

## Exercises 3.15

### Exercise 3.15.1

Let  $p$ ,  $p_i$ ,  $q$ ,  $q_i$  be density functions on  $\mathbb{R}$  and  $\alpha \in \mathbb{R}$ . Show that the cross-entropy satisfies the following properties:

- a.  $S(p_1 + p_2, q) = S(p_1, q) + S(p_2, q)$ ;
- b.  $S(\alpha p, q) = \alpha S(p, q) = S(p, q^\alpha)$ ;
- c.  $S(p, q_1 q_2) = S(p, q_1) + S(p, q_2)$ .

### Exercise 3.15.2

Show that the cross entropy satisfies the following inequality

$$S(p, q) \geq 1 - \int p(x)q(x)dx$$

### Exercise 3.15.3

Let  $p$  a fixed density. Show that the symmetric relative entropy

$$D_{KL}(p||q) + D_{KL}(q||p)$$

reaches its minimum for  $p = q$ , and the minimum is equal to zero.

### Exercise 3.15.4

Consider two exponential densities,  $p_1 = \xi^1 e^{\xi^1 x}$  and  $p_2 = \xi^2 e^{\xi^2 x}$ ,  $x \geq 0$ .

- a. Show that  $D_{KL}(p_1||p_2) = \frac{\xi^2}{\xi^1} - \ln \frac{\xi^2}{\xi^1} - 1$ .
- b. Verify  $D_{KL}(p_1||p_2) \neq D_{KL}(p_2||p_1)$ .
- c. Show that the triangle inequality doesn't hold for three arbitrary densities.

### Exercise 3.15.5

Let  $X$  be a discrete random variable. Show the inequality

$$H(X) \geq 0.$$

### Exercise 3.15.6

Prove that if  $p$  and  $q$  are the densities of two discrete random variables, then  $D_{KL}(p||q) \leq S(p, q)$

### Exercise 3.15.7

We assume the target variable  $Z$  is  $\mathcal{E}$ -measurable. What is mean squared error function in this case?

### Exercise 3.15.8

Assume that a neural network has an input-output function  $f_{w,b}$  linear in  $w$  and  $b$ . Show that the cost function (3.3.1) reaches its minimum for a unique pair  $(w^*, b^*)$ , which can be computed explicitly.

### Exercise 3.15.9

Show that the Shannon entropy can be retrieved from the Reyni entropy as

$$H(p) = \lim_{\alpha \rightarrow 1} H_\alpha(x).$$

### Exercise 3.15.10

Let  $\phi_\sigma(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}}$ . Consider the convolution operation  $(f * g)(x) := \int f(t)g(x-t)dt$ .

- Show that  $\phi_\sigma * \phi_\sigma = \phi_{\sigma\sqrt{2}}$ ;
- Find  $\phi_\sigma * \phi_{\sigma'}$  in the case  $\sigma \neq \sigma'$ .

### Exercise 3.15.11

Consider two probability densities,  $p(x)$  and  $q(x)$ . The Cauchy-Schwartz divergence is defined by

$$D_{CS}(p, q) := -\ln\left(\frac{\int p(x)q(x)dx}{\sqrt{\int p(x)^2 dx} \sqrt{\int q(x)^2 dx}}\right)$$

Show the following:

- $D_{CS}(p, q) = 0$  if and only if  $p = q$ ;
- $D_{CS}(p, q) \geq 0$ ;
- $D_{CS}(p, q) = D_{CS}(q, p)$ ;
- $D_{CS}(p, q) = -\ln \int pq dx - \frac{1}{2}H_2(p) - \frac{1}{2}H_2(q)$ , where  $H_2(\cdot)$  denotes the quadratic Reyni entropy.

### Exercise 3.15.12

- Show that for any function  $f \in L^1[0, 1]$  we have the inequality  $\|\tanh(f)\|_1 \leq \|f\|_1$ .
- Show that for any function  $f \in L^2[0, 1]$  we have the inequality  $\|\tanh(f)\|_2 \leq \|f\|_2$ .

