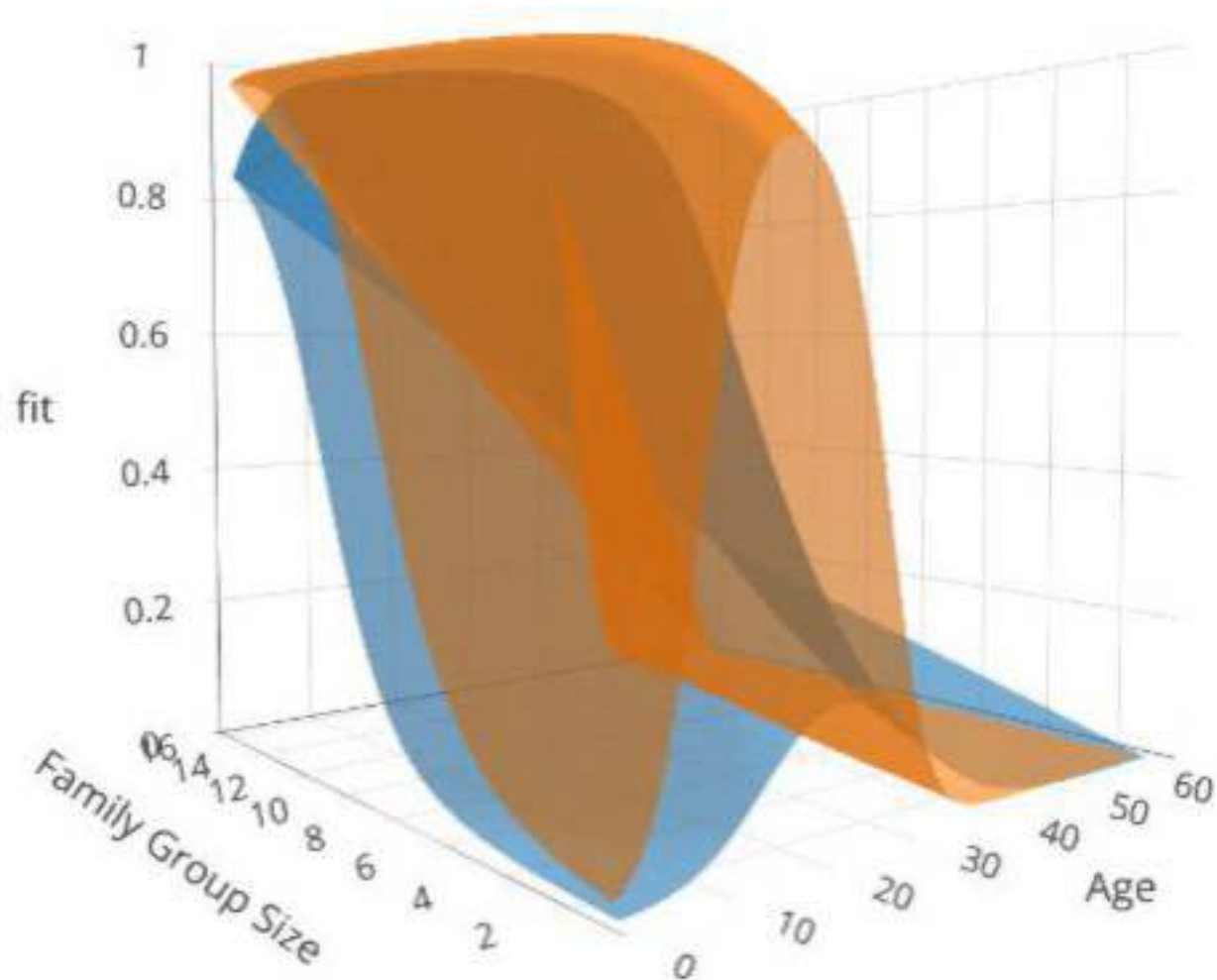
**5.1 Effect Sizes** The odds of a 14-year old female surviving were 88 times higher ( $\exp(4.4726)$ ) than a 14-y old male (95% CI: 2.9 to 2700 times).

| Factor            | Effect  | SE      | Lower 0.95 | Upper 0.95 |
|-------------------|---------|---------|------------|------------|
| Age               | 1.0784  | 1.02410 | -0.96465   | 3.12150    |
| Family_Group_Size | -1.0421 | 0.99062 | -3.01840   | 0.93412    |
| Sex - Female:Male | 4.4726  | 1.70690 | 1.06740    | 7.87780    |

