A1 Soln

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Question 1

Read Data

```
# Read the crime show ratings data
crime_show_file = "crime_show_ratings.RDS"
crime_show_data = readRDS(crime_show_file)
```

Question 1.a

Let y_i denote season rating for sample i.

Let $x_{i,2000}$ be indicator variable that is set to 1 if the decade for the sample i is 2000, 0 otherwise.

Let $x_{i,2010}$ be indicator variable that is set to 1 if the decade for the sample i is 2010, 0 otherwise.

Equation for linear model:

$$y_i = \beta_0 + \beta_1 x_{i,2000} + \beta_2 x_{i,2010} + \epsilon_i$$

Anova Assumptions:

- 1. Errors (ϵ_i) are independent
- 2. Errors are normally distributed with $E[\epsilon_i] = 0$
- 3. Errors have constant variance $var[\epsilon] = \sigma^2$

Question 1.b

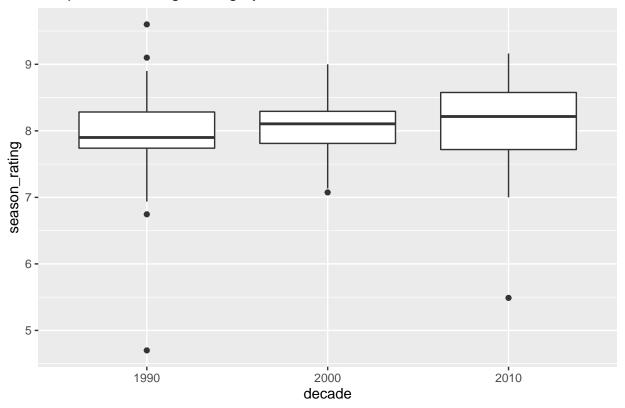
The hypotheses for ANOVA are listed and can be described as follows:

- H_0 : $\mu_{1990} = \mu_{2000} = \mu_{2010}$: The mean season rating for crime shows are the same across different decades. I.e. Different decades does not have effect over season ratings
- H_1 : at least one mean is different from the others: The mean season rating for crime shows are different across different decades. I.e. Different decades has at least some effect over season ratings

Question 1.c

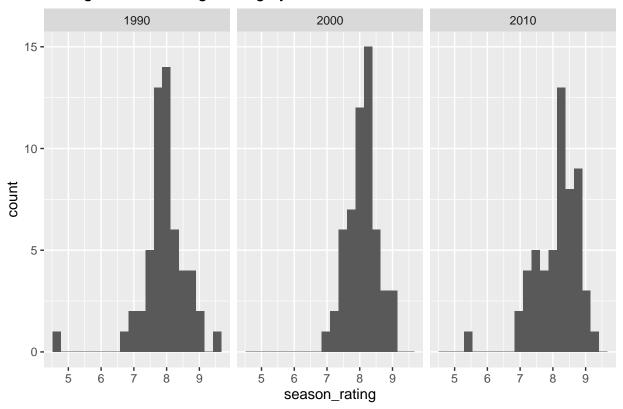
```
# Side by side box plots
crime_show_data %>% ggplot(aes(x = decade, y = season_rating)) +
    geom_boxplot() + ggtitle("Boxplots of average rating by decade for crime TV shows")
```

Boxplots of average rating by decade for crime TV shows



```
# Facetted histograms
crime_show_data %>% ggplot(aes(x = season_rating)) + geom_histogram(bins = 20) +
    facet_wrap(~decade) + ggtitle("Histograms of average rating by decade for crime TV shows")
```

Histograms of average rating by decade for crime TV shows



The box plot provides a better visualization of the data because it shows comparison across three decades' basic statistics (maximum, minimum, quartiles, median) side by side.

On the other side, with the histograms, it is harder to tell which decade has a higher rating because it only provides visualization over frequencies within each decade, and provides a relatively poor visualization for comparing between different decades.

One improvement for the box plot could be to cleaning the data before plotting. In the plot, we observe that there are some outliers, especially with 1990 and 2010. Removing those outliers may provide a even better visualization.

Accorning to the box plot, we can see that the boxes are roughly on the same level. There is no sign of extremely skewed data except for some outliers, so their means are similar to the median (all around 8). Thus it does not suggest a significant difference between the means.