### Exercise 3.15.13

Consider two distributions on the sample space  $\mathcal{X} = \{x_1, x_2\}$  given by

$$p = \begin{pmatrix} x_1 & x_2 \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix}, \quad q = \begin{pmatrix} x_1 & x_2 \\ \frac{1}{2} & \frac{2}{3} \end{pmatrix}$$

Consider the function  $\phi : \mathcal{X} \rightarrow \mathbb{R}^2$  defined by  $\phi(x_1) = (0, 1)$   $\phi(x_2) = (1, 0)$ . Find the maximum mean discrepancy between  $p$  and  $q$ .

## SOLUTIONS

### 3.15.1 (a)

The claim follows from the linearity of the integral operator. In symbols we have:

$$\begin{aligned} S(p_1 + p_2, q) &= - \int_{\mathbb{R}} (p_1(x) + p_2(x)) \ln q(x) dx = - \int_{\mathbb{R}} p_1(x) \ln q(x) dx - \int_{\mathbb{R}} p_2(x) \ln q(x) dx \\ &= S(p_1, q) + S(p_2, q). \end{aligned}$$

□

### 3.15.1 (b)

From the linearity of the integral operator, and the property  $c \ln(x) = \ln(x^c)$  we have:

$$\begin{aligned} S(\alpha p, q) &= - \int_{\mathbb{R}} \alpha p(x) \ln q(x) dx = -\alpha \int_{\mathbb{R}} p(x) \ln q(x) dx = \alpha S(p, q) \\ &= - \int_{\mathbb{R}} \alpha p(x) \ln q(x) dx = - \int_{\mathbb{R}} p(x) \ln q(x)^\alpha dx = S(p, q^\alpha). \end{aligned}$$

□

### 3.15.1 (c)

Using the addition identity for the logarithm, we get:

$$\begin{aligned} S(p, q_1 q_2) &= - \int_{\mathbb{R}} p(x) \ln q_1(x) q_2(x) dx = - \int_{\mathbb{R}} p(x) \ln q_1(x) dx - \int_{\mathbb{R}} p(x) \ln q_2(x) dx \\ &= S(p, q_1) + S(p, q_2). \end{aligned}$$

□

### 3.15.2

By the inequality  $\ln(x) \leq x - 1$ ,  $\forall x \in \mathbb{R}^+$ , and the definition of cross-entropy follows:

$$\begin{aligned} S(p, q) &= - \int_{\mathbb{R}} p(x) \ln q(x) dx \geq - \int_{\mathbb{R}} p(x) (q(x) - 1) dx \\ &\geq - \int_{\mathbb{R}} -p(x) dx - \int_{\mathbb{R}} p(x) q(x) dx = 1 - \int_{\mathbb{R}} p(x) q(x) dx. \end{aligned}$$

□

### 3.15.3

From proposition 3.5.1 follows that  $D_{KL}(p||q) \geq 0$ ,  $D_{KL}(q||p) \geq 0$ , then  $D_{KL}(p||q) + D_{KL}(q||p) \geq 0$ . Clearly the value 0 is a minimum. Let's now prove that this minimum is attained when  $p = q$ . It is well known from the cross-entropy definition  $S(p, p) = H(p)$  and  $S(q, q) = H(q)$  then:

$D_{KL}(p||q) = D_{KL}(p||p) = S(p, p) - H(p) = 0$  and  $D_{KL}(q||p) = D_{KL}(q||q) = S(q, q) - H(q) = 0$ , which in turn imply  $D_{KL}(p||q) + D_{KL}(q||p) = 0$ . □

### 3.15.4 (a)

By direct calculation we find:

$$\begin{aligned}
D_{KL}(p_1 \| p_2) &= S(p_1, p_2) - H(p_1) = - \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \ln(\xi^2 e^{-\xi^2 x}) dx - \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \ln(\xi^1 e^{-\xi^1 x}) \\
&= - \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \ln(\xi^2) dx + \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \xi^2 x dx + \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \ln(\xi^1) dx - \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} \xi^1 x dx \\
&= -(\ln(\xi^2) - \ln(\xi^1)) \int_{\mathbb{R}} \xi^1 e^{-\xi^1 x} dx + (\xi^2 - \xi^1) \int_{\mathbb{R}} \xi^1 x e^{-\xi^1 x} dx \\
&= -(\ln(\xi^2) - \ln(\xi^1)) \mathbb{E}_{X \sim \exp(\xi^1)} [1] + (\xi^2 - \xi^1) \mathbb{E}_{X \sim \exp(\xi^1)} [X] = -\ln \frac{\xi^2}{\xi^1} + (\xi^2 - \xi^1) \frac{1}{\xi^1} \\
&= -\ln \frac{\xi^2}{\xi^1} + \frac{\xi^2}{\xi^1} - 1
\end{aligned}$$

