# L5: Hardness of Estimation: Mean Estimation

## Review Learning Theory

**VC-Dimension:**  $\nu \approx$  number of parameters in model class F "dimension for model F"

Labeled data (X, y) size n

Sample Error (training error)

- separable:  $n \approx (\nu/\varepsilon) \log(\nu/\varepsilon)$  $error(X) \le \varepsilon \approx (\nu/n) \log(\nu/n)$
- non-separable:  $n \approx \nu/\varepsilon^2$  $error(X) \le \varepsilon \approx \nu/\sqrt{n}$

increases with  $\nu$  increasing

#### **Model Error**

•  $\gamma_F(X) = \min_{f \in F} error(f, X, y)$ 

decreases with  $\nu$  increasing

Total Error (test error) = Sample Error + Model Error

## **Parameter Estimation**

(Bayesian-y view)

Assume data  $X \sim g(\alpha)$  for some model g with parameters  $\alpha \in \mathbb{R}^d$ .

Distributional so g provides probability distribution e.g., each  $x \in X$  from "perfect"  $h(\alpha) + Noise$ , where Noise is random, independent of h.

The simplest case is

- $h(\alpha) = \alpha$
- $Noise = \mathcal{G}_d(0, I)$  (d-dimensional Gaussian/Normal noise)
- Goal: from  $X \sim g$ , recover  $\alpha$

Alternatively, if

- $h(\alpha) = 0$
- $Noise = \mathcal{G}_d(\alpha, I)$
- Goal: from  $X \sim g$ , recover  $\alpha$  the same problem, but now clear we are aiming to recover the **mean** which is  $E_{X \sim Noise}[X] = \alpha$

Chebyshev Inequality (Law of Large Numbers)

For *n* iid RVs  $X_1, X_2, \dots, X_n$  with  $Var[X_i] = \sigma^2$ 

$$Pr[|\bar{x} - E[X_j]| \ge \eta] \le \frac{\sigma^2}{n\eta^2}$$

Note that simple Chebyshev we have

Pr[
$$|X_j - E[X_j]| \ge \eta$$
]  $\le \frac{\sigma^2}{\eta^2}$ , but for  $Var[\bar{x}] = Var[X_j]/n = \sigma^2/n$  so  $Pr[|X - E[X]| \ge \eta] \le \frac{Var[\bar{x}]}{\eta^2} = \frac{\sigma^2}{n\eta^2}$ 

Chernoff-Hoeffding Inequality (simplified as in Azuma)

For n iid RVs  $X_1, X_2, \dots X_n$  with  $X_i \in [0, \Delta]$ 

$$Pr[|X - E[X]| \ge \eta] \le 2\exp(-\frac{2\eta^2 n}{\Lambda^2})$$

Thus as n increases our bound on the error  $\eta$  from an expected value decreases with  $1/\sqrt{n}$ . Fix either  $Pr[...] = \delta$ , and solve for  $\eta$  as a function of n.

## Trouble with High-Dimensional Mean Estimation

For each  $x \in X \sim G_d(\alpha, I)$ , then

$$E[\|x-\alpha\|^2] = \sum_{j=1}^d E[(x_j-\alpha_j)^2] = \sum_{j=1}^d E[(x_j-E[x_j])^2] = \sum_{j=1}^d Var[x_j] = d$$

Then for  $\bar{x} = \frac{1}{n} \sum_{j=1}^{d} x_i$ 

$$E[\|\bar{x}-\alpha\|^2] = \sum_{j=1}^d Var[\bar{x}_j] = d/n$$

We can also analyze the convergence

$$Pr[\|\bar{x} - \alpha\| > \eta] \le d/(n\eta^2)$$

To show this we will use the **Union Bound** that if there are k events  $E_1, \dots, E_k$ , then the probability all events are true  $Pr[E_1 \& \dots \& E_k] \le 1 - \sum_{j=1}^k Pr[E_j = FALSE]$ .

