### 1.Logistic regression

(a)首先回顾 $J(\theta)$ 的定义

$$J( heta) = -rac{1}{m} \sum_{i=1}^m \log(h_ heta(y^{(i)}x^{(i)})) \ y(i) \in \{-1,1\}, h_ heta(x) = g( heta^Tx), g(z) = 1/(1+e^{-z})$$

注意 $g^{'}(z)=g(z)(1-g(z))$ ,利用这点来求 $rac{\partial h_{ heta}(x)}{\partial heta_{k}}$ 

$$egin{aligned} rac{\partial h_{ heta}(x)}{\partial heta_k} &= rac{\partial g( heta^T x)}{\partial heta_k} \ &= g( heta^T x)(1 - g( heta^T x))rac{\partial ( heta^T x)}{\partial heta_k} \ &= g( heta^T x)(1 - g( heta^T x))x_k \ &= h_{ heta}(x)(1 - h_{ heta}(x))x_k \end{aligned}$$

利用 $\frac{\partial h_{\theta}(x)}{\partial \theta_k}$ 来求 $\frac{\partial J(\theta)}{\partial \theta_k}$ 

$$\begin{split} \frac{\partial J(\theta)}{\partial \theta_k} &= -\frac{1}{m} \sum_{i=1}^m \frac{\partial \log(h_{\theta}(y^{(i)}x^{(i)}))}{\partial \theta_k} \\ &= -\frac{1}{m} \sum_{i=1}^m \frac{1}{h_{\theta}(y^{(i)}x^{(i)}))} \frac{\partial h_{\theta}(y^{(i)}x^{(i)}))}{\partial \theta_k} \\ &= -\frac{1}{m} \sum_{i=1}^m \frac{1}{h_{\theta}(y^{(i)}x^{(i)}))} h_{\theta}(y^{(i)}x^{(i)}) (1 - h_{\theta}(y^{(i)}x^{(i)})) y^{(i)}x_k^{(i)} \\ &= -\frac{1}{m} \sum_{i=1}^m (1 - h_{\theta}(y^{(i)}x^{(i)})) y^{(i)}x_k^{(i)} \end{split}$$

这个形式方便我们求二阶偏导,继续利用 $\frac{\partial h_{\theta}(x)}{\partial \theta_{k}}$ 来求 $\frac{\partial^{2} J(\theta)}{\partial \theta_{l} \partial \theta_{k}}$ 

$$egin{aligned} rac{\partial^2 J( heta)}{\partial heta_l \partial heta_k} &= -rac{1}{m} \sum_{i=1}^m y^{(i)} x_k^{(i)} rac{\partial (1 - h_ heta(y^{(i)} x^{(i)}))}{\partial heta_l} \ &= rac{1}{m} \sum_{i=1}^m y^{(i)} x_k^{(i)} rac{\partial (h_ heta(y^{(i)} x^{(i)}))}{\partial heta_l} \ &= rac{1}{m} \sum_{i=1}^m x_k^{(i)} h_ heta(y^{(i)} x^{(i)}) (1 - h_ heta(y^{(i)} x^{(i)})) x_l^{(i)} \end{aligned}$$

接着作以下记号

$$egin{aligned} x_k &= [x_1^{(k)}, \dots, x_n^{(k)}]^T \in R^n \ & X &= \left[egin{array}{c} x_1^T \ \dots \ x_m^T \end{array}
ight] \in R^{m imes n} \end{aligned}$$

$$\Lambda = \operatorname{diag}\{h_{\theta}(y^{(1}x^{(1)})(1-h_{\theta}(y^{(1)}x^{(1)}), \ldots, h_{\theta}(y^{(m)}x^{(m)})(1-h_{\theta}(y^{(m)}x^{(m)})\} \in R^{m \times m}$$

所以Hessian矩阵H可以表达为如下形式

$$H = X^T \Lambda X$$

为了(b)题需要,这里也将 $\nabla J(\theta)$ 表示出来

$$S = egin{bmatrix} (1 - h_{ heta}(y^{(1)}x^{(1)}))y^{(1)} \ & \dots \ (1 - h_{ heta}(y^{(m)}x^{(m)}))y^{(m)} \end{bmatrix} \in R^m \ 
abla J( heta) = -rac{1}{m}X^TS$$

现在任取 $z \in R^n$ ,记 $t = X^T z$ ,那么

$$z^T H z = z^T X \Lambda X^T z = t^T \Lambda t = \sum_{i=1}^m t_i^2 h_ heta(y^{(i)} x^{(i)}) (1 - h_ heta(y^{(i)} x^{(i)})$$

