

线性回归

$$Y = f(X) + \epsilon$$

input

$$x_1^m$$
$$x_2^m$$
$$x_3^m$$

output

$$y^m$$
$$Y = M\beta \Rightarrow M^T Y = M^T M\beta \Rightarrow \beta = (M^T M)^{-1} M^{-1} Y$$

$$y = ax_1 + bx_2 + cx_3 + d + \epsilon$$

频率学派--最大似然

分号与逗号的区别：

;x -- x是个未知的常数

,x -- x是个变量

$$L(\theta) = \prod_{i=1}^m P(x_i; \theta)$$

$$\begin{aligned}\theta_{MLE} &= \arg \max_{\theta} L(\theta) \\ &= \arg \max_{\theta} \log L(\theta) \\ &= \arg \max_{\theta} \log P(x_i; \theta)\end{aligned}$$

θ 没有统计上的解释

条件概率

$$P(\theta|X) = \frac{P(\theta, X)}{P(X)} = \frac{P(\theta)P(X|\theta)}{P(X)}$$

$$\begin{aligned}\theta_{MAP} &= \arg \max_{\theta} P(\theta|X) \\ &= \arg \max_{\theta} P(\theta)P(X|\theta) \\ &= \arg \max_{\theta} \{\log P(\theta) + \log P(X|\theta)\} \\ &= \arg \max_{\theta} \{\log P(\theta) + \sum_{i=1}^m \log P(x_i|\theta)\}\end{aligned}$$

$$y = \theta^T X$$

$$y^{(i)} = \theta^T X^{(i)} + \epsilon^{(i)}$$

$$\epsilon^{(i)} = y^{(i)} - \theta^T X^{(i)} \sim N(0, \sigma^2)$$

$$\sum_{i=1}^m \log P(x_i|\theta) = m \log \frac{1}{\sqrt{2\pi}\sigma} - \frac{\sum_{i=1}^m (y^{(i)} - \theta^T X^{(i)})^2}{2\sigma^2}$$

$$\log P(\theta)=\log \frac{1}{(2\pi)^{\frac{k}{2}}}-\lambda \theta^T \theta$$

$$\begin{aligned}\theta_{MAP}&=arg\max_{\theta}\{-\sum_{i=1}^m(y^{(i)}-\theta^TX^{(i)})^2-\lambda\theta^T\theta\}\\&=arg\min_{\theta}\{\sum_{i=1}^m(y^{(i)}-\theta^TX^{(i)})^2+\lambda\theta^T\theta\}\end{aligned}$$

$\lambda\theta^T\theta$: 正则项, 这与 θ 的分布有关

$$LASSO$$

非线性回归

$$y=ax_1^2+b\log x_2+ce^{x_3}+d$$

input

$$\begin{pmatrix}x_1^2\\\log x_2\\e^{x_3}\\1\end{pmatrix}$$

$$\min_{\beta}\quad \|X\beta - Y\|^2 + \lambda\|\beta\|^2$$

凸函数

$$f(\lambda\beta_1+(1-\lambda)\beta_2)\geq \lambda f(\beta_1)+(1-\lambda)f(\beta_2)$$

伪梯度

$$\begin{aligned}y&=|x_1|+|x_2|+|x_3|\\g&=\text{sgn}(x_1,x_2,x_3)\end{aligned}$$

求解

$$\min_{\beta}\quad \|X\beta - Y\|_2^2 + \lambda\|\beta\|_2^2$$

- 求导

$$\frac{\partial (X\beta - Y)^T(X\beta - Y) + \lambda\beta^T\beta}{\partial \beta}$$

$$(X\beta - Y)^TX + \lambda\beta^T = 0$$

$$(X^TX + \lambda I)^T\beta = X^TY$$

矩阵乘法

$$C=A\cdot B\Rightarrow C_{ij}=\sum_kA_{ik}B_{kj}$$

$$D=A\cdot B\cdot C\Rightarrow D_{ij}=\sum_{k_1}\sum_{k_2}A_{ik_1}B_{k_1k_2}C_{k_2j}$$

矩阵求导

$f(X) = AX + b$

• 变量是向量，函数是向量

$$\frac{\partial(f_1, \dots, f_m)}{\partial(x_1, \dots, x_n)}$$

$$D_{ij} = \frac{\partial f_i}{\partial x_j} = A_{ij}$$

则

$$\frac{\partial f}{\partial X} = A$$

A是矩阵，x是实数 $A = f(x)$
则

$$\frac{\partial A^{-1}}{\partial x} = -A^{-1} \frac{\partial A}{\partial x} A^{-1}$$

