RESAMPLING: JACKKNIFE and BOOTSTRAP

1. JACKKNIFE

Package "bootstrap" has a <u>jackknife tool</u>. For example, here is a bias reduction for a sample variance. # We'll start with the biased version of a sample variance.

```
> variance.fn = function(X){ return( mean( (X-mean(X))^2 ) ) }
> variance.fn(mpg)
[1] 60.76274
```

Use jackknife to correct its bias.

- > install.packages("bootstrap")
- > library(bootstrap)
- > variance.JK = jackknife(mpg, variance.fn)

This gives us the jackknife estimates of standard error and bias, as well as all estimates $\widehat{m{ heta}}_{ ext{(-i)}}$

```
> names(variance.JK)
[1] "jack.se"  "jack.bias"  "jack.values" "call"
> variance.JK
$jack.se
[1] 3.741951
```

\$jack.bias [1] -0.1554034

\$jack.values

 $[1] 60.84210 \, 60.73524 \, 60.84210 \, 60.77598 \, 60.81160 \, 60.73524 \, 60.68936 \, 60.68936 \, 60.68936 \, 60.73524 \, 60.68936 \, 60.73524 \, 60.68936 \, 60.91735 \, 60.91278 \, 60.84210 \, 60.90280$ $[19]\ 60.88575\ 60.90142\ 60.91195\ 60.91735\ 60.91195\ 60.91195\ 60.90142\ 60.90280\ 60.45457\ 60.45457\ 60.52096\ 60.38306\ 60.88575\ 60.86496\ 60.91195\ 60.86746\ 60.77598\ 60.81160\ 60.86746$ $[37] \, 60.84210 \, 60.68936 \, 60.68936 \, 60.68936 \, 60.68936 \, 60.68936 \, 60.58222 \, 60.63836 \, 60.63836 \, 60.84210 \, 60.91278 \, 60.86746 \, 60.84210 \, 60.91763 \, 60.86496 \, 60.80800 \, 60.80800 \, 60.77182 \, 60.57584$ $[55]\ 60.88575\ 60.90142\ 60.91735\ 60.91195\ 60.91763\ 60.88770\ 60.90280\ 60.63836\ 60.68936\ 60.73524\ 60.68936\ 60.81160\ 60.52096\ 60.63836\ 60.58222\ 60.63836\ 60.86746\ 60.73524\ 60.68936\ 60.81160\ 60.52096\ 60.63836\ 60.58222\ 60.63836\ 60.86746\ 60.73524\ 60.68936\ 60.81160\ 60.52096\ 60.63836\ 60.58222\ 60.63836\ 60.86746\ 60.73524\ 60.68936\ 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60.51403 \, 60.51403 \, 60.51403 \, 60.63253 \, 60.37501$ $[379] \, 60.73052 \, 60.37501 \, 60.91195 \, 60.37501 \, 60.90142 \, 60.91278 \, 60.73052 \, 60.51403 \, 60.88575 \, 60.88575 \, 59.83489 \, 60.73052 \, 60.86496 \, 60.77182 \, 60.91278 \, 6$

We can now calculate the jackknife variance estimator by the jackknife formula

> n* variance.fn(mpg) - (n-1)*mean(variance.JK\$jack.values)

... or subtracting the bias from the original estimator

> variance.fn(mpg) - variance.JK\$jack.bias [1] 60.91814

... and this is precisely the usual, unbiased version of the sample variance.

> var(mpg)

[1] 60.91814

2. BOOTSTRAP

(a) A simple example: use bootstrap to estimate the standard deviation of a sample median.

> B = 1000
> n = length(mpg)
> median_bootstrap = rep(0,B)

> for (k in 1:B){
+ b = sample(n, n, replace=TRUE)
+ median_bootstrap[k] = median(mpg[b])
+ }

Number of bootstrap samples

Sample size (and the size of each bootstrap sample)

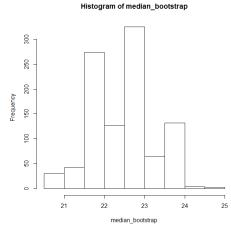
Initiate a vector of bootstrap medians

Do-loop to produce B bootstrap samples.

Take a random subsample of size n with replacement

+ median_bootstrap[k] = median(mpg[b]) # Compute the sample median of each bootstrap subsample

> hist(median bootstrap)



> median(mpg)

[1] 22.75

> sd(median_bootstrap)

[1] 0.7661376

The actual sample median is 22.75; bootstrap medians range # from 20.5 to 25.0.

