# MASTER'S DIAGNOSTIC EXAMINATION - January 2013

Student's Name
INSTRUCTIONS FOR STUDENTS:
1. The exam has an Instruction page, 8 pages of questions, and 21 pages of SAS code and output.
2. DO NOT put your NAME on the exam. Place the NUMBER assigned to you on the
UPPER RIGHT HAND CORNER of EACH PAGE of your solutions.
3. Please start your answer to EACH QUESTION on a SEPARATE sheet of paper.
4. Use only one side of each sheet of paper.
5. You must answer all four questions: Questions I, II, III and IV.
6. Be sure to attempt all parts of the four questions. It may be possible to answer a later part of a question without having solved the earlier parts.
7. Be sure to hand in all of your exam. No additional material will be accepted once the exam has ended and you have left the exam room.
8. You may use the following:
<ul> <li>Calculator which does not have capability to phone, text, or access the Web</li> <li>Pencil or pen</li> </ul>
<ul> <li>Blank paper for your solutions to the questions on this examination</li> <li>No other materials are allowed</li> </ul>
• I attest that I spent no more than 4 hours to complete the exam,
• I used only the materials described above, and
ullet I did not receive assistance from anyone during the taking of this exam.
Student's Signature
INSTRUCTIONS FOR PROCTOR:
Immediately after the student completes the exam, fax the <u>student's solutions</u> to <b>979-845-6060</b> or <b>scan</b> the solutions into a pdf file and <b>email to longneck@stat.tamu.edu</b> .
DO NOT send the exam booklet or SAS output, just send the student's solutions.
1. I certify that the time at which the student started the exam was  and the time at which the student completed the exam was
2. I certify that the student has followed all the INSTRUCTIONS FOR STUDENTS listed above.
3. I certify that the student's solutions were faxed to 979-845-6060 or
emailed to longneck@stat.tamu.edu.
cmaned to longueta@stat.tamu.euu.
Proctor's Signature

## QUESTION I.

You are the consulting statistician on a study of the genetic components of exercise. The PI plans to raise 30 mice in which a particular gene suspected to be involved in regulating exercise has been knocked out (you can think of it like turning off this gene), as well as 30 wild-type mice (mice in which the gene has not been knocked out). Each mouse will then be observed over a two week period, and the average number of minutes mouse i spends running on the exercise wheel per day,  $y_i$ , will be recorded. The variables in the study are

- the 60 response values  $y_1, y_2, \ldots, y_{60}$
- a variable indicating comparison group  $(x_i = 1 \text{ if knockout}, x_i = 0 \text{ if wild-type}, i = 1, 2, \dots, 60)$
- gender  $(g_i = 1 \text{ if female}, g_i = 0 \text{ if male}, i = 1, 2, \dots, 60)$
- average daily food consumption ( $f_i$ , a continuous number, standardized so that a value of 0 indicates average consumption, negative values indicate less than average, positive values indicate greater than average, i = 1, 2, ..., 60).

Based on your exploratory analysis of preliminary data, you and the PI agree that a sensible model is

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 g_i + \beta_3 f_i + \beta_4 (x_i \times g_i) + \epsilon_i$$

where the  $\epsilon_i$  are i.i.d. with mean 0 and s.d.  $\sigma$ .

- 1.) Interpret each of the coefficients in the above model.
- 2.) Using the following R output construct an approximate 95% confidence interval for  $\beta_1 \beta_4$ ?

Based on the interval you report, comment on your conclusions regarding whether there is a genotype effect.

Hint: If v is a column vector of length 5, then the standard error of  $v'\hat{\beta}$  is equal to the square root of  $\hat{\sigma}^2 v'(X'X)^{-1}v$ , where  $\hat{\sigma}$  is the estimated residual s.d., and X is the model matrix. In our example,  $\hat{\sigma} = 1.716$  and

$$(\boldsymbol{X}'\boldsymbol{X})^{-1} = \begin{pmatrix} 0.07 & -0.07 & -0.07 & -0.01 & 0.07 \\ -0.07 & 0.14 & 0.07 & 0.01 & -0.13 \\ -0.07 & 0.07 & 0.13 & 0.00 & -0.13 \\ -0.01 & 0.01 & 0.00 & 0.05 & 0.00 \\ 0.07 & -0.13 & -0.13 & 0.00 & 0.27 \end{pmatrix}$$

Here's the R output from fitting the above model:

#### Call:

lm(formula = y ~1 + x + g + f + x \* g)

#### Residuals:

```
Min 1Q Median 3Q Max
-4.3862 -1.0618 -0.0837 1.1362 4.2091
```

#### Coefficients:

	Estimate	Sta. Error	t value	Pr(> t )	
(Intercept)	10.6375	0.4463	23.835	< 2e-16	***
x	0.8728	0.6325	1.380	0.17319	
g	-0.9642	0.6265	-1.539	0.12955	
f	1.1632	0.3880	2.998	0.00407	**
x:g	0.0586	0.8867	0.066	0.94755	

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1

Residual standard error: 1.716 on 55 degrees of freedom Multiple R-squared: 0.221,Adjusted R-squared: 0.1643 F-statistic: 3.9 on 4 and 55 DF, p-value: 0.007369

3.) An F-test of the null hypothesis that  $\beta_1 = \beta_4 = 0$  returned a p-value of 0.15. Meanwhile, stepwise model selection tells you that the best-fitting model is the one without the treatment-by-gender interaction, with abbreviated R output:

```
Estimate Std. Error t value Pr(>|t|)
                          0.3853
                                  27.573
                                          < 2e-16 ***
(Intercept)
             10.6230
              0.9018
                          0.4404
                                   2.048
                                          0.04535 *
             -0.9349
                          0.4392
                                  -2.128
                                          0.03771 *
g
f
              1.1622
                          0.3843
                                   3.025
                                          0.00375 **
```

The PI is primarily interested in whether there is sufficient evidence of an effect of treatment, the knockout variable, on time spent exercising in mice. Based on the F-test and stepwise regression results, how would you advise the PI? Provide a detail explanation of your advice.

