



THE UNIVERSITY OF TEXAS AT AUSTIN
McCOMBS SCHOOL OF BUSINESS

Review of distributions and estimation

Lecture 8

STA 371G

The story of statistics

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- Example: I want to know the average GPA at UT, but I only have a sample of $n = 100$ GPAs.

Population	All GPAs at UT
Sample	The 100 GPAs in my sample
Parameter	Average GPA among all UT students (μ)
Statistic	Average GPA among the 100 students in my sample ($\hat{\mu}$)

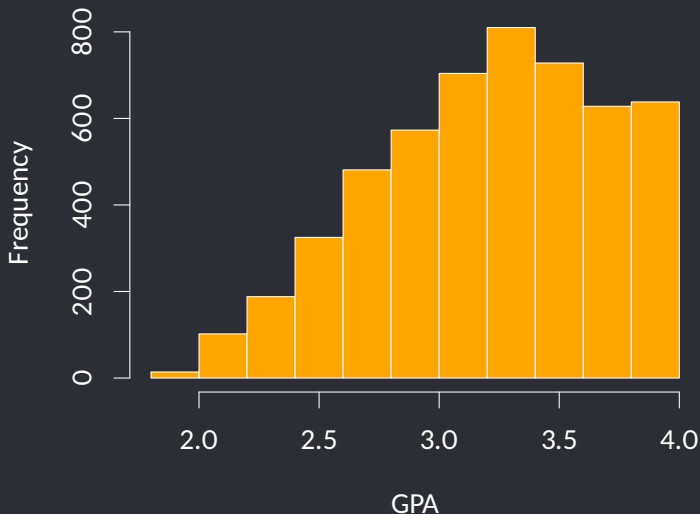
Data set

The data set `ut2000` contains information on all 5191 students that entered UT Austin in Fall 2000 and graduated within 6 years.

```
head(ut2000)
```

	SAT.V	SAT.Q	SAT.C	School	GPA	Status
1	690	580	1270	BUSINESS	3.82	G
2	530	710	1240	NATURAL SCIENCE	3.53	G
3	610	700	1310	NATURAL SCIENCE	3.37	G
4	730	700	1430	ENGINEERING	3.34	G
5	700	710	1410	NATURAL SCIENCE	3.72	G
6	540	690	1230	LIBERAL ARTS	2.69	G

```
hist(ut2000$GPA, main="", col="orange", xlab="GPA")
```



Let's take a sample

Usually, we only have access to a sample of the data. Let's pretend that we only had a sample of $n = 100$ students:

```
sample.gpas <- sample(ut2000$GPA, 100)  
mean(sample.gpas)
```

```
[1] 3.23
```

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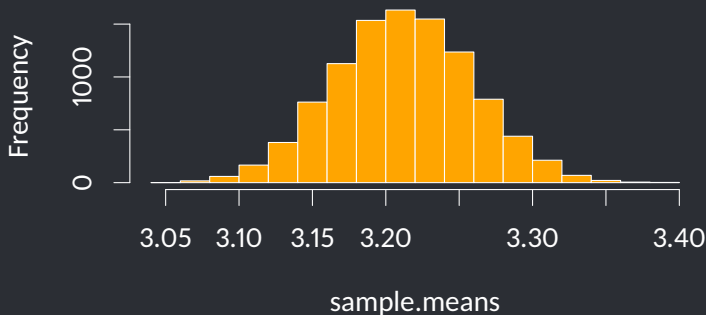
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Since we have a random sample, it's a good, but not perfect, estimate of the population GPA (3.212). But normally we don't have access to the population, so we don't know how good our estimate is!