注意 $h_{\theta}(y^{(i)}x^{(i)}) \in [0,1]$ ,所以 $h_{\theta}(y^{(i)}x^{(i)})(1-h_{\theta}(y^{(i)}x^{(i)}) \geq 0$ ,从而

$$z^T H z = \sum_{i=1}^m t_i^2 h_ heta(y^{(i)} x^{(i)}) (1 - h_ heta(y^{(i)} x^{(i)}) \geq 0$$

从而H为半正定矩阵。

(b)(c)

这里解释下计算步骤,第一步是读取数据并增加一个截距项,由以下两个函数完成。

```
%matplotlib inline
import numpy as np
#import matplotlib as mpl
#mpl.use('Agg')
import matplotlib.pyplot as plt
from numpy.linalg import inv
#from __future__ import division
def load_data():
 X = np.genfromtxt('logistic_x.txt')
 Y = np.genfromtxt('logistic_y.txt')
  return X, Y
#增加截距项
def add_intercept(X_):
 m, n = X_.shape
 X = np.zeros((m, n + 1))
  ################
 ones = np.ones((m, 1))
 X = np.append(ones, X_, axis = 1)
  ################
  return X
```

第二步是利用刚刚的公式计算梯度以及Hessian矩阵。

```
#利用之前所述的公式计算
def calc_grad(X, Y, theta):
 m, n = X.shape
 grad = np.zeros(theta.shape)
  #############
 Y_{=} = Y.reshape([-1, 1])
 d1 = (X*Y_).dot(theta)
 h = 1 / (1 + np.exp(-d1))
 S = (1 - h) * Y
  grad = -1/m * (X.T).dot(S)
  ###############
 return grad
def calc_hessian(X, Y, theta):
  m, n = X.shape
 H = np.zeros((n, n))
  ##############
 Y = Y.reshape([-1, 1])
 d1 = (X*Y).dot(theta)
 h = 1 / (1 + np.exp(-d1))
 S = np.diag(h * (1-h))
 H = X.T.dot(S).dot(X)
  ############
  return H
```

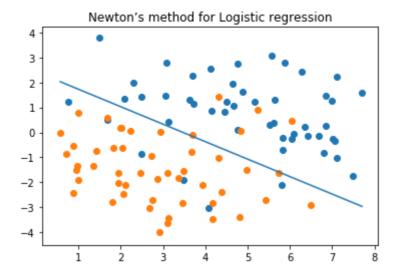
这里还有两个辅助的函数,分别是计算 $J(\theta)$ 和作图。

```
y1 = X[Y>0][:, 2]
x2 = X[Y<0][:, 1]
y2 = X[Y<0][:, 2]
#计算系数
theta = logistic_regression(X, Y)
Min = np.min(X[:, 1])
Max = np.max(X[:, 1])
x = np.array([Min, Max])
y = -(theta[0] + theta[1]*x)/theta[2]
plt.scatter(x1, y1)
plt.scatter(x2, y2)
plt.plot(x, y)
plt.title('Newton's method for Logistic regression')
###########
plt.savefig('ps1q1c.png')
return
```

### 最重要的一步就是利用如下公式迭代计算:

$$heta = heta - H^{-1} 
abla J( heta)$$

```
X_, Y = load_data()
X = add_intercept(X_)
theta = logistic_regression(X, Y)
plot(X, Y, theta)
```



全部代码可以查看my\_logistic.py这个文件。

# 2. Poisson regression and the exponential family

(a)

$$p(y;\lambda) = rac{e^{-\lambda}\lambda^y}{y!} = rac{1}{y!}e^{y\log\lambda - \lambda}$$

对比指数族的形式

$$p(y; \eta) = b(y) \exp(\eta^T T(y) - a(\eta))$$

可得

$$b(y) = rac{1}{y!}, \eta = \log \lambda, T(y) = y, a(\eta) = \lambda = e^{\eta}$$

所以泊松分布为指数族

(b)我们来计算 $g(\eta) = E[T(y); \eta]$ 

$$E[T(y);\eta] = E[y;\eta] = \lambda = e^{\eta}$$

(c)根据GLM的性质可知 $\eta=\theta^Tx$ ,那么 $\lambda=e^\eta=e^{\theta^Tx}$ 

计算对数似然函数

$$\begin{split} l &= \log p(y^{(i)}|x^{(i)};\theta) \\ &= \log(\frac{e^{-\lambda}\lambda^{y^{(i)}}}{y^{(i)}!}) \\ &= y^{(i)}\log\lambda - \lambda - \log(y^{(i)}!) \\ &= y^{(i)}\theta^Tx^{(i)} - e^{\theta^Tx^{(i)}} - \log(y^{(i)}!) \end{split}$$

关于 $\theta_i$ 求偏导可得

$$rac{\partial l}{\partial heta_{j}} = y^{(i)} x_{j}^{(i)} - e^{ heta^{T} x^{(i)}} x_{j} = (y^{(i)} - e^{ heta^{T} x^{(i)}}) x_{j}$$