#### QUESTION II.

The Storm Prediction Center (an agency of NOAA) tracks the number and characteristics of tornadoes. In this problem we will consider the variable **Killer\_tornadoes**, which is the number of tornadoes in a given year which resulted in one or more deaths. The summary statistics for the variable **Killer\_tornadoes** over 48 years of data were found using the SAS function proc means:

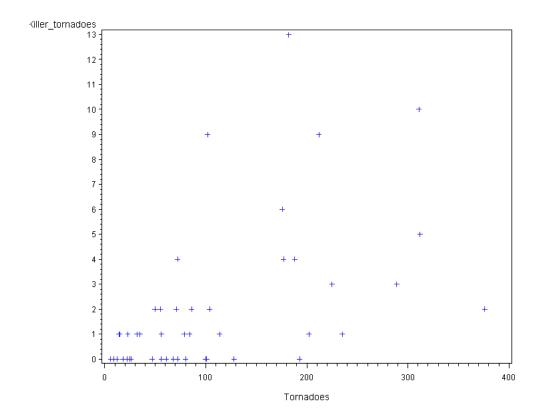
		The MEANS Proceed	lure	
	Analysis Variab	ole : Killer_torna	adoes	
N	Mean	Std Dev	Minimum	Maximum
48	2.0416667	2.9532408	0	13.0000000

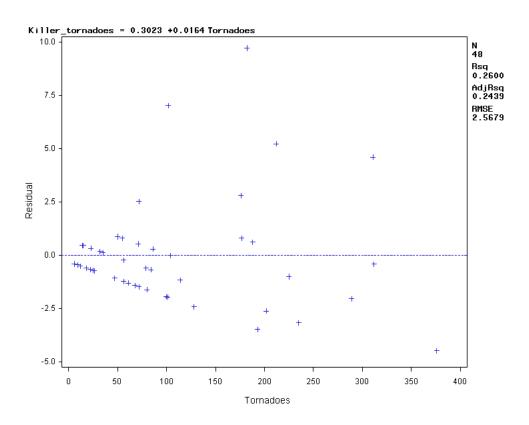
Use the above information, the SAS output, and the tables on page 5 to answer the following questions.

- 1.) Explain why  $\bar{X} \pm Z_{\alpha/2}\sqrt{\bar{X}/n}$  is a reasonable  $1-\alpha$  confidence set for  $\lambda$ , the mean of a Poisson distribution. Then assuming that the variable Killer\_tornadoes has a Poisson distribution, obtain a 95% confidence interval for the mean number of killer tornadoes.
- 2.) Is the assumption of a Poisson distribution reasonable for the variable Killer\_tornadoes? Explain why or why not based only on the summary statistics above.
- 3.) A chi-squared goodness-of-fit test was carried out by using the Poisson distribution with the estimated mean  $\hat{\mu} = 2.042$  to specify the cell probabilities for  $x = 0, 1, 2, 3, \ge 4$ . The chi-squared test statistic was computed to be  $\chi^2 = 32.35$ . What does this tell you about the assumption of the Poisson distribution for the variable Killer\_tornadoes? Explain your reasoning.
- 4.) Suppose that the researcher decides to predict the number of killer tornadoes using Tornadoes, the number of tornadoes in a year, as a predictor.

Using the 48 years of data on the pair (Tornadoes, Killer\_tornadoes), a simple linear regression model was fit with Killer\_tornadoes as the response and Tornadoes as the predictor. Some SAS output is given below. A scatter plot of the data and a plot of residuals versus the predictor are on the next page. Discuss the appropriateness of using this prediction model for the number of killer tornadoes.

		Analysis of Var	riance		
		Sum of	Mean		
Source	DF	Squares	Square	F Valu	ie Pr > F
Model	1	106.59133	106.59133	16.1	.6 0.0002
Error	46	303.32533	6.59403		
Corrected To	tal 47	409.91667			
	Root MSE	2.56788	R-Square	0.2600	
	Dependent Mean	2.04167	Adj R-Sq	0.2439	
	Coeff Var	125.77392			
		Parameter Estin	nates		
		Parameter	Standard		
Variable	DF	Estimate	Error t	: Value	Pr >  t
Intercept	1	0.30234	0.56967	0.53	0.5982
Tornadoes	1	0.01641	0.00408	4.02	0.0002





# Some Chi-Squared Percentiles

	Right-Tail Probability					
df	0.100	0.050	0.025	0.010		
1	2.71	3.84	5.02	6.63		
2	4.61	5.99	7.38	9.21		
3	6.25	7.81	9.35	11.34		
4	7.78	9.49	11.14	13.28		
5	9.24	11.07	12.83	15.09		
6	10.64	12.59	14.45	16.81		
7	12.02	14.07	16.01	18.48		
8	13.36	15.51	17.53	20.09		
9	14.68	16.92	19.02	21.67		
10	15.99	18.31	20.48	23.21		

# Some Normal Percentiles

Right-Tail Probability						
0.100	0.050	0.025	0.010			
1.282	1.645	1.960	2.326			

Some t-Distribution Percentiles

	Right-Tail Probability						
df	0.100	0.050	0.025	0.010			
1	3.078	6.314	12.706	31.821			
2	1.886	2.920	4.303	6.965			
3	1.638	2.353	3.182	4.541			
4	1.533	2.132	2.776	3.747			
6	1.440	1.943	2.447	3.143			
8	1.397	1.860	2.306	2.896			
12	1.356	1.782	2.179	2.681			
27	1.314	1.703	2.052	2.473			
53	1.298	1.674	2.006	2.399			
54	1.297	1.674	2.005	2.397			
$\infty$	1.282	1.645	1.960	2.326			

#### QUESTION III.

Consider a regression model for an experiment with a continuous response Y and two factors, A and B, each with two levels, High and Low. Consider the two models:

• Dummy Variable Model :  $E(Y) = \alpha + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$  where

$$x_1 = \left\{ \begin{array}{ll} 1 & \text{if} \quad A = \mathit{High} \\ 0 & \text{if} \quad A = \mathit{Low} \end{array} \right. \qquad x_2 = \left\{ \begin{array}{ll} 1 & \text{if} \quad B = \mathit{High} \\ 0 & \text{if} \quad B = \mathit{Low} \end{array} \right.$$