□

### 3.15.4 (b)

Suppose the equality  $D_{KL}(p \| p) = D_{KL}(q \| p)$  holds and  $\xi^1 \neq \xi^2$ , then from exercise 3.14.4.a it follows:  
 $-\ln \frac{\xi^2}{\xi^1} + \frac{\xi^2}{\xi^1} - 1 = -\ln \frac{\xi^1}{\xi^2} + \frac{\xi^1}{\xi^2} - 1 \implies \frac{\xi^2}{\xi^1} = \frac{\xi^1}{\xi^2}$ . The later implies  $\frac{\xi^1}{\xi^2} = 1$  or equivalently  $\xi^1 = \xi^2$ , which is a contradiction.

### 3.15.4 (c)

Let  $p_1 = \exp(2)$ ,  $p_2 = \exp(3)$ ,  $p_3 = \exp(4)$ . Suppose the triangle inequality holds for these three arbitrary exponential distributions. This is:

$D_{KL}(p_1 \| p_3) \leq D_{KL}(p_1 \| p_2) + D_{KL}(p_2 \| p_3)$ . By exercise 3.15.4.b we would have:

$$\begin{aligned}
D_{KL}(p_1 \| p_3) &= \frac{4}{2} - \ln \frac{4}{2} - 1 \leq D_{KL}(p_1 \| p_2) + D_{KL}(p_2 \| p_3) = \frac{3}{2} - \ln \frac{3}{2} - 1 + \frac{4}{3} - \ln \frac{4}{3} - 1 \\
2 &\leq \frac{3}{2} + \frac{4}{3} - 1 = \frac{17}{6} - 1 = \frac{11}{6} = \frac{12}{6} - \frac{1}{6} = 2 - \frac{1}{6} \text{ (contradiction!)}
\end{aligned}$$

□

### 3.15.5 (a)

Given that  $p(x)$  is a distribution, it follows that  $p(X)$  as a r.v satisfies the inequality  $0 \leq p(X) \leq 1$ . This means  $p(x) \leq 1, \forall x \in \text{sup}(X)$ . Taking natural logs on both sides of the inequality  $p(x) \leq 1$  and multiplying by  $-1$ , we obtain:  $\ln p(x) \geq 0$ ; Multiplying by  $p(x)$  and summing over the support of  $X$ , we get:

$$\mathbb{E}[-\ln p(X)] = H(X) = \sum_{x \in \text{sup}(X)} -p(x) \ln(p(x)) \geq 0.$$

□

### 3.15.6

This is an immediate consequence of exercise 3.15.5. Indeed, we have:

$$D_{KL}(p \| q) = S(p, q) - H(p) \leq S(p, q) - 0 \leq S(p, q).$$

□

### 3.15.7

If the target variable  $Z$  happens to be  $\mathcal{E}$ -measurable, then  $Y$  is independent of the sigma algebra  $\mathcal{E}$ . From this follows that  $C(\omega, b) = d(Z, Y)^2 = \mathbb{E}[(Z - \mathbb{E}[Z|\mathcal{E}])^2] = \mathbb{E}[(Z - Z)^2] = 0$ .

### 3.15.8

In this case  $f_{\omega, b}(\mathbf{x}) = \omega \cdot \mathbf{x} + b$ , defined on a compact subset of  $\mathbb{R}^n$ . Therefore, the cost function is given by:

$C(\omega, b) := \sum_{0 \leq i \leq n} (\omega \cdot \mathbf{x}^i + b - \phi(\mathbf{x}^i))^2$ . Obviously we have  $0 \leq C(\omega, b)$ . This means the function attains such

minimum inside the compact set; Let  $\mathbf{x}^i$  the  $n$ -dimensional observations, i.e  $\mathbf{x}^i = (x_1^i, \dots, x_n^i)$ . Then, the normal equations for the  $\omega_k$  (the components of the vector  $\omega$ ),  $\forall k \in [n]$  and the bias parameter  $b$  are:

$$\begin{cases} \sum_{0 \leq j \leq n} \omega_j \sum_{0 \leq i \leq n} x_j^i x_k^i + b \sum_{0 \leq i \leq n} x_k^i = \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) x_k^i, \forall k \in [n] \\ \sum_{0 \leq j \leq n} \omega_j \sum_{0 \leq i \leq n} x_j^i + nb = \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) \end{cases} \quad (1)$$

