



# MTH102 Engineering Mathematics II

Lesson 7: Introduction to continuous random variables

Term: 2024



# Outline

1 Continuous random variables

2 Mean and variance

3 Uniform distribution



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2 Mean and variance

3 Uniform distribution



## Example 1

Suppose that the office hour is from 6:00 pm. to 8:00 pm. each Tuesday. A student is coming to ask questions during the office hour. Consider the arrival time of the student, then

$$S = \{\text{the time between 6:00 pm and 8:00 pm}\}.$$

Let  $X$  be the time difference in minutes between 6:00 pm and the arrival time, then

$$X(s) \in [0, 120], \forall s \in S.$$

Find the following probabilities:

- the student arrives between 7:00 pm and 7:10 pm;
- the student arrives at 7:00 pm.

**Remark:** if a sample space  $S$  is not discrete, for example  $S = [a, b]$ , then the probability of each outcome can not be listed one by one.



# Probability density function

## Definition

Let  $X$  be a random variable over the outcome space  $S$  which is an interval or union of intervals.  $X$  is called a **continuous random variable** if there exists an integrable function  $f(x)$  satisfying the following:

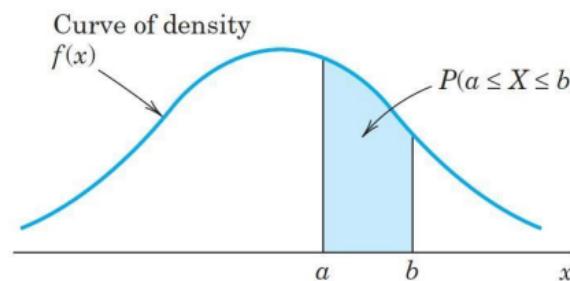
(a)  $f(x) \geq 0, x \in \mathbb{R};$

(b)  $\int_{-\infty}^{\infty} f(x)dx = 1;$

(c) If  $a \leq b$ , then

$$P(a \leq X \leq b) = \int_a^b f(x)dx.$$

$f(x)$  is called the **probability density function (pdf)** of  $X$ .





# Continuous random variable

For a continuous random variable  $X$ , the probability at a particular value  $a \in \mathbb{R}$  is zero, i.e.

$$P(X = a) = 0.$$

Therefore,

$$P(X \leq a) = P(\{X < a\} \cup \{X = a\}) = P(X < a) + P(X = a) = P(X < a).$$

Consequently,

$$P(a \leq X \leq b) = P(a < X \leq b) = P(a \leq X < b) = P(a < X < b).$$



# Cumulative distribution function

The **cumulative distribution function (cdf)** of a continuous random variable  $X$  with the pdf  $f(x)$  is given by

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t)dt, \quad x \in \mathbb{R}.$$

Although  $P(X \leq x) = P(X < x)$ , we write  $P(X \leq x)$  for consistency with the cdf of discrete random variables.

Properties:

- $F(x)$  is continuous, and for  $x$  values for which the derivative  $F'(x)$  exists,

$$F'(x) = f(x).$$

- For any  $(a, b) \subseteq S$ ,

$$P(a \leq X \leq b) = F(b) - F(a).$$



## Example 2

Let  $X$  be a continuous random variable with pdf

$$f(x) = \begin{cases} 2x, & \text{for } 0 < x < 1, \\ 0, & \text{otherwise.} \end{cases}$$

- (a) Find the cdf of  $X$  and sketch it.
- (b) Find  $P\left(\frac{1}{2} < X < \frac{3}{4}\right)$ .

$$\text{Sol: (a)} \quad F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$$

When  $0 < x < 1$ ,

$$\int_{-\infty}^x f(t) dt = \int_0^x 2t dt = 2 \times \frac{1}{2} [t^2]_0^x = x^2$$

