A3 Soln

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03 April, 2020

Question 1 Birth

1.

Write down statistical models corresponding to res and res2

Answer:

The statistical model for res is

$$Y_i \sim Binomial(\lambda_i, N_i)$$
$$h(\lambda_i) = log(\frac{\lambda_i}{1 - \lambda_i}) = X_i\beta + f(W_j; v) + \epsilon_i$$

Where

- Y_i is the response variable. It represents the number of babies that are males for group i.
- λ_i is the proportion of male babies in group i.
- N_i is the total number of babies in group i.
- $h(\lambda_i)$ is the logit link function.
- X_i , W_i are the covariates.
 - $-X_i$ contain indicator variable bygroup, numerical variables representing 12 months frequency: cos12, sin12, and 6 months frequency: cos6, sin6
 - $-W_i$ contain numeric variable timeInt, category variable bygroup and their interactions.
- β are the parameters.
- f(w;v) are the smoothing functions of timeInt interacting with bygroup, with smoothness parameter
- ϵ_i are residuals for group i.

The statistical model for res2 is

$$Y_{ij} \mid A_i, B_{ij} \sim Binomial(\lambda_{ij}, N_{ij})$$

$$h(\lambda_{ij}) = log(\frac{\lambda_{ij}}{1 - \lambda_{ij}}) = X_{ij}\beta + A_i + B_{ij} + f(W_{ij}; v) + \epsilon_{ij}$$

$$A_i \sim N(0, \sigma_A^2)$$

$$B_{ij} \sim N(0, \sigma_B^2)$$

Where

- Y_{ij} is the response variable. It represents the number of babies that are males for group ij.
- λ_{ij} is the proportion of male babies in group ij.
- N_{ij} is the total number of babies in group ij.
- $h(\lambda_{ij})$ is the logit link function.
- X_{ij} , W_i are the covariates.
 - $-X_i$ contain indicator variable bygroup, numerical variables representing 12 months frequency: cos12, sin12, and 6 months frequency: cos6, sin6
 - $-W_i$ contain numeric variable timeInt, category variable bygroup and their interactions.
- β are the parameters.
- A_i is the *i*th bygroup's deviation from the population average
- B_{ij} is the *i*th bygroup's *j*'s timeInt's deviation from the population average.
- f(w; v) are the smoothing functions of timeInt interacting with bygroup, with smoothness parameter v.
- ϵ_i are residuals for group i.

2.

Which of the two sets of results is more useful for investigating this research hypothesis?

Answer:

The results for res2 is more useful for investigating the hypothesis that stress induced by Trump's election is affecting the sex ratio at birth.

The difference between the models res and res2 is that res2 contains random effects of timeInt nested within bygroup giving the model random intercepts. This accounts for the grouped effect introduced by race and areas and the time.

Considering the statement that Rural whites voted for Trump in large numbers, and would presumably not be stressed by the results of the election, and Urban areas voted against Trump for the most part, and Americans of Hispanic origin had many reasons to be anxious following Trump's election, it seems appropriate to use the grouping effect to explain the variations caused by region and race.

This is also confirmed according to the prediction graphs (Figure 2: Predicted time trends). For res, the predicted lines fluctuate too much. The plot for res2 is smoother and illustrates the trend better.

3. TODO

Write a short report (a paragraph or two) addressing the following hypothesis: The long-term trend in sex ratios for urban Hispanics and rural Whites is consistent with the hypothesis that discrimination against Hispanics, while present in the full range of the dataset, has been increasing in severity over time.

Answer:

In order to address the hypothesis that discrimination against Hispanics, while present in the full range of the dataset, has been increasing in severity over time, we look at results from the model res2.

By looking at the prediction graphs (Figure 2: Predicted time trends), the predictions for res2 presents a smoother curve compared to that of res. From the graph, we see that rural Whites have a relatively flat curve while urban Hispanics has a downward trend. This indicates that over the timespan of 2007 to 2019, the ratio of male to female babies remains relatively the same for rural Whites and the ratio of male to female babies decreases for urban Hispanic. The random effects graph (Figure 3: bygroup:timeInt random effects) indicates that the variability explained by bygroup:timeInt has remained the same from 2013 to 2019.

Combining the two graphs, we conclude that the ratio of male to female babies decreases for urban Hispanic and remain roughly the same for rural Whites. The long-term trend is consistent with the hypothesis that discrimination against Hispanics, while present in the full range of the dataset, has been increasing in severity over time. Thus we agree with the hypothesis.

4. TODO

Write a short report addressing the following hypothesis: The election of Trump in November 2016 had a noticeable effect on the sex ratio of Hispanic-Americans roughly 5 months after the election.

Answer:

In order to address the hypothesis that the election of Trump in November 2016 had a noticeable effect on the sex ratio of Hispanic-Americans roughly 5 months after the election, we look at results from the model res.

res2 explains the general trend of the data, but does not explain the effect precisely to months. res explains month to month differences in more detail.

