



UNIVERSITY of LIMERICK

O L L S C O I L L U I M N I G H

# Time Series Analysis

## MS 4218

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## Outline

### Models for stationary time series

- ▶ General linear processes
- ▶ Moving Average (MA) processes
  - ▶ MA(1)
  - ▶ MA(2)
  - ▶ MA( $q$ )

## General linear processes

$Y_t$  is an observed time series, and  $e_t$  is unobserved white noise.

A weighted linear combination of  $e'_t$ 's where the sum of the squares of the weights is finite yields a general linear process.

$$Y_t = (\psi_0 = 1)e_t + \psi_1 e_{t-1} + \psi_2 e_{t-2} + \dots$$

If  $\psi_j = \phi^j$  and  $|\phi| < 1$ , we have a convergent series:

$$Y_t = e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots$$

## Expectation and Variance

$$E(Y_t) = E(e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots) = 0$$

$$\text{Var}(Y_t) = \text{Var}(e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots)$$

$$= \text{Var}(e_t) + \text{Var}(\phi e_{t-1}) + \text{Var}(\phi^2 e_{t-2}) + \dots$$

$$= \sigma_e^2(1 + \phi^2 + \phi^4 + \dots); \quad (a = 1, r = \frac{u_2}{u_1} = \phi^2)$$

$$= \sigma_e^2 \left( S_\infty = \frac{a}{1-r} = \frac{1}{1-\phi^2} \right)$$

$$= \frac{\sigma_e^2}{1-\phi^2}.$$

## Auto-covariance

$$\begin{aligned} \text{Cov}(Y_t, Y_{t-1}) &= \text{Cov}(\mathbf{e}_t + \phi \mathbf{e}_{t-1} + \phi^2 \mathbf{e}_{t-2} + \dots, \\ &\quad \mathbf{e}_{t-1} + \phi \mathbf{e}_{t-2} + \phi^2 \mathbf{e}_{t-3} + \dots) \end{aligned}$$

Matching only the like terms:

$$\begin{aligned} \text{Cov}(Y_t, Y_{t-1}) &= \text{Cov}(\phi \mathbf{e}_{t-1}, \mathbf{e}_{t-1}) + \text{Cov}(\phi^2 \mathbf{e}_{t-2}, \phi \mathbf{e}_{t-2}) + \dots \\ &= \phi \sigma_e^2 + \phi^3 \sigma_e^2 + \phi^5 \sigma_e^2 + \dots \\ &= \phi \sigma_e^2 (1 + \phi^2 + \phi^4 + \dots); \quad (a = 1, r = \phi^2) \\ &= \frac{\phi \sigma_e^2}{1 - \phi^2}. \end{aligned}$$

## Auto-correlation

$$\begin{aligned}
 \text{Corr}(Y_t, Y_{t-1}) &= \rho_1 = \frac{\gamma_1}{\gamma_0} \\
 &= \frac{\phi \sigma_e^2}{1 - \phi^2} \div \frac{\sigma_e^2}{1 - \phi^2} \\
 &= \phi^1.
 \end{aligned}$$

$$\text{Corr}(Y_t, Y_{t-k}) = \phi^k.$$

A constant mean and an autocovariance that depends only on the lag means the process is stationary.

## Auto-covariance again

$$\begin{aligned}
 \text{Cov}(Y_t, Y_{t-1}) &= \text{Cov}(\psi_0 \mathbf{e}_t + \psi_1 \mathbf{e}_{t-1} + \psi_2 \mathbf{e}_{t-2} + \dots, \\
 &\quad \psi_0 \mathbf{e}_{t-1} + \psi_1 \mathbf{e}_{t-2} + \psi_2 \mathbf{e}_{t-3} + \dots) \\
 &= E(\psi_0 \psi_1 \mathbf{e}_{t-1}^2 + \psi_1 \psi_2 \mathbf{e}_{t-2}^2 + \psi_2 \psi_3 \mathbf{e}_{t-3}^2 + \dots) \\
 &= \sigma_e^2 (\psi_0 \psi_1 + \psi_1 \psi_2 + \psi_2 \psi_3 + \dots) \\
 &= \sigma_e^2 \sum_{i=0}^{\infty} (\psi_i \psi_{i+1}) = \gamma_1. \\
 \text{Cov}(Y_t, Y_{t-k}) &= \sigma_e^2 \sum_{i=0}^{\infty} (\psi_i \psi_{i+k}) = \gamma_k.
 \end{aligned}$$

## Moving average (MA) processes

A general linear process with a finite number (order =  $q$ ) of non-zero  $\psi$  weights.

