

# Quant 2, Lab 5

## Matching and Weighting

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  - ▶ Doesn't solve issues with bad/good controls, selection on unobservables, etc.

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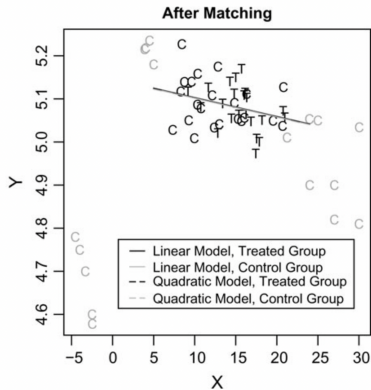
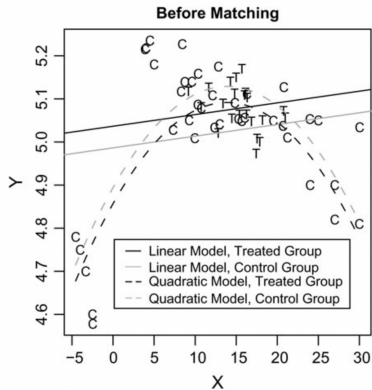
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- ▶  $\mathcal{J}(i)$  are the set of  $M$  closest control units to  $i$  in terms of  $X_i$ .

# Example

Ho et al 2007



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  - ▶ Intuitively, you have to have observations to match on. This is an empirical question

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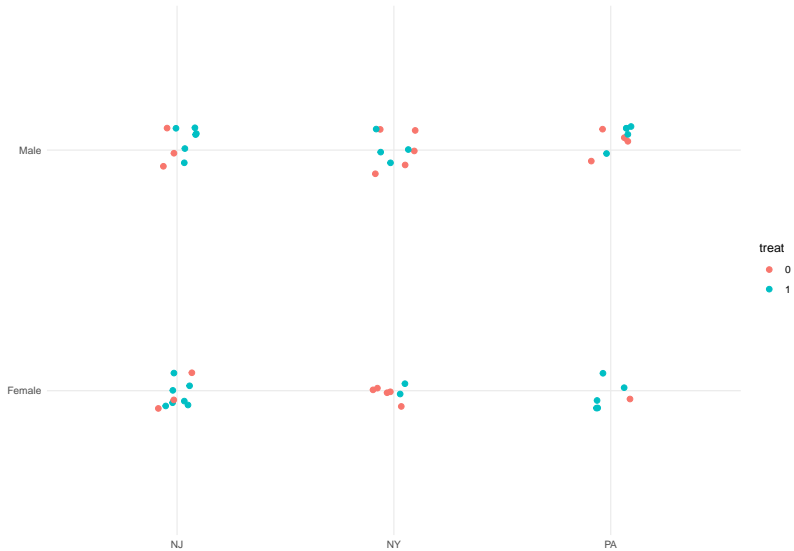
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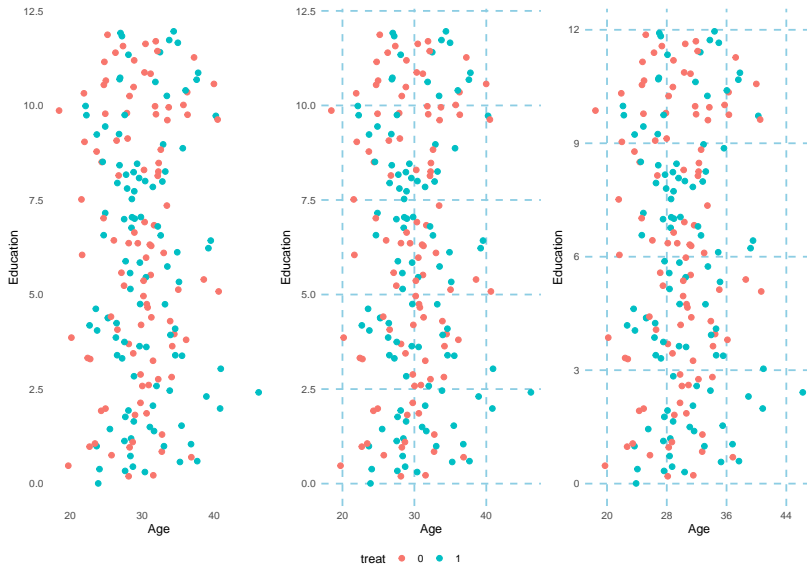
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  - ▶ Problem: Still might be bins with no control units (more likely with many covariates)

# Exact Matching: Example



# CEM Example



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  - ▶ Should we drop matches where the distance is too far (caliper)?

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  - ▶ Reduces imbalance, but if you drop treated units, estimand changes.

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  - ▶ Why use  $\hat{Y}_{j0}$  instead of observed  $Y_{j0}$ ?

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  - ▶ Dropping units can lower power (increase p-values) without a change in balance.