When  $x \leq 0$ , then

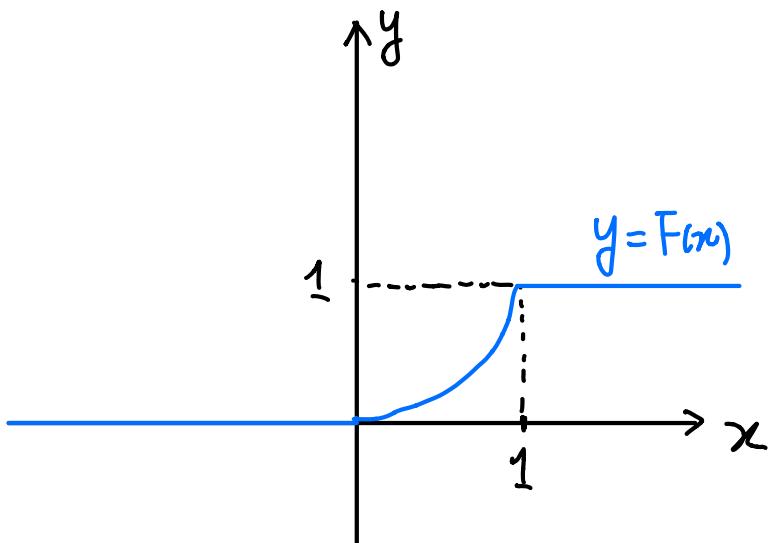
$$\int_{-\infty}^x f(t) dt = \int_{-\infty}^0 0 dt = 0.$$

When  $x \geq 1$ , then

$$\begin{aligned} \int_{-\infty}^x f(t) dt &= \int_{-\infty}^0 f(t) dt + \int_0^1 f(t) dt + \\ &\quad \int_1^x f(t) dt \\ &= \int_{-\infty}^0 0 dt + \int_0^1 2t dt + \int_1^x 0 dt \\ &= 0 + 2 \times \frac{1}{2} [t^2]_0^1 + 0 \\ &= 1 \end{aligned}$$

In conclusion,

$$F(x) = \begin{cases} 0, & \text{if } x \leq 0 \\ x^2, & \text{if } 0 < x < 1 \\ 1, & \text{if } x \geq 1 \end{cases}$$



$$\begin{aligned}
 (2) \quad P\left(\frac{1}{2} < X < \frac{3}{4}\right) &= P\left(\frac{1}{2} < X \leq \frac{3}{4}\right) \\
 &= P(X \leq \frac{3}{4}) - P(X \leq \frac{1}{2}) \\
 &= F\left(\frac{3}{4}\right) - F\left(\frac{1}{2}\right) \\
 &= \left(\frac{3}{4}\right)^2 - \left(\frac{1}{2}\right)^2, \text{ from Part (1).} \\
 &= \frac{5}{16}
 \end{aligned}$$



## Example 3

Suppose that  $X$  is a continuous random variable whose probability density function is given by

$$f(x) = \begin{cases} c(4x - 2x^2) & 0 < x < 2, \\ 0 & \text{otherwise} \end{cases}$$

- 1 Determine the value of  $c$ .
- 2 Find the cdf of  $X$ .
- 3 Find  $P(X > 1)$ .

$$\text{Sol: (1) Since } \int_{-\infty}^{+\infty} f(x) dx = 1$$

$$\Leftrightarrow \int_0^2 c(4x - 2x^2) dx = 1$$

$$\Leftrightarrow c \left( 4 \times \frac{1}{2} [x^2]_0^2 - 2 \times \frac{1}{3} [x^3]_0^2 \right) = 1$$

$$\Leftrightarrow c \left( 8 - \frac{16}{3} \right) = 1$$

$$\Leftrightarrow \frac{8}{3} c = 1$$

$$\Leftrightarrow c = \frac{3}{8}$$

(2)

$$F(x) = \int_{-\infty}^x f(t) dt$$

- When  $x \leq 0$ ,  $F(x) = \int_{-\infty}^x 0 dt = 0$

- When  $0 < x < 2$ ,  $F(x) = \int_0^x \frac{3}{8} (4t - 2t^2) dt$   
 $= \frac{3}{8} \left( 4 \times \frac{1}{2} [t^2]_0^x - 2 \times \frac{1}{3} [t^3]_0^x \right)$   
 $= \frac{3}{8} \left( 2x^2 - \frac{2}{3}x^3 \right)$

- When  $x \geq 2$ ,  $F(x) = \int_{-\infty}^0 f(t) dt + \int_0^2 f(t) dt + \int_2^\infty f(t) dt$

$$= \int_{-\infty}^0 0 dt + \int_0^2 \frac{3}{8} (4t - 2t^2) dt + \int_2^\infty 0 dt$$

$$= 0 + 1 + 0, \text{ from part (1)}$$

$$= 1.$$

Therefore,

$$F(x) = \begin{cases} 0, & \text{if } x \leq 0 \\ \frac{3}{8}(2x^2 - \frac{2}{3}x^3), & \text{if } 0 < x < 2 \\ 1, & \text{if } x \geq 2 \end{cases}$$