• Design Effects Model:  $E(Y) = \alpha^* + \beta_1^* x_1^* + \beta_2^* x_2^* + \beta_3^* x_1^* x_2^*$ ,

where

$$x_1^* = \left\{ \begin{array}{ccc} 1 & \text{if} & A = High \\ -1 & \text{if} & A = Low \end{array} \right. \qquad x_2^* = \left\{ \begin{array}{ccc} 1 & \text{if} & B = High \\ -1 & \text{if} & B = Low \end{array} \right.$$

- 1.) Express the following null hypotheses in term of the coefficients for the Dummy Variable Model:
  - a.) No interaction between A and B
  - b.) No effect due to A
  - c.) No effect due to B
- 2.) By considering the mean response for the four treatments (combinations of A and B), determine the relationship between  $\alpha$ ,  $\beta_1$ ,  $\beta_2$ ,  $\beta_3$  and  $\alpha^*$ ,  $\beta_1^*$ ,  $\beta_2^*$ ,  $\beta_3^*$ .
- 3.) Using the Dummy Variable Model, obtain expressions for the difference in mean response between
  - a.)  $A_{high}B_{high}$  and  $A_{low}B_{low}$
  - b.)  $A_{high}B_{high}$  and  $A_{high}B_{low}$
  - c.)  $A_{low}B_{high}$  and  $A_{low}B_{low}$ .
- 4.) Using the Design Effects Model, obtain expressions for the difference in mean response between
  - a.)  $A_{high}B_{high}$  and  $A_{low}B_{low}$
  - b.)  $A_{high}B_{high}$  and  $A_{high}B_{low}$
  - c.)  $A_{low}B_{high}$  and  $A_{low}B_{low}$ .
- 5.) Suppose that there is also a third factor C with two levels, High and Low. Describe how to form a Dummy Variable Model for the mean response in each of the following cases:
  - a.) Main effects only for A, B, and C.
  - b.) Main effects and all two-way interactions for A, B, and C.
  - c.) Main effects and all possible interactions for A, B, and C.
- 6.) Consider now the Dummy Variable Model for the logarithm of the mean response:
- Dummy Variable Model :  $\log(E(Y)) = \alpha + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$ ,

where

$$x_1 = \left\{ \begin{array}{ll} 1 & \text{if} \quad A = High \\ 0 & \text{if} \quad A = Low \end{array} \right. \qquad x_2 = \left\{ \begin{array}{ll} 1 & \text{if} \quad B = High \\ 0 & \text{if} \quad B = Low \end{array} \right.$$

6

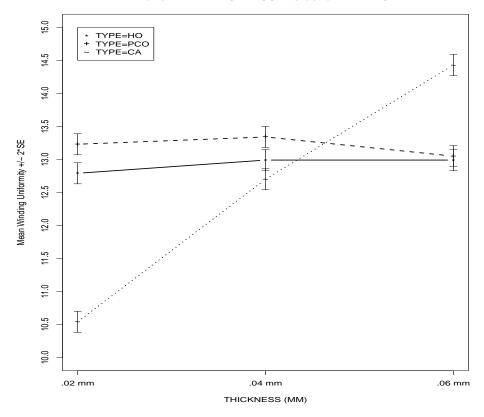
Explain what the coefficients  $\alpha, \beta_1, \beta_2, \beta_3$  represent in terms of the mean response E(Y).

## QUESTION IV.

A company is interested in the ability of a machine to consistently place electrical wire on a coil. There are three types of machines available: hand operated(HO), partially computer operated(PCO), and completely automated(CA). Three machines of each type are randomly selected from their suppliers for use in the study. The wire placed on the coils comes in one of three thicknesses: .02mm, .04mm, or .06mm. Each of the machines assembles two coils of each of the three wire thicknesses. Each wound coil is then measured for the uniformity of windings at a middle position on the coil. These measurements are given in the following table.

					TYPE	OF MA	CHINE			
			НО			PCO			CA	
MACHINE ID		1	2	3	1	2	3	1	2	3
	.02mm	12.30	13.46	12.35	13.01	13.46	13.15	10.47	10.75	10.24
		12.59	14.00	12.06	12.63	13.92	13.20	10.96	19.68	10.15
THICKNESS	$.04 \mathrm{mm}$	13.15	13.29	12.50	12.74	13.84	13.46	12.73	12.60	12.92
		13.00	13.62	12.39	12.68	13.75	13.57	12.64	12.65	12.64
	$.06\mathrm{mm}$	12.87	13.46	12.73	12.47	13.62	13.36	14.01	14.80	14.19
		12.92	13.82	12.15	12.15	13.28	13.42	14.62	14.71	14.23

#### Plot of TYPE\*THICKNESS Interaction With 2\*SE



Use the above plot, data, the attached SAS output, and the tables on page 5 to answer the following questions.

1.) Do the necessary conditions for testing hypotheses and constructing confidence intervals appear to be satisfied? Justify your answers.
$C_1$ - Normality:
$C_2$ - Equal Variance:
$C_3$ - Independence:
2.) Write a model for $y_{ijk\ell}$ , the uniformity of windings on the $\ell th$ Coil wound by the $kth$ Machine of Type $i$ using Wire Thickness $j$ .
• Make sure to include all necessary conditions on parameters and random variables in your model.
3.) At the $\alpha=.05$ level, which main effects and interactions are significant? Justify your answer by including the relevant p-values along with their pair of degrees of freedom $(df_{NUM.}, df_{DEN.})$ .
4.) The researchers were interested in estimating the difference in the mean uniformity in the windings between wires of thickness .02 mm and .06 mm wound using a Type CA machine: $\mu_{31} - \mu_{33}$
Determine the variance of the estimated mean difference: $Var(\hat{\mu}_{31} - \hat{\mu}_{33})$ , in terms of the variance components from your model.
Using the estimated variance components from the SAS output, compute the estimated standard error of the estimated mean difference: $\widehat{SE}(\hat{\mu}_{31} - \hat{\mu}_{33})$ .
• Provide supporting details for your expressions.
5.) Place the three types of machines into groups such that the machines within a group are not significantly different from any machine in the same group with respect to their mean uniformity of windings. Use an experimentwise error rate of $\alpha = .05$ .
6.) Provide a 95% confidence on the mean uniformity of windings of a wire of thickness .04 mm wound with a CA machine.
with a OH machine.
7.) The wire regulatory agency reviewed the study and were concerned that the rigidity of the different
wire samples was not included in the study. Suggest a method by which the rigidity measurements of the 54 wire samples could be used in the analysis.