This system of equations has the following matricial expression:

$$\begin{bmatrix} \sum_{0 \leq i \leq n} x_1^i x_1^i & \sum_{0 \leq i \leq n} x_2^i x_1^i & \cdots & \sum_{0 \leq i \leq n} x_m^i x_1^i & \cdots & \sum_{0 \leq i \leq n} x_1^i \\ \sum_{0 \leq i \leq n} x_1^i x_2^i & \sum_{0 \leq i \leq n} x_2^i x_2^i & \cdots & \sum_{0 \leq i \leq n} x_m^i x_2^i & \cdots & \sum_{0 \leq i \leq n} x_2^i \\ \vdots & \vdots & & \vdots & & \vdots \\ \sum_{0 \leq i \leq n} x_1^i x_k^i & \sum_{0 \leq i \leq n} x_2^i x_k^i & \cdots & \sum_{0 \leq i \leq n} x_m^i x_k^i & \cdots & \sum_{0 \leq i \leq n} x_k^i \\ \vdots & \vdots & & \vdots & & \vdots \\ \sum_{0 \leq i \leq n} x_1^i & \sum_{0 \leq i \leq n} x_2^i & \cdots & \sum_{0 \leq i \leq n} x_m^i & \cdots & n \end{bmatrix} \begin{bmatrix} \omega_1 \\ \omega_2 \\ \vdots \\ \omega_m \\ \vdots \\ b \end{bmatrix} = \begin{bmatrix} \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) x_1^i \\ \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) x_2^i \\ \vdots \\ \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) x_m^i \\ \vdots \\ \sum_{0 \leq i \leq n} \phi(\mathbf{x}^i) \end{bmatrix} \quad (2)$$

Note that the entries of the matrix in equation 2, are exactly the partial derivatives  $\partial_{\omega_k \omega_k}^2 C(\omega, b)$ ,  $\partial_{\omega_k b}^2 C(\omega, b)$ ,  $\partial_{\omega_k \omega_j}^2 C(\omega, b)$ ,  $\partial_{\omega_k b}^2 C(\omega, b)$  i.e such matrix is the hessian-matrix  $\mathcal{H}_{C(\omega, b)}$ ; Let  $\{v_m\}_{m \in [n+1]}$  be a collection of vectors in  $\mathbb{R}^n$  defined as follows:  $\forall m, 0 \leq m \leq n, v_m := (x_m^1, \dots, x_m^n)$ . For  $m = n+1$  we define:  $v_{n+1} := \mathbf{1} = (1, \dots, 1)$ , and a vector  $\varphi := (\phi(\mathbf{x}^1), \dots, \phi(\mathbf{x}^n))$ . The system of equations can be written as:

$$\begin{bmatrix} v_1 \cdot v_1 & v_1 \cdot v_2 & \cdots & v_1 \cdot v_m & \cdots & v_1 \cdot v_{n+1} \\ v_1 \cdot v_2 & v_2 \cdot v_2 & \cdots & v_2 \cdot v_m & \cdots & v_2 \cdot v_{n+1} \\ \vdots & \vdots & & \vdots & & \vdots \\ v_m \cdot v_1 & v_m \cdot v_2 & \cdots & v_m \cdot v_m & \cdots & v_m \cdot v_{n+1} \\ \vdots & \vdots & & \vdots & & \vdots \\ v_{n+1} \cdot v_1 & v_{n+1} \cdot v_2 & \cdots & v_{n+1} \cdot v_m & \cdots & v_{n+1} \cdot v_{n+1} \end{bmatrix} \begin{bmatrix} \omega_1 \\ \omega_2 \\ \vdots \\ \omega_m \\ \vdots \\ b \end{bmatrix} = \begin{bmatrix} \varphi \cdot v_1 \\ \varphi \cdot v_2 \\ \vdots \\ \varphi \cdot v_m \\ \vdots \\ \varphi \cdot v_{n+1} \end{bmatrix} \quad (3)$$

