

# CTMC

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# 1. Course Notes

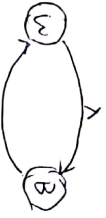
# CTMC

## - Model

Random Process  $\{X(t), t \in [0, \infty)\}$  having countable states  $S \subset \{0, 1, 2, \dots\}$  and fulfilling Markovian property.

When the system reaches state  $z$ , it stays there for a random amount of time  $T_z$ , where  $T_z$  is an exponential random variable.

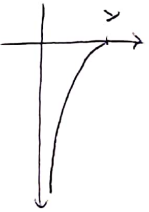
## - Representation



## - Mathematical Background

exponential distribution

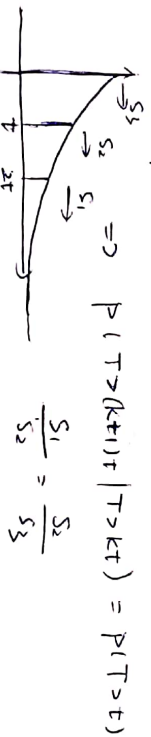
$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$



<1> Property:

① Memoryless

$$P(T > s+t | T > s) = P(T > t)$$



∴ If we sample in discrete way:  $t, 2t, \dots, kt$ .

It just looks the same as DTMC!

② Much devices in real life in fact exhibit constant failure rates

③ The exponential distribution can model the time-to-failure for complex systems consisting of many components in series (but none of the individual components contribute significantly to the total failure density)

## - Steady State of CTMC

### • Method 1. Building upon DTMCs

① Steady state of DTMC:  $\hat{\pi} = [\hat{\pi}_0, \dots, \hat{\pi}_m]$

$$\textcircled{2} \text{ Assume } 0 < \sum_{k \in S} \frac{\hat{\pi}_k}{\lambda_k} < \infty$$

means  $\lambda_k \neq 0$

recurrent

$$\textcircled{3} \pi_j = \lim_{t \rightarrow \infty} P(X(t) = j | X(0) = z)$$

$$= \frac{\hat{\pi}_j}{\sum_{k \in S} \frac{\hat{\pi}_k}{\lambda_k}}$$

★ Method 2. State-Space Based:

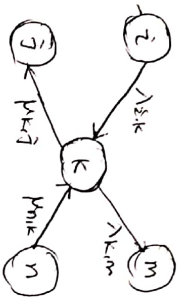
Assume: non-homogeneous CTMC



$$\frac{dP_0}{dt} = -P_0 \cdot \lambda$$

→ "rate parameter"

(1) Steady State flow



$$\frac{dP_k}{dt} = \sum_i \lambda_{ik} P_i + \sum_n \mu_{nk} P_n - \left( \sum_j \mu_{kj} + \sum_m \lambda_{km} \right) P_k$$

★ (2) Steady State Condition

$$\frac{dP_k}{dt} = 0 \quad \text{for each } k$$

(3) Generation Matrix

$$q_{ij} = \begin{cases} \lambda_{ij} & \text{if } i \neq j \\ -\lambda_i & \text{if } i = j \end{cases}$$

$$\therefore \frac{dP}{dt} = 0 \Leftrightarrow \Pi \cdot G = 0$$

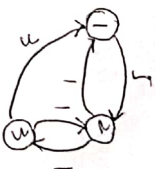
- Irreducibility and Aperiodicity

(1) A CTMC irreducible

⇒ steady state exists

Do not need aperiodical

- Example



$$G = \begin{bmatrix} -5 & 5 & 0 \\ 1 & -2 & 1 \\ 3 & 1 & -4 \end{bmatrix}$$

$$\therefore \Pi \cdot G = 0 \quad \& \quad \sum \Pi = 1$$

$$\therefore [\Pi_1, \Pi_2, \Pi_3] \begin{bmatrix} -5 & 5 & 0 \\ 1 & -2 & 1 \\ 3 & 1 & -4 \end{bmatrix} = [0, 0, 0]$$

$$\therefore \begin{bmatrix} -5 & 1 & 3 \\ 5 & -2 & 1 \\ 0 & 1 & -4 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 1 & 1 \\ 0 & 6 & 5 \\ 0 & -7 & -5 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 1 & 1 \\ 0 & 1 & -4 \\ 0 & 1 & -4 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 0 & 5 \\ 0 & 1 & -4 \\ 0 & 0 & 0 \end{bmatrix}$$