**Adjusted to: Age=14 Sex=Male**

### 5.2 Wald Statistics

| Coefficients                                  | Estimate | Std. Error | z value | Pr(> z ) |
|---|----------|------------|---------|----------|
| (Intercept)                                   | -3.8540  | 1.8404     | -2.094  | 0.03625  |
| rcs(Age, 3)Age                                | 0.5408   | 0.1931     | 2.801   | 0.00510  |
| rcs(Age, 3)Age'                               | -1.3338  | 0.4800     | -2.779  | 0.00546  |
| SexMale                                       | -0.4999  | 1.4069     | -0.355  | 0.72234  |
| rcs(Family_Group_Size, 5)Family_Group_Size    | 0.5136   | 0.5186     | 0.990   | 0.32199  |
| rcs(Family_Group_Size, 5)Family_Group_Size'   | 3.1091   | 3.1502     | 0.987   | 0.32366  |
| rcs(Family_Group_Size, 5)Family_Group_Size''  | -9.1724  | 6.1511     | -1.491  | 0.13591  |
| rcs(Family_Group_Size, 5)Family_Group_Size''' | 123.5832 | 47.0271    | 2.628   | 0.00859  |
| rcs(Age, 3)Age:SexMale                        | -0.3581  | 0.2087     | -1.716  | 0.08619  |
| rcs(Age, 3)Age':SexMale                       | 0.9591   | 0.4964     | 1.932   | 0.05334  |



**Figure 5.** Three-dimensional perspective of fit, the probability of survival in the Donner Party as a function of Age and Family Group Size, modeled with restricted cubic splines with 3 and 5 knots, respectively. Female predicted survival is orange, and the Male is blue. The striking survival of all but one female between the ages of 5 and 41 and the high death rates of all but one male in their 20s produce the two bell-shaped patterns along the Age axis. The complete survival of families of size 6 and 9, the very low survival of small families, and the intermediate survival of the largest families produce the nested cowl-like patterns along the Family Group Size axis.

### **3.6 Survival Analyses**

#### **3.6.1 Effects of Sex and Encampment Day on Survival**

The sex-based differences in survivorship by encampment day were analyzed with a Cox proportional hazard model (Table 6) and Kaplan-Meier survivorship curves (Figure 6). The proportional hazard assumption for sex, tested with the `survival::cox.zph` test, indicated only modest evidence against the constant ratio assumption ( $p = 0.1$ , Table 6). However, the Kaplan-Meier survivorship curve (Figure 6) clearly indicates that males started dying earlier and at a higher rate than females. As noted by Grayson (1997), 5 males died on day 48 (12/15/1846) and 11 males died by day 63 (12/30/1846). Fourteen males died before the first female death on day 97 (Harriet McCutcheon, age 1, 2/2/1847). There were four rescue parties, called by Donner historians the First to Fourth Relief, and their arrival at the Donner encampments is indicated in Figures 6 and 7. The endpoints of the survivorship curves are at Day 183 (April 29, 1847), when Louis Keseberg the sole survivor in the Donner encampments left for California with Relief Party 4.

