

## 8 Gaussian Process

In order to understand the Gaussian Process, we need to become conversant with manipulating multivariate normal distributions with density

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{n/2} \sqrt{|\Sigma|}} \exp \left[ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^t \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right],$$

where  $|\Sigma|$  is the determinant of a positive definite covariance matrix  $\Sigma$ .

### 8.1 Schur Complement

Let  $M$  be an invertible  $(n + m) \times (n + m)$  matrix in a block form

$$M = \begin{pmatrix} n & m \\ A & B \\ C & D \end{pmatrix} \begin{matrix} n \\ m \end{matrix}.$$

Because  $M$  is invertible, there exists a unique solution to the linear equation

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix} \begin{pmatrix} x \\ y \end{pmatrix} = \begin{pmatrix} \alpha \\ \beta \end{pmatrix}. \quad (8.1)$$

If  $D$  is invertible, then we can use Gaussian elimination to solve for  $x$  and  $y$ :

$$y = D^{-1}\beta - D^{-1}Cx \Rightarrow (M/D)x = \alpha - BD^{-1}\beta,$$

where  $M/D := A - BD^{-1}C$  is called the **Schur complement of  $D$  in  $M$** ; from the block determinant formula (Theorem A.3),

$$\det(M) = \det(D) \det(M/D),$$

we see that  $(M/D)$  is also invertible when  $M$  and  $D$  are both invertible. The solution to (8.1) is then

$$\begin{aligned} x &= (M/D)^{-1}(\alpha - BD^{-1}\beta) \\ y &= D^{-1}[\beta - C(M/D)^{-1}(\alpha - BD^{-1}\beta)]. \end{aligned}$$

Combining these two equations into a matrix form, we get

$$\begin{pmatrix} x \\ y \end{pmatrix} = \begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} \begin{pmatrix} \alpha \\ \beta \end{pmatrix},$$

where

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} (M/D)^{-1} & -(M/D)^{-1}BD^{-1} \\ -D^{-1}C(M/D)^{-1} & D^{-1} + D^{-1}C(M/D)^{-1}BD^{-1} \end{pmatrix}. \quad (8.2)$$

Similar, if  $A$  and its Schur complement  $M/A := D - CA^{-1}B$  are invertible, then

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} A^{-1} + A^{-1}B(M/A)^{-1}CA^{-1} & -A^{-1}B(M/A)^{-1} \\ -(M/A)^{-1}CA^{-1} & (M/A)^{-1} \end{pmatrix}. \quad (8.3)$$

If  $A$ ,  $M/A$ ,  $D$ , and  $M/D$  are all invertible, then we can combine (8.2) and (8.3) to get

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} (M/D)^{-1} & -A^{-1}B(M/A)^{-1} \\ -D^{-1}C(M/D)^{-1} & (M/A)^{-1} \end{pmatrix}, \quad (8.4)$$

which reproduces the formula (A.6).

**Theorem 8.1.** *If  $M = \begin{pmatrix} A & B \\ C & D \end{pmatrix}$  is a positive definite matrix, then  $A$ ,  $M/A$ ,  $D$ , and  $M/D$  are all invertible and (8.4) holds.*

*Proof.* We immediately see that  $A$  and  $D$  are invertible, because they are principal submatrices of  $M$  and, thus, must be positive definite. The block determinant formula (Theorem A.3) then implies that  $M/A$  and  $M/D$  have positive determinants and are thus invertible. In fact, we can further show that  $M/A$  and  $M/D$  are also positive definite, as follows: rewrite (8.2) as

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} I & 0 \\ -D^{-1}C & I \end{pmatrix} \begin{pmatrix} (M/D)^{-1} & 0 \\ 0 & D^{-1} \end{pmatrix} \begin{pmatrix} I & -BD^{-1} \\ 0 & I \end{pmatrix},$$

which implies that

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix} = \begin{pmatrix} I & BD^{-1} \\ 0 & I \end{pmatrix} \begin{pmatrix} M/D & 0 \\ 0 & D \end{pmatrix} \begin{pmatrix} I & 0 \\ D^{-1}C & I \end{pmatrix}.$$