Question 1.d

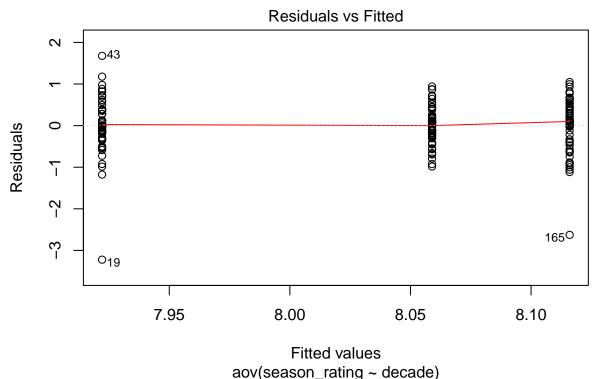
```
one_way_anova <- aov(season_rating ~ decade, data = crime_show_data)
summary(one_way_anova)</pre>
```

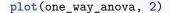
```
## Df Sum Sq Mean Sq F value Pr(>F)
## decade 2 1.09 0.5458 1.447 0.238
## Residuals 162 61.08 0.3771
```

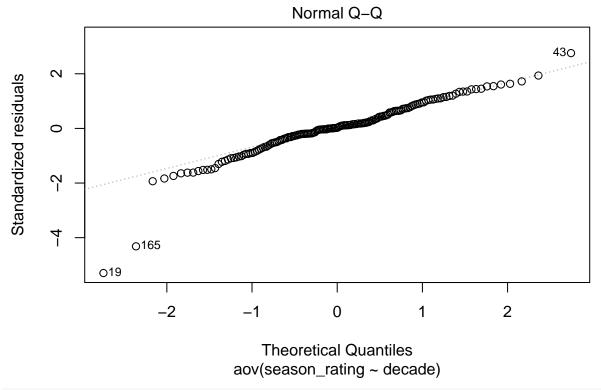
From the one way anova we can see that the p-value for the F-test is 0.238. We can interpret this as: The probability of observing the current sample given the assumption that the three decades having the same mean season ratings is 0.238.

Question 1.e

```
plot(one_way_anova, 1)
```







crime_show_data %>% group_by(decade) %>% summarise(var_rating = sd(season_rating)^2)

The first plot is the Residual vs Fitted plot. The plot shows that except for some ouliers, the residuals are roughly randomly scattered around the 0-line, and does not indicate any pattern. This shows that the data follows a linear relationship, have equal error variances, and have a few outliers.

The second plot is the normal q-q plot. From the plot, we can see that except for points 19, 165, and 43, the points form a relatively straight line, indicating that the data follows a normal distribution with a few outliers.

From the standard deviations, we calculate that the ratio of the largest within-group and biggest within-group variance estimate is $\frac{0.480}{0.203} = 2.365 < 3$. According to the rule of thumb from Dean and Voss, the assumption for equality of variances is satisfied.

Question 1.f

```
lm1 = lm(season_rating ~ decade, data = crime_show_data)
summary(lm1)
##
## Call:
## lm(formula = season_rating ~ decade, data = crime_show_data)
## Residuals:
##
       Min
                1Q
                    Median
   -3.2222 -0.2589
                    0.0135
                            0.3862
##
                                    1.6778
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                 7.9222
                            0.0828
                                    95.679
                                              <2e-16 ***
## decade2000
                 0.1368
                                     1.168
                                              0.2444
                            0.1171
## decade2010
                 0.1938
                            0.1171
                                     1.655
                                             0.0998 .
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.6141 on 162 degrees of freedom
## Multiple R-squared: 0.01756,
                                    Adjusted R-squared:
## F-statistic: 1.447 on 2 and 162 DF, p-value: 0.2382
The linear model can be expressed as
```

$$y = \beta_0 + \beta_1 x_{2000} + \beta_2 x_{2010}$$

where y is the season rating, β_i s are the coefficients, x_{2000} is the indicator variable for decade 2000, and x_{2010} is the indicator variable for decade 2010.

 β_0 is the intercept of the regression line, which is equal to the sample mean for decade 1990.

 β_1 is the amount of score increase when the indicator variable x_{2000} is set to 1.

 β_2 is the amount of score increase when the indicator variable x_{2010} is set to 1.

Then the sample mean for decade 1990 $\hat{\mu}_{1990} = \beta_0 = 7.9222$.

The sample mean for decade 2000 $\hat{\mu}_{2000} = \beta_0 + \beta_1 = 7.9222 + 0.1368 = 8.059$.

The sample mean for decade 2000 $\hat{\mu}_{2010} = \beta_0 + \beta_2 = 7.9222 + 0.1938 = 8.116$.