(3)  $P(X > 1) = 1 - P(X \leq 1)$ , from complement rule

$$= 1 - F(1)$$

$$= 1 - \frac{3}{8}(2 \times 1^2 - \frac{2}{3} \times 1^3), \text{ from Part (2)}$$

$$= \frac{1}{2}$$



## Example 4

The lifetime in hours of a certain kind of radio tube is a random variable having probability density function given by

$$f(x) = \begin{cases} 0 & x \leq 100, \\ \frac{100}{x^2} & x > 100. \end{cases}$$

What is the probability that exactly 2 of 5 such tubes in a radio set will have to be replaced within the first 150 hours of operation?

Sol: Let  $X$  be the lifetime of 1 radio tube, and  $Y$  be the number of replaced tubes among 5 tubes within the 1st 150 h of operation. Then

$$Y \sim b(5, p),$$

with

$$\begin{aligned} p &= P(X < 150) = \int_{-\infty}^{150} f(t) dt \\ &= \int_{-\infty}^{100} 0 dt + \int_{100}^{150} \frac{1}{t^2} dt \\ &= 0 + 100 \times (-1) \left[ \frac{1}{t} \right]_{100}^{150} \\ &= 100 \times \left( \frac{1}{100} - \frac{1}{150} \right) \\ &= 1 - \frac{2}{3} = \frac{1}{3} \end{aligned}$$

Therefore,

$$\begin{aligned} P(Y=2) &= \binom{5}{2} \times \left(\frac{1}{3}\right)^2 \times \left(1 - \frac{1}{3}\right)^{5-2} \\ &= \frac{5 \times 4}{2} \times \frac{1}{9} \times \left(\frac{2}{3}\right)^3 \\ &= \frac{80}{243} \end{aligned}$$

The probability that exactly 2 of 5 such tubes in a radio set have to be replaced within the 1st 150 h is  $\frac{80}{243}$ .



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# Mean of continuous random variable

In lesson 5, we have defined the mean of a discrete random variable  $X$  by

$$E[X] = \sum_x xP(X = x).$$

## Definition

If  $X$  is a continuous random variable having a probability density function  $f(x)$ , then the **mean**, or the **expectation**, of  $X$ , denoted by  $E(X)$ , is defined by

$$E[X] = \int_{-\infty}^{\infty} xf(x)dx.$$

**Remark.**  $E[X]$  may not exist (the random variable given in Example 4).



## Example 5 (continued from Example 2)

Find the mean of  $X$  when the pdf of  $X$  is

$$f(x) = \begin{cases} 2x, & \text{for } 0 < x < 1, \\ 0, & \text{otherwise.} \end{cases}$$

Sol:  $E(X) = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^1 2x^2 dx$   
 $= 2 \times \frac{1}{3} [x^3]_0^1$   
 $= \frac{2}{3}$



# Expectation of a function of a random variable

Lesson 5

In week 10, we have defined the mean of a function of a discrete random variable  $X$ , i.e. for any real-valued function  $g$ ,

$$E[g(X)] = \sum_x g(x)P(X=x).$$

## Definition

If  $X$  is a continuous random variable having a probability density function  $f(x)$ , then for any real-valued function  $g$ ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx.$$



## Example 6 (continued from Example 2)

Find  $E[e^X]$  when the pdf of  $X$  is

$$f(x) = \begin{cases} 2x, & \text{for } 0 < x < 1, \\ 0, & \text{otherwise.} \end{cases}$$

$$\begin{aligned} \text{Sol: } E(e^X) &= \int_{-b}^{+b} e^x f(x) dx = 2 \int_0^1 x e^x dx \\ &= 2 [xe^x]_0^1 - 2 \int_0^1 e^x dx = 2e - 2 [e^x]_0^1 \\ &= 2e - 2e + 2 \\ &= 2 \end{aligned}$$

$$\begin{aligned} \text{IBP: } u &= x, v' = e^x \\ u' &= 1, v = e^x \end{aligned}$$

IBP: Integration by Parts



# Properties of mean

Let  $X_1, X_2, \dots, X_n$  be random variables. The mean of a sum of random variables equals the sum of the mean of each random variable, i.e.

$$E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots + E(X_n).$$

Moreover, let  $a_1, a_2, \dots, a_n$  be constant. Then

$$E(a_1X_1 + a_2X_2 + \dots + a_nX_n) = a_1E(X_1) + a_2E(X_2) + \dots + a_nE(X_n).$$



## Example 7

Sol: Let  $X_i$  count whether the cab stops at the  $i$ -th station.