(Aside: Note that this last equation is valid whenever  $D$  is invertible and does not require that  $M/D$  is also invertible.) When  $M$  is symmetric, which is the case for any positive definite matrix,  $D$  is symmetric and  $C = B^t$ . Hence, the last equation becomes

$$\begin{pmatrix} A & B \\ B^t & D \end{pmatrix} = \begin{pmatrix} I & BD^{-1} \\ 0 & I \end{pmatrix} \begin{pmatrix} M/D & 0 \\ 0 & D \end{pmatrix} \begin{pmatrix} I & BD^{-1} \\ 0 & I \end{pmatrix}^t.$$

Since  $\begin{pmatrix} I & BD^{-1} \\ 0 & I \end{pmatrix}$  is invertible, we now see that  $\begin{pmatrix} M/D & 0 \\ 0 & D \end{pmatrix}$  has to be positive definite. Thus,  $M/D$  has to be positive definite. A similar argument shows that

$$\begin{pmatrix} A & B \\ B^t & D \end{pmatrix} = \begin{pmatrix} I & A^{-1}B \\ 0 & I \end{pmatrix}^t \begin{pmatrix} A & 0 \\ 0 & M/A \end{pmatrix} \begin{pmatrix} I & A^{-1}B \\ 0 & I \end{pmatrix},$$

which implies the  $M/A$  has to be positive definite.  $\square$

## 8.2 Marginal and Conditional Distributions of Multivariate Normal

Define the precision matrix  $\Lambda$  to be the inverse of the covariance matrix  $\Sigma$ . Let  $\mathbf{X} = (X_1, \dots, X_n)$  be a multivariate normal random vector with mean  $\boldsymbol{\mu}$ , and  $\mathbf{X}_a = (X_1, \dots, X_k)$  and  $\mathbf{X}_b = (X_{k+1}, \dots, X_n)$ , for  $1 \leq k < n$ . Then, we can decompose  $\Lambda$  in a block form as

$$\Lambda = \Sigma^{-1} = \begin{pmatrix} \Lambda_{aa} & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_{bb} \end{pmatrix}.$$

Using (8.2), we see that

$$\Lambda_{aa} = (\Sigma / \Sigma_{bb})^{-1} \Rightarrow (\Lambda_{aa})^{-1} = \Sigma / \Sigma_{bb}.$$

By exchanging the role of  $\Sigma$  and  $\Lambda$ , we also get

$$(\Sigma_{aa})^{-1} = \Lambda / \Lambda_{bb}. \quad (8.5)$$

Using this last equation, we can integrate out  $\mathbf{X}_b$  to prove

**Theorem 8.2** (Marginal Distribution). *The marginal distribution of  $\mathbf{X}_a$  is multivariate normal with mean  $\boldsymbol{\mu}_a$  and covariance matrix  $\text{Cov}(\mathbf{X}_a, \mathbf{X}_a)$ .*

*Proof.* Problem Set.  $\square$

Similarly, treating  $\mathbf{X}_b$  as a constant vector shows that

**Theorem 8.3** (Conditional Distribution). *The conditional distribution of  $\mathbf{X}_a$  given  $\mathbf{X}_b = \mathbf{x}_b$  is multivariate normal with mean*

$$E[\mathbf{X}_a | \mathbf{X}_b = \mathbf{x}_b] = \boldsymbol{\mu}_a + \Sigma_{ab} \Sigma_{bb}^{-1} (\mathbf{x}_b - \boldsymbol{\mu}_b)$$

*and variance*

$$\text{Var}[\mathbf{X}_a | \mathbf{x}_b] = \Sigma_{aa} - \Sigma_{ab} \Sigma_{bb}^{-1} \Sigma_{ba} = \Sigma / \Sigma_{bb} = \Lambda_{aa}^{-1}.$$

*Proof.* See the proof of Theorem A.17.  $\square$