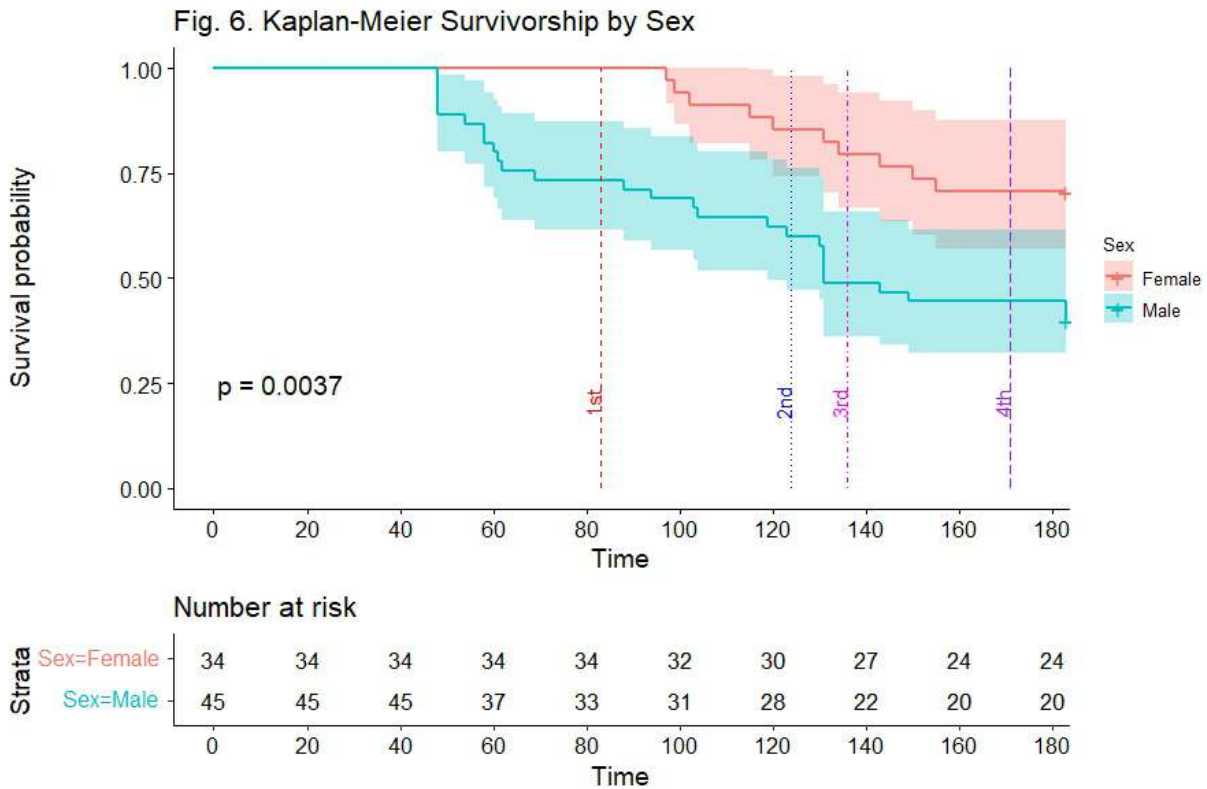


**Table 6. Cox proportional hazard model** The effects of Sex on Survival Time were analyzed.

There was a pronounced Sex effect on survival time ( $p = 0.01$ ), with the odds of a male dying being 2.6 times higher than females (1.2 to 5.5 95% CI). A `cox.zph` test in the `survival` package indicated only modest evidence against the equal hazard proportion assumption ( $\text{chisq}=2.90$ ,  $\text{df} = 1$ ,  $p = 0.0834$ ), but the Kaplan-Meier survivorship curves (Figure 6) indicates that the proportional hazard assumption was not tenable.

|                   | <b>coef</b>       | <b>exp(coef)</b> | <b>se(coef)</b> | <b>z</b> | <b>Pr(&gt; z )</b> |
|-------------------|-------------------|------------------|-----------------|----------|--------------------|
| SexMale           | 0.9683            | 2.6336           | 0.3727          | 2.598    | 0.00937            |
| <b>Factor</b>     |                   |                  |                 |          |                    |
|                   | <b>Chi-Square</b> |                  | <b>d.f.</b>     | <b>P</b> |                    |
| Family_Group_Size | 12.49             |                  | 4               | 0.014    |                    |
| Nonlinear         | 11.57             |                  | 3               | 0.009    |                    |
| TOTAL             | 12.49             |                  | 4               | 0.014    |                    |

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**Figure 6.** Kaplan-Meier Survivorship curve as a function of Sex. Males are 2.8 times more likely to die than females (Table 6). The steeper slopes for males especially early indicates that the constant hazard assumption was violated. The dates of arrival at the encampments by the four relief parties (labeled 1<sup>st</sup> to 4<sup>th</sup>) are indicated.

### 3.6.2 Effects of Employee Status on Survival

Thirteen of the 87 travelers in the Donner Party were employed as teamsters or servants. Most were teamsters hired to drive the oxen-powered wagons, but one (Antonio) was a cattle herder and Baylis & Eliza Williams were Reed family servants. Patrick Dolan, Luke Halloran, Joseph Reinhardt, and Charles Stanton were the lone travelers (Family Group Size = 1) who were not employees. Only 10 of the 13 employees were present after the first major snowfall on 28 October 1846; the others are not included in the analyses. The survivorship analyses indicated

that the employees died more rapidly than family members or non-employee bachelors (Tables 7 & 8, Figures 7 & 8).

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**Table 7. Cox proportional hazard model** Analysis of Employee vs. Family member survival.

