

Panel Regression Analysis: Micro application to production function

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Example 2. Micro-level application of panel regression analysis

RiceFarms

- Are larger farms more productive if compared to the smaller ones?
- This example:
 - explores relationship between farm size and productivity.
 - guides through the process of panel regression analysis.

Part 1. Theoretical basis

We employ the Cobb-Douglas Production function:

$$\ln y = \ln \beta_0 + \sum_{n=1}^N \beta_n \ln x_n + \sum_{k=1}^K \gamma_k \delta_k + \epsilon$$

where,

- y is the output and x_n are the inputs all in physical mass (or monetary value);
- N is the number of independent variables;
- δ_k are the shift parameters of additional dummy variables;
- β_0 , β_n , γ_n are the estimated coefficients;

Estimation strategy. Part 1.

Pooled OLS production function:

$$\begin{aligned}\ln \text{output}_{it} = & A_0 + \beta_1 \cdot \ln \text{land}_{it} + \beta_2 \cdot \ln \text{labor}_{it} \\ & + \beta_3 \cdot \ln \text{seed}_{it} + \beta_4 \cdot \ln \text{urea}_{it} \\ & + \beta_5 \cdot \ln \text{pesticide}_{it} + e_{it}\end{aligned}$$

- What are the ex-ante expectations about the regression coefficients?
 - Ideas? ...
 - Probably all β should be positive.

What about OVB?

$$\begin{aligned}\ln \text{output}_{it} = & A_0 + \beta_1 \cdot \ln \text{land}_{it} + \beta_2 \cdot \ln \text{labor}_{it} \\ & + \beta_3 \cdot \ln \text{seed}_{it} + \beta_4 \cdot \ln \text{urea}_{it} \\ & + \beta_5 \cdot \ln \text{pesticide}_{it} + e_{it}\end{aligned}$$

What omitted variables could cause bias of our estimates?

- Any!? ...
- Any!? ...
- Any OVB!? ...
- Capital, Ability, Climate, Geography...

Using the OVB formula

- Let us make an educated guess about the effect of OVB on the estimates of land-related coefficient β_1 ?
- Short: • $\text{output}_{it} = A_0^s + \beta_1^s \cdot \text{land}_{it} + \beta_2^s \cdot \text{labor}_{it} + \beta_3^s \cdot \text{seed}_{it} + \beta_4^s \cdot \text{urea}_{it} + \beta_5^s \cdot \text{pesticide}_{it} + e_{it}$
- Long: • $\text{output}_{it} = A_0 + \beta_1 \cdot \text{land}_{it} + \beta_2 \cdot \text{labor}_{it} + \beta_3 \cdot \text{seed}_{it} + \beta_4 \cdot \text{urea}_{it} + \beta_5 \cdot \text{pesticide}_{it} + \gamma$
- Aux. • $\text{Ability}_i = \pi_0 + \pi_1 \cdot \text{land}_{it} + \pi_2 \cdot \text{labor}_{it} + \pi_3 \cdot \text{seed}_{it} + \pi_4 \cdot \text{urea}_{it} + \pi_5 \cdot \text{pesticide}_{it} + e_i$
- OVB: • $\text{OVB}_{\text{land}} = \beta_1^s - \beta_1 = \pi_1 * \gamma$
- Educated guess about the bias of the estimates.
- What are the signs of π_1 and γ ?
 - $\pi_1 > 0$
 - $\gamma > 0$
 - $\text{OVB} = (+) \times (+) > 0$
- Not controlling for the OV ability_i may cause overestimation of the effect of the farm size β_1 .
- ability_i does not vary over time for each farm!

How to resolve such OVB?

- Introduce a proxy variable for ability?
 - No such data.
- Rely on the panel structure of the data.
 - Use the individual fixed effect model, for example.

Estimation strategy. Part 2.

Individual fixed effect panel regression production function.

$$\begin{aligned}\ln \text{output}_{it} = & A_0 + \beta_1 \cdot \ln \text{land}_{it} + \beta_2 \cdot \ln \text{labor}_{it} \\ & + \beta_3 \cdot \ln \text{seed}_{it} + \beta_4 \cdot \ln \text{urea}_{it} \\ & + \beta_5 \cdot \ln \text{pesticide}_{it} \\ & + \alpha_i + e_{it}\end{aligned}$$

- What are the ex-ante expectations about the regression coefficients?
 - Ideas? ...
 - Probably the same as before: all β should be positive.

Estimation strategy. Part 3. Return to scale (1)

Are larger farms more productive?

To understand this, we need to calculate how joint increase of all inputs change the output.

- If increase of all inputs by 1% increases output also by **the same 1%**, we have a **constant return to scale**.
- If increase of all inputs by 1% increases output also by **more than 1%**, we have an **increasing return to scale**.
- If increase of all inputs by 1% increases output also by **less than 1%**, we have an **decreasing return to scale**.

Estimation strategy. Part 3. Return to scale (2)

From the main equation,

$$\begin{aligned} \ln \text{output}_{it} = & A_0 + \beta_1 \cdot \ln \text{land}_{it} + \beta_2 \cdot \ln \text{labor}_{it} \\ & + \beta_3 \cdot \ln \text{seed}_{it} + \beta_4 \cdot \ln \text{urea}_{it} \\ & + \beta_5 \cdot \ln \text{pesticide}_{it} \\ & + \alpha_i + e_{it} \end{aligned}$$

- return to scale can be estimated as a sum of all coefficients:

$$\text{Rreturn to scale} = \beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5$$

- To perform a hypothesis testing about the return to scale, we need to employ:
 - HT about a linear combination of parameters, and
 - **delta method** for estimating standard errors.