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2} - \dots - \theta_q e_{t-q}.$$

This MA process is written as MA( $q$ ).

An MA process of order 1 has 1  $\theta$ .

An MA(2) has a  $\theta_1$  and  $\theta_2$ .



## MA(1) process

A first order MA process is given by:

$$Y_t = e_t - \theta e_{t-1}.$$

$$E(Y_t) = E(e_t - \theta e_{t-1})$$

$$= E(e_t) - \theta E(e_{t-1})$$

$$= 0.$$

## Variance for MA(1) process

$$\begin{aligned}
 \text{Var}(Y_t) &= \text{Var}(e_t - \theta e_{t-1}) \\
 &= \text{Var}(e_t) + (-\theta)^2 \text{Var}(e_{t-1}) \\
 &= \sigma_e^2 + \theta^2 \sigma_e^2 \\
 &= \sigma_e^2 (1 + \theta^2) \\
 &= \gamma_0.
 \end{aligned}$$

## Auto-covariance for MA(1) process

$$\text{Cov}(Y_t, Y_{t-1}) = \text{Cov}(e_t - \theta e_{t-1}, e_{t-1} - \theta e_{t-2}).$$

$$\begin{aligned} &= \text{Cov}(e_t, e_{t-1}) + \text{Cov}(e_t, -\theta e_{t-2}) \\ &\quad + \text{Cov}(-\theta e_{t-1}, e_{t-1}) + \text{Cov}(-\theta e_{t-1}, -\theta e_{t-2}) \end{aligned}$$

$$= -\theta \text{Cov}(e_{t-1}, e_{t-1})$$

$$= -\theta \sigma_e^2.$$

## Auto-covariance for MA(1) process

$$\begin{aligned}
 \text{Cov}(Y_t, Y_{t-2}) &= \text{Cov}(e_t - \theta e_{t-1}, e_{t-2} - \theta e_{t-3}) \\
 &= \text{Cov}(e_t, e_{t-2}) + \text{Cov}(e_t, -\theta e_{t-3}) \\
 &\quad + \text{Cov}(-\theta e_{t-1}, e_{t-2}) + \text{Cov}(-\theta e_{t-1}, -\theta e_{t-3}) \\
 &= 0.
 \end{aligned}$$

## Acvf MA(1)

$$\text{Cov}(Y_t, Y_{t-k}) = 0 \quad \forall \quad k > 1.$$

$$\gamma_k = \begin{cases} \sigma_e^2(1 + \theta^2) & \text{if } k = 0. \\ -\theta\sigma_e^2 & \text{if } k = 1. \\ 0 & \text{if } k > 1. \end{cases}$$

## Acf MA(1)

$$\rho_k = \frac{\gamma_k}{\gamma_0}$$

$$= \begin{cases} 1 & \text{if } k = 0. \\ -\frac{\theta\sigma_e^2}{\sigma_e^2(1+\theta^2)} = -\frac{\theta}{1+\theta^2} & \text{if } k = 1. \\ 0 & \text{if } k > 1. \end{cases}$$