This is  $\mathcal{H}_{C(\omega, b)} = G(v_1, \dots, v_{n+1})$  i.e the matrix in system 3 is a Gramm matrix, which by the postively semidefiniteness property of the Gramm matrices implies this matrix is postively defined. Then, the solution  $(\omega^*, b^*)$  to such system is unique and this solution gives indeed a minimum for  $C(\omega, b)$ . Furthermore, the values of the pair  $(\omega^*, b^*)$  are explicitly computable because the system is linear. □

### 3.15.9

Note that the Reyni entropy can be expressed as follows:  $H_\alpha(p(X)) = \frac{\ln \mathbb{E}[(p(X))^{\alpha-1}]}{1-\alpha} = \frac{\ln \int_{\sup(X)} (p(x))^{\alpha-1} dP}{\alpha-1}$ . In the last expression  $\frac{dP}{dx} = p(x)$ ; This representation is a consequence of the Radon-Nikodym's Theorem. Now, let's analyse the following function defined as a parametric integral. Let  $I(\alpha)$  defined as:

$$I(\alpha) = \int_{\sup(X)} (p(x))^{\alpha-1} dP$$

- CASE  $0 < p(x) < 1$ :

Let  $\{\alpha_n\}_{n \in \mathbb{N}} \subset \mathbb{R}$  a monotonically decreasing sequence of reals tending to 1. With this we can construct an increasing sequence of P-integrable functions of the form  $\{(p(x))^{\alpha_n-1}\}_{n \in \mathbb{N}}$ . Then, by construction the point limit will be the function  $f(x) = 1$  which is P-integrable. Applying the Lebesgue's monotone convergence theorem:

$$\lim_{k \rightarrow \infty} I(\alpha_k) = \lim_{k \rightarrow \infty} \int_{\sup(X)} (p(x))^{\alpha_k-1} dP = \int_{\sup(X)} \lim_{k \rightarrow \infty} (p(x))^{\alpha_k-1} dP = I(\lim_{k \rightarrow \infty} \alpha_k) = \lim_{\alpha \rightarrow 1^+} I(\alpha) =$$

1. This proves  $I(\alpha)$  is right-continuous

Taking  $\{\alpha_n\}_{n \in \mathbb{N}} \subset \mathbb{R}$  a monotonically increasing sequence of reals tending to 1, let's construct the sequence of decreasing P-integrable functions of the form  $\{(p(x))^{\alpha_n-1}\}_{n \in \mathbb{N}}$ . Once again the point limit function is  $f(x) = 1$ . Applying the monotone convergence

$$\lim_{k \rightarrow \infty} I(\alpha_k) = \lim_{k \rightarrow \infty} \int_{\sup(X)} (p(x))^{\alpha_k-1} dP = \int_{\sup(X)} \lim_{k \rightarrow \infty} (p(x))^{\alpha_k-1} dP = I(\lim_{k \rightarrow \infty} \alpha_k) = \lim_{\alpha \rightarrow 1^-} I(\alpha) =$$

1. This proves  $I(\alpha)$  is left-continuous.

- CASE  $p(x) > 1$ :

The same is true in this case. It is worth to note that if the sequence  $\{\alpha_n\}_{n \in \mathbb{N}}$  is chosen to be decreasing then the sequence of functions is decreasing, the contrary situation also holds.

Because the function  $f(\alpha, x) = (p(x))^{\alpha-1}$  is derivable, and its derivative is continuous (this can be proved by the monotone convergence theorem), the above can be then used to compute the limit  $\lim_{\alpha \rightarrow 1} H_\alpha(p(X))$ . By a simple application  $\alpha \rightarrow 1$  we find an indetermination of the type  $0/0$ ; Using L'Hôpital's rule we get:

$$\lim_{\alpha \rightarrow 1} H_\alpha(p(X)) = \lim_{\alpha \rightarrow 1} \frac{1}{\alpha-1} \frac{1}{\int_{\sup(X)} (p(x))^{\alpha-1} dP} \int_{\sup(X)} (p(x))^{\alpha-1} \ln(p(x)) dP = \int_{\sup(X)} -\ln(p(x)) dP = H(p).$$