Normally we only have access to one sample. But what if we had many samples, and we took the sample mean in each sample?

```
sample.means <- replicate(10000,  
  mean(sample(ut2000$GPA, 100)))  
hist(sample.means, main="", col="orange")
```



Sampling distribution of \overline{GPA}

The *sampling distribution* of \overline{GPA} is the distribution of sample means, if we took repeated samples:

$$E(\overline{GPA}) = \mu = 3.212$$

$$SD(\overline{GPA}) = \frac{\sigma}{\sqrt{n}} = \frac{0.48}{\sqrt{100}} = 0.048$$

The last value quantifies how much the sample mean will vary from sample to sample. But we normally can't compute σ since we don't have the whole population, so we estimate it by calculating the SD in the *sample* ($\hat{\sigma}$) and dividing by \sqrt{n} ; this is the *standard error of the mean*.



Confidence intervals

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Our sample statistic is $\hat{\mu}$ and our standard error is $\hat{\sigma}/\sqrt{n}$. What is the critical value?

As it turns out, the sampling distribution (of $\hat{\mu}$) is not *quite* Normal. If we standardize the sample means, the distribution of

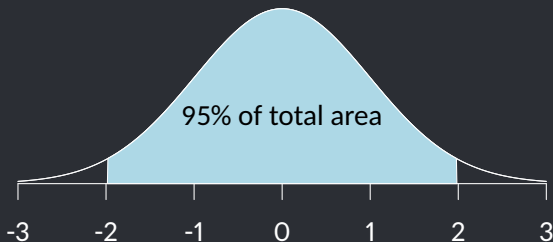
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is called a t -distribution with $n - 1$ degrees of freedom. The critical value for a 95% confidence interval is $t^* = \pm 1.984$, the value that cuts off 95% of the area under the t -distribution:



There are two helpful R functions for calculating values around t -distributions:

- `pt(x, df)` calculates $P(t < x)$ if we are looking at a distribution with df degrees of freedom.

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So the critical value for a 95% confidence interval is `qt(.975, 99)` when $n = 100$.

There are similar functions `pnorm` and `qnorm` when you are working with Normal distributions.

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There are two ways to interpret this:

- **Informally**, we are 95% confident that the population mean GPA is between 3.133 and 3.32.
- **Formally**, if we took repeated samples and found the 95% CI within each sample, 95% of the CIs would contain the population mean.

R can do this work for you!

```
t.test(sample.gpas, conf.level=0.95)
```

One Sample t-test

```
data: sample.gpas
```

```
t = 70, df = 100, p-value <2e-16
```

```
alternative hypothesis: true mean is not equal to 0
```

```
95 percent confidence interval:
```

```
3.13 3.32
```

```
sample estimates:
```

```
mean of x
```

```
3.23
```

Hypothesis tests

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- Usually, we don't have the population, and so we can't know for sure that he is wrong.
- But we do have some evidence (our sample) that we can bring to bear on the question.

Let's start by framing Sooner's claim as a null hypothesis; the alternative hypothesis is what we will believe if it turns out the null is false:

H_0	(null hypothesis)	$\mu = 3.0$
H_A	(alternative hypothesis)	$\mu \neq 3.0$

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$$\begin{array}{ll|l} H_0 & \text{(null hypothesis)} & \mu = 3.0 \\ H_A & \text{(alternative hypothesis)} & \mu \neq 3.0 \end{array}$$

A hypothesis test will produce a p -value, which represents the probability:

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In other words, if Sooner is correct, how likely is it that in our sample we would see a sample mean that is so far away from his hypothesized value?

R can run hypothesis tests for us:

```
t.test(sample.gpas, mu=3)
```

One Sample t-test

```
data: sample.gpas
```

```
t = 5, df = 100, p-value = 5e-06
```

```
alternative hypothesis: true mean is not equal to 3
```

```
95 percent confidence interval:
```

```
3.13 3.32
```

```
sample estimates:
```

```
mean of x
```

```
3.23
```

So:

$$P(\text{seeing data this extreme in our sample} \mid H_0 \text{ is true}) = 5 \times 10^{-6}.$$



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$\alpha = .05$ is a good “default” to use unless you have a reason to set it higher or lower.