Teamsters died at a rate 4.5 times that of Family members (95% CI: 2.1 to 9.3). A survival::cox.zph test indicated a clear violation of the equal proportion assumption (chisq=6.45, df = 1, p = 0.01).

|                      | coef   | exp(coef) | se(coef) | z     | Pr(> z ) |
|----------------------|--------|-----------|----------|-------|----------|
| Teamster_Hired_Hands | 1.5068 | 4.5124    | 0.3679   | 4.096 | 4.21e-05 |

|                      | exp(coef) | exp(-coef) | lower .95 | upper .95 |
|----------------------|-----------|------------|-----------|-----------|
| Teamster_Hired_Hands | 4.512     | 0.2216     | 2.194     | 9.281     |

Concordance= 0.63 (se = 0.038 )  
 Likelihood ratio test= 13.32 on 1 df, p=3e-04  
 Wald test = 16.77 on 1 df, p=4e-05  
 Score (logrank) test = 20.02 on 1 df, p=8e-06

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The Cox proportional hazards test (survival::cox.zph) indicated that the hazard ratio of employees relative to non-employees was not constant with time, violating the Cox constant proportional hazard assumption (p = 0.01), so an additional covariate, Survival\_Time, was added to the Cox model. A plot of Schoenfeld residuals indicated a reasonable fit to the proportional hazards model with the relative risk ratios dropping with time. The results are shown in Table 8 and Figures 7 & 8.

**Table 8. Cox time-dependent hazard model.** The model determined that employees initial risk of dying was 480 times that of family members, but the relative risk between employees and family members declined by 4% per day as shown by the Kaplan-Meier survivorship curves (Figure 9).

**Model:**

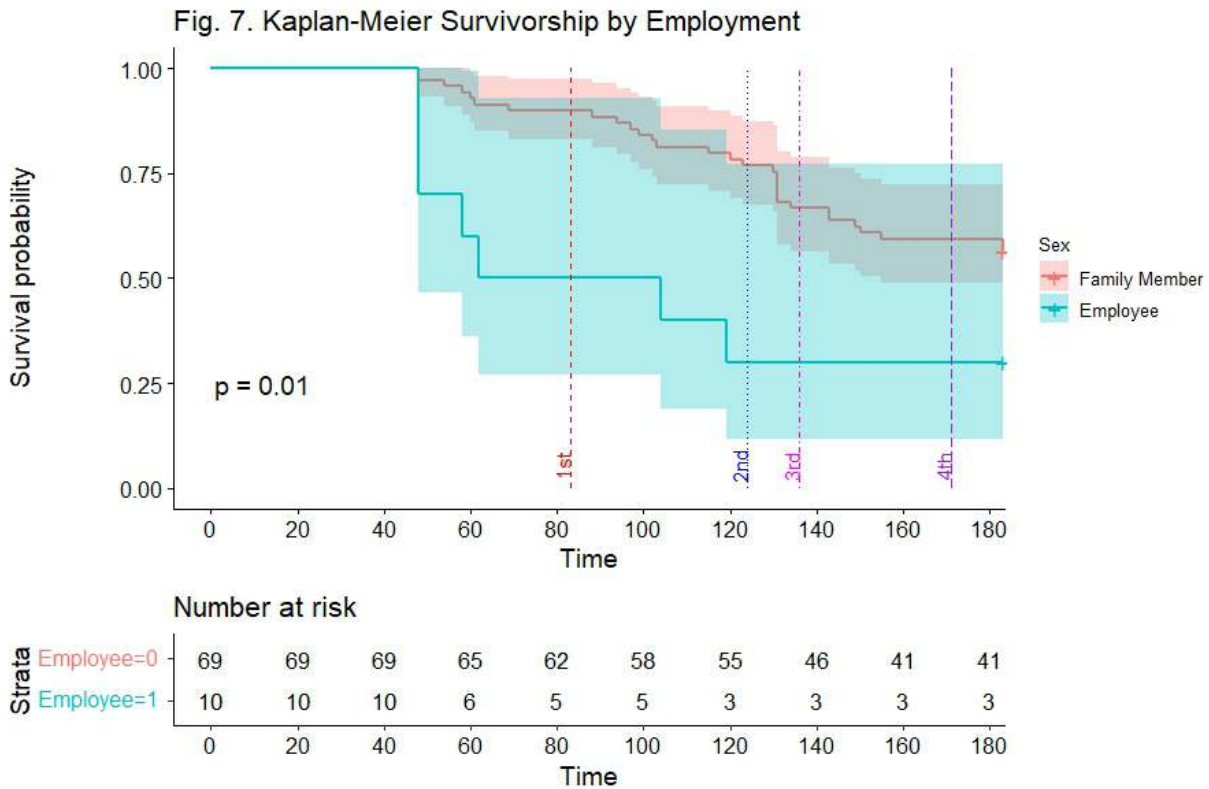
```
Donner$SurvTime_Employee <- with(Donner, Survival_Time *
Employee)
coxph(formula = Surv(Survival_Time, Death) ~ Employee +
SurvTime_Employee, data = Donner)
n= 79, number of events= 37
```

