

EPIDEMIOLOGY

Emergence and global spread of *Listeria monocytogenes* main clinical clonal complex

Alexandra Moura^{1,2,*}, Noémie Lefrancq^{3†}, Thierry Wirth^{4,5}, Alexandre Leclercq^{1,2}, Vitor Borges⁶, Brent Gilpin⁷, Timothy J. Dallman⁸, Joachim Frey⁹, Eelco Franz¹⁰, Eva M. Nielsen¹¹, Juno Thomas¹², Arthur Pightling¹³, Benjamin P. Howden^{14,15}, Cheryl L. Tarr¹⁶, Peter Gerner-Smidt¹⁶, Simon Cauchemez³, Henrik Salje^{3‡}, Sylvain Brisson^{17‡}, Marc Lecuit^{1,2,18*‡}, *Listeria CC1 Study Group*

The bacterial foodborne pathogen *Listeria monocytogenes* clonal complex 1 (*Lm*-CC1) is the most prevalent clonal group associated with human listeriosis and is strongly associated with cattle and dairy products. Here, we analyze 2021 isolates collected from 40 countries, covering *Lm*-CC1 first isolation to present days, to define its evolutionary history and population dynamics. We show that *Lm*-CC1 spread worldwide from North America following the Industrial Revolution through two waves of expansion, coinciding with the transatlantic livestock trade in the second half of the 19th century and the rapid growth of cattle farming and food industrialization in the 20th century. In sharp contrast to its global spread over the past century, transmission chains are now mostly local, with limited inter- and intra-country spread. This study provides an unprecedented insight into *L. monocytogenes* phylogeography and population dynamics and highlights the importance of genome analyses for a better control of pathogen transmission.

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INTRODUCTION

Listeria monocytogenes (*Lm*) is a foodborne bacterial zoonotic pathogen that can cause listeriosis, a severe infection with a high case fatality rate in immunocompromised individuals (1, 2). Molecular studies have shown the clonal population structure of *Lm* (3, 4) and the worldwide distribution of clonal complex 1 (*Lm*-CC1, initially called epidemic clone ECI) (5, 6), a cosmopolitan clonal group defined by multilocus sequence typing (MLST; fig. S1), which was first isolated from an Italian soldier with meningitis during the first world war (WWI) (7, 8). Notably, *Lm*-CC1 is the most prevalent clinical clonal complex in several countries (9–15) and actually corresponds to 20% of all of *Lm* clinical isolates deposited at the National Center for Biotechnology Information (NCBI) (fig. S2). *Lm*-CC1 belongs to *Lm* major lineage I and evolved from a subgroup of serotype 4b ancestry (fig. S1) (3, 4).

While there is no proven interhuman horizontal transmission of listeriosis, it was only in 1983 that the foodborne transmission of human listeriosis was formally established (16). Since then, *Lm*-CC1 has been reported in different food matrices, including dairy products (17, 18), which can be heavily contaminated and constitute a major source of human listeriosis (19, 20). Previous studies have also demonstrated the hypervirulence of *Lm*-CC1 (9), and its higher efficiency in gut colonization and fecal shedding, compared to hypovirulent *Lm* clones (17, 21, 22). Moreover, increasing evidence shows that bovines, which are frequent *Lm* asymptomatic carriers (23–28) and contribute to *Lm* enrichment in soils (24), are the main source of disease (29) and constitute a reservoir for *Lm*-CC1 (21, 22). In addition to *Lm* subclinical infections that may contaminate milk (22, 25), the long-term persistence of *Lm* in cattle manure-amended soils (30) also poses serious risks of transmission to fresh produce.

Understanding the global evolution of *Lm*-CC1, which is now spread over all continents (6), as well as its emergence and dissemination across different spatial levels is critical to understand *Lm* population dynamics and to develop better control strategies, particularly in countries with aging and/or immunosuppressed populations who are most at risk for severe infection. However, the complex movement of livestock and food products associated with asymptomatic intestinal colonization complicates traditional epidemiological investigations aiming at deciphering *Lm* epidemiology by linking isolates in space and time. Here, we took a population biology approach to fill this knowledge gap and conducted the largest genomic *Lm*-CC1 study to date, combining genomic and evolutionary approaches to decipher its evolutionary history and pattern of emergence and spread.

RESULTS

Lm-CC1 is composed of three sublineages of uneven prevalence

We analyzed 2021 genomes, including 1230 newly sequenced isolates, originating from 40 countries in six continents and diverse sources (Fig. 1A, dataset S1, and fig. S3), including those from human

¹Institut Pasteur, Université de Paris, Inserm U1117, Biology of Infection Unit, Paris, France. ²Institut Pasteur, National Reference Center and WHO Collaborating Center Listeria, 75015 Paris, France. ³Institut Pasteur, Université de Paris, Mathematical Modelling of Infectious Diseases Unit, CNRS UMR 2000, Paris, France. ⁴Institut Systématique Evolution Biodiversité (ISYEB), Muséum National d'Histoire Naturelle, CNRS, Sorbonne Université, Université des Antilles, EPHE, Paris, France. ⁵PSL University, EPHE, Paris, France. ⁶Department of Infectious Diseases, National Institute of Health Dr. Ricardo Jorge, Lisbon, Portugal. ⁷Christchurch Science Centre, Institute of Environmental Science and Research Limited, Christchurch, New Zealand. ⁸Public Health England, London, UK. ⁹Vetsuisse Faculty, University of Bern, Bern, Switzerland. ¹⁰Centre for Infectious Disease Control, National Institute for Public Health and the Environment (RIVM), Bilthoven, Netherlands. ¹¹Statens Serum Institut, Copenhagen, Denmark. ¹²Division of the National Health Laboratory Service, National Institute for Communicable Diseases, Johannesburg, South Africa. ¹³Biostatistics and Bioinformatics, Center for Food Safety and Applied Nutrition, U.S. Food and Drug Administration, College Park, MD, USA. ¹⁴Microbiological Diagnostic Unit Public Health Laboratory, Department of Microbiology and Immunology, The Doherty Institute for Infection and Immunity, University of Melbourne, Melbourne, Victoria, Australia. ¹⁵Infectious Diseases Department, Austin Health, Heidelberg, Victoria, Australia. ¹⁶Centers for Disease Control and Prevention, Atlanta, GA, USA. ¹⁷Institut Pasteur, Université de Paris, Biodiversity and Epidemiology of Bacterial Pathogens, Paris, France. ¹⁸Necker-Enfants Malades University Hospital, Division of Infectious Diseases and Tropical Medicine, APHP, Institut Imagine, Paris, France.
*Corresponding author. Email: amoura@pasteur.fr (A.M.); marc.lecuit@pasteur.fr (M.L.)
†Present address: Department of Genetics, University of Cambridge, Cambridge, UK.
‡These authors shared senior authorship.

