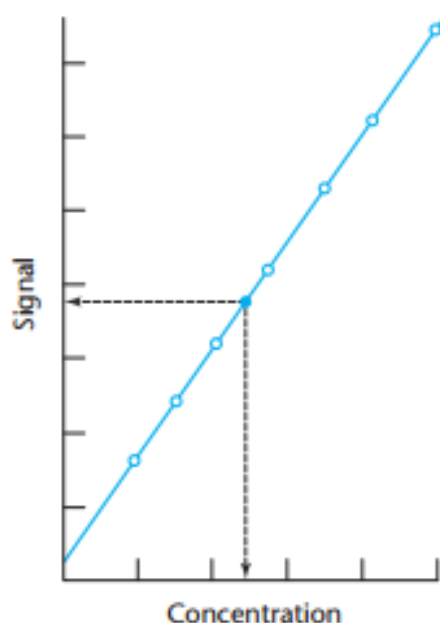


## MA4605 2016 – Analytical Chemistry - Lecture 6A

### 5.2 Calibration graphs in instrumental analysis

The usual procedure is as follows, the analyst takes a series of materials (normally at least three or four, and possibly several more) in which the concentration of the analyte is known. These calibration standards are measured in the analytical instrument under the same conditions as those subsequently used for the test (i.e. the 'unknown') materials.

Once the calibration graph has been established the analyte concentration in any test material can be obtained by interpolation.



**Figure 5.1** Calibration procedure in instrumental analysis: ○ calibration points; ● test sample.

This general procedure raises several important statistical questions:

- 1 Is the calibration graph linear? If it is a curve, what is the form of the curve? (recall: Polynomial Regression)
- 2 Bearing in mind that each of the points on the calibration graph is subject to errors, what is the best straight line (or curve) through these points?
- 3 Assuming that the calibration plot is actually linear, what are the errors and confidence limits for the slope and the intercept of the line?
- 4 When the calibration plot is used for the analysis of a test material, what are the errors and confidence limits for the determined concentration?
- 5 What is the ***limit of detection*** of the method? That is, what is the least concentration of the analyte that can be detected with a predetermined level of 'confidence'?

Before tackling these questions in detail, we must consider a number of aspects of plotting calibration graphs.

Firstly, it is usually essential that the calibration standards cover the whole range of concentrations required in the subsequent analyses.

With the important exception of the '***Method of Standard Additions***' (next section ) concentrations of test materials are normally determined by interpolation and not by extrapolation.

## **Blank**

Secondly, it is crucially important to include the value for a 'blank' in the calibration curve.

**Important:** The blank contains ***no deliberately added*** analyte, but does contain the same solvent, reagents, etc., as the other test materials, and is subjected to exactly the same sequence of analytical procedures.

The instrument signal given by the blank will sometimes not be zero. This signal is subject to errors like all the other points on the calibration plot, so it is wrong in principle to subtract the blank value from the other standard values before plotting the calibration graph.

This is because when two quantities are subtracted, the error in the final result cannot also be obtained by simple subtraction. Subtracting the blank value from each of the other instrument signals before plotting the graph thus gives incorrect information on the errors in the calibration process.

It should be noted that the calibration curve is always plotted with the instrument signals on the vertical (y) axis and the standard concentrations on the horizontal (x) axis.

**Important:** This is because many procedures (such as Ordinary Least Squares) assume that all the errors are in ***the fitted values*** and that the standard concentrations (x-values) are error-free.

## **Limits of detection**

**Important:** One of the principal benefits of using instrumental methods of analysis is that they are capable of ***detecting and determining trace and ultra-trace quantities of analytes.***

These benefits have led to the appreciation of the importance of very low concentrations of many materials, for example in biological and environmental samples, and thus to the development of many further techniques in which lower limits of detection are a major criterion of successful application. It is therefore evident that statistical methods for assessing and comparing limits of detection are of importance.

***Important: In general terms, the limit of detection of an analyte may be described as that concentration which gives an instrument signal (y) significantly different from the 'blank' or 'background' signal.***

This description gives the analyst a good deal of freedom to decide the exact definition of the limit of detection, based on a suitable interpretation of the phrase 'significantly different'.

The Blank Signal is the expected analyte concentration expected to be found when replicates of a blank sample containing no analyte are tested.

*(The blank signal is equivalent to the regression intercept).*

### Formal Definition of Limits of Detection

There is still no full agreement between researchers, publishers, and professional and statutory bodies on this point. But there is an increasing trend to define the limit of detection as the analyte concentration giving a signal equal to the blank signal,  $y_B$ , plus three standard deviations of the blank,  $s_B$ :

(Recall: Confidence intervals for fitted values)

$$\text{Limit of detection} = y_B + 3s_B \quad (5.12)$$

(Remark: There are other definitions on Limits of Detection)

The Limit of Detection is the lowest analyte concentration likely to be reliably distinguished from the blank and at which detection is feasible.

The significance of this last definition is illustrated in more detail in Figure 5.7.

