Question 6

a)
$$W_t = Y_t - Y_{t-12}$$
$$= \beta_0 + \beta_1 t + s_t + e_t$$
$$- \beta_0 - \beta_1 (t - 12) - s_{t-12} - e_{t-12}$$
$$= 12\beta_1 + e_t - e_{t-12} \qquad (\text{since } s_t = s_{t-12})$$

The mean function is:

$$E(W_t) = 12\beta_1 + E(e_t) - E(e_{t-12})$$

= 12\beta_1 + 0 - 0 = 12\beta_1.

The autocovariance is given by

$$Cov(W_t, W_{t-k})$$
= $Cov(12\beta_1 + e_t - e_{t-12}, 12\beta_1 + e_{t-k} - e_{t-k-12})$
= $Cov(e_t - e_{t-12}, e_{t-k} - e_{t-k-12})$
= $Cov(e_t, e_{t-k}) - Cov(e_t, e_{t-k-12})$
- $Cov(e_{t-12}, e_{t-k}) + Cov(e_{t-12}, e_{t-k-12})$

- $\operatorname{Cov}(e_t, e_{t-k}) = \sigma_e^2$ when k = 0.
- $Cov(e_t, e_{t-k-12}) = 0$ always.
- $Cov(e_{t-12}, e_{t-k}) = \sigma_e^2$ when k = 12.
- $Cov(e_{t-12}, e_{t-k-12} = \sigma_e^2 \text{ when } k = 0.$

$$Cov(W_t, W_{t-k}) = \begin{cases} 2\sigma_e^2 & \text{when } k = 0\\ -\sigma_e^2 & \text{when } k = 12\\ 0 & \text{otherwise} \end{cases}$$

Thus, the mean is constant and the autocovariance depends only on k, i.e., we can write γ_k .

The autocorrelation is given by

$$\rho_{t,t-k} = \frac{\gamma_{t,t-k}}{\sqrt{\gamma_{t,t} \gamma_{t-k,t-k}}}$$

$$= \frac{\gamma_k}{\sqrt{\gamma_0 \gamma_0}} \qquad \text{(series is stationary)}$$

$$= \frac{\gamma_k}{\gamma_0}$$

$$= \begin{cases} \frac{2\sigma_e^2}{2\sigma_e^2} = 1 & \text{when } k = 0 \\ \frac{-\sigma_e^2}{2\sigma_e^2} = -0.5 & \text{when } k = 12 \\ \frac{0}{2\sigma_e^2} = 0 & \text{otherwise} \end{cases}$$

Thus, applying ∇_{12} , i.e., seasonal differencing, eliminates linear and seasonal trend.

b) The mean function would be the same since $E(X_t) = 0$ just like for e_t .

However, unlike e_t , covariance may be non-zero for various lags, k (for e_t , it is only non-zero at lag k = 0).

Thus, using the fact that the autocovariance for X_t is $Cov(X_t, X_t - k) = \gamma_k^*$ (i.e., it only depends on the lag), the autocovariance for Y_t is

$$Cov(W_t, W_{t-k})$$
= $Cov(X_t, X_{t-k}) - Cov(X_t, X_{t-k-12})$
 $- Cov(X_{t-12}, X_{t-k}) + Cov(X_{t-12}, X_{t-k-12})$
= $\gamma_{t-(t-k)}^* - \gamma_{t-(t-k-12)}^*$
 $- \gamma_{t-12-(t-k)}^* + \gamma_{t-12-(t-k-12)}^*$
= $\gamma_k^* - \gamma_{k+12}^* - \gamma_{k-12}^* + \gamma_k^*$
= $2\gamma_k^* - \gamma_{k+12}^* - \gamma_{k-12}^*$.

This still depends only on k (i.e., stationary) is more complicated than in part (a). Note that the ACF is:

$$\rho_{k} = \frac{\gamma_{k}}{\gamma_{0}} = \frac{2\gamma_{k}^{*} - \gamma_{k+12}^{*} - \gamma_{k-12}^{*}}{\gamma_{0}}$$

$$= \frac{2\gamma_{k}^{*} - \gamma_{k+12}^{*} - \gamma_{k-12}^{*}}{2\gamma_{0}^{*} - \gamma_{12}^{*} - \gamma_{k-12}^{*}}$$

$$= \frac{2\gamma_{k}^{*} - \gamma_{k+12}^{*} - \gamma_{k-12}^{*}}{2\gamma_{0}^{*} - 2\gamma_{12}^{*}}$$
(Since $\gamma_{k} = \gamma_{-k}$ by definition)

c) Here we have $Y_t = \beta_0 + \beta_1 t + \beta_2 t^2 + s_t + e_t$ and therefore:

$$W_t = Y_t - Y_{t-12}$$

$$= \beta_0 + \beta_1 t + \beta_2 t^2 + s_t + e_t$$

$$- \beta_0 - \beta_1 (t - 12) - \beta_2 (t - 12)^2 - s_{t-12} - e_{t-12}$$

$$= 12\beta_1 + 24\beta_2 t - 144\beta_2 + e_t - e_{t-12}$$

$$= 12\beta_1 - 144\beta_2 + 24\beta_2 t + e_t - e_{t-12}$$

This is clearly non-stationary since

$$E(W_t) = 12\beta_1 - 144\beta_2 + 24\beta_2 t.$$

However, the autocovariance is still the same as in part (a) since $12\beta_1 - 144\beta_2 + 24\beta_2 t$ is non-random and, hence, disappears from covariance.

d) Here we have

$$\nabla W_t = 12\beta_1 - 144\beta_2 + 24\beta_2 t + e_t - e_{t-12}$$
$$-12\beta_1 + 144\beta_2 - 24\beta_2 (t-1) - e_{t-1} + e_{t-13}$$
$$= 24\beta_2 + \underbrace{e_t - e_{t-12} - e_{t-1} + e_{t-13}}_{\text{sum of stationary series}}.$$

This is clearly stationary as it is the sum of stationary series plus a constant.

This question shows that applying $\nabla \nabla_{12}$, i.e., seasonal differencing and then differencing, eliminates quadratic and seasonal trend.