The following SAS program was used to analyze the data. The output is contained in the following pages.

```
ods html;ods graphics on;
OPTIONS LS=90 PS=55 nocenter nodate;
TITLE 'SAS OUTPUT FOR PROBLEM IV';
DATA MANU;
INPUT TYPE $ MACH $ THICK $ Y @@;
TRT=COMPRESS (TYPE) || COMPRESS (THICK);
CARDS;
HO M1 .02 12.59 HO M2 .02 14.00 HO M3 .02
                                          12.06
HO M1 .04 13.15 HO M2 .04 13.29 HO M3 .04 12.50
HO M1 .04 13.00 HO M2 .04 13.62 HO M3 .04 12.39
HO M1 .06 12.87 HO M2 .06 13.46 HO M3 .06 12.73
HO M1 .06 12.92 HO M2 .06 13.82 HO M3 .06 12.15
PCO M1 .02 13.01 PCO M2 .02 13.46 PCO M3 .02 13.15
PCO M1 .02 12.63 PCO M2 .02 13.92 PCO M3 .02 13.20
PCO M1 .04 12.74 PCO M2 .04 13.84 PCO M3 .04 13.46
PCO M1 .04 12.68 PCO M2 .04 13.75 PCO M3 .04 13.57
PCO M1 .06 12.47 PCO M2 .06 13.62 PCO M3 .06 13.36
PCO M1 .06 12.15 PCO M2 .06 13.28 PCO M3 .06 13.42
CA M1 .02 10.47 CA M2 .02 10.75 CA M3 .02
CA M1 .02 10.96 CA M2 .02 10.68 CA M3 .02 10.15
CA M1 .04 12.73 CA M2 .04 12.60 CA M3 .04 12.92
CA M1 .04 12.64 CA M2 .04 12.65 CA M3 .04 12.64
CA M1 .06 14.01 CA M2 .06 14.80 CA M3 .06 14.19
CA M1 .06 14.62 CA M2 .06 14.71 CA M3 .06 14.23
RUN;
PROC GLM ORDER=DATA;
CLASS TYPE MACH THICK;
MODEL Y = TYPE THICK TYPE*THICK MACH(TYPE) THICK*MACH(TYPE);
RANDOM MACH(TYPE) THICK*MACH(TYPE)/TEST;
LSMEANS TYPE THICK THICK*TYPE/STDERR PDIFF ADJUST=TUKEY;
RUN;
PROC MIXED CL ALPHA=.05 COVTEST;
CLASS TYPE MACH THICK:
MODEL Y = TYPE THICK TYPE*THICK;
RANDOM MACH(TYPE) THICK*MACH(TYPE);
LSMEANS TYPE THICK TYPE*THICK/ ADJUST=TUKEY;
RUN;
PROC GLM:
CLASS TRT;
MODEL Y = TRT;
MEANS TRT/HOVTEST=BF;
OUTPUT OUT=ASSUMP R=RESID P=MEANS;
PROC GPLOT; PLOT RESID*MEANS/VREF=0;
PROC UNIVARIATE DEF=5 PLOT NORMAL;
VAR RESID;
RUN:
ods graphics off; ods html close;
```

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## SAS OUTPUT FOR PROBLEM IV

## The GLM Procedure

Class Level Information						
Class	Levels	Values				
TYPE	3	HO PCO CA				
MACH	3	M1 M2 M3				
THICK	3	.02 .04 .06				

Number of Observations Read	54	
Number of Observations Used	54	

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## SAS OUTPUT FOR PROBLEM IV

## The GLM Procedure

**Dependent Variable: Y** 

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	26	59.31313333	2.28127436	51.34	<.0001
Error	27	1.19980000	0.04443704		
<b>Corrected Total</b>	53	60.51293333			

R-Square	Coeff Var	Root MSE	Y Mean
0.980173	1.634679	0.210801	12.89556

Source	DF	Type I SS	Mean Square	F Value	Pr > F
TYPE	2	3.83981111	1.91990556	43.21	<.0001
THICK	2	15.59923333	7.79961667	175.52	<.0001
TYPE*THICK	4	30.27515556	7.56878889	170.33	<.0001
MACH(TYPE)	6	8.46555556	1.41092593	31.75	<.0001
MACH*THICK(TYPE)	12	1.13337778	0.09444815	2.13	0.0507

Source	DF	Type III SS	Mean Square	F Value	Pr > F
TYPE	2	3.83981111	1.91990556	43.21	<.0001
THICK	2	15.59923333	7.79961667	175.52	<.0001
TYPE*THICK	4	30.27515556	7.56878889	170.33	<.0001
MACH(TYPE)	6	8.46555556	1.41092593	31.75	<.0001
MACH*THICK(TYPE)	12	1.13337778	0.09444815	2.13	0.0507

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## SAS OUTPUT FOR PROBLEM IV

#### The GLM Procedure

Source	Type III Expected Mean Square
TYPE	Var(Error) + 2 Var(MACH*THICK(TYPE)) + 6 Var(MACH(TYPE)) + Q(TYPE,TYPE*THICK)
THICK	Var(Error) + 2 Var(MACH*THICK(TYPE)) + Q(THICK,TYPE*THICK)
TYPE*THICK	Var(Error) + 2 Var(MACH*THICK(TYPE)) + Q(TYPE*THICK)
MACH(TYPE)	Var(Error) + 2 Var(MACH*THICK(TYPE)) + 6 Var(MACH(TYPE))
MACH*THICK(TYPE)	Var(Error) + 2 Var(MACH*THICK(TYPE))

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#### SAS OUTPUT FOR PROBLEM IV

The GLM Procedure Tests of Hypotheses for Mixed Model Analysis of Variance

**Dependent Variable: Y** 

	Source	DF	Type III SS	Mean Square	F Value	Pr > F
*	TYPE	2	3.839811	1.919906	1.36	0.3256
	Error	6	8.465556	1.410926		

Error: MS(MACH(TYPE))

\* This test assumes one or more other fixed effects are zero.