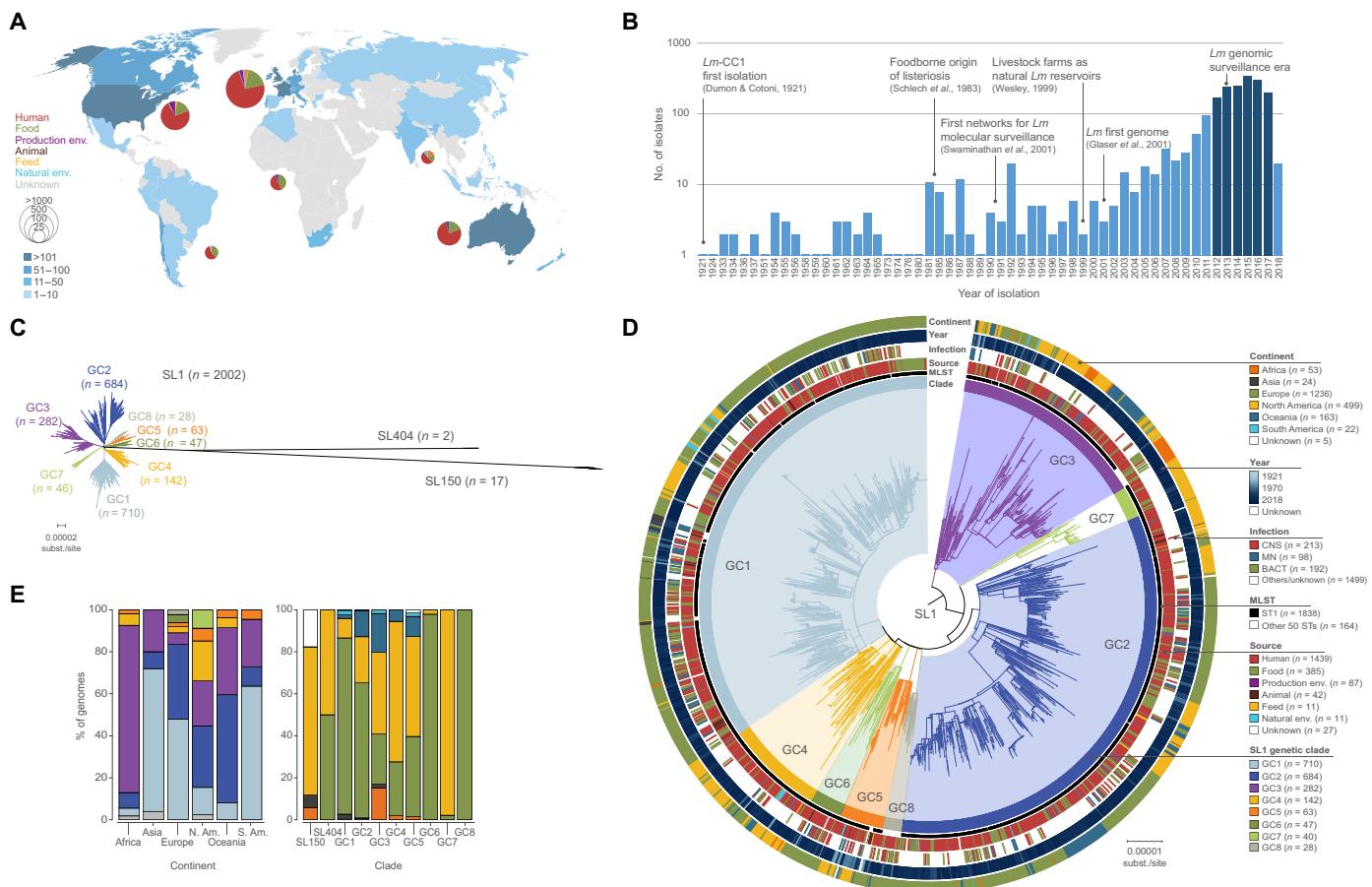


Fig. 1. Geographical and temporal distribution of the isolates used in this study ($N=2021$) and phylogenetic analyses. (A) Geographical distribution and source distribution. Sampled countries are colored in blue, with hue gradient according to the number of isolates. Pie charts are proportional to the number of isolates sampled in each continent and represent the repartition of sample source types, using the source color key indicated in (D). Eight isolates had unknown sampling location and are not shown in the map. (B) Temporal distribution of isolates collected in this study. Darker blue bars indicate the period for which exhaustive clinical sampling was obtained for seven countries spanning three continents (2012–2017; US, FR, UK, DK, NL, AU, and NZ). (C) Unrooted maximum likelihood phylogenetic tree of 2021 *Lm*-CC1 genomes. The tree was generated from analysis (GTR+F+G4 model, 1000 ultrafast bootstraps) of a 1.29-Mb recombination-purged core genome alignment. (D) Midpoint-rooted maximum likelihood phylogenetic tree of 2002 SL1 genomes based on a recombination-purged core genome alignment of 1.29 Mb. The four external rings indicate the world region, year, type of infection, and source type, respectively. The two inner rings indicate ST1 isolates and the eight SL1 GCs identified in this study, respectively. (E) Percentage of genomes by world region (left) and clade (right). Partitions are colored by world regions and clade, using the same color code as in (D).

patients ($n=1452$, 72%), food and food processing environments ($n=477$, 24%), animal and farm environments ($n=54$, 3%), natural environments ($n=11$, 0.5%), or unknown sources ($n=27$, 1%). We covered a time span of 98 years, from the first *Lm*-CC1 isolation to the present time (1921–2018), and included all contemporary clinical isolates collected between 2012 and mid-2017 within the surveillance framework of seven countries over three continents (Fig. 1, A and B).

Lm-CC1 genome sizes ranged from 2.77 to 3.25 million base pairs (Mbp), with an average number of 2879 ± 77 coding sequences and G+C content of 37.7 to 38.3% (fig. S4). On the basis of MLST (4), 58 sequence types (STs) could be distinguished, with ST1 representing 91% ($n=1838$) of isolates. On the basis of core genome MLST (cgMLST) (31), we identified within *Lm*-CC1 867 cgMLST types, 92% of which were country specific (fig. S5). Rarefaction analysis based on cgMLST resampling did not reach an asymptote (fig. S5), indicating that despite the high number of sequences

obtained in this study, a substantial amount of *Lm*-CC1 diversity remains undetected.

