

# Transfer Report

Kevin O'Brien

September 21, 2015

# 1 Model Formulation and Formal Testing

? formulates a model for un-replicated observations for a method comparison study as a mixed model.

$$\begin{aligned} Y_{ij} &= \mu_j + S_i + \epsilon_{ij} \quad i = 1, 2 \dots n \quad j = 1, 2 \\ S &\sim N(0, \sigma_s^2) \quad \epsilon_{ij} \sim N(0, \sigma_j^2) \end{aligned} \quad (1)$$

As with all mixed models, the variance of each observation is the sum of all the associated variance components.

$$\begin{aligned} var(Y_{ij}) &= \sigma_s^2 + \sigma_j^2 \\ cov(Y_{i1}, Y_{i2}) &= \sigma_s^2 \end{aligned} \quad (2)$$

? offers maximum likelihood estimators, commonly known as Grubbs estimators, for the various variance components:

$$\begin{aligned} \hat{\sigma}_s^2 &= \sum \frac{(x_i - \bar{x})(y_i - \bar{y})}{n-1} = Sxy \\ \hat{\sigma}_1^2 &= \sum \frac{(x_i - \bar{x})^2}{n-1} = S^2x - Sxy \\ \hat{\sigma}_2^2 &= \sum \frac{(y_i - \bar{y})^2}{n-1} = S^2y - Sxy \end{aligned} \quad (3)$$

The standard error of these variance estimates are:

$$\begin{aligned} var(\sigma_1^2) &= \frac{2\sigma_1^4}{n-1} + \frac{\sigma_S^2\sigma_1^2 + \sigma_S^2\sigma_2^2 + \sigma_1^2\sigma_2^2}{n-1} \\ var(\sigma_2^2) &= \frac{2\sigma_2^4}{n-1} + \frac{\sigma_S^2\sigma_1^2 + \sigma_S^2\sigma_2^2 + \sigma_1^2\sigma_2^2}{n-1} \end{aligned} \quad (4)$$

?presents confidence intervals for the relative precisions of the measurement methods,  $\Delta_j = \sigma_S^2/\sigma_j^2$  (where  $j = 1, 2$ ), as well as the variances  $\sigma_S^2, \sigma_1^2$  and  $\sigma_2^2$ .

$$\Delta_1 > \frac{C_{xy} - t(|A|/n - 2))^{\frac{1}{2}}}{C_x - C_{xy} + t(|A|/n - 2))^{\frac{1}{2}}} \quad (5)$$

where

$$\begin{aligned}
C_x &= (n-1)S_x^2 \\
C_{xy} &= (n-1)S_{xy} \\
C_y &= (n-1)S_y^2 \\
A &= C_x \times C_y - (C_{xy})^2
\end{aligned}$$

$t$  is the  $100(1 - \alpha/2)\%$  quantile of Student's  $t$  distribution with  $n - 2$  degrees of freedom.  $\Delta_2$  can be found by changing  $C_y$  for  $C_x$ . A lower confidence limit can be found by calculating the square root. This inequality may also be used for hypothesis testing.

For the interval estimates for the variance components, ? presents three relations that hold simultaneously with probability  $1 - 2\alpha$  where  $2\alpha = 0.01$  or  $0.05$ .

$$\begin{aligned}
|\sigma^2 - C_{xy}K| &\leq M(C_x C_y)^{\frac{1}{2}} \\
|\sigma_1^2 - (C_x - C_{xy})K| &\leq M(C_x(C_x + C_y - 2C_{xy}))^{\frac{1}{2}} \\
|\sigma_2^2 - (C_y - C_{xy})K| &\leq M(C_y(C_x + C_y - 2C_{xy}))^{\frac{1}{2}}
\end{aligned} \tag{6}$$

