#### Quant 2, Lab 4

Pitfalls of control strategies continued: Effective Samples, Specification Error/Double ML

Sylvan Zheng

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```
set.seed(12)
N <- 1000
X <- rnorm(N)
D <- X + rbinom(N, size = 1, prob = 0.10) * rnorm(N)
Y <- D + X + rnorm(N)
df <- data.frame(X = X, D = D, Y = Y)</pre>
```

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- ▶ Then, does this unit help identify the effect of D|X on Y?

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```

For most of the sample, X = D

113

##

887

```
table(df$X == df$D)
##
## FALSE TRUE
```

```
models <- list(
   feols(Y ~ D + X, data = df),
   feols(Y ~ D + X, data = df %>% filter(df$X != df$D))
)
```

Dependent Variable:	Υ	
Model:	(1)	(2)
Variables		
Constant	-0.0192	0.0570
	(0.0336)	(0.0961)
D	1.019***	1.015***
	(0.0936)	(0.0901)
Χ	0.9519***	0.9995***
	(0.1022)	(0.1384)
Fit statistics		
Observations	1,000	113
		·

IID standard-errors in parentheses

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  - If we assume linearity of the treatment assignment in  $X_i$ , then easy to construct
  - Run the regression  $D_i = X_i \gamma + e_i$
  - ► Take residual  $\hat{e}_i = D_i X_i \hat{\gamma}$  and square it

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- ► Eg, original sample is "representative". Is effective sample "representative?"
- Interpretation on the research setting. Controlling for confounders almost mechanically makes the effective sample non representative.

# Turning Personal Experience into Political Attitudes: The Effect of Local Weather on Americans' Perceptions about Global Warming

Patrick J. Egan New York University Megan Mullin Temple University

How do people translate their personal experiences into political attitudes? It has been difficult to explore this question using observational data, because individuals are typically exposed to experiences in a selective fashion, and self-reports of exposure may be biased and unreliable. In this study, we identify one experience to which Americans are exposed nearly at random—their local weather—and show that weather patterns have a significant effect on people's beliefs about the evidence for global warming.

Outcome variable getwarmord (climate change attitudes)

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```
out.d <- feols(ddt_week ~ educ_hsless + educ_coll + educ_postgrad +
    educ_dk + party_rep + party_leanrep + party_leandem +
    party_dem + male + raceeth_black + raceeth_hisp +
    raceeth_notwbh + raceeth_dkref + age_1824 + age_2534 +
    age_3544 + age_5564 + age_65plus + age_dk + ideo_vcons +
    ideo_conservative + ideo_liberal + ideo_vlib + ideo_dk +
    attend_1 + attend_2 + attend_3 + attend_5 + attend_6 +
    attend_9 | doi + state + wbnid_num, d)

# Extract the residuals and take their square
d$wts <- residuals(out.d)^2</pre>
```

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- "Nominal" weight: just the (normalized) number of observations per state

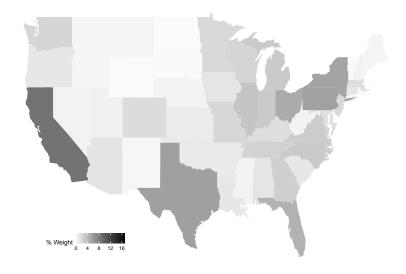
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```
nom_map <- theme_state_map(d %>%
    group_by(state) %>%
    summarize(nom = n() * 100 / nrow(d)) %>%
    ggplot(aes(map_id = state)) +
    geom_map(aes(fill = nom), map = state_map) +
    labs(title = "Nominal Sample"))
```

#### nom\_map

Nominal Sample

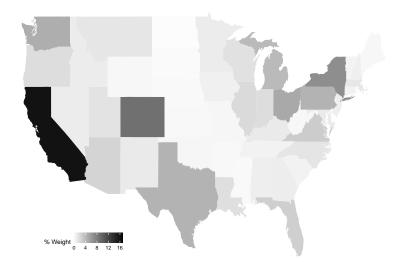


➤ To characterize the "effective" contribution of each state, use the effective sample weight instead

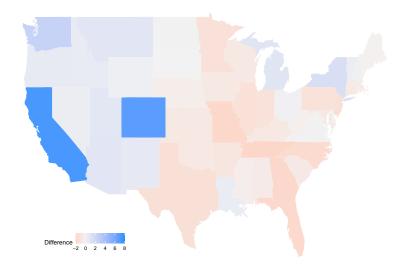
```
eff_map <- theme_state_map(d %>%
    group_by(state) %>%
    summarize(eff = sum(wts) * 100 / sum(d$wts)) %>%
    ggplot(aes(map_id = state)) +
    geom_map(aes(fill = eff), map = state_map) +
    labs(title = "Effective Sample"))
```

eff\_map

Effective Sample



Difference Effective and Nominal Weight



- ▶ Linear model with K covariates. In matrix form:  $y = X'\beta + \varepsilon$
- ► FWL gives a formula for the OLS estimate of the k<sup>th</sup> coefficient.

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Equivalent to the following:

Regress the individual variable  $X_k$  on all the other covariates and take the residuals

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  - The part of Y unexplained by  $X_{not_k}$  " ~ "The part of  $X_k$  unexplained by  $X_{not_k}$ "

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  - ► "The part of Y unexplained by  $X_{not_k}$ " ~ "The part of  $X_k$  unexplained by  $X_{not_k}$ "
  - Note that to get  $\hat{\beta}_k$  it is enough to regress the non-residualized y on residualized  $X_k$  (why?), but the SE won't be right

#### FWL in R