此处求最大值,用随机梯度上升法,更新规则为

$$heta_j = heta_j + (y^{(i)} - e^{ heta^T x^{(i)}}) x_j = heta_j - (e^{ heta^T x^{(i)}} - y^{(i)}) x_j$$

(d)T(y)=y, 所以

$$p(y; \eta) = b(y) \exp(\eta^T T(y) - a(\eta)) = b(y) \exp(\eta^T y - a(\eta))$$

$$E[T(y); \eta] = E[y; \eta]$$

$$egin{aligned} l &= \log p(y|X; heta) \ &= \log \left( b(y) \exp(\eta^T y - a(\eta)) 
ight) \ &= \eta^T y - a(\eta) + \log b(y) \end{aligned}$$

因为 $\eta = \theta^T x$ , 所以

$$\frac{\partial l}{\partial \theta_i} = yx_j - \frac{\partial a(\eta)}{\partial \eta} \frac{\partial \eta}{\partial \theta_i} = (y - \frac{\partial a(\eta)}{\partial \eta})x_j$$

接下来只要证明 $\frac{\partial a(\eta)}{\partial \eta}=h(x)=E[y;\eta]$ 即可,利用 $p(y;\eta)$ 为概率密度函数

$$\int_{-\infty}^{+\infty} b(y) \exp(\eta^T y - a(\eta)) dy = 1 \ \int_{-\infty}^{+\infty} b(y) \exp(\eta^T y) dy = \exp(a(\eta))$$

西边关于*n*求偏导可得

$$egin{aligned} \int_{-\infty}^{+\infty} y b(y) \exp(\eta^T y)) dy &= \exp(a(\eta)) rac{\partial a(\eta)}{\partial \eta} \ rac{\partial a(\eta)}{\partial \eta} &= \int_{-\infty}^{+\infty} y b(y) \exp(\eta^T y - a(\eta)) dy = E[y; \eta] \end{aligned}$$

所以

$$\frac{\partial l}{\partial \theta_i} = (y - h(x))x_j$$

从而梯度上升法的更新规则为

$$\theta_i = \theta_i + \alpha(y - h(x))x_i = \theta_i - \alpha(h(x) - y)x_i$$

# 3.Gaussian discriminant analysis

(a)先计算P(x)

$$\begin{split} P(x) &= P(y=1)P(x|y=1) + P(y=-1)P(x|y=-1) \\ &= \phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \mathrm{exp}\Big( -\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1) \Big) + (1-\phi) \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \mathrm{exp}\Big( -\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1}) \Big) \end{split}$$

利用贝叶斯公式计算P(y|x),分y=1,y=-1计算

$$\begin{split} P(y|x) &= \frac{P(x|y)P(y)}{P(x)} \\ &= \frac{P(x|y)P(y)}{\phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \mathrm{exp}\Big(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\Big) + (1-\phi)\frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \mathrm{exp}\Big(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\Big)} \end{split}$$

所以

$$\begin{split} P(y=1|x) &= \frac{P(x|y=1)P(y=1)}{\phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right) + (1-\phi) \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)} \\ &= \frac{\frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right) \phi}{\phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right) + (1-\phi) \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)} \\ &= \frac{1}{1 + \frac{1-\phi}{\phi} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1}) + \frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)} \\ P(y=-1|x) &= \frac{P(x|y=-1)P(y=-1)}{\phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right) + (1-\phi) \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)} \\ &= \frac{(1-\phi) \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)}{\phi \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)} \\ &= \frac{1}{1+\frac{\phi}{1-\phi} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1) + \frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1})\right)} \end{split}$$

可以看到,指数部分都有一样的式子,现在计算这个式子

$$\begin{split} -\frac{1}{2}(x-\mu_{-1})^T \Sigma^{-1}(x-\mu_{-1}) + \frac{1}{2}(x-\mu_{1})^T \Sigma^{-1}(x-\mu_{1}) &= \frac{1}{2} \Big( x^T \Sigma^{-1} x - 2 \mu_{1}^T \Sigma^{-1} x + \mu_{1}^T \Sigma^{-1} \mu_{1} - x^T \Sigma^{-1} x + 2 \mu_{-1}^T \Sigma^{-1} x - \mu_{-1}^T \Sigma^{-1} \mu_{-1} \Big) \\ &= \frac{1}{2} \Big( 2 (\mu_{-1}^T \Sigma^{-1} - \mu_{1}^T \Sigma^{-1}) x + \mu_{1}^T \Sigma^{-1} \mu_{1} - \mu_{-1}^T \Sigma^{-1} \mu_{-1} \Big) \\ &= (\mu_{-1}^T \Sigma^{-1} - \mu_{1}^T \Sigma^{-1}) x + \frac{1}{2} (\mu_{1}^T \Sigma^{-1} \mu_{1} - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) \end{split}$$

