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## 1 LME Introduction

Henderson [1984a] shows that the solution to the mixed model equations

$$\begin{pmatrix} X'R^{-1}X & X'R^{-1}Z \\ Z'R^{-1}X & Z'R^{-1}Z + G^{-1} \end{pmatrix} \begin{pmatrix} \beta \\ \mathbf{b} \end{pmatrix} = \begin{pmatrix} X'R^{-1}y \\ Z'R^{-1}y \end{pmatrix} \quad (1)$$

gives a solution for which the estimate  $\mathbf{b}$  of the fixed effects is equal to that obtained by generalized least squares.

$$\hat{\beta} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y} \quad (2)$$

? showed that the BLUP of  $\mathbf{b}$  is

$$\hat{\mathbf{b}} = (DZ'\mathbf{V})(\mathbf{y} - \mathbf{X}\hat{\beta}) \quad (3)$$

### 1.0.1 Henderson's equations

Because of the dimensionality of  $\mathbf{V}$  (i.e.  $n \times n$ ) computing the inverse of  $\mathbf{V}$  can be difficult. As a way around the this problem ????? offered a more simpler approach of jointly estimating  $\hat{\beta}$  and  $\hat{\mathbf{b}}$ . ? made the (ad-hoc) distributional assumptions  $y|\mathbf{b} \sim N(\mathbf{X}\beta + Z\mathbf{b}, \Sigma)$  and  $\mathbf{b} \sim N(0, D)$ , and proceeded to maximize the joint density of  $y$  and  $\mathbf{b}$

$$\left| \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix} \right|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} \begin{pmatrix} \mathbf{b} \\ y - \mathbf{X}\beta - Z\mathbf{b} \end{pmatrix}' \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix}^{-1} \begin{pmatrix} \mathbf{b} \\ y - \mathbf{X}\beta - Z\mathbf{b} \end{pmatrix} \right\}, \quad (4)$$

with respect to  $\beta$  and  $b$ , which ultimately requires minimizing the criterion

$$(y - X\beta - Zb)' \Sigma^{-1} (y - X\beta - Zb) + b' D^{-1} b. \quad (5)$$

This leads to the mixed model equations

$$\begin{pmatrix} X' \Sigma^{-1} X & X' \Sigma^{-1} Z \\ Z' \Sigma^{-1} X & X' \Sigma^{-1} X + D^{-1} \end{pmatrix} \begin{pmatrix} \beta \\ b \end{pmatrix} = \begin{pmatrix} X' \Sigma^{-1} y \\ Z' \Sigma^{-1} y \end{pmatrix}. \quad (6)$$

Using these equations, obtaining the estimates requires the inversion of a matrix of dimension  $p + q \times p + q$ , considerably smaller in size than  $V$ . ? shows that these mixed model equations do not depend on normality and that  $\hat{\beta}$  and  $\hat{b}$  are the BLUE and BLUP under general conditions, provided  $D$  and  $\Sigma$  are known.

? points out that although ? initially referred to the estimates  $\hat{\beta}$  and  $\hat{b}$  from (12) as “joint maximum likelihood estimates”, ? later advised that these estimates should not be referred to as “maximum likelihood” as the function being maximized in (11) is a joint density rather than a likelihood function. ? remarks that it is clear that Henderson used joint estimation for computational purposes, without recognizing the theoretical implications.

### Henderson’s equations

? made the (ad-hoc) distributional assumptions  $y|b \sim N(X\beta + Zb, \Sigma)$  and  $b \sim N(0, D)$ , and proceeded to maximize the joint density of  $y$  and  $b$

$$\left| \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix} \right|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} \begin{pmatrix} b \\ y - X\beta - Zb \end{pmatrix}' \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix}^{-1} \begin{pmatrix} b \\ y - X\beta - Zb \end{pmatrix} \right\}, \quad (7)$$

with respect to  $\beta$  and  $b$ , which ultimately requires minimizing the criterion

$$(y - X\beta - Zb)' \Sigma^{-1} (y - X\beta - Zb) + b' D^{-1} b. \quad (8)$$

This leads to the solutions

$$\begin{pmatrix} X' \Sigma^{-1} X & X' \Sigma^{-1} Z \\ Z' \Sigma^{-1} X & X' \Sigma^{-1} X + D^{-1} \end{pmatrix} \begin{pmatrix} \beta \\ b \end{pmatrix} = \begin{pmatrix} X' \Sigma^{-1} y \\ Z' \Sigma^{-1} y \end{pmatrix}. \quad (9)$$

? points out that although ? initially referred to the estimates  $\hat{\beta}$  and  $\hat{b}$  from (12) as “joint maximum likelihood estimates” ? later advised that these estimates should not be referred to as “maximum likelihood” as the function being maximized in (11) is a joint density rather than a likelihood function.

### 1.0.2 Henderson’s equations

Because of the dimensionality of  $V$  (i.e.  $n \times n$ ) computing the inverse of  $V$  can be difficult. As a way around this problem ????? offered a more simpler approach of jointly estimating  $\hat{\beta}$  and  $\hat{b}$ . ? made the (ad-hoc) distributional assumptions  $y|b \sim N(X\beta + Zb, \Sigma)$  and  $b \sim N(0, D)$ , and proceeded to maximize the joint density of  $y$  and  $b$

$$\left| \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix} \right|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} \begin{pmatrix} b \\ y - X\beta - Zb \end{pmatrix}' \begin{pmatrix} D & 0 \\ 0 & \Sigma \end{pmatrix}^{-1} \begin{pmatrix} b \\ y - X\beta - Zb \end{pmatrix} \right\}, \quad (10)$$

with respect to  $\beta$  and  $b$ , which ultimately requires minimizing the criterion

$$(y - X\beta - Zb)' \Sigma^{-1} (y - X\beta - Zb) + b' D^{-1} b. \quad (11)$$

This leads to the mixed model equations

$$\begin{pmatrix} X' \Sigma^{-1} X & X' \Sigma^{-1} Z \\ Z' \Sigma^{-1} X & Z' \Sigma^{-1} Z + D^{-1} \end{pmatrix} \begin{pmatrix} \beta \\ b \end{pmatrix} = \begin{pmatrix} X' \Sigma^{-1} y \\ Z' \Sigma^{-1} y \end{pmatrix}. \quad (12)$$

Using these equations, obtaining the estimates requires the inversion of a matrix of dimension  $p + q \times p + q$ , considerably smaller in size than  $V$ . ? shows that these mixed model equations do not depend on normality and that  $\hat{\beta}$  and  $\hat{b}$  are the BLUE and BLUP under general conditions, provided  $D$  and  $\Sigma$  are known.

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