□

### 3.15.10 (a)

This is a consequence of exercise 3.15.9.b. By taking  $\sigma_1 = \sigma_2 = \sigma$  we get that  $\varphi_{\sigma_1}(x) \star \varphi_{\sigma_2}(x) = \varphi_{\sigma'}(x)$ , with  $\sigma' = \sqrt{2\sigma^2}$

### 3.15.10 (b)

We have  $\phi_\sigma(x) := \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}}$ ; Computing  $(\phi_{\sigma_1} \star \phi_{\sigma_2})(x) = \int_{\mathbb{R}} \phi_{\sigma_1}(t) \phi_{\sigma_2}(t-x) dt$  we get:

$$\begin{aligned}
 (\phi_{\sigma_1} \star \phi_{\sigma_2})(x) &= \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi\sigma_1^2\sigma_2^2}\sqrt{2\pi}} e^{-\frac{t^2\sigma_2^2 - (t-x)\sigma_1^2}{2\sigma_1^2\sigma_2^2}} dt = \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi\sigma_1^2\sigma_2^2}\sqrt{2\pi}} e^{-\frac{t^2(\sigma_1^2 + \sigma_2^2) + 2tx\sigma_1^2 - x^2\sigma_1^2}{2\sigma_1^2\sigma_2^2}} dt \\
 &= \int_{\mathbb{R}} \frac{\sqrt{\sigma_1^2 + \sigma_2^2}}{\sqrt{2\pi\sigma_1^2\sigma_2^2}\sqrt{2\pi}\sqrt{\sigma_1^2 + \sigma_2^2}} e^{-(\sigma_1^2 + \sigma_2^2) \frac{t^2 - 2tx\frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2} + \frac{x^2(\sigma_1^2)^2}{(\sigma_1^2 + \sigma_2^2)^2} + x^2\sigma_1^2 - \frac{x^2(\sigma_1^2)^2}{(\sigma_1^2 + \sigma_2^2)^2}}}{2\sigma_1^2\sigma_2^2}} dt \\
 &= \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi(\sigma_1^2 + \sigma_2^2)}} e^{-\frac{x^2}{2(\sigma_1^2 + \sigma_2^2)}} \frac{1}{\sqrt{2\pi\frac{\sigma_1^2\sigma_2^2}{\sigma_1^2 + \sigma_2^2}}} e^{-\frac{(t - x\frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2})^2}{2\frac{\sigma_1^2\sigma_2^2}{\sigma_1^2 + \sigma_2^2}}} dt \\
 &= \frac{1}{\sqrt{2\pi(\sigma_1^2 + \sigma_2^2)}} e^{-\frac{x^2}{2(\sigma_1^2 + \sigma_2^2)}} \int_{\mathbb{R}} \frac{1}{\sqrt{2\pi\frac{\sigma_1^2\sigma_2^2}{\sigma_1^2 + \sigma_2^2}}} e^{-\frac{(t - x\frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2})^2}{2\frac{\sigma_1^2\sigma_2^2}{\sigma_1^2 + \sigma_2^2}}} dt \\
 &= \frac{1}{\sqrt{2\pi(\sigma_1^2 + \sigma_2^2)}} e^{-\frac{x^2}{2(\sigma_1^2 + \sigma_2^2)}} \cdot 1 = \phi_{\sigma'}(x), \text{ with } (\sigma')^2 = \sigma_1^2 + \sigma_2^2
 \end{aligned}$$

□

### 3.15.11 (a)

- CASE  $p = q$ :

Making the substitution in the formula for the Cauchy-Schwartz- divergence, we find:

$$D_{CS}(p, p) = -\ln\left(\frac{\int p(x)p(x)dx}{\sqrt{\int p(x)^2dx}\sqrt{\int p(x)^2dx}}\right) = -\ln(1) = 0.$$

- CASE  $D_{CS}(p, q) = 0$ :

If  $D_{CS}(p, q) = 0$ , then applying exponentials to both sides we obtain:

$$\frac{\int p(x)q(x)dx}{\sqrt{\int p(x)^2dx}\sqrt{\int q(x)^2dx}} = 1 \implies \|\int p(x)q(x)dx\| = \sqrt{\int p(x)^2dx}\sqrt{\int q(x)^2dx}. \text{ This is the case of}$$

equality in the Cauchy-Bunyakovsky-Schwarz inequality. Then:  $q = \lambda p, \lambda \neq 0$ ; Let's now prove  $\lambda = 1$ . Plugging this we get:

$$\frac{\lambda \int p(x)p(x)dx}{\lambda^2 \sqrt{\int p(x)^2dx}\sqrt{\int p(x)^2dx}} = 1 = \frac{\lambda}{\lambda^2} \implies \lambda = 1 \implies p = q.$$