$$\begin{split} ⪻[(\bar{x}_j - \alpha_j)^2 > (\eta')^2] \leq \frac{Var[X_j]}{n(\eta')^2} = \frac{1}{n(\eta')^2} = \delta' \\ &\text{So applying the union bound on } d \text{ coordinates, with } \delta = \delta' \cdot d, \\ &\text{Setting } \eta = \eta' \cdot \sqrt{d} \text{ so } \eta^2 = (\eta')^2 d, \text{ we have } \\ ⪻[\|\bar{x} - \alpha\|^2 > \eta^2] \leq \delta \end{split}$$

Solving for  $\eta = \eta' \cdot \sqrt{d}$  and  $\eta' = \frac{1}{\sqrt{n\delta}}$ , so  $\eta = \frac{\sqrt{d}}{\sqrt{n\delta}}$ 

Or  $n = d/(\eta^2 \delta)$ 

#### Two Mean Example:

Consider two mean estimations in  $\mathbb{R}^d$ 

$$X_1 \sim \mathcal{G}_d(\alpha,I)$$
 and  $X_2 \sim \mathcal{G}_d(\alpha',I)$ 

where we are promised that  $|\alpha_1 - \alpha_1'| = 2$  and  $\alpha_j = \alpha_j'$  for j > 1.

As n increases, we can get estimates of  $\alpha_1$  and  $\alpha'_1$  to concentrate to values  $\eta$  much less than 2.

But each  $x \in X_1$  has  $E[\|x - \alpha\|^2] = d$ . and  $E[\|\bar{x} - \alpha\|^2] = d/n$ 

Moreover  $Pr[\|\bar{x} - \alpha\| \ge \sqrt{d/n\delta}] \le \delta$ .

Let  $\bar{x}_1 = \frac{1}{|X_1|} \sum_{x \in X_1} x$  and similar for  $\bar{x}_2$ .

 $E[\|\bar{x}_1 - \alpha\|^2] = d/|X_1|.$ 

To get  $\|\bar{x}_1 - \alpha\| \le \eta$ , we need approximately  $d/\eta^2$  samples.

#### **Outliers:**

One outlier can significantly affect sample mean  $\bar{x}$ 

In d=1, the median is a good estimate for  $\alpha$  and resistant to outliers.

What is analog in high dimensions?

- coordinate-wise median:  $v = (v_1, ..., v_d)$  has  $v_j$  as median of j th coordinates.
- L1 median (geometric median): v minimize sum of distances to  $x \in X$  centerpoint: v so no halfspace containing v contains more than |X|/(d+1) points Turkey median:  $v = \arg\max_{v \in \mathbb{R}^d} \min_{h \in H, v \in h} \frac{|X \cap h|}{|X|}$

Turkey median works well (uses  $d/\eta^2$  samples, even with outliers), but best algorithms take about  $|X|^{d-1}$  time to compute.

## Robust High-Dimensional Mean Estimation

With  $n=d/\eta^2$  samples, one can use new approaches - that allow for  $\eta$ -fraction of outliers, rest from  $\mathcal{G}_d(0,I)$  - in  $\operatorname{poly}(nd/\eta)$  time - find a point  $\hat{v} \in \mathbb{R}^d$  so  $\|v-\hat{v}\| \leq \tilde{O}(\eta)$ 

## Careful Pruning

Start with Sample Mean  $\bar{x}$ , and center data.

Compute top principal vector u, project data along  $u: X_u = \{x_u = \langle x, u \rangle \mid x \in X\}$ . Compare CDF to that of 1-d normal. Prune extreme points that are too far off. [\*] Repeat until no sign of outliers.

[\*] Vershynin (2011):  $\Sigma \in \mathbb{R}^{n \times d}$  so each entry iid  $\Sigma_{i,j} \sim \mathcal{N}(0,1)$ . With probability at least  $1-2\exp(-t^2/2)$ 

$$\sqrt{n} - \sqrt{d} - t \leq s_{\min}\left(\Sigma\right) \leq \|\Sigma\|_2 \leq \sqrt{n} + \sqrt{d} + t$$

## Median of Means

Decompose X into k components randomly (e.g., for k = 3, 5, or 7)

Compute mean of each  $\bar{x}_1, \dots, \bar{x}_k$ .

Return coordinate-wise median of set  $\{\bar{x}_1, \dots, \bar{x}_k\}$ .

Works quite well if |X| is large enough to split.