	Source	DF	Type III SS	Mean Square	F Value	Pr > F
*	THICK	2	15.599233	7.799617	82.58	<.0001
	TYPE*THICK	4	30.275156	7.568789	80.14	<.0001
	MACH(TYPE)	6	8.465556	1.410926	14.94	<.0001
	Error	12	1.133378	0.094448		

Error: MS(MACH\*THICK(TYPE))

\* This test assumes one or more other fixed effects are zero.

Source	DF	Type III SS	Mean Square	F Value	Pr > F
MACH*THICK(TYPE)	12	1.133378	0.094448	2.13	0.0507
Error: MS(Error)	27	1.199800	0.044437		

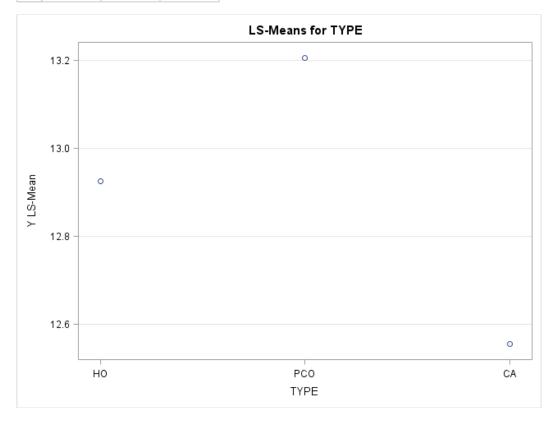
SAS Output Page 5 of 20

## SAS OUTPUT FOR PROBLEM IV

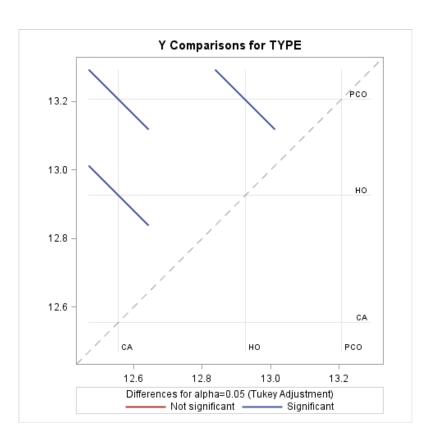
## Least Squares Means Adjustment for Multiple Comparisons: Tukey

TYPE	Y LSMEAN	Standard Error	Pr >  t	LSMEAN Number
НО	12.925556	0.0496863	<.0001	1
PCO	13.2061111	0.0496863	<.0001	2
CA	12.5550000	0.0496863	<.0001	3

Least Squares Means for effect TYPE Pr >  t  for H0: LSMean(i)=LSMean(j) Dependent Variable: Y					
i/j	1	2	3		
1		0.0013	<.0001		
2	0.0013		<.0001		
3	<.0001	<.0001			



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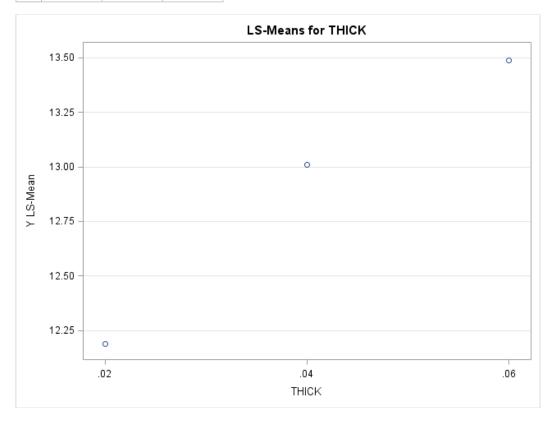
SAS Output Page 7 of 20

## SAS OUTPUT FOR PROBLEM IV

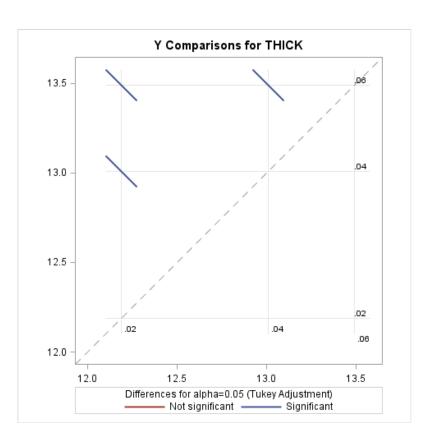
## Least Squares Means Adjustment for Multiple Comparisons: Tukey

THICK	Y LSMEAN	Standard Error	Pr >  t	LSMEAN Number
.02	12.1877778	0.0496863	<.0001	1
.04	13.0094444	0.0496863	<.0001	2
.06	13.4894444	0.0496863	<.0001	3

Least Squares Means for effect THICK Pr >  t  for H0: LSMean(i)=LSMean(j) Dependent Variable: Y					
i/j	1	2	3		
1		<.0001	<.0001		
2	<.0001		<.0001		
3	<.0001	<.0001			



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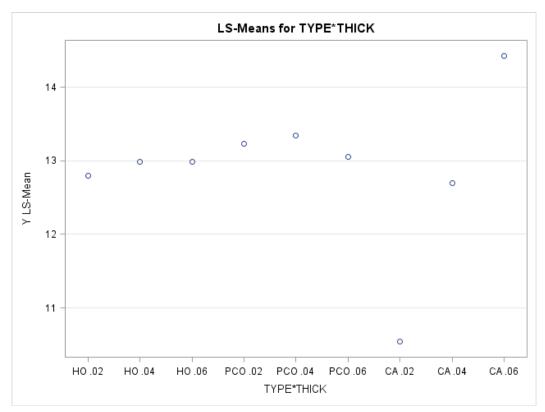
## SAS OUTPUT FOR PROBLEM IV