5 months after November 2016 (including November), is March 2017. From the prediction graphs (Figure 2: Predicted time trends), after the vertical line at March 2017,

Question 2

1.

Write a down the statistical model corresponding to the gamm4 calls above, explaining in words what all of the variables are.

Answer:

The model corresponding to gamItaly and gamHubei is

$$Y_i \mid A_i \sim Poisson(\lambda_i)$$

$$h(\lambda_i) = log(\lambda_i) = X_i\beta + A_i + f(W_i; v) + \epsilon_i$$

$$A_i \sim N(0, \sigma_A^2)$$

Where

- Y_i is the response variable. It represents the number of deaths for group i.
- λ_i is the mean number of deaths for group i.
- $h(\lambda_i)$ is the log link function.
- X_i , W_i are the covariates.
 - $-X_i$ contain indicator variable weekday.
 - W_i contain numeric variable timeInt.
- β are the parameters.
- A_i is the *i*th timeIid's deviation from the population average. In this case, every day has its own random intercept.
- f(w;v) are the smoothing functions of timeInt interacting with bygroup, with smoothness parameter v.
- ϵ_i are residuals for group i.

The difference between the two models are

- Y_i corresponds to death cases in different regions: gamItaly for Italy, and gamHubei for Hubei.
- f(w; v) has different number of knots. gamItaly's smoothing function has up to 40 knots where as gamHubei's smoothing function has up to 100 knots.

2.

Write a paragraph describing, in non-technical terms, what information the data analysis presented here is providing. Write text suitable for a short 'Research News' article in a University of Toronto news publication, assuming the audience knows some basic statistics but not much about non-parametric modelling.

Answer:

The log standard deviation for timeInt random effect in Italy is 0.10172. This means that each day explains 1.2256 of the variance in death cases in Italy.

The log standard deviation for timeInt random effect in Hubei is 0.41303. This means that each day explains 2.2843 of the variance in death cases in Hubei.

On a typical Friday in Italy, it is likely to have 2.7183 death cases. Comparing to Fridays, Monday to Thursday and Saturday tend to have more death cases, and Sunday tend to have less death cases.

On a typical Friday in Hubei, it is likely to have 0.2247 death cases. Comparing to Fridays, only Sunday tends to have more death cases, and the other days tend to have less death cases.

According to the prediction graphs, Italy's death cases seems to be increasing in the future, because the predicted lines and its 95% confidence interval both indicate an upward growth. Hubei's death cases seems to be decreasing in the future. However, the 95% confidence interval is not consistent with the point estimate of the prediction. This indicates that although we predict Hubei's death cases will decrease, we cannot make this claim with great confidence, and it is possible that Hubei's death cases will increase in the future.

3. TODO

Explain, for each of the tests below, whether the test is a valid LR test and give reasons for your decision.

Answer:

lmtest::1rtest(Hubei2\$mer, gamHubei\$mer) is a valid LR test because Hubei2 is nested within gamHubei.
Hubei2 is the special case when gamHubei removes the weekday covariate and uses mean instead.

nadiv::LRTest(logLik(Hubei2\$mer), logLik(gamHubei\$mer), boundaryCorrect = TRUE) is a valid LR test because Hubei2 is nested within gamHubei. Hubei2 is the special case when gamHubei removes the weekday covariate and uses mean instead. By doing so, the parameter is dropped from the full model (gamHubei) was on the boundary of its parameter space.

lmtest::1rtest(Hubei3, gamHubei\$mer) is a valid LR test because Hubei3 is nested within gamHubei. It
is equivalent to setting the random effect timeIid to 0 in gamHubei.

nadiv::LRTest(logLik(Hubei3), logLik(gamHubei\$mer), boundaryCorrect = TRUE) TODO

lmtest::lrtest(Hubei4, gamHubei\$mer) is a valid LR test because Hubei4 is nested within gamHubei. It
is equivalent to using straight lines as f(timeInt) in gamHubei.

nadiv::LRTest(logLik(Hubei4), logLik(gamHubei\$mer), boundaryCorrect = TRUE) TODO

lmtest::lrtest(Hubei2\$mer, Hubei3) is not a valid LR test because Hubei2 and Hubei3 are not nested.
Hubei2 contains a random effect of timeIid, which Hubei3 does not. Hubei3 has weekday as one of its covariates, but Hubei2 does not. Thus the two models are not nested, and the LR test is inappropriate.

nadiv::LRTest(logLik(Hubei2\$mer), logLik(Hubei3), boundaryCorrect = TRUE) is not a valid LR
test because Hubei2 and Hubei3 are not nested. Hubei2 contains a random effect of timeIid, which Hubei3

does not. Hubei3 has weekday as one of its covariates, but Hubei2 does not. Thus the two models are not nested, and the LR test is inappropriate.