• 变量是矩阵，函数是标量

$$\frac{\partial f}{\partial x} = \begin{bmatrix} \frac{\partial f}{\partial x_{11}} & \cdots & \frac{\partial f}{\partial x_{1n}} \\ \vdots & \ddots & \vdots \\ \frac{\partial f}{\partial x_{n1}} & \cdots & \frac{\partial f}{\partial x_{nn}} \end{bmatrix}$$

• 迹
迹
tr 是矩阵对角线元素之和

$$\frac{\partial tr[AXB]}{\partial x}$$

$$[AXB]_{hk} = \sum_i \sum_j A_{hi} X_{ij} B_{jk}$$

A是个矩阵
A的迹 $tr[A] = \sum_{i=1}^n a_{ii}$

$$tr[AXB] = \sum_h \sum_i \sum_j A_{hi} X_{ij} B_{jh}$$

$$\therefore \frac{\partial tr[AXB]}{\partial x_{ij}} = \sum_h A_{hi} B_{jh} = \sum_h B_{jh} A_{hi} = [BA]_{ji} = [(BA)^T]_{ij}$$

$$\therefore \frac{\partial tr[AXB]}{\partial x} = A^T B^T$$

$$\frac{\partial tr[AX^T B]}{\partial x} = \frac{\partial tr[B^T X A^T]}{\partial x} = BA$$

$$\frac{\partial tr[AXBXC]}{\partial x} = A^T C^T X^T B^T + B^T X^T A^T C^T$$

• 范式

$$\|A\|_F^2 = \sum_i \sum_j a_{ij}^2$$

$$[AA^T]_{ii} = \sum_k A_{ik}^2$$

$$tr[AA^T] = \sum_i \sum_j A_{ij}^2 = \|A\|_F^2$$

$$\begin{aligned} \frac{\partial \|B - XA\|_F^2}{\partial x} &= \frac{\partial tr[(B - XA)^T(B - XA)]}{\partial tr[B^T B + A^T X^T X A - A^T X^T B - B^T X A]} \\ &= \frac{\partial x}{2XAA^T - 2BA^T} \end{aligned}$$

令 $2XAA^T - 2BA^T = 0$ 得:

$$X = BA^T(AA^T)^{-1}$$

logistic regression⇒分类

$$q(X = x) = P(Y = 1|X = x)$$

$$\frac{q}{1 - q} \in (0, +\infty) \quad \log \frac{q}{1 - q} \in (-\infty, +\infty)$$

$$\log \frac{q}{1 - q} = ax_1 + bx_2 + cx_3 + d = \beta x$$

$$q = \frac{e^{\beta x}}{1 + e^{\beta x}}$$

$$P(Y_1 = y_1, \dots, Y_n = y_n | X_1 = x_1, \dots, X_n = x_n)$$

$$\begin{aligned} &= \prod_{i: y_i=1}^n q(x_i, \beta) \cdot \prod_{i: y_i=0}^n (1 - q(x_i, \beta)) \\ &= \sum_{i=1}^n y_i \log q(x_i; \beta) + (1 - y_i) \log(1 - q(x_i; \beta)) \\ &= \sum_{i=1}^n \{y_i \beta^T x_i - \log(1 + e^{\beta^T x_i})\} \end{aligned}$$

GDA（高斯判别分析）⇒二分类

基本概念

• 概率密度

$$(X, Y) \Rightarrow \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

• 边缘密度函数

$$f_X(x) \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) \int_{-\infty}^{\infty} f(x, y) dx$$

- 独立
if $f(x, y) = f_X(x) \cdot f_Y(y) \quad \forall x, y$
then X和Y相互独立
- 条件概率

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)}$$

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

- 贝叶斯定理

$$\begin{aligned} P(A|B) &= \frac{P(A, B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)} \\ &= \frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|\bar{A})P(\bar{A})} \\ P(y|x) &= \frac{P(x|y)P(y)}{P(x)} \end{aligned}$$

$$\arg \max_y P(y|x) = \arg \max_y P(x|y)P(y)$$

$$\begin{cases} P_0 = P(Y = 0|X = x) \sim P(X = x|Y = 0)P(Y = 0) \\ P_1 = P(Y = 1|X = x) \sim P(X = x|Y = 1)P(Y = 1) \end{cases}$$

$$(x^{(i)}, y^{(i)})$$

$$\prod_{i=1}^m P(x^{(i)}, y^{(i)}) = \prod_{i=1}^m P(x^{(i)}|y^{(i)})P(y^{(i)})$$