So  $X_i \sim \text{Bernoulli}(p)$ , with

$$p = P(\{\text{at least 1 passenger get off at station } i\}) = 1 - \left(\frac{4}{5}\right)^4$$

i.e.  $P(X_i = 1) = p$ ,  $P(X_i = 0) = 1 - p$ . And  $E(X_i) = p$ , for  $i = 1, \dots, 5$ .

A cab drives 4 passengers to their destinations. There are 5 potential cab stops during the route. If there is no one saying that he/she will get off when the cab is approaching one stop, the driver will continue driving without stop. Assume that for each passenger, it is equally likely that he/she will get off at any of the 5 cab stops. In average, how many times will the cab take stops?

Then the number of times in average that the cab takes stops is

$$\begin{aligned} E(X_1 + \dots + X_5) &= E(X_1) + \dots + E(X_5) = 5 E(X_1) = 5p = 5 \times \left[1 - \left(\frac{4}{5}\right)^4\right] \\ &= \frac{369}{125} \approx 2.952 \end{aligned}$$



# Definition of variance

## Definition

The variance of a random variable  $X$ , denoted by  $\text{Var}(X)$ , is defined as the mean of the function of  $X$  with  $g(X) = (X - E[X])^2$ , i.e.

$$\text{Var}(X) = E[(X - E[X])^2] = \int_{-\infty}^{\infty} (x - E[X])^2 f(x) dx.$$

In the practice, it is more convenient to compute the variance via the following equivalent formula

$$\text{Var}(X) = E[X^2] - (E[X])^2 = \int_{-\infty}^{\infty} x^2 f(x) dx - (E[X])^2.$$

- The variance is always nonnegative.
- The square root of  $\text{Var}(X)$ , i.e.  $\sqrt{\text{Var}(X)}$ , is called the *standard deviation* of  $X$ .



## Example 8 (continued from Example 2)

Find the variance and standard deviation of  $X$  when the pdf of  $X$  is

$$f(x) = \begin{cases} 2x, & \text{for } 0 < x < 1, \\ 0, & \text{otherwise.} \end{cases}$$

Sol:  $\mathbb{E}(X^2) = \int_{-\infty}^{+\infty} x^2 f(x) dx = 2 \int_0^1 x^3 dx = \frac{2}{4} [x^4]_0^1 = \frac{1}{2}$

So  $\text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = \frac{1}{2} - \left(\frac{2}{3}\right)^2, \text{ by Example 5}$   
 $= \frac{1}{18}$

And then stand deviation of  $X = \sqrt{\text{Var}(X)} = \sqrt{\frac{1}{18}} = \frac{\sqrt{2}}{6}$



# Property of variance

Let  $X$  be a random variable,  $a, b$ , and  $c$  be constants. Consider  $Y = aX + b$  as a linear function of  $X$ . Then

- $\text{Var}(c) = 0.$
- 

$$\text{Var}(Y) = \text{Var}(aX + b) = a^2 \text{Var}(X).$$



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# Uniform random variable

A random variable  $X$  is said to be *uniformly* distributed over the interval  $(a, b)$  if its probability density function is given by

$$f(x) = \begin{cases} \frac{1}{b-a} & a < x < b, \\ 0 & \text{otherwise.} \end{cases}$$

The cdf of  $X$  is

$$F(x) = \begin{cases} 0 & x \leq a, \\ \frac{x-a}{b-a} & a < x < b, \\ 1 & x \geq b. \end{cases}$$