```
set.seed(123)
N <- 1000
X <- rnorm(N, mean = 0, sd = 1)
# Generate binary treatment D, making D and X correlated
D <- rbinom(N, size = 1, prob = plogis(X))
Y <- 2 * D + 0.5 * X + rnorm(N, mean = 0, sd = 1)
model_ols <- lm(Y ~ D + X)</pre>
```

#### FWI in R

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N <- 1000
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Y <- 2 * D + 0.5 * X + rnorm(N, mean = 0, sd = 1)
model_ols <- lm(Y ~ D + X)</pre>
```

```
coeftable(model_ols)[, 1:2] %>% kable()
```

	Estimate	Std. Error
(Intercept)	-0.0292034	0.0450014
D	2.0576326	0.0682031
Χ	0.4307956	0.0343768

#### FWL in R

```
resid_Y <- residuals(lm(Y ~ X))
resid_D <- residuals(lm(D ~ X))
model_fwl <- lm(resid_Y ~ resid_D)</pre>
```

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	Estimate	Std. Error
(Intercept)	0.000000	0.0310951
resid_D	2.057633	0.0681689

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- Simulation:

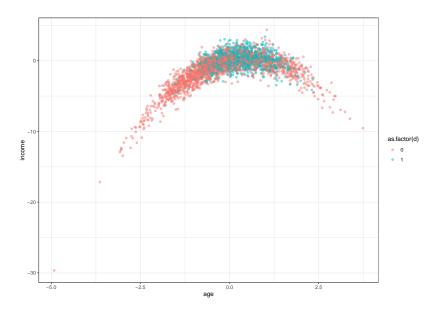
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```
set.seed(6)
N <- 3000
effect <- 0.2
age <- rnorm(N, 0, 1)
age2 <- -(age)^2 + age
d <- rbinom(N, size = 1, prob = plogis(age2))
income <- -(age)^2 + age + rnorm(N) + d * effect</pre>
```



▶ Using a linear specification for control leads to bias in estimate

etable(feols(income ~ age + d, data = dat), tex = T, fitstat = c

Dependent Variable: Model:	income (1)
Variables	
Constant	-1.294***
	(0.0403)
age	0.9361***
	(0.0328)
d	1.038***
	(0.0687)
Fit statistics	
Observations	3,000

IID standard-errors in parentheses

Signif. Codes: \*\*\*: 0.01, \*\*: 0.05, \*: 0.1

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- How do we know what specification to use?
- "Classical" Machine Learning: flexible algorithms to estimate nonlinear relationships

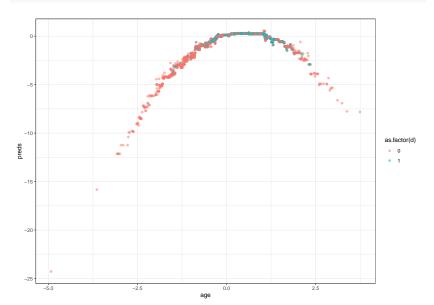
- ▶ But, a problem quickly arises.
- How do we know what specification to use?
- "Classical" Machine Learning: flexible algorithms to estimate nonlinear relationships
  - XGBoost, Random Forest, Lasso/ElasticNet... (Quant 3 for more on this)

# Classical Machine Learning

```
pacman::p_load(mlr3, xgboost, mlr3learners)
# Use XGBoost algorithm
learner <- lrn("regr.xgboost", eta = 0.1, nrounds = 35)</pre>
# Set up a task to predict 'income'
task <- as_task_regr(</pre>
    select(dat, income, age),
    target = "income"
# Fit to the data
learner$train(task)
```

# Classical Machine Learning

dat\$preds <- learner\$predict\_newdata(dat)\$response</pre>



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- ▶ Basic idea is to use ML to model both y ~ x and d ~ x and use the residuals to retrieve a consistent estimate for  $\theta$

- Ok, so we can predict Y given X in a nonlinear way.
- How do we use this to retrieve a good estimate of Y ~ D | X?
- ▶ Basic idea is to use ML to model both y ~ x and d ~ x and use the residuals to retrieve a consistent estimate for  $\theta$
- Sounds familiar? (FWL)

```
# Model y as a function of age
y.x <- lrn("regr.xgboost", eta = 0.1, nrounds = 35)
y.x.task <- as_task_regr(
    select(dat, income, age),
    target = "income"
)
y.x$train(y.x.task)</pre>
```

d.x\$train(d.x.task)

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y.x <- lrn("regr.xgboost", eta = 0.1, nrounds = 35)
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    select(dat, income, age),
    target = "income"
y.x$train(y.x.task)
# Model D as a function of age
d.x \leftarrow lrn("regr.xgboost", eta = 0.1, nrounds = 35)
d.x.task <- as_task_regr(</pre>
    select(dat, d, age),
    target = "d"
```

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    select(dat, d, age),
    target = "d"
d.x$train(d.x.task)
```

```
# Calculate residuals
d.x.resid <- dat$d - d.x$predict_newdata(dat)$response
y.x.resid <- dat$income - y.x$predict_newdata(dat)$response</pre>
```

```
lm(y.x.resid ~ d.x.resid)

##

## Call:
## lm(formula = y.x.resid ~ d.x.resid)
##

## Coefficients:
## (Intercept) d.x.resid
## -0.03956 0.24633
```