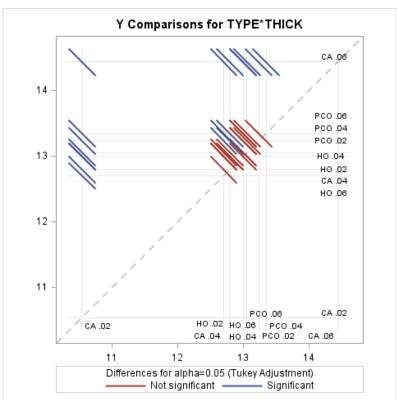
Least Squares Means Adjustment for Multiple Comparisons: Tukey

TYPE	THICK	Y LSMEAN	Standard Error	Pr >  t	LSMEAN Number
НО	.02	12.7933333	0.0860591	<.0001	1
но	.04	12.9916667	0.0860591	<.0001	2
НО	.06	12.9916667	0.0860591	<.0001	3
PCO	.02	13.2283333	0.0860591	<.0001	4
PCO	.04	13.3400000	0.0860591	<.0001	5
PCO	.06	13.0500000	0.0860591	<.0001	6
CA	.02	10.5416667	0.0860591	<.0001	7
CA	.04	12.6966667	0.0860591	<.0001	8
CA	.06	14.4266667	0.0860591	<.0001	9

	Least Squares Means for effect TYPE*THICK Pr >  t  for H0: LSMean(i)=LSMean(j) Dependent Variable: Y										
i/j	1	2	3	4	5	6	7	8	9		
1		0.7808	0.7808	0.0310	0.0033	0.4882	<.0001	0.9961	<.0001		
2	0.7808		1.0000	0.5913	0.1434	0.9999	<.0001	0.3117	<.0001		
3	0.7808	1.0000		0.5913	0.1434	0.9999	<.0001	0.3117	<.0001		
4	0.0310	0.5913	0.5913		0.9899	0.8614	<.0001	0.0045	<.0001		
5	0.0033	0.1434	0.1434	0.9899		0.3323	<.0001	0.0004	<.0001		
6	0.4882	0.9999	0.9999	0.8614	0.3323		<.0001	0.1324	<.0001		
7	<.0001	<.0001	<.0001	<.0001	<.0001	<.0001		<.0001	<.0001		
8	0.9961	0.3117	0.3117	0.0045	0.0004	0.1324	<.0001		<.0001		
9	<.0001	<.0001	<.0001	<.0001	<.0001	<.0001	<.0001	<.0001			

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## SAS OUTPUT FOR PROBLEM IV

#### **The Mixed Procedure**

Model Information							
Data Set	WORK.MANU						
Dependent Variable	Υ						
Covariance Structure	Variance Components						
Estimation Method	REML						
Residual Variance Method	Profile						
Fixed Effects SE Method	Model-Based						
Degrees of Freedom Method	Containment						

Class Level Information							
Class	Levels	Values					
TYPE	3	CA HO PCO					
MACH	3	M1 M2 M3					
THICK	3	.02 .04 .06					

Dimensions						
<b>Covariance Parameters</b>	3					
Columns in X	16					
Columns in Z	36					
Subjects	1					
Max Obs Per Subject	54					

Number of Observations	
Number of Observations Read	54
Number of Observations Used	54
Number of Observations Not Used	0

Iteration History									
Iteration	Evaluations	-2 Res Log Like	Criterion						
0	1	79.60478912							
1	1	33.50991503	0.00000000						

## Convergence criteria met.

Covariance Parameter Estimates									
Cov Parm Estimate Standard Error Z Value Pr > Z Alpha Lower Uppe									
MACH(TYPE)	0.2194	0.1359	1.61	0.0532	0.05	0.08676	1.2528		
MACH*THICK(TYPE)	0.02501	0.02021	1.24	0.1079	0.05	0.008092	0.3330		
Residual	0.04444	0.01209	3.67	0.0001	0.05	0.02778	0.08233		

Fit Statistics	
-2 Res Log Likelihood	33.5

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AIC (smaller is better)	39.5
AICC (smaller is better)	40.1
BIC (smaller is better)	40.1

Type 3 Tests of Fixed Effects									
Effect Num DF Den DF F Value P									
TYPE	2	6	1.36	0.3256					
THICK	2	12	82.58	<.0001					
TYPE*THICK	4	12	80.14	<.0001					

	Least Squares Means										
Effect	TYPE	THICK	Estimate	Standard Error	DF	t Value	Pr >  t				
TYPE	CA		12.5550	0.2800	6	44.84	<.0001				
TYPE	НО		12.9256	0.2800	6	46.17	<.0001				
TYPE	PCO		13.2061	0.2800	6	47.17	<.0001				
THICK		.02	12.1878	0.1721	12	70.81	<.0001				
THICK		.04	13.0094	0.1721	12	75.58	<.0001				
THICK		.06	13.4894	0.1721	12	78.37	<.0001				
TYPE*THICK	CA	.02	10.5417	0.2981	12	35.36	<.0001				
TYPE*THICK	CA	.04	12.6967	0.2981	12	42.59	<.0001				
TYPE*THICK	CA	.06	14.4267	0.2981	12	48.39	<.0001				
TYPE*THICK	НО	.02	12.7933	0.2981	12	42.91	<.0001				
TYPE*THICK	НО	.04	12.9917	0.2981	12	43.58	<.0001				
TYPE*THICK	НО	.06	12.9917	0.2981	12	43.58	<.0001				
TYPE*THICK	PCO	.02	13.2283	0.2981	12	44.37	<.0001				
TYPE*THICK	PCO	.04	13.3400	0.2981	12	44.75	<.0001				
TYPE*THICK	PCO	.06	13.0500	0.2981	12	43.77	<.0001				