To better understand the phylogenetic diversity of *Lm*-CC1, we built maximum likelihood phylogenies and identified three sublineages (SL1, SL404, and SL150, named on the basis of their smallest ST number). These sublineages have highly uneven frequency (Fig. 1, C and D, and fig. S6), with SL1 ($n=2002$, isolated worldwide) representing 99.1% of the isolates, while 0.1% are SL404 ($n=2$, found in Europe and North America) and 0.8% represent SL150 ($n=17$, found in North America, Africa, and Asia). Within SL1, we further identified eight distinct genetic clades (GCs), which we named GC1 to GC8 by decreasing prevalence (Fig. 1 and fig. S6). The average genetic distance was 1166 ± 134 whole-genome single-nucleotide polymorphisms (wgSNPs) (and 478 ± 20 cgMLST alleles) between *Lm*-CC1 sublineages and 76 ± 16 wgSNPs (and 40 ± 9 cgMLST alleles) within SL1 clades (table S1 and fig. S7). The finding that SL1 is by far the major sublineage in *Lm*-CC1 is consistent

with its increased virulence and/or transmission [as seen at *Lm* species level for hypervirulent clones (32)] or may indicate that SL404 and SL150 are restricted to some yet unknown ecological niches. Within SL1, all different GCs were well represented, with strong spatial structure: GC1 is the most prevalent clade in Europe (48%, 593 of 1237), Asia (68%, 17 of 25), and South America (64%, 14 of 22); GC2 is the most prevalent clade in North America (29%, 150 of 512) and Oceania (52%, 84 of 163), while GC3 is the most prevalent clade in Africa (80%, 43 of 54) (Fig. 1E and fig. S3).

The *Lm*-CC1 pangenome is diverse

Analysis of *Lm*-CC1 pangenome identified 10,789 orthologous coding sequences (BlastP identity cutoff of $\geq 95\%$), 2649 of which (92% of the average isolate genome content) were present in at least 95% of isolates (core genome) (Fig. 2). The accessory genome included 8140 gene families, of which 2844 (35%) were unique to one isolate, and was enriched in transcription, replication/repair, and cell wall functions, as well as in gene families of unknown function (Fig. 2). Plasmids were present in 6% (120 of 2021) of isolates and were more prevalent in GC7 (83%; Fig. 2). The origins of most accessory genes (64%) and plasmids (89%) were attributable to the Listeriaceae family and other Firmicutes taxa (Fig. 2F). Intact prophages were present in 62% isolates (1263 of 2021) and were distributed across the breadth of *Lm*-CC1 phylogeny, except in SL404 (Fig. 2). In contrast to *Listeria* pathogenic islands LIPI-1 (33) and LIPI-3 (34), which were present in all isolates, the *Listeria* genomic island LGI2 (35), previously identified in *Lm*-CC1 isolates encoding resistance to cadmium and arsenic, was present in 14% (277 of 2021) isolates and only in GC3 (80%; 225 of 283), GC5 (60%, 38 of 63), and SL150 (82%, 14 of 17; Fig. 2). Sublineage-specific genes were detected (Fig. 2C, tables S2 and S3), and pangenome-wide association analyses identified 24 genes that are associated with a clinical origin (table S4). The impact of these traits on isolates' differential ecology or virulence remains to be studied, yet the presence of human isolates in all sublineages and clades shows that pathogenic isolates are not restricted to a specific *Lm*-CC1 clade.

Emergence and worldwide spread of *Lm*-CC1 main sublineage (SL1) occurred in the last 200 years

To understand *Lm*-CC1 evolution and spread, we performed temporal and phylogeographic analyses on the full dated dataset (1972 *Lm*-CC1 genomes), as well as on a subset of 200 genomes representative of *Lm*-CC1 genetic, temporal, and geographic diversity (see Materials and Methods; figs. S8 to S10). To control for the overrepresentation of recent isolates and geographic locations, sensitivity analyses of the evolutionary rate estimations were also performed on normalized subsets (see Materials and Methods). We estimate a core genome substitution rate of 1.95×10^{-7} substitutions per site per year [95% confidence interval (CI), 1.75×10^{-7} to 2.15×10^{-7} ; fig. S8], consistent with previous findings (31). We estimate that *Lm*-CC1 originated about 1800 years ago (date, 197 CE; 95% CI, 860 BCE to 1045 CE; Fig. 3B) and infer that its last common ancestor evolved in North America (fig. S10 and table S5), long before European colonization and the introduction of cattle in the Americas at the end of the 15th century (36). Although the low number of genomes available for Asia, Africa, and South America could bias this estimation, the estimated origin was also supported by the measures of population variability, which significantly showed higher genetic diversity within North America as compared to other world

regions ($P < 10^{-10}$, pairwise Wilcox test with Bonferroni correction for multiple comparisons; fig. S7 and table S1), and by the basal position of North American *Lm*-CC1 isolates in the phylogeny (Fig. 3B and fig. S10). The primary natural reservoir of *Lm*-CC1 and the events that led to its emergence remain unknown.

Demographic analyses performed using the Bayesian Skyline Plot method (37) (Fig. 3A) show that *Lm*-CC1 effective population size experienced two waves of expansion: the first in the late 1880s and the second in the 1930s, coinciding with the first and second ages of globalization, respectively. Tajima's and Fu and Li's D statistics (38, 39) also supported a recent *Lm*-CC1 population expansion and SL1 emergence ($D < 0$; table S1). SL1 emerged approximately 160 years ago (date, 1857; 95% CI, 1821 to 1888), thus closely following the start of the Industrial Revolution (Fig. 4). The first SL1 introductions into Europe occurred around 1868 (GC6/GC8 ancestor; 95% CI, 1827 to 1890), 1871 (GC3/GC7 ancestor; 95% CI, 1838 to 1905), and 1889 (GC2; 95% CI, 1852 to 1909), concomitant with the 1870 North Atlantic Meat trade agreement (40). Under this agreement, surplus cattle in North America were shipped to Europe, which had experienced severe livestock shortages due to widespread disease outbreaks (contagious bovine pleuropneumonia and foot and mouth disease), leading to an unprecedented man-made 1000-fold increase in cattle movement from North America to Europe (41). Within the same period, intracontinental diversification also took place, likely driven by cattle movements across North America and railway expansion in North America and Europe. The first SL1 introductions that occurred in Oceania (1903, GC2) followed the "Great Drought" of 1895–1903, which severely affected livestock (42).

