Joseph Blubaugh Stat 630 Hw08

1) 6.1.7

$$P(\theta) = \frac{e^{-\theta}\theta^x}{x!}$$

$$a)X = Poisson(\lambda) = \frac{e^{-\lambda}\lambda^x}{x!}$$

$$L(\theta|x_1, ..., x_n) = \prod_{i=1}^n \frac{e^{-\theta}\theta^{x_i}}{x_i!}$$

$$= \frac{e^{-\theta}\theta^x}{x!}, where, x = \sum_{i=1}^n x_i$$

$$b)f_{\theta} = h(s)g_{\theta}(T(s))$$

$$f_{\theta}(x_1, ..., x_n) = \prod_{i=1}^n \frac{e^{-\theta}\theta^{x_i}}{x!}$$

$$= \frac{e^{-\theta}\theta^{T}(x_1, ..., x_n)}{T(x_1, ..., x_n)} 1$$

$$where, T(x_1, ..., x_n) = \sum_{i=1}^n x_1, and$$

$$g_{\theta}(T) = \frac{e^{-\theta}\theta^T}{T!}, and$$

$$h(x_1, ..., h_n) = 1$$

2) 6.1.19

$$f_{\theta} = \frac{\theta^{a_0}}{\gamma(a_0)} x^{a_0 - 1} e^{-\theta x}$$

$$f_{\theta} = h(x) g_{\theta}(T(x))$$

$$= \prod_{i=1}^{n} \frac{\theta^{a_0}}{\gamma(a_0)} x^{a_0 - 1} e^{-\theta x} 1$$

$$where, h = 1$$

$$T(x_1, ..., x_n) = \sum_{i=1}^{m} x_i$$

$$g_{\theta}(T) = \frac{\theta^{(a_0)}}{\gamma(a_0)} \gamma^{a_0 - 1} e^{-\theta \gamma}$$

3) 6.2.4

$$Y = Poisson(\theta) = \frac{e^{-\theta}\theta^x}{n!}$$

$$l(\theta|x_1, ..., x_n) = \sum_{i=1}^n log(f_{\theta}(x_i))$$

$$log(f_{\theta}(x) = -\theta + xlog(\theta) + log(\frac{1}{n}))$$

$$\frac{dl(\theta|x_1, ..., x_n)}{d\theta} = \frac{x}{\theta} - n = 0$$

$$\theta = \frac{x}{n}$$

$$b)Bias_{\theta}(T) = E(T) - \psi(\theta)$$

$$= \theta - \frac{x}{n}$$

$$variance = \theta$$

$$MSE = \theta - \theta - \frac{x}{n} = \frac{x}{n}$$

4) 6.2.5

$$f = a_0 log(\theta) + (a_0 - 1) log(x) - \theta x + log(\frac{1}{\gamma a_0})$$

$$f' = \frac{a_0}{\theta} - x = 0$$

$$\theta = \frac{a_0}{x}$$

$$b) bias = E_{\theta}(T) - \psi(\theta)$$

$$= \frac{a_0}{\theta} - \frac{a_0}{\theta} = 0$$

$$c) E(x) = a_0 \theta, \theta = \frac{a_0}{E(x)}$$

$$\hat{\theta} = \frac{a_0}{\bar{x}}$$

5) 6.2.6

a) These are independent evnts of successes or failurs so we can use the negative binomial distribution

$$P_x(x) = {r-1-x \choose x} \theta^r (1-\theta)^x \text{ where, } r = 1$$

$$= {x \choose x} \theta(\theta^x) = \theta^{x+1}$$

$$\log(\theta^{x+1}) = (x+1)\log(\theta)$$

$$\frac{dl(\theta^{x+1})}{d\theta} = \frac{x+1}{\theta} = 0$$

$$b)E(x) = \frac{(1-\theta)}{\theta}, \neq \frac{x+1}{\theta}$$

6) 6.2.8

$$Weibull(\beta) = \beta(1+x)^{-\beta-1}$$
$$l(\beta|x) = log(\beta) + (-\beta - 1)log(x+1)$$
$$\frac{dl(\beta|x)}{x\beta} = \frac{1}{\beta} + \frac{(-\beta - 1)}{1+x}$$

7) 6.2.12

a) $mle = \hat{\sigma}^2 = \frac{n-1}{n}s^2 = \frac{1}{n}\sum_{i=1}^n (x_i - \bar{x})^2$ any distribution with a mean μ and variance σ^2 ... They are the same

$$L(\mu, \sigma | x_1, ... x_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma}} e^{.5(\frac{x_i - \mu}{\sigma})^2}$$

$$l(\mu, \sigma) = \log(L(\mu, \sigma | x_1, ..., x_n))$$

$$= \frac{-n}{2} \log(2\pi) - n\log(\sigma) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

$$\frac{dl(\mu, \sigma)}{d\sigma} = \frac{-n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 = 0$$

$$\sigma^2 = \frac{1}{n} \sum_{i=1}^n i = 1^n (x_i - \hat{x})^2$$

8) 6.2.19

$$AA = \theta, Aa = 2\theta(1-\theta), aa = (1-\theta)^2$$

a) Bernoulli Distribution

$$b)L(\theta|x_1, x_2, x_3) = \theta^{x_1} 2\theta (1 - \theta)^{x_2} (1 - \theta)^{2x_3}$$
$$l(\theta|x_1, x_2, x_3) = x_1 log(\theta) + x_2 log(2\theta - \theta^2) + 2x_3 log(1 - \theta)$$
$$score = \frac{x_1}{\theta} + \frac{x_2}{2\theta - \theta^2} + \frac{2x_3}{1 - \theta} = 0$$

9) 6.3.15

- a) yes
- b) no

10) 6.3.24

$$\psi(\theta) = aT_1 + (1 - a)T_2$$

$$Bias_{\theta}(T) = E_{\theta}(T) - \psi(\theta)$$

$$0 = E_{\theta}(T) - \psi(\theta)$$

$$\psi(\theta) = E_{\theta}(T)$$

$$b)var_{\theta}(T_1) = aT_1, var_{\theta}(T_2) = T_2(1-a)$$

c) when a = .5, both T1 and T2 have the same weight and will minimize $T_1 + (1-a)T_2$ when T_1 and T_2 are unknown

10) additional

a) $x_1...x_n$, $exponential(\lambda) = \lambda e^{-\lambda x}$ $T_c = \frac{c}{\sum_{i=1}^n x_i}$, when C = 1 the estimator will have the smallest MSE because the expected mean $exponential(\lambda) = \frac{1}{\lambda}$