|                   | <b>coef</b>      | <b>exp(coef)</b>  | <b>se(coef)</b>  | <b>z</b>         | <b>Pr(&gt; z )</b> |
|-------------------|------------------|-------------------|------------------|------------------|--------------------|
| Employee          | 6.16800          | 477.22945         | 1.26154          | 4.889            | 1.01e-06           |
| SurvTime_Employee | -0.04558         | 0.95545           | 0.01434          | -3.177           | 0.00149            |
|                   | <b>exp(coef)</b> | <b>exp(-coef)</b> | <b>lower .95</b> | <b>upper .95</b> |                    |
| Employee          | 477.2294         | 0.002095          | 40.263           | 5656.5203        |                    |
| SurvTime_Employee | 0.9554           | 1.046632          | 0.929            | 0.9827           |                    |

```
Concordance= 0.621 (se = 0.035 )
Likelihood ratio test= 27.03 on 2 df, p=1e-06
Wald test = 32.13 on 2 df, p=1e-07
Score (logrank) test = 63.63 on 2 df, p=2e-14
```

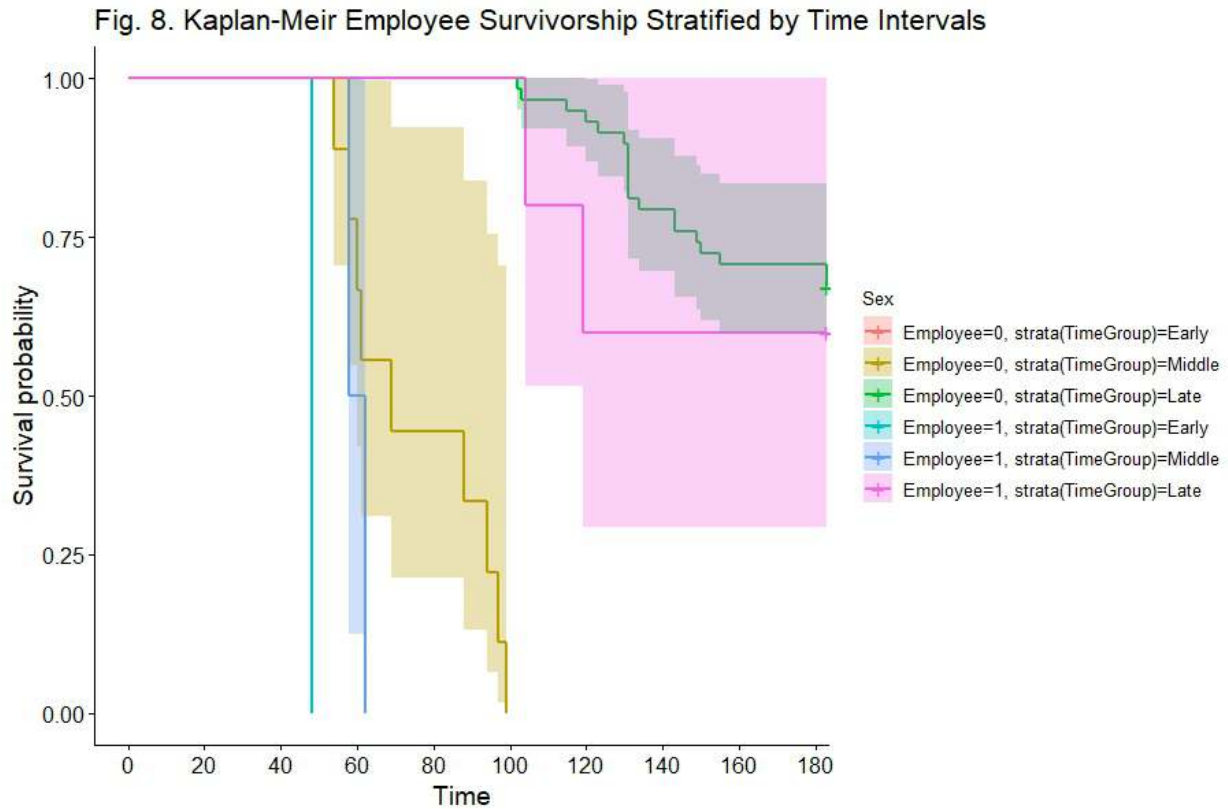
|                   | <b>chisq</b> | <b>df</b> | <b>p</b>    |
|-------------------|--------------|-----------|-------------|
| Employee          | 4.97         | 1         | 0.026       |
| SurvTime_Employee | 2.70         | 1         | 0.100       |
| GLOBAL            | 5.14         | 2         | 0.077=2e-14 |

Table 8 and Figures 7 & 8 reveal that the proportional hazard ratio between Employees and non-Employees is not constant with time. This is evident in the lack of parallel slopes in the Kaplan-Meier survivorship curves in Figure 7. Kaplan-Meier curves stratified by time period (Figure 8) show the changes through time intervals in the hazard function. While Employees have a lower survival in each time period with no non-employees dying in the first interval, there was a clear convergence of death rates during the last time interval 100 days after the Donner Party was snowed in.