	Differences of Least Squares Means										
Effect	TYPE	THICK	_TYPE	_THICK	Estimate	Standard Error	DF	t Value	Pr >  t	Adjustment	Adj P
TYPE	CA		НО		-0.3706	0.3959	6	-0.94	0.3855	Tukey	0.6396
TYPE	CA		PCO		-0.6511	0.3959	6	-1.64	0.1512	Tukey	0.2996
TYPE	НО		PCO		-0.2806	0.3959	6	-0.71	0.5052	Tukey	0.7676
THICK		.02		.04	-0.8217	0.1024	12	-8.02	<.0001	Tukey-Kramer	<.0001
THICK		.02		.06	-1.3017	0.1024	12	-12.71	<.0001	Tukey-Kramer	<.0001
THICK		.04		.06	-0.4800	0.1024	12	-4.69	0.0005	Tukey-Kramer	0.0014
TYPE*THICK	CA	.02	CA	.04	-2.1550	0.1774	12	-12.15	<.0001	Tukey-Kramer	<.0001
TYPE*THICK	CA	.02	CA	.06	-3.8850	0.1774	12	-21.90	<.0001	Tukey-Kramer	<.0001
TYPE*THICK	CA	.02	НО	.02	-2.2517	0.4216	12	-5.34	0.0002	Tukey-Kramer	0.0037
TYPE*THICK	CA	.02	НО	.04	-2.4500	0.4216	12	-5.81	<.0001	Tukey-Kramer	0.0018
TYPE*THICK	CA	.02	НО	.06	-2.4500	0.4216	12	-5.81	<.0001	Tukey-Kramer	0.0018
TYPE*THICK	CA	.02	PCO	.02	-2.6867	0.4216	12	-6.37	<.0001	Tukey-Kramer	0.0008
TYPE*THICK	CA	.02	PCO	.04	-2.7983	0.4216	12	-6.64	<.0001	Tukey-Kramer	0.0005
TYPE*THICK	CA	.02	PCO	.06	-2.5083	0.4216	12	-5.95	<.0001	Tukey-Kramer	0.0015

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TYPE*THICK	C 4	04	C 4	00	4 7000	0.4774	40	0.75	. 0004	Tules Massas	. 0004
	CA	.04	CA	.06	-1.7300	0.1774	12	-9.75	<.0001	Tukey-Kramer	<.0001
TYPE*THICK	CA	.04	НО	.02	-0.09667	0.4216	12	-0.23	0.8225	Tukey-Kramer	1.0000
TYPE*THICK	CA	.04	НО	.04	-0.2950	0.4216	12	-0.70	0.4975	Tukey-Kramer	0.9978
TYPE*THICK	CA	.04	НО	.06	-0.2950	0.4216	12	-0.70	0.4975	Tukey-Kramer	0.9978
TYPE*THICK	CA	.04	PCO	.02	-0.5317	0.4216	12	-1.26	0.2313	Tukey-Kramer	0.9252
TYPE*THICK	CA	.04	PCO	.04	-0.6433	0.4216	12	-1.53	0.1530	Tukey-Kramer	0.8245
TYPE*THICK	CA	.04	PCO	.06	-0.3533	0.4216	12	-0.84	0.4184	Tukey-Kramer	0.9927
TYPE*THICK	CA	.06	НО	.02	1.6333	0.4216	12	3.87	0.0022	Tukey-Kramer	0.0391
TYPE*THICK	CA	.06	НО	.04	1.4350	0.4216	12	3.40	0.0052	Tukey-Kramer	0.0835
TYPE*THICK	CA	.06	НО	.06	1.4350	0.4216	12	3.40	0.0052	Tukey-Kramer	0.0835
TYPE*THICK	CA	.06	PCO	.02	1.1983	0.4216	12	2.84	0.0148	Tukey-Kramer	0.1974
TYPE*THICK	CA	.06	PCO	.04	1.0867	0.4216	12	2.58	0.0242	Tukey-Kramer	0.2867
TYPE*THICK	CA	.06	PCO	.06	1.3767	0.4216	12	3.27	0.0068	Tukey-Kramer	0.1039
TYPE*THICK	НО	.02	НО	.04	-0.1983	0.1774	12	-1.12	0.2855	Tukey-Kramer	0.9599
TYPE*THICK	НО	.02	НО	.06	-0.1983	0.1774	12	-1.12	0.2855	Tukey-Kramer	0.9599
TYPE*THICK	НО	.02	PCO	.02	-0.4350	0.4216	12	-1.03	0.3225	Tukey-Kramer	0.9743
TYPE*THICK	НО	.02	PCO	.04	-0.5467	0.4216	12	-1.30	0.2191	Tukey-Kramer	0.9144
TYPE*THICK	НО	.02	PCO	.06	-0.2567	0.4216	12	-0.61	0.5540	Tukey-Kramer	0.9992
TYPE*THICK	НО	.04	НО	.06	-866E-17	0.1774	12	-0.00	1.0000	Tukey-Kramer	1.0000
TYPE*THICK	НО	.04	PCO	.02	-0.2367	0.4216	12	-0.56	0.5849	Tukey-Kramer	0.9995
TYPE*THICK	НО	.04	PCO	.04	-0.3483	0.4216	12	-0.83	0.4248	Tukey-Kramer	0.9934
TYPE*THICK	НО	.04	PCO	.06	-0.05833	0.4216	12	-0.14	0.8923	Tukey-Kramer	1.0000
TYPE*THICK	НО	.06	PCO	.02	-0.2367	0.4216	12	-0.56	0.5849	Tukey-Kramer	0.9995
TYPE*THICK	НО	.06	PCO	.04	-0.3483	0.4216	12	-0.83	0.4248	Tukey-Kramer	0.9934
TYPE*THICK	НО	.06	PCO	.06	-0.05833	0.4216	12	-0.14	0.8923	Tukey-Kramer	1.0000
TYPE*THICK	PCO	.02	PCO	.04	-0.1117	0.1774	12	-0.63	0.5409	Tukey-Kramer	0.9989
TYPE*THICK	PCO	.02	PCO	.06	0.1783	0.1774	12	1.01	0.3347	Tukey-Kramer	0.9779
TYPE*THICK	PCO	.04	PCO	.06	0.2900	0.1774	12	1.63	0.1281	Tukey-Kramer	0.7715

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# SAS OUTPUT FOR PROBLEM IV

#### The GLM Procedure

Class Level Information				
Class	Levels	Values		
TRT	9	CA.02 CA.04 CA.06 HO.02 HO.04 HO.06 PCO.02 PCO.04 PCO.06		

Number of Observations Read	54	
Number of Observations Used	54	

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## SAS OUTPUT FOR PROBLEM IV