## Maximum correlation for MA(1)

$$\rho_1 = -\frac{\theta}{1 + \theta^2}.$$

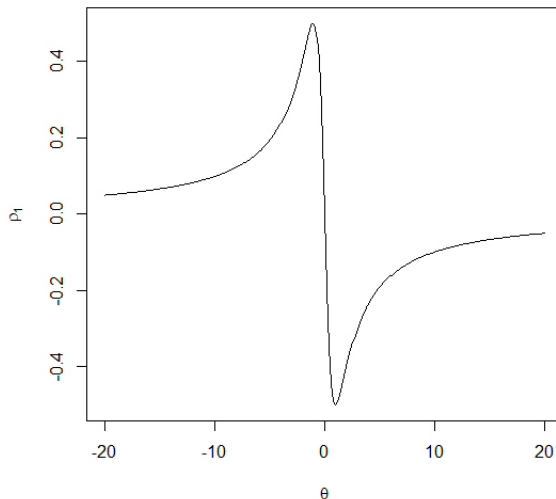
$$\frac{d}{d\theta}\rho_1 = \frac{\theta^2 - 1}{(1 + \theta^2)^2}$$

$$= 0 \text{ when } \theta = \pm 1.$$

$$\theta = -1 \Rightarrow \rho = \frac{1}{2}.$$

$$\theta = +1 \Rightarrow \rho = -\frac{1}{2}.$$

## Lag 1 auto-correlation of an MA(1) process for different $\theta$ values





## Non-uniqueness (and later invertibility)

2 MA(1) processes with parameters a.  $\theta \neq 0$  and b.  $\frac{1}{\theta}$  have same value of  $\rho_1$ .

$$a.\rho_1 = -\frac{\theta}{1 + \theta^2}.$$

$$b.\rho_1 = -\frac{1}{\theta} \div \left(1 + \frac{1}{\theta^2}\right)$$

$$= -\frac{1}{\theta} \times \frac{\theta^2}{1 + \theta^2}$$

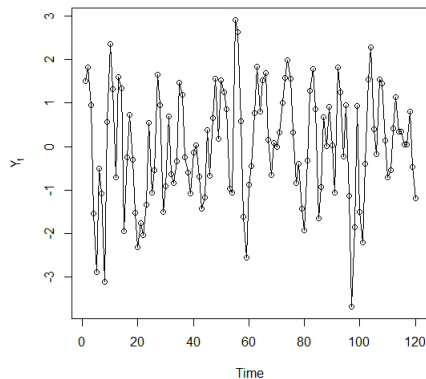
$$= -\frac{\theta}{1 + \theta^2}.$$

## ma1.2.s data in library(TSA)

$e_t$ 's are white noise and  $\theta = -0.9$ .

```
data(ma1.2.s)
ma1.2.s
plot(ma1.2.s,
      ylab=expression(Y[t]), type='o')
```

## Timeplot of ma1.2.s



Relatively smooth.

## Sample and theoretical correlation at lag 1 in ma1.2.s data

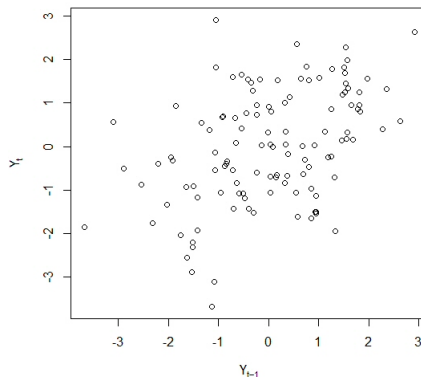
```
n<-length(ma1.2.s)
r1<-cor(ma1.2.s[-1],ma1.2.s[-n]); r1 = 0.428
```

$$\begin{aligned}\rho_1 &= -\frac{\theta}{1+\theta^2} \\ &= -\frac{-0.9}{1+(-0.9)^2} = 0.4972.\end{aligned}$$

```
plot(y=ma1.2.s,x=zlag(ma1.2.s),
     ylab=expression(Y[t]),
     xlab=expression(Y[t-1]),type='p')
```

## Plot of $Y_t$ v $Y_{t-1}$ for MA(1) with $\theta = -0.9$

$$\rho_1 = 0.4972; r_1 = 0.428.$$



Moderately strong upward trend.