目标函数 \Rightarrow

$$\sum_{i=1}^m \log P(x^{(i)}|y^{(i)}) + \sum_{i=1}^m \log P(y^{(i)})$$

$$P(X = x|Y = 0) \sim N(\mu_0, \Sigma)$$

$$P(X = x|Y = 1) \sim N(\mu_1, \Sigma)$$

$$P(y^{(i)}) = \Phi^{y^{(i)}} (1 - \Phi)^{1-y^{(i)}}$$

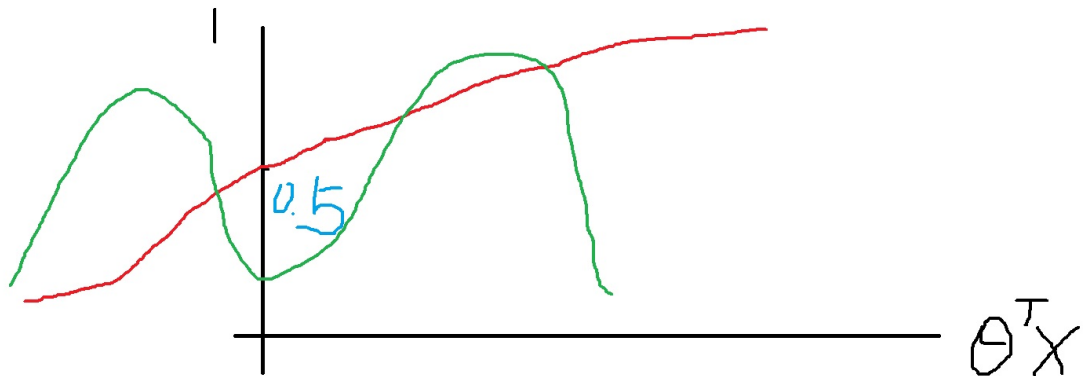
$$\Phi = \frac{1}{m} \sum_{i=1}^m \mathbb{S}\{y^{(i)} = 1\}$$

$$\mu_0 = \frac{\sum_{i=1}^m \mathbb{S}\{y^{(i)} = 0\} x^{(i)}}{\sum_{i=1}^m \mathbb{S}\{y^{(i)} = 0\}}$$

$$\mu_1 = \frac{\sum_{i=1}^m \mathbb{S}\{y^{(i)} = 1\} x^{(i)}}{\sum_{i=1}^m \mathbb{S}\{y^{(i)} = 1\}}$$

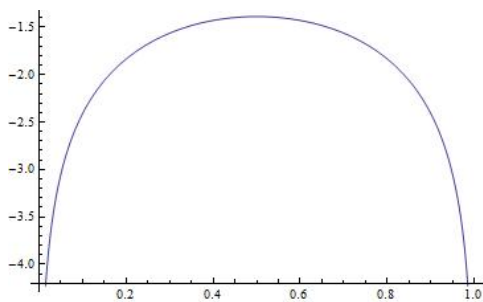
$$\Sigma = \frac{1}{m} \sum_{i=1}^m (x^{(i)} - \mu_{y^{(i)}})(x^{(i)} - \mu_{y^{(i)}})^T$$

$$\mathbb{S}(\text{表达式}) = \begin{cases} 1 & \text{表达式为真} \\ 0 & \text{表达式为假} \end{cases}$$



Naive Bayes

$$f(x) = \log x + \log(1 - x) \quad x \in (0, 1)$$



$$f(x) = a \log x + b \log(1 - x) \quad a > 0, b > 0$$

令

$$f'(x) = \frac{a}{x} - \frac{b}{1-x} = 0 \Rightarrow x = \frac{a}{a+b}$$

$$f(x_1, x_2) = a \log x_1 + b \log x_2 \quad a > 0, b > 0, x_1, x_2 > 0, x_1 + x_2 = 1$$

则

$$x_1 = \frac{a}{a+b}, x_2 = \frac{b}{a+b}$$

$$f(x_1, \dots, x_n) = \sum_{i=1}^n a_i \log x_i \quad x_i > 0, \sum x_i = 1$$

则

$$x_i = \frac{a_i}{\sum_j a_j}$$

Input:

$$(x^{(i)}, y^{(i)})_{i=1,2,\dots,m}$$

$$x^{(i)} = (x_1^{(i)}, \dots, x_n^{(i)}) \quad x_j^{(i)} \in \{0, 1\}$$

$$y^{(i)} \in \{1, \dots, k\}$$

$$\begin{aligned} & \log \prod_{i=1}^m P(x^{(i)}, y^{(i)}) \\ &= \sum_{i=1}^m \log P(x^{(i)}, y^{(i)}) \\ &= \sum_{i=1}^m \log P(y^{(i)}) + \sum_{i=1}^m \log P(x^{(i)} | y^{(i)}) \\ &= \sum_{i=1}^m \log P(y^{(i)}) + \sum_{i=1}^m \sum_{j=1}^n \log P(x_j^{(i)} | y^{(i)}) \end{aligned}$$

$$\begin{aligned} \sum_{i=1}^m \log P(y^{(i)}) &= \sum_{j=1}^k \sum_{i: y^{(i)}=j}^m \log P(y^{(i)} = j) \\ &= \sum_{j=1}^k \alpha_j \log P_j \\ \sum P_j &= 1 \quad P_j \sim \alpha_j \quad P_j = \frac{\alpha_j}{\sum \alpha_i} \end{aligned}$$