#### The GLM Procedure

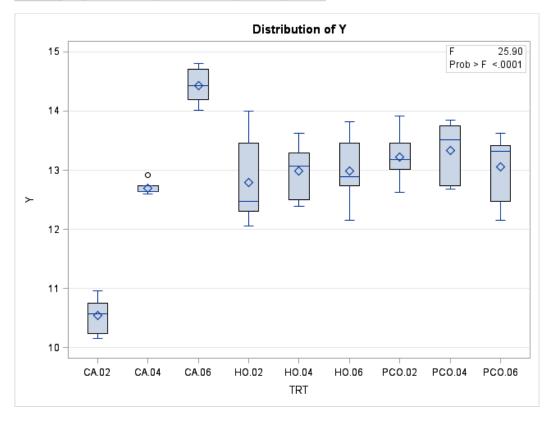
## **Dependent Variable: Y**

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	8	49.71420000	6.21427500	25.90	<.0001
Error	45	10.79873333	0.23997185		
<b>Corrected Total</b>	53	60.51293333			

R-Square	Coeff Var	Root MSE	Y Mean
0.821547	3.798745	0.489869	12.89556

Source	DF	Type I SS	Mean Square	F Value	Pr > F
TRT	8	49.71420000	6.21427500	25.90	<.0001

Source	DF	Type III SS	Mean Square	F Value	Pr > F
TRT	8	49.71420000	6.21427500	25.90	<.0001



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## SAS OUTPUT FOR PROBLEM IV

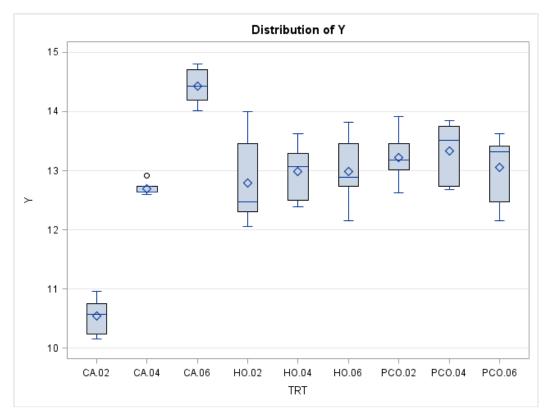
#### The GLM Procedure

Brown and Forsythe's Test for Homogeneity of Y Variance ANOVA of Absolute Deviations from Group Medians							
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F		
TRT	8	0.8672	0.1084	0.94	0.4948		
Error	45	5.1969	0.1155				

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## SAS OUTPUT FOR PROBLEM IV

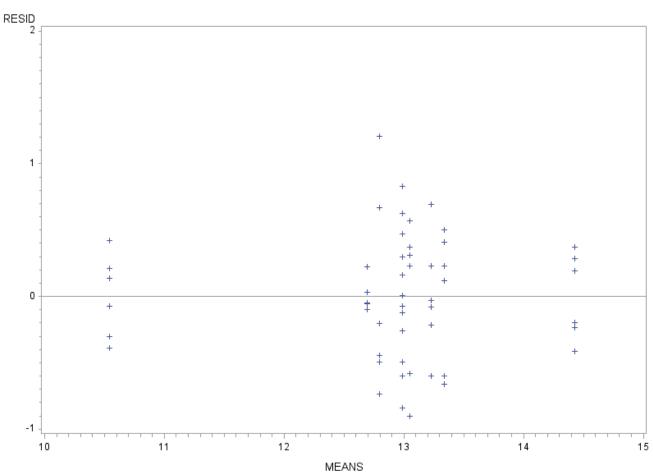
#### The GLM Procedure



Level of		Υ			
TRT	N	Mean	Std Dev		
CA.02	6	10.5416667	0.31211643		
CA.04	6	12.6966667	0.11741664		
CA.06	6	14.4266667	0.32413989		
HO.02	6	12.7933333	0.76413786		
HO.04	6	12.9916667	0.47173792		
HO.06	6	12.9916667	0.58348665		
PCO.02	6	13.2283333	0.43466846		
PCO.04	6	13.3400000	0.50616203		
PCO.06	6	13.0500000	0.59282375		

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## SAS OUTPUT FOR PROBLEM IV

The UNIVARIATE Procedure Variable: RESID

Moments						
N	54	Sum Weights	54			
Mean	0	Sum Observations	0			
Std Deviation	0.4513864	Variance	0.20374969			
Skewness	0.13033828	Kurtosis	-0.1764522			
Uncorrected SS	10.7987333	Corrected SS	10.7987333			
Coeff Variation		Std Error Mean	0.06142591			

Basic Statistical Measures						
Location Variability						
Mean	0.00000	Std Deviation	0.45139			
Median	-0.03750	Variance	0.20375			
Mode	-0.05667	Range	2.10667			
		Interquartile Range	0.60000			

Note: The mode displayed is the smallest of 2 modes with a count of 2.

Tests for Location: Mu0=0						
Test	Statistic p Value			lue		
Student's t	t	0	Pr >  t	1.0000		
Sign	М	-1	Pr >=  M	0.8919		
Signed Rank	S	-4.5	Pr >=  S	0.9695		

Tests for Normality							
Test	Statistic		p Value				
Shapiro-Wilk	w	0.98865	Pr < W	0.8877			
Kolmogorov-Smirnov	D	0.054676	Pr > D	>0.1500			
Cramer-von Mises	W-Sq	0.028584	Pr > W-Sq	>0.2500			
Anderson-Darling	A-Sq	0.190357	Pr > A-Sq	>0.2500			

Quantiles (Definition 5)			
Quantile	Estimate		
100% Max	1.206667		
99%	1.206667		
95%	0.691667		
90%	0.570000		
75% Q3	0.298333		
50% Median	-0.037500		
25% Q1	-0.301667		
10%	-0.600000		

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5%	-0.733333
1%	-0.900000
0% Min	-0.900000

Extreme Observations						
Lowest		Highest				
Value	Obs	Value	Obs			
-0.900000	34	0.628333	11			
-0.841667	18	0.666667	2			
-0.733333	6	0.691667	23			
-0.660000	28	0.828333	17			
-0.601667	12	1.206667	5			