所以

$$\begin{split} P(y=1|x) &= \frac{1}{1 + \frac{1-\phi}{\phi} \exp\left((\mu_{-1}^T \Sigma^{-1} - \mu_1^T \Sigma^{-1})x + \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1})\right)} \\ &= \frac{1}{1 + \exp\left((\mu_{-1}^T \Sigma^{-1} - \mu_1^T \Sigma^{-1})x + \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) + \ln(\frac{1-\phi}{\phi})\right)} \\ P(y=-1|x) &= \frac{1}{1 + \frac{\phi}{1-\phi} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1) + \frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_{-1})\right)} \\ &= \frac{1}{1 + \exp\left(-(\mu_{-1}^T \Sigma^{-1} - \mu_1^T \Sigma^{-1})x - \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) - \ln(\frac{1-\phi}{\phi})\right)} \end{split}$$

综合两式, P(y|x)可以写为

$$\begin{split} P(y|x) &= \frac{1}{1 + \exp\left(y\left((\mu_{-1}^T \Sigma^{-1} - \mu_1^T \Sigma^{-1})x + \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) + \ln(\frac{1-\phi}{\phi})\right)\right)} \\ &= \frac{1}{1 + \exp\left(-y\left((\mu_1^T \Sigma^{-1} - \mu_{-1}^T \Sigma^{-1})x - \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) - \ln(\frac{1-\phi}{\phi})\right)\right)} \end{split}$$

\$

$$\theta = (\mu_1^T \Sigma^{-1} - \mu_{-1}^T \Sigma^{-1})^T = \Sigma^{-1} (\mu_1 - \mu_{-1})$$
  
$$\theta_0 = -\frac{1}{2} (\mu_1^T \Sigma^{-1} \mu_1 - \mu_{-1}^T \Sigma^{-1} \mu_{-1}) - \ln(\frac{1 - \phi}{\phi})$$

从而

$$p(y|x;arphi,\Sigma,\mu_{-1},\mu_1) = rac{1}{1+\exp(-y( heta^Tx+ heta_0))}$$

(b)(c)

(c)是(b)的一般情形, 所以直接处理(c)

先观察P(x|y)的形式,可以得到如下公式

$$P(x|y) = rac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \mathrm{exp} \Big( -rac{1}{2} (x - \mu_y)^T \Sigma^{-1} (x - \mu_y) \Big)$$

接着计算 $\log P(x,y)$ 

$$\begin{split} \log & P(x,y) = \log P(x|y) P(y) \\ & = \log \Big( \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \Big( -\frac{1}{2} (x - \mu_y)^T \Sigma^{-1} (x - \mu_y) \Big) \phi^{1\{y=1\}} (1 - \phi)^{^{1\{y=-1\}}} \Big) \\ & = \log \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} - \frac{1}{2} (x - \mu_y)^T \Sigma^{-1} (x - \mu_y) + 1\{y = 1\} \log \phi + 1\{y = -1\} \log (1 - \phi) \end{split}$$

对数似然函数为

$$\begin{split} \ell(\varphi,\mu_{-1},\mu_{1},\Sigma) &= \log \prod_{i=1}^{m} p(x^{(i)},y^{(i)};\varphi,\mu_{-1},\mu_{1},\Sigma) \\ &= \sum_{i=1}^{m} \log p(x^{(i)},y^{(i)};\varphi,\mu_{-1},\mu_{1},\Sigma) \\ &= \sum_{i=1}^{m} \left( \log \frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}} - \frac{1}{2} (x^{(i)} - \mu_{y^{(i)}})^{T} \Sigma^{-1} (x^{(i)} - \mu_{y^{(i)}}) + 1\{y^{(i)} = 1\} \log \phi + 1\{y^{(i)} = -1\} \log (1 - \phi) \right) \end{split}$$

关于  $\phi$  求梯度

$$\begin{split} \frac{\partial \ell}{\partial \phi} &= \sum_{i=1}^m \left( \frac{1\{y^{(i)} = 1\}}{\phi} - \frac{1\{y^{(i)} = -1\}}{1 - \phi} \right) = 0 \\ &\sum_{i=1}^m 1\{y^{(i)} = 1\}(1 - \phi) - 1\{y^{(i)} = -1\}\phi = 0 \\ &\sum_{i=1}^m 1\{y^{(i)} = 1\} = m\phi \\ &\phi = \frac{1}{m} \sum_{i=1}^m 1\{y^{(i)} = 1\} \end{split}$$