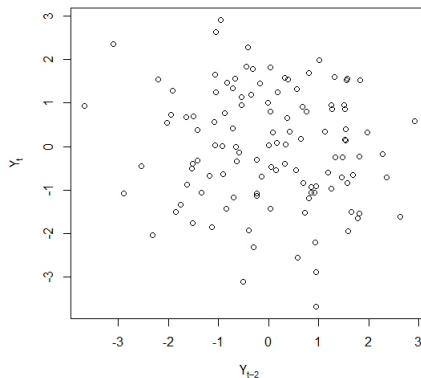
## Sample and theoretical correlation at lag 2 in ma1.2.s data

```
n<-length(ma1.2.s)
r2<-cor(ma1.2.s[-(1:2)],ma1.2.s[-((n-1):n)])
r2=-0.1152471
```

$$\rho_2 = 0.$$

```
plot(y=ma1.2.s,x=zlag(ma1.2.s,2),
     ylab=expression(Y[t]),
     xlab=expression(Y[t-2]),type='p')
```

## Plot of $Y_t$ v $Y_{t-2}$ for MA(1) with $\theta = -0.9$

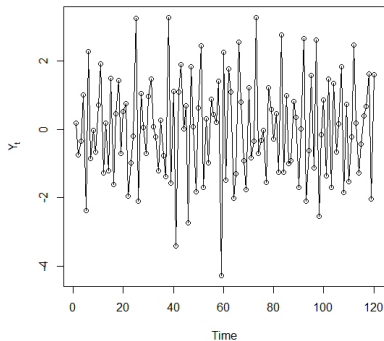


Zero auto-correlation at lag 2:  $\rho_2 = 0$ ;  $r_2 = -0.1152471$ .

## Time plot of MA(1) process: ma1.1.s data in library(TSA)

```
data (ma1.1.s)
```

```
plot (ma1.1.s, ylab=expression(Y[t]), type='o')
```



Jagged plot, and here,  $\theta = +0.9$



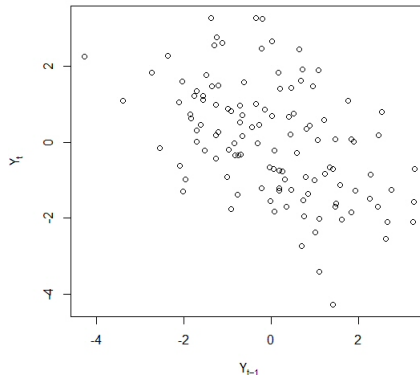
## Sample and theoretical correlation at lag 1 in *ma1.1.s* data

```
n<-length(ma1.1.s)
r1<-cor(ma1.2.s[-1],ma1.2.s[-n]); r1 = -0.476
```

$$\begin{aligned}\rho_1 &= -\frac{\theta}{1+\theta^2} \\ &= -\frac{0.9}{1+0.9^2} = -0.4972.\end{aligned}$$

```
plot(y=ma1.2.s,x=zlag(ma1.2.s),
     ylab=expression(Y[t]),
     xlab=expression(Y[t-1]),type='p')
```

## Plot of $Y_t$ v $Y_{t-1}$ for MA(1) with $\theta = +0.9$



$\rho_1 = -0.4972$  and  $r_2 = -0.476$ .  
Moderately strong downward trend.

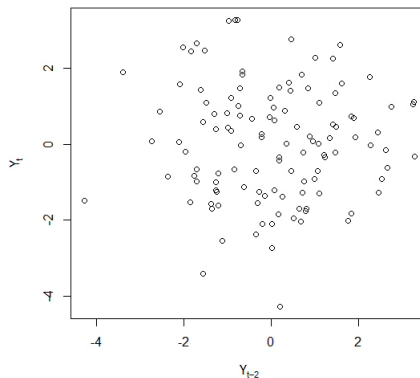
## Sample and theoretical correlation at lag 2 in ma1.1.s data

```
n<-length(ma1.1.s)
r2<-cor(ma1.1.s[-(1:2)],ma1.1.s[-((n-1):n)])
r2=-0.009831165.
```

$\rho_2 = 0.$

```
plot(y=ma1.1.s,x=zlag(ma1.1.s,2),
     ylab=expression(Y[t]),
     xlab=expression(Y[t-2]),type='p')
```

## Plot of $Y_t$ v $Y_{t-2}$ for MA(1) with $\theta = +0.9$



Zero auto-correlation at lag 2.  $r_2 = -0.009831165$ .

