

1. (a) Since the P -values of the Wald tests of $H_0 : \beta_{\mathbf{k618}} = 0$ and $H_0 : \beta_{\mathbf{hc}} = 0$ equal 0.3424 and 0.5877, respectively, we could consider eliminating these variables. To determine whether this is appropriate, we should carry out a simultaneous test of $H_0 : \beta_{\mathbf{k618}} = \beta_{\mathbf{hc}} = 0$.
 - (b) Since there are the same number of unique profiles of the predictors as the number of observations, we need to use the Hosmer-Lemeshow test for goodness of fit. Since the H-L $X^2 = 24.7058$ with a P -value = 0.0917, we conclude that there is strong evidence of lack of fit for this model.
 - (c) The odds of women who went to college participating in the labor force are estimated to be $e^{\hat{\beta}_{wc}} = e^{0.8072} = 2.24$ times the odds of a woman who did not go to college.
 - (d)
 - Estimated sensitivity: $\hat{P}(\hat{y} = 1|y = 1) = 290/428 = 0.6776$
 - Estimated specificity: $\hat{P}(\hat{y} = 0|y = 0) = 220/325 = 0.6769$
 - Estimated probability of correct classification: $(290 + 220)/753 = 0.6773$
 - (e) The marginal model plot for **age** has similar smooth lines indicating that the model linear in age is appropriate. The marginal model plot for **lwg** has a curve for the smoothed observations being highly curved whereas the smoothed predicted values are much more linear. This indicates that the model with a linear term in **lwg** is not appropriate and some nonlinear model in **lwg** should be used.
2. (a) The response variable **pubs** is ordered, so a cumulative odds model is reasonable. The test for the proportional odds assumption has $X^2 = 5.7017$ with a P -value = 0.3363 indicating that the proportional odds assumption is reasonable.
 - (b)
 - Proportional odds model: Reject $H_0 : \beta_{\mathbf{mar}} = 0$ at level 0.05 since $X^2 = 3.9481$ with a P -value = 0.0469 and conclude that **mar** is useful in the model, given that the other predictors are present.
 - Baseline category model: Do not reject $H_0 : \beta_{\mathbf{mar},1} = \beta_{\mathbf{mar},2} = 0$ at level 0.05 since $X^2 = 4.3792$ with a P -value = 0.1120 and conclude that **mar** is not useful in the model, given that the other predictors are present.
 - (c) To make $P(\mathbf{pubs} = 0)$ to be as large as possible, we need $\text{logit}(P(Y \leq 0))$ to be large. To make the linear predictor as large as possible, choose the largest value of predictors with positive coefficients and the smallest value of variables with negative coefficients. Thus, we would choose **fem** = 1, **mar** = 0, **kid5** = 3, **phd** = 0.75, **ment** = 0.
 - (d) $\log(\hat{\pi}_2/\hat{\pi}_1) = -0.351 + 0.092x - (-0.531 + 0.062x) = 0.180 + 0.030x$. Now $\hat{\pi}_1 > \hat{\pi}_2$ iff $\log(\hat{\pi}_2/\hat{\pi}_1) = 0.18 + 0.03x < 0$. This occurs for $x < -0.18/0.03 = -6$. Since $x \geq 0$, $\hat{\pi}_1 < \hat{\pi}_2$ for all values of x .
3. (a) Since $G^2 = 9.0862$ with a P -value = 0.0106, we reject $H_0 : \beta_{P \times G} = 0$ and conclude that there is strong evidence that the ORs between **gender** and **accept** differ for the three programs.
 - (b) Using the homogeneous association model, reject $H_0 : \beta_G = 0$ since the Wald $X^2 = 23.176$ with a P -value < 0.0001. There is strong evidence of partial association between **gender** and **accept**, controlling for **program**.
 - (c)
 - i. Homogeneous association model: $\text{logit}(\hat{\pi}(f, p)) - \text{logit}(\hat{\pi}(m, c)) = 0.5184 + 0.6612 + 0.00653 - (0.5184 + 0 - 3.5477) = 4.21543$. Then $\widehat{OR} = e^{4.21543} = 67.72$.
 - ii. Model with interaction: $\text{logit}(\hat{\pi}(f, p)) - \text{logit}(\hat{\pi}(m, c)) = 0.5339 + 0.2334 - 0.0429 + 0.8389 - (0.5339 + 0 - 3.3162 + 0) = 4.3456$. Then $\widehat{OR} = e^{4.3456} = 77.14$.