关于 $\mu_1, \mu_{-1}$ 求梯度

$$\begin{split} \nabla_{\mu_1}\ell &= -\sum_{i=1}^m \Sigma^{-1}(x^{(i)} - \mu_{y^{(i)}})1\{y^{(i)} = 1\} = 0 \\ &\sum_{i=1}^m (x^{(i)} - \mu_1)1\{y^{(i)} = 1\} = 0 \\ &\mu_1 = \frac{\sum_{i=1}^m 1\{y^{(i)} = 1\}x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 1\}} \\ \nabla_{\mu_{-1}}\ell &= -\sum_{i=1}^m \Sigma^{-1}(x^{(i)} - \mu_{y^{(i)}})1\{y^{(i)} = -1\} = 0 \\ &\sum_{i=1}^m (x^{(i)} - \mu_{-1})1\{y^{(i)} = -1\} = 0 \\ &\mu_{-1} = \frac{\sum_{i=1}^m 1\{y^{(i)} = -1\}x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = -1\}} \end{split}$$

求 $\sum$ 的时候利用一些技巧性,我们不求 $\sum$ 的极大似然估计,而是求 $\sum^{-1}$ 的极大似然估计,然后再求出 $\sum$ 的极大似然估计,利用如下两个式子

$$egin{aligned} 
abla_A \det |A| &= \det |A| (A^{-1})^T \ 
abla_A (x^T A y) &= 
abla_A \mathrm{trace}(x^T A y) &= x y^T \end{aligned}$$

那么

$$egin{aligned} 
abla_{\sum^{-1}}\ell &= 
abla_{\sum^{-1}} \left(rac{m}{2} \log \left| \sum^{-1} 
ight| 
ight) - rac{1}{2} 
abla_{\sum^{-1}} \sum_{i=1}^m (x^{(i)} - \mu_y^{(i)})^T \Sigma^{-1} (x^{(i)} - \mu_y^{(i)}) &= 0 \ &rac{m}{2} rac{1}{\left| \sum^{-1} 
ight|} \left| \sum^{-1} 
ight| \sum^{-1} \sum_{i=1}^m (x^{(i)} - \mu_y^{(i)}) (x^{(i)} - \mu_y^{(i)})^T &= 0 \ &\sum = rac{1}{m} \sum_{i=1}^m (x^{(i)} - \mu_y^{(i)}) (x^{(i)} - \mu_y^{(i)})^T \end{aligned}$$

所以结论成立。

## 4.Linear invariance of optimization algorithms

(a)我们计算 $\nabla_z g(z), \nabla_z^2 g(z)$ 。

先计算 $\nabla_z g(z)$ ,

$$\begin{split} \frac{\partial g(z)}{\partial z_i} &= \sum_{k=1}^n \frac{\partial f(Az)}{\partial (Az)_k} \frac{\partial (Az)_k}{\partial z_i} \\ &= \sum_{k=1}^n \frac{\partial f(Az)}{\partial x_k} \frac{\partial (Az)_k}{\partial z_i} \\ &= \sum_{k=1}^n \frac{\partial f(Az)}{\partial x_k} A_{ki} \\ &= (A^T)_i \nabla_x f(Az) \\ &(A^T)_i \& \, \pi A^T \, \text{的 第 } i \text{行} \end{split}$$

所以

$$abla_z g(z) = A^T 
abla_x f(Az)$$

接着计算 $\nabla^2_z g(z)$ 

$$egin{aligned} rac{\partial^2 g(z)}{\partial z_j \partial z_i} &= rac{\partial \Big( \sum_{k=1}^n rac{\partial f(Az)}{\partial x_k} A_{ki} \Big)}{\partial z_j} \ &= \sum_{k=1}^n \sum_{l=1}^n rac{\partial^2 f(Az)}{\partial (Az)_l \partial x_k} rac{\partial (Az)_l}{\partial x_j} A_{ki} \ &= \sum_{k=1}^n \sum_{l=1}^n rac{\partial^2 f(Az)}{\partial (Az)_l \partial x_k} rac{\partial (Az)_l}{\partial x_j} A_{ki} \ &= \sum_{k=1}^n \sum_{l=1}^n rac{\partial^2 f(Az)}{\partial x_l \partial x_k} A_{lj} A_{ki} \end{aligned}$$

从而

$$abla_z^2 g(z) = A^T 
abla_x^2 f(Az) A^T$$

接着利用数学归纳法来证明结论。

n=0时,

$$z^{(0)} = A^{-1}x^{(0)} = 0$$

所以n=0时结论成立。假设 $n\leq i$ 时, $z^{(n)}=A^{-1}x^{(n)}$ ,那么n=i+1时

$$\begin{split} z^{(i+1)} &= z^{(i)} - (\nabla_z^2 g(z^{(i)}))^{-1} \nabla_z g(z) \\ &= A^{-1} x^{(i)} - A^{-1} (\nabla_x^2 f(Az^{(i)}))^{-1} (A^T)^{-1} A^T \nabla_x f(Az^{(i)}) \\ &= A^{-1} x^{(i)} - A^{-1} (\nabla_x^2 f(Az^{(i)}))^{-1} \nabla_x f(Az^{(i)}) \\ &= A^{-1} (x^{(i)} - (\nabla_x^2 f(Az^{(i)}))^{-1} \nabla_x f(Az^{(i)})) \\ &= A^{-1} (x^{(i)} - (\nabla_x^2 f(x^{(i)}))^{-1} \nabla_x f(x^{(i)})) \\ &= A^{-1} x^{(i+1)} \end{split}$$