In the following decades and after WWI, multiple CC1 introductions continued from North America into Europe (GC1, GC4, GC5, and GC8), Oceania (GC1 and GC4), and Asia (GC3) and from Europe to Africa (GC3) (Fig. 4, A and B), although the location of the emergence of GC2, GC4, and GC5 clades is uncertain (table S5). The rate of intercontinental bacterial movement declined after 1930s (Fig. 4C), concomitant with the protectionist trade policies that followed the "Great Depression", which led to a sharp reduction of livestock exports from the United States during the first half of the 20th century (43). A second wave of SL1 expansion occurred after this period, likely driven by a new increase in intercontinental movements favored by the industrialization of food production and globalization of the food and cattle trade (Fig. 3A and fig. S11). Other important human pathogens that have a zoonotic reservoir such as *Escherichia coli* O157:H7 (44) and *Campylobacter jejuni* ST-61 (45) have been estimated to have most recent common ancestors (MRCA) at similar times and to have also undergone population expansions in the context of animal trade or intensive cattle farming, respectively.

A stabilization of *Lm*-CC1 population is observed after 1984 (Fig. 3A), coincident with major advances in infectious disease prevention in dairy cattle (46) and with the relative decrease of the dairy cattle population in Western countries, in particular Europe (fig. S11). It also coincides with the time when human listeriosis foodborne origin was formally proven (16), which led to the implementation of surveillance programs in North America and Europe (47–49), in particular in the dairy sector following cheese- and milk-related *Lm*-CC1 outbreaks (50). Whether these findings can be observed in other dairy-associated *Lm* clonal complexes, such as CC6 (lineage I) or CC37 and CC101 (lineage II) (17, 18), will deserve future studies.

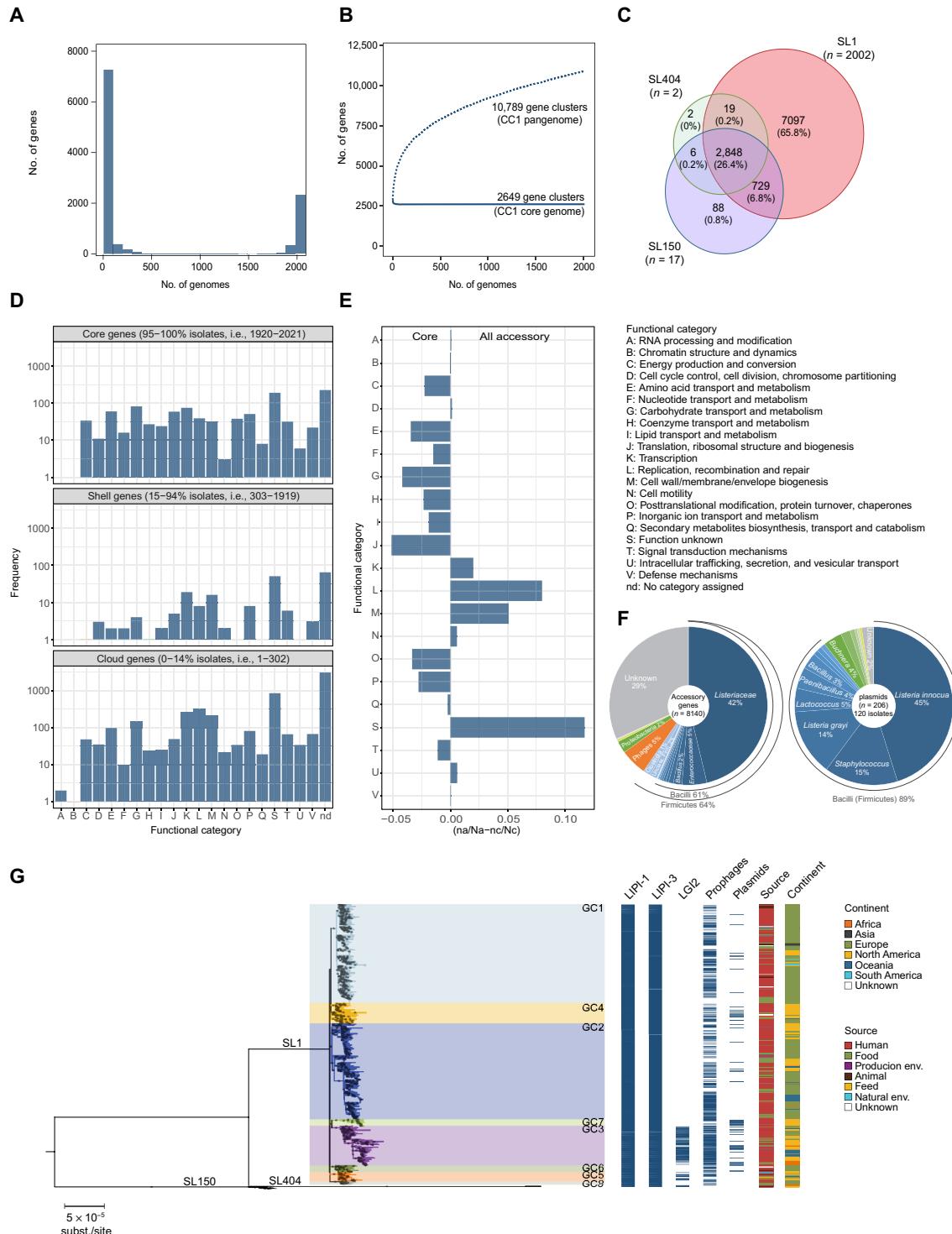


Fig. 2. *Lm*-CC1 pangenome analysis. (A) Frequency of sampled gene families. (B) Pan- and core-gene families sampled. (C) Venn diagram showing the number of gene families present in at least one sublineage member. (D) Distribution of the functional categories of the clusters of orthologous group (COG) category in core versus accessory genome, calculated as the difference between the ratio of each category (*n*) and the total number of hits (*N*) among each gene pool set, as in $(n_{\text{accessory}}/N_{\text{accessory}} - n_{\text{core}}/N_{\text{core}})$. (E) Functional category distribution. (F) Taxonomic classification of accessory genes and plasmids, based on eggNOG-mapper v.2 and MOB-suite v.2, respectively. (G) Distribution of *Listeria* genomic islands, prophages, and plasmids across *Lm*-CC1 phylogeny. The midpoint-rooted maximum likelihood phylogenetic tree (GTR+F+G4 model, 1000 ultra-fast bootstraps) was inferred from the 1.29-Mb recombination-purged core genome alignment of 2021 *Lm*-CC1 genomes.