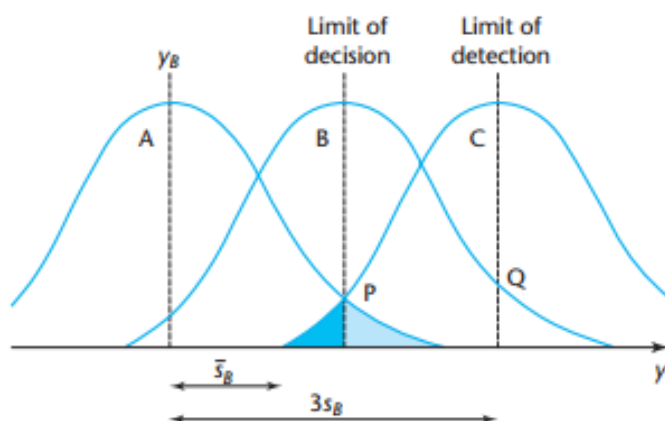


Figure 5.7 Definitions of the limit of decision and the limit of detection.

An analyst studying trace concentrations is confronted with two problems: it is important to avoid claiming the presence of the analyte when it is actually absent, but it is equally important to avoid reporting that the analyte is absent when it is in fact present.

(The situation is analogous to the occurrence of Type I and Type II errors in significance tests, i.e. False Positives etc)

(**IMPORTANT**) There may be occasions when an analyst is anxious to avoid at all costs the possibility of reporting the absence of the analyte when it is in fact present, but is relatively unworried about the opposite error. It is clear that whenever a limit of detection is cited in a paper or report, the definition used to obtain it must also be provided

### **Limits of Determination ( not examinable )**

Some attempts have been made to define a further limit, the 'limit of quantitation' (or 'limit of determination'), which is regarded as the lower limit for precise quantitative measurements, as opposed to qualitative detection.

A value of  $y_B + 10 s_B$ , has been suggested for this limit, but it is not very widely used.

### Example 5.3.1

Standard aqueous solutions of fluorescein are examined in a fluorescence spectrometer, and yield the following fluorescence intensities (in arbitrary units):

Fluorescence intensities:	2.1	5.0	9.0	12.6	17.3	21.0	24.7
Concentration, $\text{pg ml}^{-1}$	0	2	4	6	8	10	12

Determine the correlation coefficient,  $r$ .

### Example 5.7.1

Estimate the limit of detection for the fluorescein determination studied in the previous sections.

We use equation (5.12) with the values of  $y_B (= a)$  and  $s_B (= s_{y/x})$  previously calculated. The value of  $y$  at the limit of detection is found to be  $1.52 + 3 \times 0.4329$ , i.e. 2.82. Use of the regression equation then yields a detection limit of  $0.67 \text{ pg ml}^{-1}$ . Figure 5.8 summarizes all the calculations performed on the fluorescein determination data.

```
> summary(fit)
```

Call:

```
lm(formula = Int ~ Conc)
```

Residuals:

1	2	3	4	5	6	7
0.58214	-0.37857	-0.23929	-0.50000	0.33929	0.17857	0.01786

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	1.5179	0.2949	5.146	0.00363	**
Conc	1.9304	0.0409	47.197	8.07e-08	***

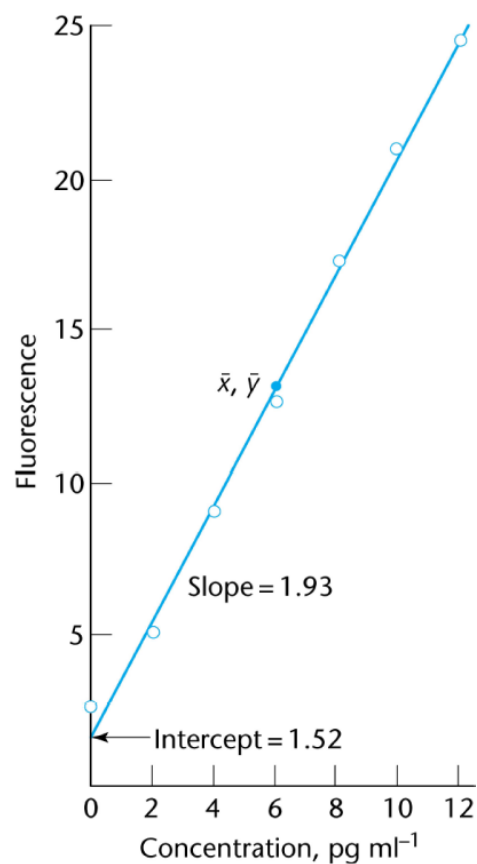
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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4328 on 5 degrees of freedom

Multiple R-squared: 0.9978, Adjusted R-squared: 0.9973

F-statistic: 2228 on 1 and 5 DF, p-value: 8.066e-08





## Limits of detections

- Set the blank signal, i.e. the concentration  $x = 0$ , in the fitted regression model.
- Compute from fitted regression the corresponding response, i.e.  $y = a$ .
- Add to  $a$  three times estimated standard error of  $y$ , i.e.  $3s_{y/x}$ :

$$a + 3s_{y/x}.$$

- Any measurement below the above value is questionable in representing a positive concentration.
- The corresponding value  $x_0$  is the lowest value of concentration that can be trusted as non-zero.