其中倒数第二步是因为 $x^{(i)}=Az^{(i)}$ ,所以n=i+1时结论成立,从而牛顿法满足invariant to linear reparameterizations

(b)对于梯度下降法, 继续利用

$$abla_z g(z) = A^T 
abla_x f(Az)$$

假设 $z^{(i)} = A^{-1}x^{(i)}$ ,那么

$$egin{aligned} z^{(i+1)} &= z^{(i)} - lpha 
abla_z g(z) \ &= A^{-1} x^{(i)} - lpha A^T 
abla_x f(Az) \ &= A^{-1} x^{(i)} - lpha A^T 
abla_x f(x^{(i)}) \end{aligned}$$

但是

$$egin{aligned} x^{(i+1)} &= x^{(i)} - lpha 
abla_x f(x^{(i)}) \ A^{-1} x^{(i+1)} &= A^{-1} x^{(i)} - lpha A^{-1} 
abla_x f(x^{(i)}) \end{aligned}$$

所以 $z^{(i+1)}$ 与 $A^{-1}x^{(i+1)}$ 不相等,从而梯度下降法不满足invariant to linear reparameterizations

### 5. Regression for denoising quasar spectra

(a)

(i)

$$J( heta) = rac{1}{2} \sum_{i=1}^m w^{(i)} \Big( heta^T x^{(i)} - y^{(i)} \Big)^2$$

记

$$X = [x^{(1)}, \dots, x^{(m)}]^T \ y = [y^{(1)}, \dots, y^{(m)}]^T \ W = rac{1}{2} ext{diag}\{w^{(1)}, \dots, w^{(m)}\}$$

那么

$$J(\theta) = (X\theta - y)^T W (X\theta - y)$$

(ii)

$$J(\theta) = (X\theta - y)^T W (X\theta - y)$$

$$= \theta^T X^T W X \theta - 2y^T W^T X \theta + y^T y$$

$$\nabla J(\theta) = 2X^T W X \theta - 2X^T W y = 0$$

$$\theta = (X^T W X)^{-1} X^T W y$$

(iii)

$$p(y^{(i)}|x^{(i)}; heta) = rac{1}{\sqrt{2\pi}\sigma^{(i)}} \mathrm{exp}\Big(-rac{(y^{(i)}- heta^Tx^{(i)})^2}{2(\sigma^{(i)})^2}\Big)^2$$

概率似然函数为

$$\prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma^{(i)}} \exp\Big(-\frac{(y^{(i)} - \theta^T x^{(i)})^2}{2(\sigma^{(i)})^2}\Big)^2 = \Big(\prod_{i=1}^{m} \frac{1}{\sqrt{2\pi}\sigma^{(i)}}\Big) \exp\Big(-\sum_{i=1}^{m} \frac{(y^{(i)} - \theta^T x^{(i)})^2}{2(\sigma^{(i)})^2}\Big)^2$$

对于固定的 $\sigma^{(i)}$ ,最大化这个概率似然函数等价于最小化

$$\frac{1}{2} \sum_{i=1}^{m} \frac{(y^{(i)} - \theta^T x^{(i)})^2}{(\sigma^{(i)})^2}$$

记 $w^{(i)}=rac{1}{\left(\sigma^{(i)}
ight)^{2}}$ ,这个是式子可以转化为

$$rac{1}{2} \sum_{i=1}^m w^{(i)} (y^{(i)} - heta^T x^{(i)})^2$$

(b)

(i)

第一步还是读取数据以及增加截距。

```
%matplotlib inline
import numpy as np
#import matplotlib as mpl
#mpl.use('Agg')
import matplotlib.pyplot as plt
from numpy.linalg import inv
def load_data():
  train = np.genfromtxt('quasar_train.csv', skip_header=True, delimiter=',')
  test = np.genfromtxt('quasar_test.csv', skip_header=True, delimiter=',')
  wavelengths = np.genfromtxt('quasar_train.csv', skip_header=False, delimiter=',')[0]
  return train, test, wavelengths
def add_intercept(X_):
 X = None
  ######################
 X_{-} = X_{-}.reshape((-1, 1))
  m, n = X_.shape
  ones = np.ones((m, 1))
 X = np.append(ones, X_, axis = 1)
  ###################
  return X
```