令 α_j 代表那些满足 $y^{(i)} = j$ 的样本个数

拉普拉斯平滑

$$P_j = \frac{\alpha_j + 1}{m + k} > 0$$

$$P(x_j | y) = \frac{\text{那些个 } y^{(i)} = y \text{ 且 } x_j^{(i)} = x \text{ 的样本个数}}{\text{那些个 } y^{(i)} = y \text{ 的样本个数}}$$

拉普拉斯平滑

$$P(x_j|y)=\frac{1+\text{那些个}y^{(i)}=y\text{且}x_j^{(i)}=x\text{的样本个数}}{2+\text{那些个}y^{(i)}=y\text{的样本个数}}$$

$$\begin{aligned}&P(Y=1|(x_1,x_2,\cdots,x_n))\\&=\frac{P(Y=1)\cdot P(x_1,\cdots,x_n|Y=1)}{P(x_1,\cdots,x_n)}\end{aligned}$$

EM

在凸区域上, 满足 $f(\lambda x+(1-\lambda)y)\leq \lambda f(x)+(1-\lambda)f(y)$ 则为凸函数

$$f(y)\geq f(x)+Df(x)^T(y-x)$$

$$f(\sum_{i=1}^n\lambda_ix_i)\leq \sum\lambda_if(x_i)\quad \sum\lambda_i=1$$

设 $p(x)$ 是概率密度函数, 则

$$\int_{-\infty}^{\infty}p(x)dx=1$$

$$E(x)=\int_{-\infty}^{\infty}x\cdot p(x)dx$$

令 $y=f(x)$ 是凸函数

$$f(\int x\cdot p(x)dx)\leq \int_{-\infty}^{\infty}f(x)p(x)dx$$

$$f(E(X))\leq E(f(X))$$

$$\int_{-\infty}^{+\infty}f(x)dx=1$$

则

$$\int_0^{+\infty}f(x)\log xdx\leq \log \int_0^{+\infty}f(x)xdx$$

GMM-高斯混合模型

$$\begin{aligned}l(\theta)&=\sum_{i=1}^m\log(P(x_i;\theta))\\&=\sum_{i=1}^m\log\sum_{z_i\in\Omega}P(x_i,z_i;\theta)\quad\text{边缘概率}\\&=\sum_{i=1}^m\log\sum_{z_i}Q(z^i)\frac{P(x_i,z_i;\theta)}{Q(z^i)}\end{aligned}$$

令 $\Phi(z^i)=\frac{P(x_i,z_i;\theta)}{Q(z^i)},\sum_{z^i\in\Omega}Q(z^i)=1$, 则:

$$\sum_{i=1}^m\log\sum_{z_i}Q(z^i)\frac{P(x_i,z_i;\theta)}{Q(z^i)}\geq\sum_{i=1}^m\sum_{z_i\in\Omega}Q(z^i)\log\frac{P(x_i,z_i;\theta)}{Q(z^i)}$$

取等条件:

$$\frac{P(x_i, z_i; \theta)}{Q(z^i)} = C \quad C \text{ 是常数}$$

即

$$P(x_i, z_i; \theta) \sim Q(z^i) \quad \sum_{z_i} Q(z_i) = 1$$

$$Q(z_i) = \frac{P(x_i, z_i; \theta)}{\sum_{z_i \in \Omega} P(x_i, z_i; \theta)} = \frac{P(x_i, z_i; \theta)}{P(x_i; \theta)} = P(z_i | x_i; \theta)$$

E-step: $\theta(z_i)$ 如何选取

M-step: 寻找 θ , 优化目标函数

$$\begin{aligned} z(x) = 1 &\Rightarrow \mu_1, \Sigma_1 \\ z(x) = 2 &\Rightarrow \mu_2, \Sigma_2 \end{aligned}$$

$$\begin{aligned} Q_i(z^{(i)} = j) &= P(z^{(i)} = j | x^{(i)}; \mu, \Sigma) \\ &= \frac{P(z^{(i)} = j; x^{(i)}; \mu, \Sigma)}{P(x^{(i)}; \mu, \Sigma)} \\ &= \frac{P(z^{(i)} = j) \cdot P(x^{(i)} | z^{(i)} = j)}{\sum_{z^{(i)}} P(x^{(i)}, z^{(i)}; \mu, \Sigma)} \end{aligned}$$