## MA(2) process

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2}.$$

$$E(Y_t) = E(e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2})$$

$$= E(e_t) - \theta_1 E(e_{t-1}) - \theta_2 E(e_{t-2})$$

$$= 0.$$

## MA(2) Variance

$$\begin{aligned}
 \gamma_0 &= \text{Var}(\mathbf{e}_t - \theta_1 \mathbf{e}_{t-1} - \theta_2 \mathbf{e}_{t-2}) \\
 &= \text{Var}(\mathbf{e}_t) + \text{Var}(-\theta_1 \mathbf{e}_{t-1}) + \text{Var}(-\theta_2 \mathbf{e}_{t-2}) \\
 &= \sigma_e^2 + \theta_1^2 \sigma_e^2 + \theta_2^2 \sigma_e^2 \\
 &= \sigma_e^2 (1 + \theta_1^2 + \theta_2^2).
 \end{aligned}$$

## MA(2) Auto-covariance

$$\begin{aligned}
 \gamma_1 &= \text{Cov}(e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2}, e_{t-1} - \theta_1 e_{t-2} - \theta_2 e_{t-3}) \\
 &= \text{Cov}(-\theta_1 e_{t-1}, e_{t-1}) + \text{Cov}(-\theta_2 e_{t-2}, -\theta_1 e_{t-2}) \\
 &= \sigma_e^2(-\theta_1 + \theta_1 \theta_2).
 \end{aligned}$$

## MA(2) Auto-covariance cont.

$$\gamma_2 = \text{Cov}(\mathbf{e}_t - \theta_1 \mathbf{e}_{t-1} - \theta_2 \mathbf{e}_{t-2}, \\ \mathbf{e}_{t-2} - \theta_1 \mathbf{e}_{t-3} - \theta_2 \mathbf{e}_{t-4})$$

$$= \text{Cov}(-\theta_2 \mathbf{e}_{t-2}, \mathbf{e}_{t-2})$$

$$= -\theta_2 \sigma_e^2.$$

$$\text{Cov}(Y_t, Y_{t-k}) = 0 \quad \forall \quad k > 2.$$



## MA(2) Auto-covariance summary

$$\gamma_k = \begin{cases} (1 + \theta_1^2 + \theta_2^2)\sigma_e^2 & \text{if } k = 0. \\ (-\theta_1 + \theta_1\theta_2)\sigma_e^2 & \text{if } k = 1. \\ -\theta_2\sigma_e^2 & \text{if } k = 2. \\ 0 & \text{if } k > 2. \end{cases}$$

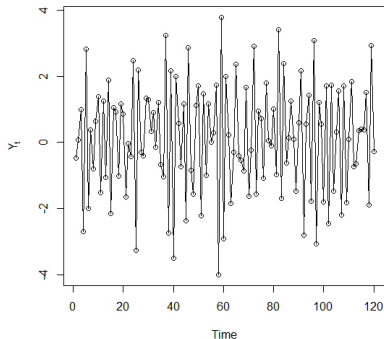
## Acf MA(2)

$$\rho_k = \frac{\gamma_k}{\gamma_0}$$

$$= \begin{cases} 1 & \text{if } k = 0. \\ \frac{-\theta_1 + \theta_1\theta_2}{1 + \theta_1^2 + \theta_2^2} & \text{if } k = 1. \\ \frac{-\theta_2}{1 + \theta_1^2 + \theta_2^2} & \text{if } k = 2. \\ 0 & \text{if } k > 2. \end{cases}$$

## Time plot of MA(2) process: data ma2.s with $\theta_1 = 1$ & $\theta_2 = -0.6$

```
data (ma2.s)
plot (ma2.s, ylab=expression(Y[t]), type='o')
```



It oscillates across midline each time.