Recent SL1 transmission chains are mostly local

To further analyze more recent strain transmission dynamics, we compared the genetic diversity of SL1 isolates from 2010 to 2018 ($n = 1266$) across different spatial scales. To avoid oversampling isolates from

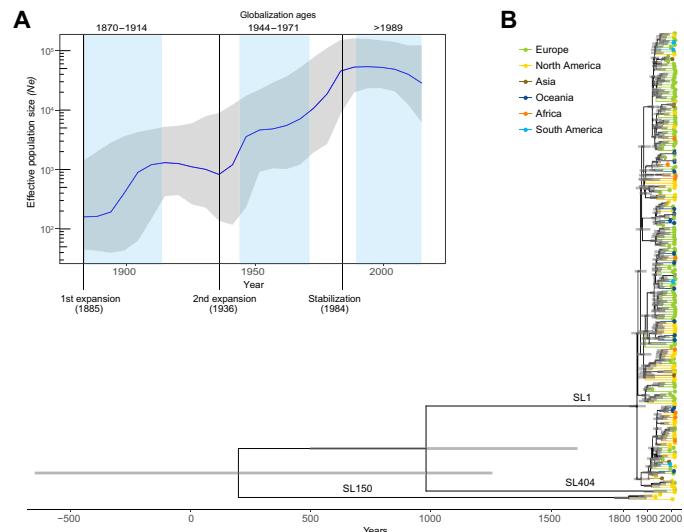


Fig. 3. Bayesian temporal and demographic analyses on a representative 200 isolate dataset of *Lm*-CC1. (A) Bayesian skyline plot (BSP) with the estimation of *Lm*-CC1 effective population size (N_e). The y axis refers to the predicted number of individuals (log scale), and the x axis refers to the time scale (in years). The median population size is marked in blue, with its 95% high posterior density (HPD) in gray. Blue vertical panels delimitate the three globalization ages (1870–1914, 1944–1971, and 1989–present). (B) Bayesian time-calibrated tree. Nodes represent the estimated mean divergence times, and gray bars represent the 95% HPD CIs of node age. Scale indicates time (in years). Terminal branches and tips are colored by continents, as indicated in the key panel.

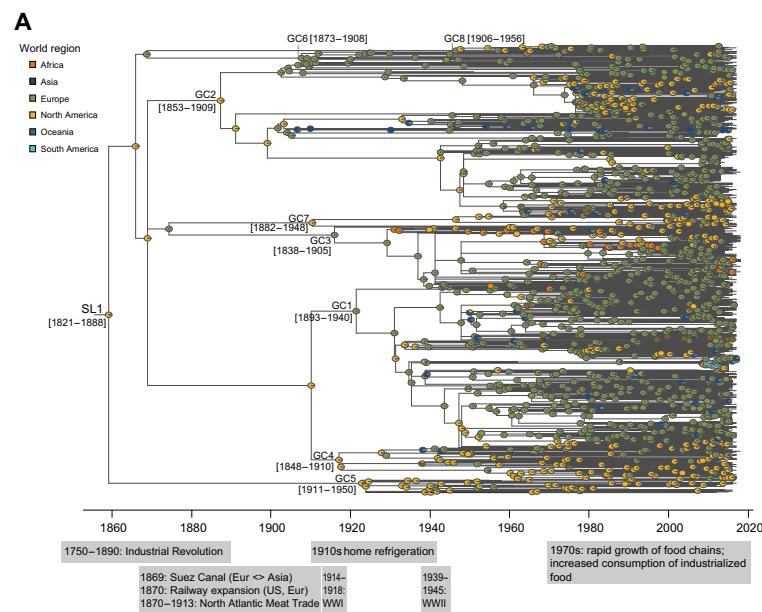


Fig. 4. Phylogeography of sublineage SL1. (A) Time-calibrated phylogeny based on the 1956 SL1 genomes (full view of *Lm*-CC1 in fig. S10). Pies at the nodes represent the probability of ancestral geographical locations, estimated using PastML using the MPPA method with an F81-like model. (B) Inferred spread of SL1 populations across continents. The first introductions of each phylogroup are represented by arrows from their estimated world region origin. Gray labels and arrows denote clades with uncertain ancestral origin among datasets. (C) Proportion of intercontinental transitions per 10-year bins, normalized by the total number of phylogenetic branches per bin.

outbreak investigations, we excluded all nonclinical isolates from confirmed outbreaks ($n = 91$ isolates from 19 outbreaks). We find that pairs of isolates present within the same 2-year period and the same country are 18.7 times (95% CI, 4.7 to 190.7) more likely to have their MRCA within the past 5 years than pairs of isolates coming from other intracontinental countries >1000 km apart (Fig. 5A). Furthermore, we observe no difference in the probability of having a recent MRCA in isolates coming from nearby intracontinental countries (<1000 km) than from further apart. Isolates coming from different continents are about 100 times less likely to have an MRCA within the past 5 years (0.2; 95% CI, 0.01 to 2.9) than isolates from the same countries (18.7; 95% CI, 4.7 to 190.7) (Fig. 5A). This strong local spatial structure persists for very long time periods, with complete mixing of isolates within a continent appearing only after 50 years (Fig. 5A). At a finer spatial scale, available for France (“départements”, subregional administrative division in France; fig. S12), a strong local spatial structure is also evident, with the proportion of genetically close pairs of clinical cases being higher between isolates coming from the same French department (4.4%; 95% CI, 1 to 10.6%) than between isolates coming from different departments (0.2%; 95% CI, 0.04 to 0.5%), with no effect of distance between them (Fig. 5B). As expected, in densely urban areas with no farming, such as the city of Paris, clinical strains are significantly less likely to share a recent MRCA than in rural areas or other departments [0.0% (95% CI, 0.0 to 4.4%) versus 3.9% (95% CI, 1.0 to 9.5%)] (Fig. 5C). This result is consistent with urban infections being driven by unrelated *Lm* introductions originating from across the country. Spatial dependence between French isolates persists for 20 years (fig. S13), with on average 20 (1/0.05) different sources of human infection present at any one time per department (Fig. 5B).