The bootstrap estimator of SD(median) is 0.7661376.

(b) Bootstrap confidence interval

The histogram above shows the distribution of a sample median. We can find its $(\alpha/2)$ -quantile and $(1-\alpha/2)$ -quantile. They estimate the true quantiles that embrace the sample median with probability $(1-\alpha)$.

```
> alpha=0.05; quantile( median_bootstrap, alpha/2 )
2.5%
21
> quantile( median_bootstrap, 1-alpha/2 )
97.5%
24
```

A bootstrap confidence interval for the population median is [21, 24].

(c) R package "boot"

```
# There is a special bootstrap tool in R package "boot".
```

To use it for the same problem, we define a function that computes a sample median for a given subsample...

```
> median.fn = function(X,subsample){
+ return(median(X[subsample]))
+ }
```

For example, here is a sample median for a small subsample of size 10 without replacement...

```
> median.fn( mpg, sample(n,10) )
[1] 24.95
```

If the subsample is the whole sample, with all its indices, we get our original sample median...

```
> median.fn( mpg, ) [1] 22.75
```

Then we apply this function to many bootstrap samples (R is their number) and summarize the obtained statistics...

```
> library (boot)
> boot( mpg, median.fn, R=10000 )
```

```
ORDINARY NONPARAMETRIC BOOTSTRAP

Call:
boot(data = mpg, statistic = median.fn, R = 10000)

Bootstrap Statistics:
    original bias std. error
t1* 22.75 -0.1406 0.7687722
```

The bootstrap estimate of the standard deviation of a sample median is 0.7687722, and the bootstrap estimate of the bias is -0.1406. If we run the same command again, we might get somewhat different estimates, because the bootstrap method is based on random resampling.

```
> boot( mpg, median.fn, R=10000 )
original bias std. error
t1* 22.75 -0.13368 0.7727318
```

```
> boot( mpg, median.fn, R=10000 )
      original bias std. error
t1* 22.75 -0.13437 0.7617304
```

(d) Another problem – find a bootstrap estimate of the standard deviation of each regression coefficient

```
# Define a function returning regression intercept and slope...
> slopes.fn = function( dataset, subsample ){
+ reg = lm( mpg ~ horsepower, data=dataset[subsample, ] )
+ beta = coef(reg)
+ return(beta)
+ }
> boot( Auto, slopes.fn, R=10000 )
Bootstrap Statistics:
                        bias
                                 std. error
      original
t1* 39.9358610  0.0329999048  0.857890744
t2* -0.1578447 -0.0003929433 0.007444107
# We got bootstrap estimates, Std(intercept) = 0.85789, Std(slope) = 0.007444.
# Actually, this can be obtained easily, and without bootstrap...
> summary( lm( mpg ~ horsepower, data=Auto ) )
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
                                                <2e-16 ***
(Intercept) 39.935861
                           0.717499
                                       55.66
                                                <2e-16 ***
horsepower -0.157845
                           0.006446
                                      -24.49
```

But.... Why are the bootstrap estimates of standard errors higher than the ones obtained by regression formulae? Random variation due to bootstrapping? Let's try bootstrap a few more times.

```
> boot( Auto, slopes.fn, R=10000 )
                               std. error
      original
                       bias
t1* 39.9358610 0.0366632516 0.864274619
t2* -0.1578447 -0.0003676081 0.007459809
> boot( Auto, slopes.fn, R=10000 )
                      bias
                              std. error
      original
t1* 39.9358610 0.024793987 0.866532993
t2* -0.1578447 -0.000309597 0.007501155
> boot( Auto, slopes.fn, R=10000 )
Bootstrap Statistics:
                       bias
                               std. error
      original
t1* 39.9358610 0.0411404502 0.868404735
t2* -0.1578447 -0.0004208256 0.007500111
```

<u>A comment about discrepancy with the standard estimates</u>. Bootstrap estimates are pretty close to each other. They should be, because each of them is based on 10,000 bootstrap samples. However, the bootstrap standard errors of slopes are still a little higher than the estimates obtained by the standard regression analysis. The book explains this phenomenon by a number of assumptions that the standard regression

requires. In particular, nonrandom predictors X and a constant variance of responses $Var(Y) = \sigma^2$. Bootstrap is a nonparametric procedure, it is not based on any particular model. Thus, it can account for the variance of predictors. Here, "horsepower" is the predictor, and it is probably random.

Since some of the standard regression assumptions are questionable here, the bootstrap method for estimating standard errors is more reliable.