□



### 3.15.11 (b)

By hypothesis both  $p, q$  are densities, then they are both nonnegative. This implies:

$$\int p(x)q(x)dx = \int |p(x)q(x)|dx \leq \|p\|_2^2 \|q\|_2^2 \implies 0 < \frac{\int p(x)q(x)dx}{\|p\|_2^2 \|q\|_2^2} \leq 1; \text{ Remember that } \ln(x) \leq 0 \text{ if } 0 < x < 1 \text{ then, taking logarithms to both sides in the inequality and multiplying by } -1 \text{ yields:}$$

$$D_{CS}(p, q) = -\ln\left(\frac{\int p(x)q(x)dx}{\sqrt{\int p(x)^2 dx} \sqrt{\int q(x)^2 dx}}\right) \geq 0 \quad \square$$

### 3.15.11 (c)

This property follows easily from the definition of the CS-divergence.

### 3.15.11 (d)

Applying the properties of the logarithms we obtain:

$$\begin{aligned} D_{CS}(p, q) &= -\ln\left(\frac{\int p(x)q(x)dx}{\sqrt{\int p(x)^2 dx} \sqrt{\int q(x)^2 dx}}\right) = -\ln \int p(x)q(x)dx - \frac{1}{2} \ln \int p(x)^2 dx - \frac{1}{2} \ln \int q(x)^2 dx \\ &= -\ln \int p(x)q(x)dx - \frac{1}{2} H_2(p) - \frac{1}{2} H_2(q) \end{aligned}$$

□

### 3.15.12 (a)

The claim follows from the inequality  $|u| \geq |\tanh(u)|$ , then making the substitution  $u = f(x)$ , with  $f \in \mathcal{L}_1[0, 1]$  we get  $|f(x)| \geq |\tanh(f(x))|$ . On the other hand, the function  $\tanh$  is derivable  $\forall u \in \mathbb{R}$ . Furthermore it's derivative is given by the identity:  $\frac{d}{du} \tanh(u) = 1 - \tanh^2(u), \forall u \in \mathbb{R}$ . Observe that  $\frac{d}{du} \tanh(u) \leq 1, \forall u \in \mathbb{R}$ . Applying the Intermediate Value Theorem yields:

- CASE  $u > 0$

$$\begin{aligned} 0 < \frac{\tanh(u) - \tanh(0)}{u - 0} &= \frac{\tanh(u)}{u} = \frac{d}{du} \tanh(u) \Big|_{u=\theta}, \text{ where } \theta \in (0, u) \\ &= 1 - \tanh^2(\theta) \leq 1 \implies \tanh(u) \leq u = |u|, \forall u > 0 \end{aligned}$$

- CASE  $u < 0$

$$\begin{aligned} 0 < \frac{\tanh(u) - \tanh(0)}{u - 0} &= \frac{\tanh(u)}{u} = \frac{d}{du} \tanh(u) \Big|_{u=\theta}, \text{ where } \theta \in (u, 0) \\ &= -1 + \tanh^2(\theta) \leq 1 \implies \tanh(u) \geq u = -|u|, \forall u < 0 \end{aligned}$$

Both cases can be summarized as  $|u| \geq |\tanh(u)|$ . If  $u = f(x)$  this implies  $|f(x)| \geq |\tanh(f(x))|$ . Then, by the monotonicity of the integral operator we get the desired result.

□

### 3.15.12 (b)

In 3.15.12.a was proved the inequality  $|u| \geq |\tanh(u)|$ ,  $\forall u \in \mathbb{R}$  so if  $f \in \mathcal{L}_2[0, 1]$ . Then,  $|f(x)| \geq |\tanh(f(x))|$ ; Taking squares the latter inequality yields:  $|f(x)|^2 \geq |\tanh(f(x))|^2$  and once again, due to the monotonicity of the integral operator we get the desired inequality.