initial parameters are then computed as

$$\pi_{m} = N_{m}/N_{\text{tot}}$$

$$\mu_{m} = N_{m}^{-1} \sum_{n=1}^{N} \sum_{t=1}^{T^{(n)}} o_{t}^{(n)} h_{m}(o_{t}^{(n)}) \qquad (15)$$

$$\sigma_{m}^{2} = N_{m}^{-1} \sum_{n=1}^{N} \sum_{t=1}^{T^{(n)}} (o_{t}^{(n)} - \mu_{m})^{2} h_{m}(o_{t}^{(n)}) \qquad (16)$$

$$\sigma_m^2 = N_m^{-1} \sum_{n=1}^N \sum_{t=1}^{T^{(n)}} (o_t^{(n)} - \mu_m)^2 h_m(o_t^{(n)})$$
 (16)

This approximation is then improved upon by utilizing the expectation-maximization procedure described by Bilmes [22].

$$\pi'_{m} = N_{\text{tot}}^{-1} \sum_{n=1}^{N} \sum_{t=0}^{T^{(n)}} \chi_{m}(o_{t}^{(n)}, \mathbf{E}, \boldsymbol{\pi})$$

$$\mu'_{m} = (\pi_{m} N_{\text{tot}})^{-1} \sum_{n=1}^{N} \sum_{t=0}^{T^{(n)}} o_{t}^{(n)} \chi_{m}(o_{t}^{(n)}, \mathbf{E}, \boldsymbol{\pi})$$

$$\sigma'_{m}^{2} = (\pi_{m} N_{\text{tot}})^{-1} \sum_{n=1}^{N} \sum_{t=0}^{T^{(n)}} (o_{t}^{(n)} - \mu'_{m})^{2} \chi_{m}(o_{t}^{(n)}, \mathbf{E}, \boldsymbol{\pi}) (17)$$

where the function  $p(m|o, \mathbf{E}, \pi)$  is given by the fuzzy membership function:

$$\chi_m(o, \mathbf{E}, \boldsymbol{\pi}) = \frac{\pi_m \, \varphi(o|\mathbf{e}_m)}{\sum_{l=1}^M \pi_l \, \varphi(o|\mathbf{e}_l)}$$
(18)

This iterative procedure is terminated when the change in the parameters  $\{\pi, \mu, \sigma^2\}$  falls below a certain relative threshold, such as  $||\pi^{(n)} - \pi^{(n-1)}||/||\pi^{(n)}|| < 10^{-4}$ .

Once initial state observable emission parameters E are determined, an initial transition matrix is estimated using an iterative likelihood maximization approach that enforces detailed balance [23]. First, a matrix of fractional transition counts  $C \equiv (c_{ij})$  is estimated using the membership function:

$$c_{ij} = \sum_{n=1}^{N} \sum_{t=1}^{T^{(n)}} \chi_i(o_{t-1}^{(n)}, \mathbf{E}, \pi) \chi_j(o_t^{(n)}, \mathbf{E}, \pi)$$
(19)

A symmetric  $M \times M$  matrix  $\mathbf{X} \equiv (x_{ij})$  is initialized by

$$x_{ij} = x_{ji} = c_{ij} + c_{ji} (20)$$

and a vector of row sums

$$x_i = \sum_{j=1}^{M} x_{ij} \tag{21}$$

Then, the iterative procedure described in Algorithm 1 of [23] is applied. For each update iteration, we first update the diagonal elements of X:

$$x'_{ii} = \frac{c_{ii}(x_i - x_{ii})}{c_i - c_{ii}}$$

$$x'_{i} \equiv \sum_{j=1}^{M} x'_{ij}$$
(22)

could we use the notation  $X_i$  and  $C_i$  to happened to  $\gamma_{ti}$  for  $t=T_n$ . Check this?] indicate  $Z_i$   $X_{ij}$ ? You shouldn't need to update them then

followed by the off-diagonal elements:

$$x'_{ij} = x'_{ji} = \frac{-b + \sqrt{b^2 - 4ac}}{2a}$$

$$x'_{i} = \sum_{j=1}^{M} x'_{ij}$$
(23)

where the quantities a, b, and c are computed from X and Cas

$$a \equiv c_{i} - c_{ij} + c_{j} - c_{ji}$$

$$b \equiv c_{i}(x_{j} - x_{ji}) + c_{j}(x_{i} - x_{ij}) - (c_{ij} + c_{ji})(x_{i} - x_{ij} + x_{j} - x_{ji})$$

$$c \equiv -(c_{ij} + c_{ji})(x_{i} - x_{ij})(x_{j} - x_{ji})$$
(24)

[JDC: Can we merge the updates for  $x'_{ii}$  and  $x'_{ij}$  to simplify this description?] Once a sufficient number of iterations have been completed to compute a stable estimate of X (such as the relative convergence criteria  $||\mathbf{X}^{(n)} - \mathbf{X}^{(n-1)}||/||\mathbf{X}^{(n)}|| <$  $10^{-4}$ , the maximum likelihood transition matrix estimate T is computed as

$$T_{ij} = \frac{x_{ij}}{x_i}. (25)$$

Note that the equilibrium probability vector  $\pi$  computed during the Gaussian mixture model fitting is not respected during this step.

## B. Fitting a maximum likelihood HMM

The HMM model parameters  $\Theta \equiv \{T, E\}$  are fit to the observed data O through use of the expectation-maximization (EM) algorithm [24]. This is an iterative procedure, where the model parameters are subsequently refined through successive iterations.

During each iteration, the Baum-Welch algorithm [12] is used to compute, for each trace n,  $\Xi^{(n)} \equiv (\xi_{tij}^{(n)})$ , which represents the probability that the system transitions from hidden state i at time t-1 to hidden state j at time t, and  $\gamma_{ti}^{(n)}$ , the probability that the system was in state i at time t. This is accomplished by first executing the forward algorithm, which proceeds (suppressing the superscript (n)) as

$$\alpha_{tj} = \begin{cases} \rho_i \varphi(o_0 | \mathbf{e}_j) & t = 0\\ \varphi(o_t | \mathbf{e}_j) \sum_{i=1}^{M} \alpha_{(t-1)i} T_{ij} & t = 1, \dots, T_p \end{cases}$$
(26)

followed by the backward algorithm,

$$\beta_{ti} = \begin{cases} 1 & t = T_n \\ \sum_{j=1}^{M} T_{ij} \varphi(o_{t+1} | \mathbf{e}_j) \beta_{(t+1)j} & t = T_n - 1, \dots, 0 \end{cases}$$
(27)

The  $M \times M \times T_n$  matrix  $\Xi$  is then computed for  $t = 0, \dots, (T_n - T_n)$ 1) as

$$\xi_{tij} = \alpha_{ti} \varphi(o_{t+1}|\mathbf{e}_i) T_{ij} \beta_{(t+1)j} / \sum_{i=1}^{M} \alpha_{Ti}$$

$$\gamma_{ti} = \sum_{j=1}^{M} \xi_{tij}$$
(28)

$$\gamma_{ti} = \sum_{j=1}^{M} \xi_{tij} \tag{29}$$

In practice, the logarithms of these quantities are computed instead to avoid numerical underflow. [JDC: Not sure what