*(Remark: In this graphic, the intercept estimate is denoted “a” rather than “b<sub>0</sub>”)*

*(Also : We'll use  $S_b$  as the notation for Residual standard error for Blank Signal). Our notation is therefore :*

$Y_b + 3S_b$
--------------

**Important:** You are also to compute the value for  $X$  that yields  $Y_b + 3S_b$

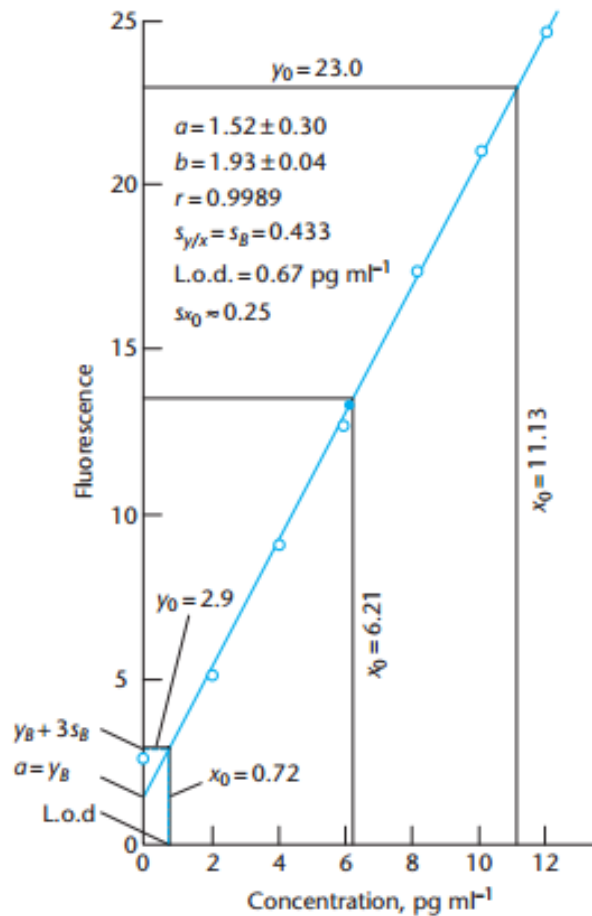


Figure 5.8 Summary of the calculations using the data in Example 5.3.1.

Intercept:

$$Y_b = b_0 = 1.52$$

Limit of detection (in terms of signal)

$$Y_b + 3S_b = 1.52 + (3 \times 0.433) = 2.819$$

What value of X (i.e. concentration) yields a signal of 2.819?

$$1.5179 + (1.9304 \times 0.674) = 2.819$$

Answer:  $Y = 2.819$  units and  $X = 0.674$  units

### **Remark: Censored Data**

A left censored value is one that is known only to be less than some value, e.g.  $< 5$  ppm.

A right censored value is one that is known only to be more than some value.

#### **11.2.1 Type I Left Singly Censored Data**

The benzene data presented in Table ?? illustrate type I left singly censored data with a single censoring level of 2 ppb. There are  $N = 36$  observations, with  $c = 33$  censored observations and  $n = 3$  uncensored observations. The trichloroethylene data presented in Table ?? have a single censoring level of 5 ppb, with  $N = 24$ ,  $c = 10$ , and  $n = 14$ . The Skagit data set (stored as Skagit.NH3\_N.df in ENVSTATS) contains 395 monthly measurements of ammonia nitrogen ( $\text{NH}_3\text{-N}$ ) concentrations (mg/l) in the Skagit River (Marblemount, Washington station) made from January 1978 through December 2010. Table 11.1 shows the 60 observations made between January 1978 and December 1982, ordered from smallest to largest, for which all censored observations were censored at 0.01 mg/l. Notice that there are samples with observed values of 0.01 mg/l, which will be treated differently from the 16 samples reported as  $< 0.01$  mg/l.

Table 11.1:  $\text{NH}_3\text{-N}$  concentrations, as mg/l, measured in the Skagit River, Washington State, between January 1978 and December 1982.

<b><math>\text{NH}_3\text{-N}</math> concentration (mg/l)</b>								
<0.01	<0.01	<0.01	<0.01	<0.01	<0.01	<0.01	<0.01	<0.01
<0.01	<0.01	<0.01	<0.01	<0.01	<0.01	<0.01	0.01	0.01
0.01	0.01	0.01	0.01	0.01	0.01	0.01	0.01	0.01
0.01	0.02	0.02	0.02	0.02	0.02	0.02	0.02	0.02
0.02	0.02	0.02	0.02	0.02	0.02	0.02	0.03	0.03
0.03	0.03	0.03	0.03	0.04	0.04	0.04	0.04	0.04
0.04	0.05	0.05	0.05	0.06	0.47			