**Figure 7.** Kaplan-Meier Survivorship curves as a function of Employee vs. Non-employee.

Employees are initially 480 times more likely to die than family members, but the relative risk ratio declines 4% per day (Table 7).



**Figure 8.** Stratified Kaplan-Meier Survivorship curves as a function of Employee vs.

Non-employee status in three time periods with breaks at days 50 and 100. During the Early period only Employees died. In the Middle period both groups died at high rates with Employee death rates far exceeding Non-employees. During the Final period, both groups died at a lower rate than the middle period with the Employees initially dying at a higher rate. During the Late period, there was extensive overlap in 95% confidence intervals.

### 3.6. 3 Effects of Sex on the Survivorship of the Forlorn Hope

On 16 December 1846 17 members of the Donner Party left the encampments to cross the crest of the Sierras to get help. Only fourteen had snowshoes, and two without horseshoes, unable to walk in the huge drifts, returned to the encampments on the first day. Because of the dire fate of the fifteen who continued, the group has become known as the Forlorn Hope, a term usually used

for troops at the vanguard of a military charge with a low chance of survival. Of the fifteen members of the Forlorn Hope, all 5 women survived but only 2 of the 10 men survived. Of the 8 men who died, 7 were cannibalized.

Using Fisher's exact hypergeometric test, the probability of observing such a Sex-based difference in survivorship (100% to 20%) by chance is just 0.006993. Usually the relative survival of females versus males would be assessed using the odds ratio  $O_F / O_M$ , where the odds of female survival are  $O_F = P_F / (1 - P_F)$ , and where  $P_F$  is the probability of female survival. Because the probability of female survival is 1, neither an odds ratio nor 95% confidence limit for the odds ratio can be calculated for the sex-based odds of survival for the Forlorn Hope.

#### **4. DISCUSSION**

Why present these analyses? First, the Donner Party Tragedy is a major historical event. It has been the subject of more than a dozen books, numerous scholarly articles, and was the subject of a 1992 Ric Burns PBS documentary. A 2009 movie "Donner Party," was based on the Forlorn Hope rescue mission. Cannibalism plays a prominent role in the movie, but strangely the movie cuts the Forlorn Hope from 15 to 10 members.

The key focus of most tales of the Donner Party is cannibalism. However, the Breen and Reed families did not resort to cannibalism, and both families survived intact. Rarick (2008, p. 239) attributes the remarkable survival of the 9-member Breen family to their more abundant supply of beef. The only two Donner survivors 40 or older were Margaret (40) and Patrick (51) Breen. Rarick (2008, p. 239) argues that the perfect survival of the 6-member Reed family was due not to food availability but to the indomitable Margaret Reed who had to beg for food after her

husband James Reed was driven out of the Donner Party for killing the teamster James Snyder on 5 October 1846.

A second reason for this study is that it shows vividly how social data can be analyzed statistically, a tradition that dates back to Quetelet's 1835 *Physique Sociale* (Social Physics), described by Gallagher (2020). This study confirms the conclusions of Grayson (1990, 1994, 1997, 2018) and Rarick (2008) that three factors largely control the survivorship of the Donner Party: Age, Sex and Family Group Size. The effects of Family Group Size on survival are strongly confounded with Employee status since the 13 employees were classified as singletons in the Family Group Size analyses. Employees—teamsters, servants, and a cattle herders—died earlier and at a higher rate than family members (Tables 6 & 7, Figures 7 & 8). Christmas (2008, p. 74) argues that the high employee mortality was due to their strenuous work and food deprivation by their employers, “they were likely the first to go hungry, and the least prepared for it.” Rarick (2008, p. 155-156) describes how two employees, Milt Elliot and Eliza Williams, weren't even allowed to die in the same cabin with the Reeds, their former employers whom they had formerly considered family.