The case-wise differences and means are  $D_i = Y_{i1} - Y_{i2}$  and  $A_i = (Y_{i1} + Y_{i2})/2$  respectively. Both  $D_i$  and  $A_i$  follow a bivariate normal distribution with  $E(D_i) = \mu_D = \mu_1 - \mu_2$  and  $E(A_i) = \mu_A = (\mu_1 + \mu_2)/2$ . The variance matrix  $\Sigma$  is

$$\Sigma = \begin{bmatrix} \sigma_1^2 + \sigma_2^2 & \frac{1}{2}(\sigma_1^2 - \sigma_2^2) \\ \frac{1}{2}(\sigma_1^2 - \sigma_2^2) & \sigma_S^2 + \frac{1}{4}(\sigma_1^2 + \sigma_2^2) \end{bmatrix} \tag{7}$$

? demonstrates how the Grubbs estimators for the error variances can be calculated

using the difference values, providing a worked example on a data set.

$$\begin{aligned}\hat{\sigma}_1^2 &= \sum (y_{i1} - \bar{y}_1)(D_i - \bar{D}) \\ \hat{\sigma}_2^2 &= \sum (y_{i2} - \bar{y}_2)(D_i - \bar{D})\end{aligned}\tag{8}$$

## 1.1 Morgan Pitman

The test of the hypothesis that the variance of both methods are equal is based on the correlation value  $\rho_{D,A}$  which is evaluated as follows;

$$\rho(D, A) = \frac{\sigma_1^2 - \sigma_2^2}{\sqrt{(\sigma_1^2 + \sigma_2^2)(4\sigma_S^2 + \sigma_1^2 + \sigma_2^2)}}\tag{9}$$

The correlation constant takes the value zero if, and only if, the two variances are equal. Therefore a test of the hypothesis  $H : \sigma_1^2 = \sigma_2^2$  is equivalent to a test of the hypothesis  $H : \rho(D, A) = 0$ . This corresponds to the well-known  $t$  test for a correlation coefficient with  $n - 2$  degrees of freedom.

? describes the Morgan-Pitman test as identical to the test of the slope equal to zero in the regression of  $Y_{i1}$  on  $Y_{i2}$ , adding that this result can be shown using straightforward algebra.

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As with all mixed models, the variance of each observation is the sum of all the associated variance components.

$$\begin{aligned} \text{var}(Y_{ij}) &= \sigma_s^2 + \sigma_j^2 \\ \text{cov}(Y_{i1}, Y_{i2}) &= \sigma_s^2 \end{aligned} \quad (12)$$

? offers maximum likelihood estimators, commonly known as Grubbs estimators, for the various variance components:

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The standard error of these variance estimates are:

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?presents confidence intervals for the relative precisions of the measurement methods,  $\Delta_j = \sigma_S^2/\sigma_j^2$  (where  $j = 1, 2$ ), as well as the variances  $\sigma_S^2, \sigma_1^2$  and  $\sigma_2^2$ .

$$\Delta_1 > \frac{C_{xy} - t(|A|/n - 2))^{\frac{1}{2}}}{C_x - C_{xy} + t(|A|/n - 2))^{\frac{1}{2}}} \quad (15)$$

where

$$\begin{aligned} C_x &= (n-1)S_x^2 \\ C_{xy} &= (n-1)S_{xy} \\ C_y &= (n-1)S_y^2 \\ A &= C_x \times C_y - (C_{xy})^2 \end{aligned}$$

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The case-wise differences and means are  $D_i = Y_{i1} - Y_{i2}$  and  $A_i = (Y_{i1} + Y_{i2})/2$  respectively. Both  $D_i$  and  $A_i$  follow a bivariate normal distribution with  $E(D_i) = \mu_D = \mu_1 - \mu_2$  and  $E(A_i) = \mu_A = (\mu_1 + \mu_2)/2$ . The variance matrix  $\Sigma$  is

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$$\begin{aligned} \hat{\sigma}_1^2 &= \sum (y_{i1} - \bar{y}_1)(D_i - \bar{D}) \\ \hat{\sigma}_2^2 &= \sum (y_{i2} - \bar{y}_2)(D_i - \bar{D}) \end{aligned} \quad (18)$$

## 2.1 Paired sample T-test

? discusses the use of the well known paired sample  $t$  test to test for inter-method bias;  $H : \mu_D = 0$ . The test statistic is distributed a  $t$  random variable with  $n - 1$  degrees of freedom and is calculated as follows;

$$t^* = \bar{D} / \frac{S_D}{\sqrt{n}} \quad (19)$$

where  $\bar{D}$  and  $S_D$  is the average of the differences of the  $n$  observations.