DISCUSSION

Understanding pathogen evolutionary history is essential to better interpret the population dynamics and biodiversity of microbial

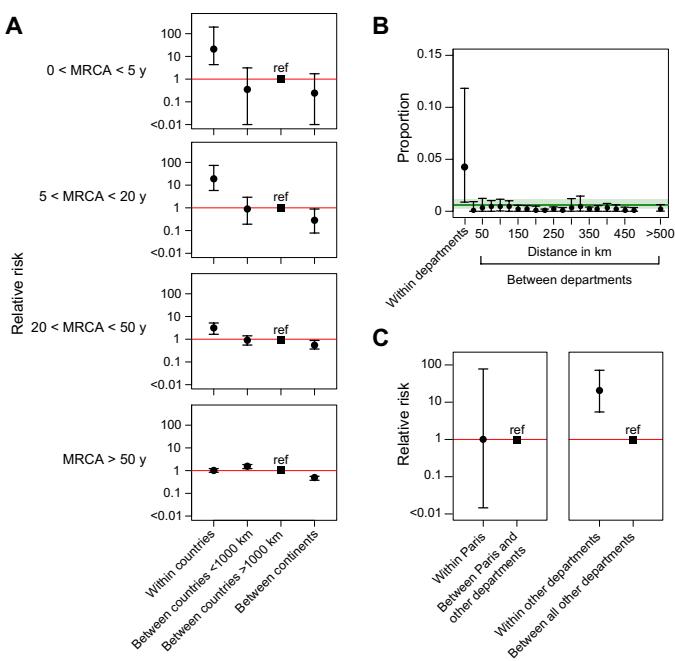


Fig. 5. Transmission dynamics of sublineage SL1: Evidence of local expansion after global spread. (A) Each point summarizes the relative risk that a pair of isolates has an MRCA within a defined time frame and between different spatial scales (within the same country, within the same continent, or within different continents), relative to the risk that a pair of isolates from countries separated by >1000 km has an MRCA in the same range (set as the reference value, “ref”). Error bars represent the 95% CIs, based on 100 bootstrap time-calibrated trees. (B) Proportion of pairs of isolates within the same country (France) sharing an MRCA of 5 or less years in function of the spatial distance within and between administrative departments (shown in the map). The green line indicates the mean proportion of genetically close strains regardless of the geographical location. (C) Relative risk for a pair of isolates to share an MRCA of 5 or less years when (left) both are coming from Paris compared to when coming from another department ($P = 0.43$) and (right) when coming from the same department in France, except Paris, compared to when coming from different departments ($P < 0.001$, see Materials and Methods for details).

infectious agents, and for effective disease surveillance. Here, we have shown that *Lm*-CC1 has spread worldwide following the Industrial Revolution through two waves of expansion coinciding with the transatlantic livestock trade in the second half of the 19th century and the rapid growth of cattle farming and food industrialization in the 20th century, respectively. The emergence of *Lm*-CC1 main sublineage SL1 circa 1857 (95% CI, 1821 to 1888) is concomitant with the emergence of other important widespread human pathogens that have a cattle reservoir, such as *E. coli* O157:H7 (1890; 95% CI, 1845 to 1925) (44) and *C. jejuni* ST-61 (1859; 95% CI, 1692 to 1943) (45), which have also undergone population expansions facilitated by intensive inter- and intracontinental animal trades (44) and livestock production (45). On the basis of our current dataset, North America is the most likely origin of *Lm*-CC1, yet this may be challenged when more genomes representing early diverged branches from undersampled regions are available.

Our results also show that, by the time *Lm* was first recognized as a human foodborne pathogen in 1983, *Lm*-CC1 had spread worldwide since long, and that genotypes are now firmly established at a local level, with decades-long localized persistence. These results are consistent with the establishment of separate, locally entrenched

sources of *Lm*-CC1 with limited flow of bacteria either within or between countries, in line with cgMLST analyses in which 92% of clusters are country specific.

In the absence of interhuman horizontal transmission, this observation likely represents persistent infection sources, i.e., individual herds and/or production facilities, in which *Lm* can reside for several years (27, 51). Outbreak investigations performed at the local scale, including in farm environments, would therefore likely improve the identification of contaminating sources, which remain unknown in about 80% of clusters of human cases (52). Identifying and eradicating sources along the food chain, from the farm to the fork, could lead to significant long-term reductions in the transmission of the *Lm*-CC1.

Our isolates’ dataset exhibits an overrepresentation of Western Europe samples due to the relative scarcity of available genomes from Asia, Africa, and South America, and a relative lack of isolates from natural and animal reservoirs, which may miss other clades and past and current transmission chains in those regions. This limitation reflects the general global disparities in pathogen genomic sequencing. Nevertheless, this study sheds unprecedented light onto the evolutionary history, epidemiology, and population dynamics of *Lm*-CC1, providing critical clues on its worldwide spread. Similar approaches targeting other major globally distributed clonal complexes will allow clarifying their transmission dynamics and uncovering epidemiological specificities of *Lm* clones. Deciphering the dynamics and drivers of *Lm* sublineages across time and space will inform infection control policies to reinforce detection and genome analyses at both local and global levels to ultimately reduce the burden of listeriosis.

MATERIALS AND METHODS

Bacterial isolates and genome sequencing

A total of 2021 high-quality *Lm*-CC1 genomes collected by this study group ($n = 1230$) and from NCBI repositories ($n = 791$, as of 14 March 2018) were analyzed. These were part of an initial dataset of 2154 CC1 genomes, from which 133 were discarded because of low sequencing coverage (<40× after read trimming, $n = 62$) or low assembly quality (>200 contigs and/or N50 < 20 kb, $n = 71$) (31). The 2021 isolates originated from human ($n = 1453$; 72%) and animal hosts ($n = 44$; 2%), food ($n = 387$; 19%), food processing environments ($n = 88$; 4%), feed ($n = 11$; 0.5%), natural environments ($n = 11$; 0.6%), or unknown sources ($n = 27$; 1%) (Fig. 1 and dataset S1). Isolates were sampled in 40 countries from six continents, between 1921 and 2018 (Fig. 1 and dataset S1). Between 2012 and mid-2017, exhaustive sampling was obtained for seven countries in three continents in the context of listeriosis national surveillance programs in Australia ($n = 75$), Denmark ($n = 42$), France ($n = 395$), The Netherlands ($n = 53$), New Zealand ($n = 34$), the United Kingdom ($n = 106$), and the United States ($n = 317$). Sequencing reads were obtained using Illumina sequencing platforms (Illumina, San Diego, USA) and 2 × 50 bp ($n = 110$), 2 × 75 bp ($n = 2$), 2 × 100 bp ($n = 233$), 2 × 125 bp ($n = 9$), 2 × 150 bp ($n = 1,145$), 2 × 250 bp ($n = 351$), and 2 × 300 bp ($n = 138$) paired-end runs (dataset S1).