Question 2

Read Data

```
# Read the crime show ratings data
smokeFile = "smokeDownload.RData"
if (!file.exists(smokeFile)) {
   download.file("http://pbrown.ca/teaching/303/data/smoke.RData",
        smokeFile)
}
(load(smokeFile))
## [1] "smoke"
                      "smokeFormats"
smokeFormats[smokeFormats[, "colName"] == "chewing_tobacco_snuff_or",
   c("colName", "label")]
##
## 151 chewing_tobacco_snuff_or
                                                                                    label
## 151 RECODE: Used chewing tobacco, snuff, or dip on 1 or more days in the past 30 days
# get rid of 9, 10 year olds and missing age and race
smokeSub = smoke[which(smoke$Age > 10 & !is.na(smoke$Race)),
   ]
smokeSub$ageC = smokeSub$Age - 16
smokeModel = glm(chewing_tobacco_snuff_or ~ ageC + RuralUrban +
   Race + Sex, data = smokeSub, family = binomial(link = "logit"))
knitr::kable(summary(smokeModel)$coef, digits = 3)
```

	Estimate	Std. Error	z value	$\Pr(> z)$
(Intercept)	-2.700	0.082	-32.843	0.000
ageC	0.341	0.021	16.357	0.000
RuralUrbanRural	0.959	0.088	10.934	0.000
Raceblack	-1.557	0.172	-9.068	0.000
Racehispanic	-0.728	0.104	-6.981	0.000
Raceasian	-1.545	0.342	-4.515	0.000
Racenative	0.112	0.278	0.404	0.687
Racepacific	1.016	0.361	2.814	0.005
SexF	-1.797	0.109	-16.485	0.000

```
logOddsMat = cbind(est = smokeModel$coef, confint(smokeModel,
    level = 0.99))
oddsMat = exp(logOddsMat)
oddsMat[1, ] = oddsMat[1, ]/(1 + oddsMat[1, ])
rownames(oddsMat)[1] = "Baseline prob"
knitr::kable(oddsMat, digits = 3)
```

	est	0.5 %	99.5 %
Baseline prob	0.063	0.051	0.076
ageC	1.407	1.334	1.485
RuralUrbanRural	2.610	2.088	3.283
Raceblack	0.211	0.132	0.320
Racehispanic	0.483	0.367	0.628

-			
	est	0.5 %	99.5 %
Raceasian	0.213	0.077	0.466
Racenative	1.119	0.509	2.163
Racepacific	2.761	0.985	6.525
SexF	0.166	0.124	0.218

Question 2.a

The statistical model that corresponds to smokeModel is

$$Y_i \sim Binomial(N_i, \mu_i)$$
$$h(\mu_i) = log(\frac{\mu_i}{1 - \mu_i}) = X_i^T \beta$$

where

- μ_i is the sample mean $E(Y_i)$
- n is the population size
- $h(\mu_i)$ is the logit link function

Within X_i , the variables are composed of age centered around 16 and indicator variables for Region (Rural / Urban), Race (Black, Hispanic, Asian, Native, Pacific), and Sex (Male, Female).

Question 2.b

For the baseline prob row,

- The value under est refers to the probability of observing the subset of individuals who are 16-year-old urban white males that have used chewing tobacco, snuff, or dip on 1 or more days in the past 30 days.
- The value under 0.5% and 99.5% represents the 99% confidence interval for the estimated probability

Question 2.c TODO

```
newData = data.frame(Sex = rep(c("M", "F"), c(3, 2)), Race = c("white",
    "white", "hispanic", "black", "asian"), ageC = 0, RuralUrban = rep(c("Rural",
    "Urban"), c(1, 4)))
smokePred = as.data.frame(predict(smokeModel, newData, se.fit = TRUE,
    type = "link"))[, 1:2]
smokePred$lower = smokePred$fit - 3 * smokePred$se.fit
smokePred$upper = smokePred$fit + 3 * smokePred$se.fit
smokePred
           fit
                   se.fit
                              lower
## 1 -1.740164 0.05471340 -1.904304 -1.576024
## 2 -2.699657 0.08219855 -2.946253 -2.453062
## 3 -3.427371 0.10692198 -3.748137 -3.106605
## 4 -6.053341 0.19800963 -6.647370 -5.459312
## 5 -6.041103 0.35209311 -7.097383 -4.984824
expSmokePred = exp(smokePred[, c("fit", "lower", "upper")])
knitr::kable(cbind(newData[, -3], 1000 * expSmokePred/(1 + expSmokePred)),
    digits = 1)
```

Sex	Race	RuralUrban	fit	lower	upper
M	white	Rural	149.3	129.6	171.4
M	white	Urban	63.0	49.9	79.2
M	hispanic	Urban	31.5	23.0	42.8
\mathbf{F}	black	Urban	2.3	1.3	4.2
F	asian	Urban	2.4	0.8	6.8