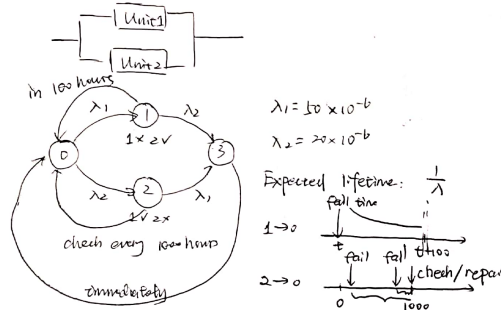
$$\rightarrow \begin{bmatrix} 1 & 0 & 5 \\ 0 & 1 & -4 \\ 0 & 0 & 32 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 0 & 5 \\ 0 & 1 & -4 \\ 0 & 0 & 32 \end{bmatrix} \rightarrow \begin{bmatrix} 1 & 0 & 5 \\ 0 & 1 & -4 \\ 0 & 0 & 0 \end{bmatrix}$$

$$\Pi_3 = \frac{1}{32}, \quad \Pi_2 = \frac{24}{32}, \quad \Pi_1 = \frac{7}{32}$$

### A more realistic scenario

- Recap

- Model: A two-unit system



① Approximation :

for state 3  $\rightarrow$  state 0: a large  $\mu$  (much larger than  $\lambda_1, \lambda_2$ )

for state 1  $\rightarrow$  state 0:  $\left\{ \begin{array}{l} \text{constant } \frac{1}{100} \text{ per hour} \\ \text{exponential with mean 100 hours} \end{array} \right.$  so small!

for state 2  $\rightarrow$  state 0: exponential with mean 100 hours

- Transient Analysis :

<1> evaluate the instantaneous effect of periodic repair

for  $p_2 \rightarrow p_0$



## 2. Introduction

### 2.1. Definition

A continuous-time Markov chain  $X(t)$  is defined by two components: a **jump chain**, and a set of **holding time parameters**  $\lambda_i$ . The jump chain consists of a countable set of states  $S \subset 0, 1, 2, \dots$  along with transition probabilities  $p_{ij}$ . We assume  $p_{ii}=0$ , for all non-absorbing states  $i \in S$ . We assume

1. if  $X(t) = i$ , the time until the state changes has  $\text{Exponential}(\lambda_i)$  distribution
2. if  $X(t) = i$ , the next state will be  $j$  with probability  $p_{ij}$

### 2.2. Inherent Property

Because of **Exponential Distribution**, The process inherently satisfies the Markov Property because of the **Memoryless** Property of Exponential Distribution.

for all  $0 \leq t_1 < t_2 < \dots < t_n < t_{n+1}$ , we have

$$\begin{aligned} P(X(t_{n+1}) = j \mid X(t_n) = i, X(t_{n-1}) = i_{n-1}, \dots, X(t_1) = i_1) \\ = P(X(t_{n+1}) = j \mid X(t_n) = i) \end{aligned}$$

## 2.3. Transition Matrix

For a continuous-time Markov chain, we define the transition matrix  $P(t)$ . The  $(i, j)$ th entry of the transition matrix is given by

$$P_{ij}(t) = P(X(t) = j \mid X(0) = i)$$

So the transition matrix satisfies the following properties:

1.  $P(0)$  is equal to the identity matrix,  $P(0) = I$
2. the rows of the transition matrix must sum to 1

$$\sum_{j \in S} p_{ij}(t) = 1, \quad \text{for all } t \geq 0$$

3. for all  $s, t \geq 0$ :

$$P(s + t) = P(s)P(t)$$

## 3. Stationary and Limiting Distributions

### 3.1. Stationary Distribution

Let  $X(t)$  be a continuous-time Markov chain with transition matrix  $P(t)$  and state space  $S = 0, 1, 2, \dots$ . A probability distribution  $\pi$  on  $S$ , i.e, a vector  $\pi = [\pi_0, \pi_1, \pi_2, \dots]$ , where  $\pi_i \in [0, 1]$  and

$$\sum_{i \in S} \pi_i = 1$$

is said to be a **stationary distribution** for  $X(t)$  if

$$\pi = \pi P(t), \text{ for all } t \geq 0$$

If a CTMC is irreducible  $\Rightarrow$  Steady/Staionary Distributed exists

## 3.2. Limiting Distributions

The probability distribution  $\pi = [\pi_0, \pi_1, \pi_2, \dots]$  is called the limiting distribution of the continuous-time Markov chain  $X(t)$  if for all  $i, j \in S$

$$\pi_j = \lim_{t \rightarrow \infty} P(X(t) = j \mid X(0) = i)$$

and we have

$$\sum_{j \in S} \pi_j = 1$$

## 3.3. Method to obtain Limiting distributions

Let  $X(t), t \geq 0$  be a continuous-time Markov chain with an irreducible positive recurrent jump chain. Suppose that the unique stationary distribution of the jump chain is given by

$$\tilde{\pi} = [\tilde{\pi}_0, \tilde{\pi}_1, \tilde{\pi}_2, \dots].$$

Further assume that (means **no recurrent state**)

$$0 < \sum_{k \in S} \frac{\tilde{\pi}_k}{\lambda_k} < \infty.$$

Then

$$\pi_j = \lim_{t \rightarrow \infty} P(X(t) = j \mid X(0) = i) = \frac{\frac{\tilde{\pi}_j}{\lambda_j}}{\sum_{k \in S} \frac{\tilde{\pi}_k}{\lambda_k}}.$$

for all  $i, j \in S$ . That is,  $\pi = [\pi_0, \pi_1, \pi_2, \dots]$  is the **limiting distribution** of  $X(t)$

.

## 4. Generation Matrix

### 4.1. Definition

For a continuous-time Markov chain, we define the **generator matrix**  $G$ . The  $(i, j)$ th entry of the transition matrix is given by

$$g_{ij} = \begin{cases} \lambda_i p_{ij} & \text{if } i \neq j \\ -\lambda_i & \text{if } i = j \end{cases}$$

### 4.2. Forward and Backward Equations

Forward and Backward Equations describes the relations between generation matrix and transition matrix. It is apparent through the derivative of exponential distribution.

The **forward equations** state that

$$P'(t) = P(t)G,$$

which is equivalent to

$$p'_{ij}(t) = \sum_{k \in S} g_{ik} p_{kj}(t), \text{ for all } i, j \in S.$$

The **backward equations** state that

$$P'(t) = GP(t),$$

which is equivalent to

$$p'_{ij}(t) = \sum_{k \in S} g_{ik} p_{kj}(t), \text{ for all } i, j \in S.$$

### 4.3. The second method to Obtain Stationary Distribution

Consider a continuous Markov chain  $X(t)$  with the state space  $S$  and the generator Matrix  $G$ . The probability distribution  $\pi$  on  $S$  is a stationary distribution for  $X(t)$  if and only if it satisfies

$$\pi G = 0.$$

That means for each state: **flow in= flow out**

## 4.4. Transition Rate Diagram

A continuous-time Markov chain can be shown by its **transition rate diagram**. In this diagram:

1. the values  $g_{ij}$  are shown on the edges. The values of  $g_{ii}$  are not usually shown because they are implied by the other values, i.e.

$$g_{ii} = - \sum_{j \neq i} g_{ij}.$$