The Kaplan-Meier survivorship curves clearly indicate that starvation and cold temperatures created striking increases in male mortality after about 50 days and males died at a high rate more than a month before notable female mortality (Figure 6). Employees died at a higher rate than non-employees (Figures 7 & 8). The stratified Kaplan-Meier survivorship curves (Figure 8) reveal that most of the disparity in death rates between Employees and non-Employees occurred before day 100, after which there were just five employees left. The survival of three of the last



five employees may have been due to their leaving the encampments. Noah James and Baylis Williams left with the First Relief Group, and Jean Pierre Trudeau left with the Third Relief Group. The first of two remaining employees to die was Milt Elliot, a Reed teamster who tried to walk out with two others on 4 January but returned to the encampments on January 8<sup>th</sup> and died on 8 February (day 103). The final employee to die was John Denton, a Donner teamster who died on 24 February (day 119) as part of the First Relief journey back to California when the other members of the First Relief Party left him behind on the trail.

While there were four relief parties and a fifth salvage party, there is no simple pattern between a Relief Party's arrival at the encampments and a change in death rates (Figures 6 & 7). The Reliefs often arrived with little food and took with them what little food they had brought for the return trip to California with rescued travelers. Those left behind were often sick and starving and continued to die. As documented by Grayson (2018), some of the travelers rescued by the First and Second Relief parties died on the journey to California.

A third reason for statistically analyzing the Donner Party is that the patterns are of interest to physiologists, anthropologists, archaeologists, and those just curious about how starvation and family ties affect survivorship. Philbrick (2000) described the tragedy of the Nantucket Whaleship Essex, sunk after being rammed by a sperm whale on 20 November 1820, a tragedy that inspired Melville's *Moby Dick*. He cites (p. 167) the Donner Party to explain why 7 of 11 (63.6%) white Nantucket whalers survived on the Essex whaleboats while only 1 of 6 (16.7%) black whalers survived. All but one black whaler, the first to die, were cannibalized, and the first four whalers eaten were black. Philbrick argued that there was no historical evidence that the

black whalers were deprived of food. He argued that black whalers came from the wharves of Boston and may have been unhealthier and had a lower fat content than the white Nantucket whalers. The one black whaler to survive, William Bond, was the steward in the Essex officers' quarters and had plentiful food. Not calculated by Philbrick were the odds of such a racial disparity in survival. The odds of a black Essex whaler dying were 7.6 times that of a white whaler, but statistics offer only modest evidence to reject the chance null hypothesis (Fisher's exact test, two-sided  $p = 0.13$ , 95% CI for the odds ratio: 0.56 to 470, which includes 1.0 the chance expected value.) .

Brown (2009, p. 137) calculated the metabolic rate [MR] of Sarah Fosdick (aged 22, MR ~ 3100 kcal/d [cal/d in Brown]) and her father Franklin Graves (aged 57, MR ~ 3600 kcal/d) to explain why she may have survived while he did not. Grayson (2018) presents a thorough review of the possible reasons for the sex-based differential mortality, emphasizing differences in height, weight, and fat content as well as differences in behavior (e.g., male sex-based suicide rates are higher than females). Grayson (1990, 1994, 1997, and 2018) presents statistics and graphical displays to assess the roles of age, sex and family size on survivorship but used neither nonlinear regression nor formal survivorship analyses to reveal the curvilinear patterns which appear to be important keys to Donner survivorship.

The survival of every woman but one aged 4 to 36 years is perhaps the most striking pattern in the data. Eleanor Eddy, the only woman in that age span to die, died at age 25 on 7 February 1847 a mere 3 days after the death of her 1-y old daughter Margaret and 53 days after Eleanor's

husband Edward left her to lead the 17-traveler Forlorn Hope group, leaving her alone with two children and little food.

While the effects of family group size and kinship interactions are important, the Family Group Size effect appear to be strongly influenced by the perfect survival of the Reed and Breen families as shown in Figure 5 and 6, rather than a monotonic pattern in which greater kinship links yield higher survival, one of Grayson's (1990, 1994) conclusions. Teamsters and servants make up the bulk of the single-member families, and Rarick (2008) noted that they were more likely to die from starvation due to their greater exertion reaching the final Donner Pass encampments. Both Rarick (2008) and Brown (2009) argue that the larger body mass of teamsters relative to women and non-employees accounts for the relatively higher male mortality. The Kaplan-Meier survivorship curve (Figure 8) shows that singleton groups, dominated by teamsters and servants, have a tremendously high early mortality, 480 times that of family members (Table 8).

A fourth reason for analyzing the Donner Party data is that these data offer an interesting case study to introduce restricted cubic splines, GAMs, and survivorship analyses to intermediate and upper level statistics classes. Wood (2017, p. 136) commented on his use of a case study with just 23 subjects to introduce the Cox proportional hazard model, "This is clearly a small sample from a statistical point of view, but not from a human point of view, and it is important to try to determine whether there really is evidence for a difference between the treatments." The Donner Party is one of the few historical data, where each datum tells a story. Scholars have documented

the background and fate of each of the 87 Donner travelers, but only Grayson has tested statistical hypotheses with the data.