Sequence analysis

Whole-genome sequencing reads were available for 1988 of 2021 isolates. Reads were corrected and trimmed from adapter sequences and nonconfident bases (minimum read length of 30 bases and

minimum quality Phred score of 20, i.e., 99% base call accuracy) using fqCleaner v.3.0 (<https://gitlab.pasteur.fr/GIPhy/fqCleanER>). Assemblies were obtained from paired-end trimmed reads of ≥ 75 bp ($n = 1878$ isolates) by using SPAdes v.3.11.0 (53) with the automatic k-mer, --only-assembler and --careful options. For paired-end trimmed reads of 50 bp ($n = 111$), assemblies were built using CLC Assembly Cell v.5.0.0 (Qiagen, Denmark), with estimated library insert sizes ranging from 50 to 850 bp. Contigs smaller than 500 nucleotides were discarded from both SPAdes- and CLC-generated assemblies.

Pangenome analysis

Gene prediction and annotation were carried out from the draft assemblies using Prokka v.1.12 (54). Functional and taxonomic classification was carried out with eggNOG-mapper v2 (55) using DIAMOND (Double Index Alignment of Next-generation sequencing Data). The presence of plasmids, intact prophages, and *Listeria* genomic regions was inferred from the assemblies using MOB-suite v.2.0.1 (56), PHASTER (<https://phaster.ca/>) (57), and BIGSdb-*Lm* (<http://bigsdb.pasteur.fr/listeria/>) (31, 58), respectively. Pangenome analyses were carried out using Roary v.3.12 (59) with an amino acid identity cutoff of 95% and splitting homologous groups containing paralogs into groups of true orthologs. Venn diagrams were obtained using Venny 2.1. Pangenome-wide association analyses were performed using treeWAS v.1.0 (60), to control for phylogenetic structure, using the *Lm*-CC1 core genome maximum likelihood phylogeny (see the “Phylogenetic analyses” section) and a significance threshold of $P < 10^{-5}$.

In silico molecular typing

Polymerase chain reaction (PCR) serogrouping (5 loci) (61), MLST (7 loci) (4), and cgMLST (1748 loci) (31) profiles were extracted from draft assemblies using the BIGSdb-*Lm* platform (<http://bigsdb.pasteur.fr/listeria/>) as previously described (31). Profiles were compared using the single linkage clustering method implemented in BioNumerics v.7.6 (Applied-Maths, Belgium). cgMLST profiles were classified into cgMLST types (CTs) and sublineages (SLs) using previously defined cutoffs (7 and 150 allelic mismatches, respectively, out of 1748 loci) (31). Rarefaction curves were computed with vegan v.2.5-6 (62) R package, estimated with the rarefaction function (Joshua Jacobs, joshuajacobs.org/R/rarefaction) using 100 random samples per point.

Phylogenetic analyses

Core genome multiple sequence alignments were built from the 1748 cgMLST loci concatenated sequences (31). Briefly, individual allele sequences were translated into amino acids, aligned separately with MUSCLE v.3.8.31 (63), and back-translated into nucleotide sequence alignment. Concatenation of the 1748 loci alignments resulted in a multiple sequence alignment of 1.57 Mb.

In parallel, wgSNP-based alignments were built from trimmed reads and NCBI assemblies using the Snippy v.4.1.0 pipeline (<https://github.com/tseemann/snippy>). The closed CC1 genome F2365 (accession no. NC_002973.6), from the 1985 Canadian cheese outbreak, was used as reference in read mapping, resulting in an alignment of 2.29 Mb.

Gubbins v.2.2.0 (64) was used to detect recombination regions in both core and whole-genome alignments, using default parameters and a minimum of three base substitutions required to identify recombination. Alignment regions positive for recombination were

then completely removed from the original alignments, resulting in recombination-free core- and whole-genome alignments of 1.29 and 2.28 Mb, respectively. Maximum likelihood phylogenies were obtained from the recombination-purged alignments using IQ-tree v.1.6.7.2 (65) under the determined best-fit nucleotide substitution model (GTR+F+G4, as determined by ModelFinder) and ultrafast bootstrapping of 1000 replicates. Trees were visualized and annotated with ggtree v.1.14.6 (66) and iTol v.4.2 (67).

To measure the degree of genetic variation within sublineages, GCs, and geographic locations, the pairwise allelic and SNP distance matrices were calculated from the cgMLST profiles and multiple sequence alignments, respectively. SNP distances were computed, taking into account only the ATGC polymorphic positions, extracted from the alignments using SNP-sites v.2.4.1 (68). The nucleotide diversity and the Tajima's and Fu and Li's D statistics per alignment were calculated using the R package PopGenome v.2.6.1 (69).

Demographic and spatiotemporal analysis

Substitution rates and demographic changes over time were estimated using a Bayesian approach in BEAST v1.10.4 (70). Two demographic models were tested: the coalescent Bayesian skyline model (non-parametric), which allows a wide range of demographic scenarios, avoiding the biases of prespecified parametric models in the estimates of demographic history (37), and the coalescent constant model (parametric), which assumes that the populations have remained constant through time. To ensure feasible computational running times, the full dated dataset ($n = 1972$ genomes with country and year of isolation) was down-sampled to a subset of 200 isolates randomly selected out of 421 isolates representative of genomic and geographic diversity of the full dataset (one isolate per country per cluster of 99% core genome similarity; dataset S1). To control for the overrepresentation of recent isolates and geographic locations, isolates were also divided into bins of 10 years based on the isolation dates and randomly subsampled 10 times, allowing only a maximum of 10 isolates per bin and, when possible, equal representation of continents. Sampling times positively correlated with the genetic divergence ($P < 0.05$, F-statistic test; fig. S8), as observed using TempEst v1.5.1 (71). To assess the significance of the temporal signatures observed, 10 randomized tip date datasets were used as controls. BEAST estimations were made using the nucleotide evolutionary model GTR+Γ4 and a default gamma prior distribution of 1, under an uncorrelated relaxed clock model, to allow each branch of the phylogenetic tree to have its own evolutionary rate (72). Runs were performed in triplicates, each consisting of Monte Carlo Markov chains of 400 million generations, with a 25% burn-in. Parameter values were sampled every 10,000 generations. The effective sample size values were confirmed to be higher than 200 for all parameters using Tracer v.1.7 (73). The tested models were compared by marginal likelihood, and stronger support was obtained for the skyline demographics model (Bayes factor = 10.179). The time of the MRCA and 95% highest posterior densities (95% HPDs) were inferred from the nodes of the maximum clade credibility tree. Estimations of the effective population size along the years were computed using Tracer v.1.7 (73).