这一部分直接利用最小二乘法计算结果。

### 作图函数

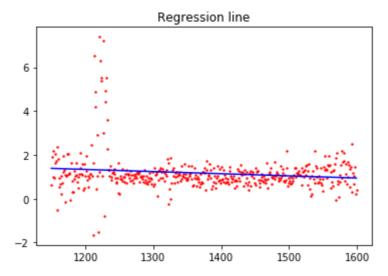
### 产生结果

```
raw_train, raw_test, wavelengths = load_data()

## Part b.i

lr_est, theta = LR_smooth(raw_train[0], wavelengths)
print('Part b.i) Theta=[%.4f, %.4f]' % (theta[0], theta[1]))
plot_b(wavelengths, raw_train[0], [lr_est], ['Regression line'], 'ps1q5b1.png')
```

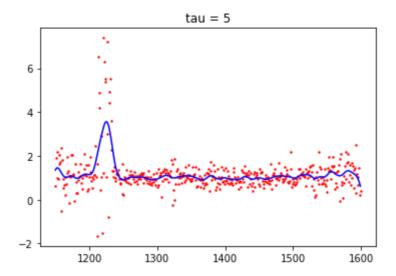
Part b.i) Theta=[2.5134, -0.0010]



# (ii)利用a中的公式 $heta=(X^TWX)^{-1}X^TWy$ 计算

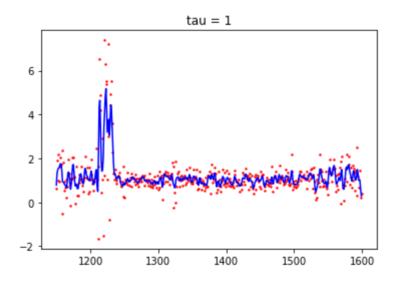
### 计算结果

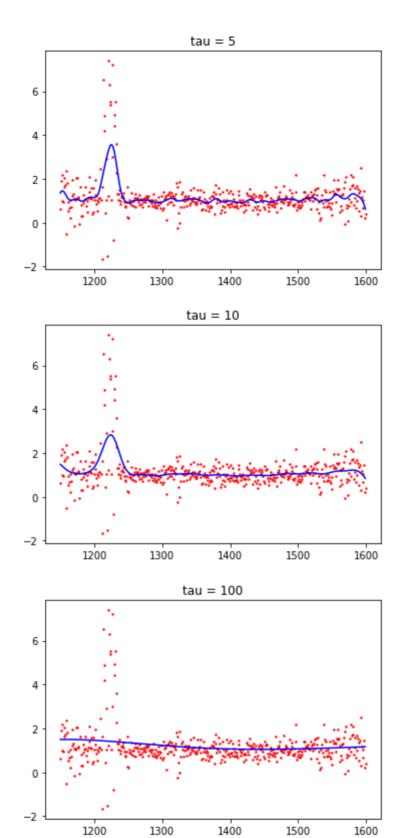
```
## Part b.ii
lwr_est_5 = LWR_smooth(raw_train[0], wavelengths, 5)
plot_b(wavelengths, raw_train[0], [lwr_est_5], ['tau = 5'], 'ps1q5b2.png')
```

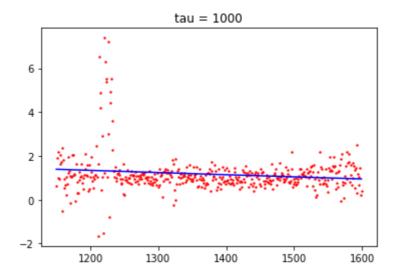


### 对不同的参数作图。

```
### Part b.iii
lwr_est_1 = LWR_smooth(raw_train[0], wavelengths, 1)
lwr_est_10 = LWR_smooth(raw_train[0], wavelengths, 10)
lwr_est_100 = LWR_smooth(raw_train[0], wavelengths, 100)
lwr_est_1000 = LWR_smooth(raw_train[0], wavelengths, 1000)
plot_b(wavelengths, raw_train[0],
    [lwr_est_1, lwr_est_5, lwr_est_10, lwr_est_1000],
    ['tau = 1', 'tau = 5', 'tau = 10', 'tau = 1000'],
    'ps1q5b3.png')
```







(c)

(i)

第一步是利用局部加权回归来使得数据更平滑一些。

```
def smooth_data(raw, wavelengths, tau):
    smooth = None
    #############

smooth = []
    for spectrum in raw:
        smooth.append(LWR_smooth(spectrum, wavelengths, tau))
    ############

return np.array(smooth)

smooth_train, smooth_test = [smooth_data(raw, wavelengths, 5) for raw in [raw_train, raw_test]]
```