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$$t^* = \bar{D} / \frac{S_D}{\sqrt{n}} \quad (20)$$

where  $\bar{D}$  and  $S_D$  is the average of the differences of the  $n$  observations.

### 3 Thompson 1963

? defines  $\Delta_j$  to be a measure of the relative precision of the measurement methods, with  $\Delta_j = \sigma_S^2/\sigma_j^2$  (where  $j = 1, 2$ ). Confidence intervals for  $\Delta_j$  are also presented.

$$\Delta_1 > \frac{C_{xy} - t(\frac{|A|}{n-1})^{\frac{1}{2}}}{C_x - C_{xy} + t(\frac{|A|}{n-1})^{\frac{1}{2}}}, \quad (21)$$

where

$$C_x = (n-1)S_x^2,$$

$$C_{xy} = (n-1)S_{xy},$$

$$C_y = (n-1)S_y^2,$$

$$A = C_x \times C_y - (C_{xy})^2.$$

The value  $t$  is the  $100(1 - \alpha/2)\%$  quantile of Student's  $t$  distribution with  $n - 2$  degrees of freedom. The ratio  $\Delta_2$  can be found by interchanging  $C_y$  and  $C_x$ . A lower confidence limit can be found by calculating the square root. The inequality in equation 1.10 may also be used for hypothesis testing.

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? contains tables for  $K$  and  $M$ .



? offers a formal simultaneous hypothesis test for the mean and variance of two paired data sets. Using simple linear regression of the differences of each pair against the sums, a line is fitted to the model, with estimates for intercept and slope ( $\hat{\beta}_0$  and  $\hat{\beta}_1$ ). The null hypothesis of this test is that the mean ( $\mu$ ) and variance ( $\sigma^2$ ) of both data sets are equal if the slope and intercept estimates are equal to zero (i.e  $\sigma_1^2 = \sigma_2^2$  and  $\mu_1 = \mu_2$  if and only if  $\beta_0 = \beta_1 = 0$  )

A test statistic is then calculated from the regression analysis of variance values (?) and is distributed as ‘F’ random variable. The degrees of freedom thereof are  $\nu_1 = 2$  and  $\nu_2 = n - 2$  (where n is the number of pairs). The critical value is chosen for  $\alpha\%$  significance with those same degrees of freedom. ? amends this methodology for use in method comparison studies, using the averages of the pairs, as opposed to the sums, and their differences. This approach can facilitate simultaneous usage of test with the Bland-Altman methodology. Bartko’s test statistic take the form:

$$F.test = \frac{(\Sigma d^2) - SSReg}{2MSReg} \quad (22)$$

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Averages	1	0.04	0.04	0.74	0.4097
Residuals	10	0.60	0.06		

Table 1: Regression ANOVA of case-wise differences and averages for Grubbs Data

For the Grubbs data,  $\Sigma d^2 = 5.09$ ,  $SSReg = 0.60$  and  $MSreg = 0.06$  Therefore the test statistic is 37.42, with a critical value of 4.10. Hence the means and variance of the Fotobalk and Counter chronometers are assumed to be simultaneously equal.

Importantly, this methodology determines whether there is both inter-method bias and precision present, or alternatively if there is neither present. It has previously been

demonstrated that there is a inter-method bias present, but as this procedure does not allow for separate testing, no conclusion can be drawn on the comparative precision of both methods.

### 3.1 Formal Testing

The Bland Altman plot is a simple tool for inspection of the data, but in itself it offers no formal testing procedure in this regard. To this end, the approach proposed by ? is a formal test on the Pearson correlation coefficient of casewise differences and means ( $\rho_{AD}$ ). According to the authors, this test is equivalent to a well established tests for equality of variances, known as the ‘Pitman Morgan Test’ (??).

For the Grubbs data, the correlation coefficient estimate ( $r_{AD}$ ) is 0.2625, with a 95% confidence interval of (-0.366, 0.726) estimated by Fishers ‘r to z’ transformation (?). The null hypothesis ( $\rho_{AD} = 0$ ) would fail to be rejected. Consequently the null hypothesis of equal variances of each method would also fail to be rejected.