## Sample and theoretical correlation at lag 1 in ma2.s data

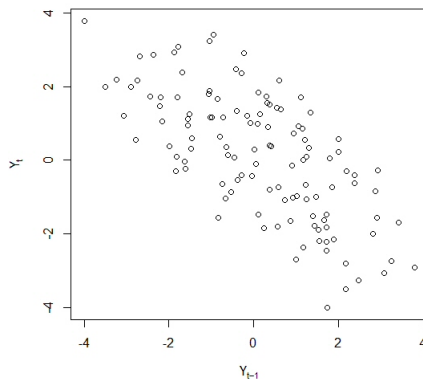
```
n<-length(ma2.s)
r1<-cor(ma2.s[-1],ma2.s[-n])
r1 = -0.684
```

$$\begin{aligned}\rho_1 &= \frac{-\theta_1 + \theta_1\theta_2}{1 + \theta_1^2 + \theta_2^2} \\ \rho_1 &= \frac{-1 + (1)(-0.6)}{1 + 1^2 + (-0.6)^2} \\ &= -0.678.\end{aligned}$$

```
plot(y=ma2.s,x=zlag(ma2.s),
     ylab=expression(Y[t]),
     xlab=expression(Y[t-1]),type='p')
```

## Plot of $Y_t$ v $Y_{t-1}$ for MA(2)

$$\rho_1 = -0.678; r_1 = -0.684.$$



## Sample and theoretical correlation at lag 2 in ma2.s data

```
n<-length (ma2.s)
r2<-cor (ma2.s[-(1:2)] , ma2.s[-((n-1):n)])
r2 = 0.251
```

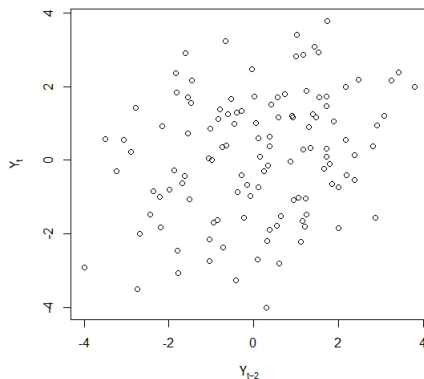
$$\rho_2 = \frac{-\theta_2}{1 + \theta_1^2 + \theta_2^2} = \frac{-(-0.6)}{1 + 1^2 + (-0.6)^2}$$

$$\rho_2 = 0.254.$$

```
plot (y=ma2.s, x=zlag (ma2.s, 2),
      ylab=expression (Y[t]),
      xlab=expression (Y[t-2]), type='p')
```

## Plot of $Y_t$ v $Y_{t-2}$ for MA(2)

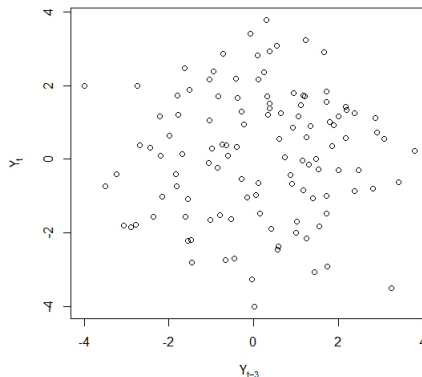
$$\rho_2 = 0.254; r_2 = 0.251.$$



## Plot of $Y_t$ v $Y_{t-3}$ for MA(2)

```
r3<-cor (ma2.s [ - (1:3) ] , ma2.s [ - ( (n-2) :n) ] )
r_3=0.0391.
```

$$\rho_3 = 0.$$





## General MA( $q$ ) process: Acf 0 after lag $q$

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2} - \cdots - \theta_q e_{t-q}.$$

$$\gamma_0 = (1 + \theta_1^2 + \theta_2^2 + \cdots + \theta_q^2) \sigma_e^2.$$

$$\rho_k = \begin{cases} \frac{-\theta_k + \theta_1 \theta_{k+1} + \cdots + \theta_{q-k} \theta_q}{1 + \theta_1^2 + \theta_2^2 + \cdots + \theta_q^2} & \text{if } k = 1, 2, \dots, q \\ 0 & \text{if } k > q. \end{cases}$$

## Next

### Models for stationary time series cont.

- ▶ Auto-Regressive (AR) processes
  - ▶ AR(1)
  - ▶ AR(2)
  - ▶ AR( $p$ )
- ▶ Mixed ARMA processes
- ▶ Invertibility