Only a few of the methods described here are presented in introductory statistics books. Ramsey & Schafer (2013 and two previous editions) use a pared 45-sample Donner dataset to introduce binary logistic regression. Dalgaard's (2008, p 251-258) introductory statistics text introduces Cox and Kaplan-Meier survival analyses and nonlinear curve fitting but covers neither splines nor GAMs. These methods are presented clearly in advanced texts like Harrell (2015), Wood (2017), and Andrews (2021). Harrell (2015) is particularly good on restricted cubic spline regression, for which he wrote the `rms::rcs` and `rms::Predict` functions. Wood (2017) provides detailed descriptions of restricted cubic splines, GAMs with his `mgcv` package, and the Cox proportional hazards model. Zuur et al. (2009) present ecological examples of GAMs. Andrews (2021) in a concise chapter on nonlinear regression covers both restricted cubic splines and GAMs.

Restricted cubic spline regression permits the user to choose the location of knots or joining points in the curved regressions. In this paper, only the default knot locations provided by Harrell's `rms::rcs` function were used, e.g., quartiles for 4-knot analyses. Harrell's (2015) book provides examples of selecting specific knot locations. Choosing the appropriate number of knots or the GAM basis function  $k$  is very important. Simply choosing the model with the lowest AIC appears to lead to overfitting with too many knots. The 4-AIC threshold used here reduces the number of knots and while the AIC was higher, the fit appeared to be more easily interpreted. The most severe case of overfitting encountered in this study was fitting both Age and Family

Group Size using rcs to model survivorship (Table 5 and Figure 5). Using the AIC criterion with the 4-AIC threshold produced knot sizes of 6 for both parameters with an AIC of 56.33. The Wald statistics had high p-values, and the 3-d figure had little explanatory value. A  $k$ -fold cross validation analyses based on the minimization of log loss found 3 knots for Age and 5 knots for Family Group Size. The results shown in Figure 5 are clearly interpretable and display the major patterns described by the regression analyses. The take-home messages are: 1) Don't use AIC minimization alone to choose knot size for rcs as it tends to select too many knots leading to overfitting, 2)  $k$ -fold cross validation appears to work well to find the optimal  $k$ 's for both restricted cubic splines and GAMs, 3) The goal of this study was explanation, not prediction, and the AIC criterion using a 4-unit threshold or  $k$ -fold cross validation produced lower knot sizes resulting in more easily interpreted graphical displays.

Much of the R coding for this paper was aided by OpenAI's GPT-4. GPT-4 through its training through September 2021 appears fully aware of how to program using Harrell's R rms package, in particular his Glm, rcs and Predict functions. GPT-4 reported on more than one occasion that it hadn't been trained on Harrell's (2015) text but was aware of his R packages. GPT-4 is also adept at coding GAM analyses using Wood's mgcv function (Wood 2017, 2019), including  $k$ -fold cross-validation with the caret package. GPT-4 also wrote the code for the Cox and Kaplan-Meier survivorship functions.

To demonstrate another of GPT-4's skills, I prompted GPT-4 with the abstract for this paper and asked it to write a villanelle in iambic pentameter summarizing the results. It did (see online Appendix).

## ACKNOWLEDGMENTS

Many thanks to Donald Grayson whose Donner scholarship inspired this paper and who provided the latest Donner demographic data. Thanks to Frank Harrell whose book and 2021 4-d course introduced me to restricted cubic spline regression.

## APPENDIX

A villanelle composed by Open AI's GPT-4 in iambic pentameter summarizing the results.

[https://raw.githubusercontent.com/DonnerCurves/donner-data-analysis/main/Appendix\\_villanelle.pdf](https://raw.githubusercontent.com/DonnerCurves/donner-data-analysis/main/Appendix_villanelle.pdf)

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## DATA AVAILABILITY

### Data

<https://raw.githubusercontent.com/EugeneGall/donner-data-analysis/main/Donner.csv>

### R Code

[https://raw.githubusercontent.com/EugeneGall/donner-data-analysis/main/Donner\\_Gallagher\\_Public.R](https://raw.githubusercontent.com/EugeneGall/donner-data-analysis/main/Donner_Gallagher_Public.R)

### This manuscript

[https://raw.githubusercontent.com/EugeneGall/donner-data-analysis/main/Donner\\_Gallagher\\_MS.pdf](https://raw.githubusercontent.com/EugeneGall/donner-data-analysis/main/Donner_Gallagher_MS.pdf)