There has no been no further mention of this particular test in the subsequent article published by Bland and Altman, although ? refers to Spearmans’ rank correlation coefficient.

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$$F.test = \frac{(\Sigma d^2) - SSReg}{2MSReg} \quad (24)$$

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Averages	1	0.04	0.04	0.74	0.4097
Residuals	10	0.60	0.06		

Table 2: Regression ANOVA of case-wise differences and averages for Grubbs Data

For the Grubbs data,  $\Sigma d^2 = 5.09$ ,  $SSReg = 0.60$  and  $MSreg = 0.06$  Therefore the test statistic is 37.42, with a critical value of 4.10. Hence the means and variance of the Fotobalk and Counter chronometers are assumed to be simultaneously equal.

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## 5 Bartko's Regression and Ellipse

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$$F.test = \frac{(\Sigma D^2) - SSReg}{2MSReg} \quad (25)$$

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
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Table 3: Regression ANOVA of case-wise differences and averages for Grubbs Data

For the Grubbs data,  $\Sigma D^2 = 5.09$ ,  $SSReg = 0.60$  and  $MSreg = 0.06$  Therefore the test statistic is 37.42, with a critical value of 4.102821 (calculate using r code  $qf(0.95, 2, 10)$ ). Hence the means and variance of the Fotobalk and Counter chronometers are assumed to be simultaneously equal.

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## 5.1 Bartko's Ellipse

? offers a graphical complement to the Bland-Altman plot, in the form of a bivariate confidence ellipse. ? provides the relevant calculations.



GrubbsBartko.jpeg

Figure 1: Bartko's Ellipse For Grubbs Data

## 5.2 Bartko's Bradley-Blackwood Test

This is a regression based approach that performs a simultaneous test for the equivalence of means and variances of the respective methods.

$$D = (X_1 - X_2) \tag{26}$$



$$M = (X_1 + X_2)/2 \quad (27)$$

The Bradley Blackwood Procedure fits D on M as follows:

$$D = \beta_0 + \beta_1 M \quad (28)$$

Both beta values, the intercept and slope, are derived from the respective means and standard deviations of their respective data sets.

We determine if the respective means and variances are equal if both beta values are simultaneously equal to zero. The Test is conducted using an F test, calculated from the results of a regression of D on M.

We have identified this approach to be examined to see if it can be used as a foundation for a test perform a test on means and variances individually.

Russell et al have suggested this method be used in conjunction with a paired t-test , with estimates of slope and intercept.

subsection-t-test

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$$D = (X_1 - X_2) \quad (29)$$

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The Bradley Blackwood Procedure fits D on M as follows:

$$D = \beta_0 + \beta_1 M \quad (31)$$

Both beta values, the intercept and slope, are derived from the respective means and standard deviations of their respective data sets.

We determine if the respective means and variances are equal if both beta values are simultaneously equal to zero. The Test is conducted using an F test, calculated from the results of a regression of D on M.

Russell et al have suggested this method be used in conjunction with a paired t-test, with estimates of slope and intercept. Bradley and Blackwood have developed a regression based approach assessing the agreement.

The Bradley Blackwood test is a simultaneous test for bias and precision. They propose a regression approach which fits D on M, where D is the difference and average of a pair of results.

## 5.4 Pitman & Morgan Test

This test assess the equality of population variances. Pitman's test tests for zero correlation between the sums and products.

Correlation between differences and means is a test statistics for the null hypothesis of equal variances given bivariate normality.

## 6 Bartko's Regression and Ellipse

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sums, a line is fitted to the model, with estimates for intercept and slope ( $\beta_0$  and  $\beta_1$ ). The null hypothesis of this test is that the mean ( $\mu$ ) and variance ( $\sigma^2$ ) of both data sets are equal if the slope and intercept estimates are equal to zero (i.e  $\sigma_1^2 = \sigma_2^2$  and  $\mu_1 = \mu_2$  if and only if  $\beta_0 = \beta_1 = 0$  )

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Figure 2: Bartko's Ellipse For Grubbs Data