The mean and variance of  $X$  are

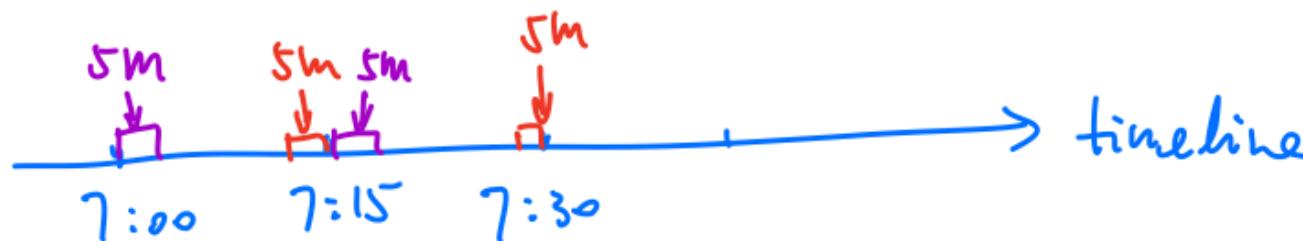
$$E(X) = \frac{a+b}{2}, \quad \text{Var}(X) = \frac{(b-a)^2}{12}.$$



## Example 9

Buses arrive at a specified stop at 15-minute intervals starting at 7:00 am. That is, they arrive at 7:00, 7:15, 7:30, 7:45 and so on. If a passenger arrives at the stop at a time that is uniformly distributed between 7:00 and 7:30, find the probability that he waits

- 1 less than 5 minutes for a bus;
- 2 more than 10 minutes for a bus.



Sol : Suppose  $X$  be the number of minutes past 7 that the passenger arrives at the stop. Then  $X \sim \text{Uniform}[0, 30]$ .

Since the buses arrive every 15 m starting at 7:00 Am,

$\{\text{passenger will wait less than } 5 \text{ m}\} = \{\text{he arrives}$

$\text{during } 7:10 \text{ to } 7:15\} \cup \{\text{he arrives during } 7:25 \text{ to } 7:30\}$ .

$$(a) \quad P(\{\text{passenger will wait less than } 5 \text{ m}\})$$

$$= P(10 < X < 15) + P(25 < X < 30)$$

$$= F(15) - \bar{F}(10) + F(30) - \bar{F}(25)$$

$$= \frac{15}{30} - \frac{10}{30} + \frac{30}{30} - \frac{25}{30}$$

$$= \frac{1}{3}$$

$$(b) \quad \{\text{passenger will wait more than } 10 \text{ m}\}$$

$$= \{\text{passenger arrives at stop during } 7:00 \text{ to } 7:05\} \cup$$

$$\{ \text{---} 7:15 \text{ to } 7:20 \}$$

$$\text{So } P(\{\text{passenger will wait more than } 10 \text{ m}\})$$

$$= P(0 < X < 5) + P(15 < X < 20)$$

$$= F(5) - \bar{F}(0) + F(20) - \bar{F}(15) = \frac{5}{30} - 0 + \frac{20}{30} - \frac{15}{30} = \frac{1}{3}.$$



## Example 10

A stick of length 1 is split randomly at a point  $X$  that is uniformly distributed over  $(0, 1)$ . Suppose a point  $p$  on the stick is located  $p$  unit of distance away from the stick's left endpoint and  $0 \leq p \leq 1$ . Determine the expected length of the piece with the stick's left endpoint that contains the point  $p$ .

Sol : It is known  $X \sim \text{Uniform}(0, 1)$ .

Let  $Y$  be the length of the piece that contains point  $p$ .

Consider the indicator function  $\mathbb{1}_A$  for an event  $A$  as follows:

$$\mathbb{1}_A = \begin{cases} 1, & \text{if } A \text{ occurs} \\ 0, & \text{if } A^c \text{ occurs.} \end{cases}$$

Then let  $A = \{X > p\}$ , we have

$$Y = X \mathbb{1}_{\{X > p\}}.$$

Therefore,

$$\mathbb{E}(Y) = \mathbb{E}(X \mathbb{1}_{\{X > p\}})$$

$$= \int_p^{+\infty} x f(x) dx$$

$$= \int_p^1 x \cdot 1 dx, \quad \text{since } X \sim \text{Unif}(0,1)$$

$$f(x) = \begin{cases} 1, & \text{if } x \in (0,1) \\ 0, & \text{otherwise} \end{cases}$$

In conclusion, the expected length of piece that contains the point  $p$  is  $\frac{1}{2}(1-p^2)$ .



# Exercise

Let  $X$  be the result of an experiment, and  $X$  is uniformly distributed over  $(2, 5)$ . The experiment has been carried out 3 times independently, what is the probability that it happens at least twice that the result is larger than 3?