Phylogeography analyses were then extended to the 1972 CC1 genomes for which country and year of isolation were available. Time-calibrated phylogenies were inferred from the maximum likelihood core genome trees (obtained with IQ-tree, as described above) using either BactDating v1.0.1 (74), TreeTime v0.5.2 (75),

or Treedater v0.3.0 (76), assuming a relaxed clock model and the estimated substitution rate of $1.954 \times 10^{-7} \pm 2.015 \times 10^{-8}$ substitutions per site per year (obtained with BEAST as described above). Cophenetic correlations between BEAST and the three alternative large-scale dating methods were evaluated, and better R^2 coefficient scores were obtained for Treedater (fig. S9). For this reason, the latter dated tree was used in further downstream analyses.

Ancestral geographic reconstruction was performed with PastML (77) using the marginal posterior probabilities approximation (MPPA) method with an F81-like model, and estimated ancestral state probabilities were mapped onto the full time-calibrated phylogeny using the R package ape v5.3 (78).

SL1 global transmission dynamics

To infer the transmission dynamics at a recent time scale (Fig. 5A and fig. S13), we focused on the *Lm*-CC1 main sublineage, and we analyzed the genetic similarity of SL1 isolates from 2010 to 2018 ($n = 1266$) across different temporal and spatial scales, as described before (79). To avoid oversampling isolates from outbreak investigations, we excluded all nonclinical isolates from confirmed outbreaks ($n = 91$ isolates from 19 outbreaks). We computed the probability P_1 that a pair of isolates that satisfy a given location criteria that were sampled within 2 years of each other had an MRCA in a specific range (0 to 5 years, 5 to 20 years, 20 to 50 years, and >50 years), relative to the probability P_{ref} that a pair of isolates that satisfy the reference location criterion sampled within 2 years of each other had an MRCA within that particular range. The location criteria used were as follows: (i) within countries (both isolates come from the same country), (ii) between countries ≤ 1000 km (isolates come from distinct countries, separated by less than 1000 km, from the same continent), (iii) between countries >1000 km (isolates come from distinct countries, separated by more than 1000 km, from the same continent; used as reference), and (iv) between continents (isolates come from distinct continents). Spatial relationships between isolates were calculated using the centroid coordinates of the countries or regions of origin.

We estimated these probabilities using

$$P_1 = \frac{\# \text{ pairs } \{\text{MRCA} \in \text{window} \& \text{sampled within 2 years \& given location criteria}\}}{\# \text{ pairs } \{\text{sampled within 2 years \& given location criteria}\}}$$

$$P_{\text{ref}} = \frac{\# \text{ pairs } \{\text{MRCA} \in \text{window} \& \text{sampled within 2 years \& distant countries}\}}{\# \text{ pairs } \{\text{sampled within 2 years \& distant countries}\}}$$

Last, the relative risk (RR) was given by

$$\text{RR} = \frac{P_1}{P_{\text{ref}}}$$

To measure uncertainty, we used a combination of bootstrapping observations and sampling trees from the Treedater v0.3.0 package (76) to incorporate both sampling and tree uncertainty. Over repeated resamples, we first selected a random tree and calculate the evolutionary distance separating all pairs of sequences. Then, we resampled all the isolates with replacement and recalculate RR each time. The 95% CIs are the 2.5 and 97.5% quantiles from the resultant distribution from 1000 resampling events.

SL1 local transmission dynamics

To assess the SL1 local transmission dynamics, we used available data from France. We computed the proportion of closely related

pairs of French isolates (defined as having an MRCA of <5 years) as a function of the spatial distance within and between administrative departments (Fig. 5B)

$$p(\text{location}) = \frac{\# \text{ pairs } \{\text{MRCA} < 5 \text{ years \& sampled within 2 years \& given location}\}}{\# \text{ pairs } \{\text{sampled within 2 years \& given location}\}}$$

The different location criteria used are as follows: (i) within department: both isolates come from the same department; (ii) between departments: isolates come from different departments, separated by a distance from 50 to >500 km. The French departments are shown in the map in fig. S12. As shown by Salje *et al.* (79), the reciprocal of $p(\text{within department})$ represents the lower limit of the number of sources of human infection circulating within a department. To assess uncertainty, we used the bootstrapping approach as described above.

To explore possible differences between departments, we computed the relative risk that a pair of isolates share an MRCA of less than 5 years when both come from the same department compared to when coming from different departments. We looked at two different groups of departments: (i) Paris alone (Fig. 5C, left): within Paris (both isolates come from Paris) and between Paris and other departments (for each pair of isolates, one of them come from Paris, and the other one from another department); (ii) other departments, except Paris (Fig. 5C, right): with other departments (both isolates come from the same department, excluding Paris) and between all other departments (isolates come from two different departments, excluding Paris). For each group, to compute the relative risk RR, we used the same approach as explained above. We estimated

$$P_1 = \frac{\# \text{ pairs } \{\text{MRCA} < 5 \text{ years \& sampled within 2 years \& same department}\}}{\# \text{ pairs } \{\text{sampled within 2 years \& same department}\}}$$

$$P_{\text{ref}} = \frac{\# \text{ pairs } \{\text{MRCA} < 5 \text{ years \& sampled within 2 years \& different departments}\}}{\# \text{ pairs } \{\text{sampled within 2 years \& different departments}\}}$$

Last, the relative risk is given by

$$\text{RR} = \frac{P_1}{P_{\text{ref}}}$$

To determine uncertainty, we used the same bootstrapping approach as described above. To assess the statistical significance of each RR, we performed a one-tailed test. We set the null hypothesis (H_0) as $\text{RR} \leq 1$ and alternative hypothesis (H_1) as $\text{RR} > 1$. For each group, composed of N bootstrap events, we computed

$$p = \frac{\sum_{i=1}^N I(\text{RR}_i \leq 1)}{N}$$

SUPPLEMENTARY MATERIALS

Supplementary material for this article is available at <https://science.org/doi/10.1126/sciadv.abj9805>

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Alexandra MouraNoémie LefrancqThierry WirthAlexandre LeclercqVitor BorgesBrent GilpinTimothy J. DallmanJoachim FreyEelco FranzEva M. NielsenJuno ThomasArthur PightlingBenjamin P. HowdenCheryl L. TarrPeter Gerner-SmidtSimon CauchemezHenrik SaljeSylvain BrisseMarc Lecuit

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