(ii) 这部分比较难懂,详细解释下各个步骤,第一步将数据分为左边和右边,波长小于1200的为left,大于1300的为right,利用如下函数完成这部分工作。

第二步要计算距离, 定义如下函数

```
def dist(a, b):
    dist = 0
    ##############
    dist = ((a - b)**2).sum()
    ############
    return dist
```

```
#### Part c.ii

left_train, right_train = split(smooth_train, wavelengths)

left_test, right_test = split(smooth_test, wavelengths)
```

最后一步是最复杂的,首先计算距离,这里的距离定义为

$$d(f_1,f_2) = \sum_i (f_1(\lambda_i) - f_2(\lambda_i))^2$$

对于此题来说, $f_1(\lambda), f_2(\lambda)$ 分别对应了测试数据以及训练数据,利用下式计算距离矩阵。

```
d = (right_train - right_test)**2

#求和
d1 = d.sum(axis=1)
```

然后要找到距离最近的k个点,这里为3个点,我的思路是先对数据进行排序,然后找到前三个数据的索引

```
#找到排名前3的作为neighb_k(f_right)
tempd = d1.copy()
tempd.sort()
#找到索引
index = (d1==tempd[0])|(d1==tempd[1])|(d1==tempd[2])
```

接着计算h, 为距离的最大值, 根据题目中的公式对数据进行转换

```
h = d1.max()

d1 = d1/h
```

最后一步根据kernel以及题目中的公式计算即可

```
d1 = 1 - d1
#求lefthat
a = (d1[index].dot(left_train[index]))
b = d1[index].sum()

lefthat = a/b
```

以上全部综合起来就是如下函数:

```
def func_reg(left_train, right_train, right_test):
 m, n = left_train.shape
 #m=200,n=50
 lefthat = np.zeros(n)
 #############################
  #right_train 200*300
 #right_test 1*300
 #left_train 200*50
 #求题目中的d(f1,f2),先求每个点的距离,200*300矩阵
 d = (right_train - right_test)**2
 #按照行求和200*1
 d1 = d.sum(axis=1)
  #找到排名前3的作为neighb_k(f_right)
 tempd = d1.copy()
 tempd.sort()
 #找到索引
 index = (d1==tempd[0]) | (d1==tempd[1]) | (d1==tempd[2])
 #h为d1中的最大值
 h = d1.max()
 d1 = d1/h
 #ker (1-t)
 d1 = 1 - d1
 #求lefthat
 a = (d1[index].dot(left_train[index]))
 b = d1[index].sum()
 lefthat = a/b
 ##############################
 return lefthat
```

### 作图函数

```
#将左边右边区分开来,左边<1200,右边>=1300
def plot_c(Yhat, Y, X, filename):
    plt.figure()
    ##########
    plt.plot(X[:50],Yhat)
    plt.plot(X,Y)
    plt.show()
    ##########
plt.savefig(filename)
    return
```

### 产生结果

### #### Part c.ii

```
left_train, right_train = split(smooth_train, wavelengths)
left_test, right_test = split(smooth_test, wavelengths)
```

train\_errors = [dist(left, func\_reg(left\_train, right\_train, right)) for left, right in zip(left\_train, right\_train)] print('Part c.ii) Training error: %.4f' % np.mean(train\_errors))

Part c.ii) Training error: 1.0664

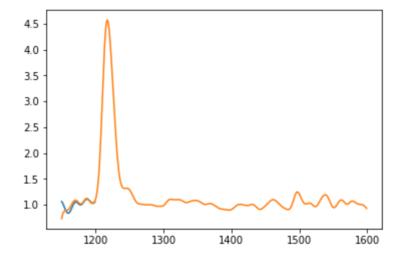
(iii)

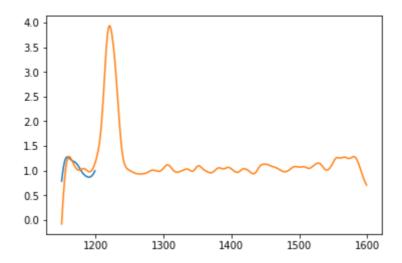
#### ### Part c.iii

test\_errors = [dist(left, func\_reg(left\_train, right\_train, right)) for left, right in zip(left\_test, right\_test)] print('Part c.iii) Test error: %.4f' % np.mean(test\_errors))

left\_1 = func\_reg(left\_train, right\_train, right\_test[0])
plot\_c(left\_1, smooth\_test[0], wavelengths, 'ps1q5c3\_1.png')
left\_6 = func\_reg(left\_train, right\_train, right\_test[5])
plot\_c(left\_6, smooth\_test[5], wavelengths, 'ps1q5c3\_6.png')

Part c.iii) Test error: 2.7100





全部代码可以查看my\_quasars.py这个文件。