Sol: Let  $Y$  be the number of times that the result is larger than 3. Since  $X \sim \text{Uniform}(2, 5)$ , we have

$Y \sim b(3, p)$ , with

$$\begin{aligned} p &= P(X > 3) = 1 - P(X \leq 3) \\ &= 1 - F(3) \\ &= 1 - \frac{3-2}{5-2} \\ &= \frac{2}{3} \end{aligned}$$

Then

$$\begin{aligned} P(Y \geq 2) &= 1 - P(Y < 2) \\ &= 1 - P(Y \leq 1) \\ &= 1 - P(Y = 0) - P(Y = 1) \\ &= 1 - \binom{3}{0} \left(\frac{2}{3}\right)^0 \left(\frac{1}{3}\right)^3 - \binom{3}{1} \left(\frac{2}{3}\right) \left(\frac{1}{3}\right)^2 \\ &= 1 - \cancel{\frac{1}{3}} \times 1 \times \frac{1}{27} - 3 \times \frac{2}{3} \times \frac{1}{9} \\ &= 1 - \cancel{\frac{1}{3}} \frac{7}{27} \\ &= \cancel{\frac{1}{3}} \frac{20}{27} \end{aligned}$$

In conclusion, the probability is  ~~$\frac{1}{3}$~~   $\frac{20}{27}$  that at least three the result is larger than 3.



## Example 11

A bakery sells rolls in units of a dozen. The demand  $X$  (in 1000 units) for rolls has a uniform distribution on  $[2, 4]$ . It costs \$2 to make a unit that sells for \$5 on the first day when the rolls are fresh. Any leftover units are sold on the second day for \$1. How many units should be made to maximize the expected value of the profit?

Sol : Demand  $X \sim \text{Uniform}(2, 4)$ .

then  $X$  has density  $f(t) = \begin{cases} \frac{1}{2}, & \text{if } 2 < t < 4 \\ 0, & \text{otherwise} \end{cases}$

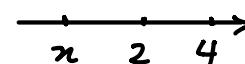
Let  $x$  be the units of rolls made. Then we have the profit

$$Y = \begin{cases} (5-2)x + (1-2)(x-X), & \text{if } X < x \\ (5-2)x, & \text{if } X \geq x \end{cases}$$

$$= \begin{cases} 4X - x, & \text{if } X < x \\ 3x, & \text{if } X \geq x \end{cases}$$

$$\mathbb{E}(Y) = \int_{-\infty}^x (4t-x) f(t) dt + \int_x^{+\infty} 3x f(t) dt$$

• When  $x \leq 2$ ,



$$\int_{-\infty}^x (4t-x) f(t) dt + \int_x^{+\infty} 3x f(t) dt$$

$$= 0 + \int_2^4 \frac{3x}{2} dt = \frac{3}{2}x(4-2) = 3x$$

, it attains the maximum 6(\$), when  $x = 2$  (dozen).  
1000

• When  $2 < x < 4$ ,

$$\int_{-\infty}^x (4t-x) f(t) dt + \int_x^{+\infty} 3x f(t) dt$$

$$= \int_2^x \frac{4t-x}{2} dt + \int_x^4 \frac{3x}{2} dt$$

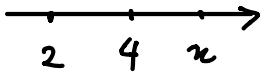
$$= 2 \times \frac{1}{2} [t^2]_2^x - \frac{x}{2}(x-2) + \frac{3}{2}x(4-x)$$

$$= -x^2 + 7x - 4 = -(x - \frac{7}{2})^2 + \frac{53}{4}$$

it attains the maximum  $\frac{53}{4}$  (\$), when  $x = \frac{7}{2}$  (dozen).  
1000

• When  $x \geq 4$

$$\int_{-\infty}^x (4t - x) f(t) dt + \int_x^{+\infty} 3x f(t) dt$$
$$= \int_2^4 \frac{4t-x}{2} dt + 0$$
$$= 2 \times \frac{1}{2} [t^2]_2^4 - \frac{x}{2} (4-2)$$
$$= 12 - x.$$



it attains the maximum  $\delta (\$)$ , when  $x = 4$  (dozen).

In comparison the 3 cases above, we have the maximum profits maybe gained are respectively  $6 (\$) < \delta (\$) < \frac{33}{4} (\$)$

when  $x = 2, 4, \frac{7}{2}$  (dozen).

In conclusion,  ~~$\frac{3500}{2}$  units~~ <sup>3500 dozen</sup> should be made to maximize the expected value of profit.