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- ▶ This estimator benefits by adjusting for unusually high or low values of  $\pi(X)$

## True vs. Estimated p-scores

- ▶ True propensity scores are only known sometimes (e.g., randomized experiments). In most non-experimental settings, the p-score is unknown and must be estimated
- ▶ When estimating, we have two cases:
  - ▶ If  $X$  is discrete, we know that  $\hat{\pi}(X)$  can be an exact approximation (why?)
  - ▶ If  $X$  is not discrete (or high-dimensional), how should we approximate it?
- ▶ We need to estimate  $\pi(X)$  in a way that is flexible and will converge to the truth in the limit – e.g., semi-parametric estimation of  $\pi$ .
  - ▶ Note a linear model of  $\pi$  will inherently be wrong b/c probabilities are bounded between 0 and 1
  - ▶ Practical implication: logit estimation of  $\pi(X)$  is reasonable, allowing for flexible specification of  $X$
  - ▶ As dimension of  $X$  grows, ML / lasso style models grow in value

## True vs. Estimated p-scores

- ▶ Important result: even if you know the true function  $\pi(X)$ , better to use the estimated function than the truth (Imbens, Hirano and Ridder (2002))
  - ▶ Intuition: the deviations from the “true” propensity score  $\hat{\pi}(X) - \pi(X)$  are informative for the estimation of the treatment effects (a la extra moment restrictions in GMM)
- ▶ Clear tension – as dimension of controls increases, the noisiness in  $\pi$  grows as well

# True vs. Estimated p-scores

```
set.seed(123)
ht.est <- function(y, d, w) {
  n <- length(y)
  (1 / n) * sum((y * d * w) - (y * (1 - d) * w))
}
n <- 200
x <- rbinom(n, size = 1, prob = 0.5)
dprobs <- 0.5 * x + 0.4 * (1 - x)
d <- rbinom(n, size = 1, prob = dprobs)
y <- 5 * d - 10 * x + rnorm(n, sd = 5)
true.w <- ifelse(d == 1, 1 / dprobs, 1 / (1 - dprobs))
pprobs <- predict(glm(d ~ x))
est.w <- ifelse(d == 1, 1 / pprobs, 1 / (1 - pprobs))
ht.est(y, d, est.w)
```

```
## [1] 5.029735
```

```
ht.est(y, d, true.w)
```

```
## [1] 5.740815
```

# True vs. Estimated p-scores

```
sims <- 10000
true.holder <- rep(NA, sims)
est.holder <- rep(NA, sims)
for (i in 1:sims) {
  x <- rbinom(n, size = 1, prob = 0.5)
  dprobs <- 0.5 * x + 0.4 * (1 - x)
  d <- rbinom(n, size = 1, prob = dprobs)
  y <- 5 * d - 10 * x + rnorm(n, sd = 5)
  true.w <- ifelse(d == 1, 1 / dprobs, 1 / (1 - dprobs))
  pprobs <- predict(glm(d ~ x))
  est.w <- ifelse(d == 1, 1 / pprobs, 1 / (1 - pprobs))
  est.holder[i] <- ht.est(y, d, est.w)
  true.holder[i] <- ht.est(y, d, true.w)
}
var(est.holder)
```

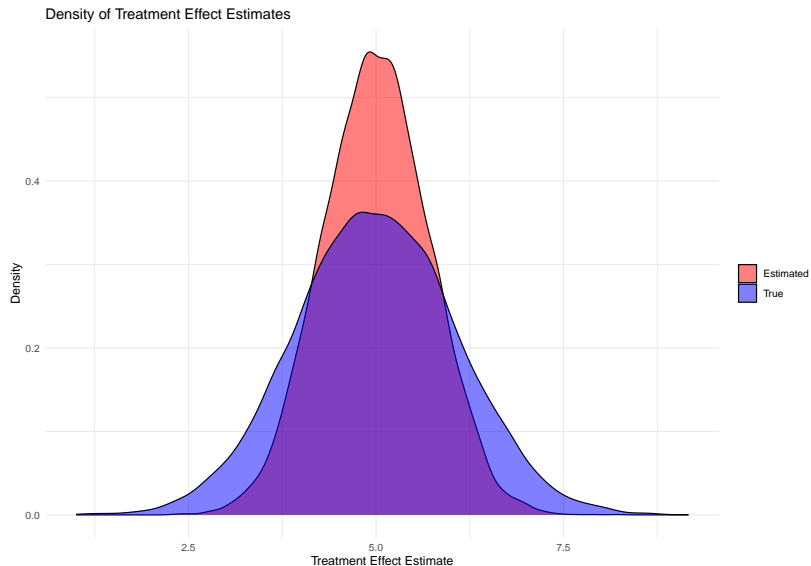
```
## [1] 0.5062535
```

```
var(true.holder)
```

```
## [1] 1.147964
```



# True vs. Estimated p-scores



So??

- ▶ Why is the estimated propensity score more efficient than the true PS?
- ▶ Removing chance variations using  $\hat{\pi}(X_i)$  adjusts for any small imbalances that arise because of a finite sample.
- ▶ True PS only adjusts for the expected differences between samples.
- ▶ Only true if propensity score model is correctly specified!!