Estimation strategy. Part 3. Return to scale (3)

HT about a linear combination of parameters:

- $H_0 - \beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5 = 1$ (also can be $= 0$ or any number)
- $H_1 - \beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5 \neq 1$ (also can be $\neq 0$ or any number)

We compute standard errors using **delta method** (`car::deltaMethod()`).

And perform HT using F statistics.

Part 2. Exploratory data analysis and data description

We operate a farm-level data with following variables:

- **output** - gross output of rice in kg
- **land** - the total area cultivated with rice, measured in hectares
- **seed** - seed in kilogram
- **urea** - urea in kilogram
- **pesticide** - urea in kilogram
- **labor** - total labor in hours (excluding harvest labor)

Data preparation

Load data and calculate rice yields and logs of all variables.

```
1 library(tidyverse)
2 library(plm)
3 library(modelsummary)
4 library(performance)
5 library(lmtest)
6
7 farm_dta <- read_csv("farm_panel.csv")
8
9 glimpse(farm_dta)
```

Rows: 1,026

Columns: 15

```
$ id      <dbl> 101001, 101001, 101001, 101001, 101001, 101001, 101017, 101...
$ time    <dbl> 1999, 2000, 2001, 2002, 2003, 2004, 1999, 2000, 2001, 2002,...
$ output  <dbl> 7980, 4083, 2650, 4500, 16300, 17424, 3840, 2800, 950, 240,...
$ land    <dbl> 3.000, 2.000, 1.000, 2.000, 3.572, 3.572, 1.420, 1.420, 0.4...
$ labor   <dbl> 2915, 2155, 1075, 2091, 3889, 3519, 810, 855, 460, 109, 230...
$ hiredlabor <dbl> 2875, 2110, 980, 2081, 3889, 3519, 670, 805, 380, 40, 210, ...
$ famlabor <dbl> 40, 45, 95, 10, 1, 1, 140, 50, 80, 69, 20, 1, 108, 63, 57, ...
$ seed    <dbl> 90, 40, 100, 60, 105, 105, 50, 20, 15, 7, 15, 15, 5, 10, 10...
$ urea    <dbl> 900, 600, 700, 600, 400, 400, 120, 100, 150, 50, 100, 100, ...
$ pest    <dbl> 6000, 3000, 5000, 5000, 10200, 10200, 0, 0, 900, 0, 2000, 2...
$ varieties <chr> "mixed", "trad", "high", "high", "high", "high", "trad", "h...
$ status  <chr> "owner", "owner", "owner", "owner", "share", "share", "mixe...
$ bimas   <chr> "mixed", "mixed", "mixed", "mixed", "no", "no", "mixed", "m...
$ ...     <chr> "-----" "-----" "-----" "-----"
```

Computing yields and logs

```
1 farm_dta_log <- farm_dta %>%  
2   mutate(  
3     l_output = log(output),  
4     l_land = log(land),  
5     l_seed = log(seed),  
6     l_urea = log(urea),  
7     l_pest = log(pest),  
8     l_labor = log(labor),  
9     yields_mt_ha = output / land / 1000  
10  )
```

Summary statistics

```

1 n_inf <- function(x) sum(is.infinite(x))
2 n_missing <- function(x) sum(is.na(x)|is.nan(x))
3 datasummary(
4   l_output + l_land + l_seed + l_urea + l_pest + l_labor + yields_mt_ha ~
5   N + n_missing + n_inf + Mean + SD + Median + Min + Max,
6   data = farm_dta_log)

```

	N	n_missing	n_inf	Mean	SD	Median	Min	Max
l_output	1026	0.00	0.00	6.73	0.99	6.79	3.74	9.95
l_land	1026	0.00	0.00	-1.30	0.95	-1.25	-4.61	1.67
l_seed	1026	0.00	0.00	2.37	0.94	2.30	0.00	7.13
l_urea	1026	0.00	0.00	3.98	1.16	4.09	0.00	7.13
l_pest	1026	0.00	713.00					11.04
l_labor	1026	0.00	0.00	5.56	0.85	5.53	2.83	8.47
yields_mt_ha	1026	0.00	0.00	3.38	1.68	3.16	0.40	27.50

- Any problems with data?

Any problems with data?

- Any? ...
- `pest`, when transformed with logs, produces `-Inf` values.
 - Why is that so?
 - Any? ...
 - Because there are zero values of pesticides application $\ln 0 = -\infty$.
- How to resolve the $\ln 0$ problem?

-Infinity in logs: lazy solution

- Before log transformation, substitute any zero with a small value, for example 0.0001;

```
1 farm_dta_log <- farm_dta_log %>%  
2   mutate(l_pest_lazy = ifelse(is.infinite(l_pest), log(0.0001), l_pest))
```

-Infinity in logs: smart solution

- Introduce reverse dummy variables for each variable with log of zero, see: Battese (1997);
- Substitute negative infinity with zero. . . .

```
1 farm_dta_log <- farm_dta_log %>%  
2   mutate(pest_revdum = ifelse(is.infinite(l_pest), 1, 0),  
3         l_pest_smart = ifelse(is.infinite(l_pest), 0, l_pest))
```

Summary statistics after data cleaning

```
1 datasummary(l_pest + l_pest_lazy + l_pest_smart + pest_revstum ~
2             N + n_missing + n_inf + Mean + SD + Median + Min + Max,
3             data = farm_dta_log)
```

	N	n_missing	n_inf	Mean	SD	Median	Min
l_pest	1026	0.00	713.00				
l_pest_lazy	1026	0.00	0.00	-4.37	7.35	-9.21	-9.21
l_pest_smart	1026	0.00	0.00	2.04	3.15	0.00	0.00
pest_revstum	1026	0.00	0.00	0.69	0.46	1.00	0.00

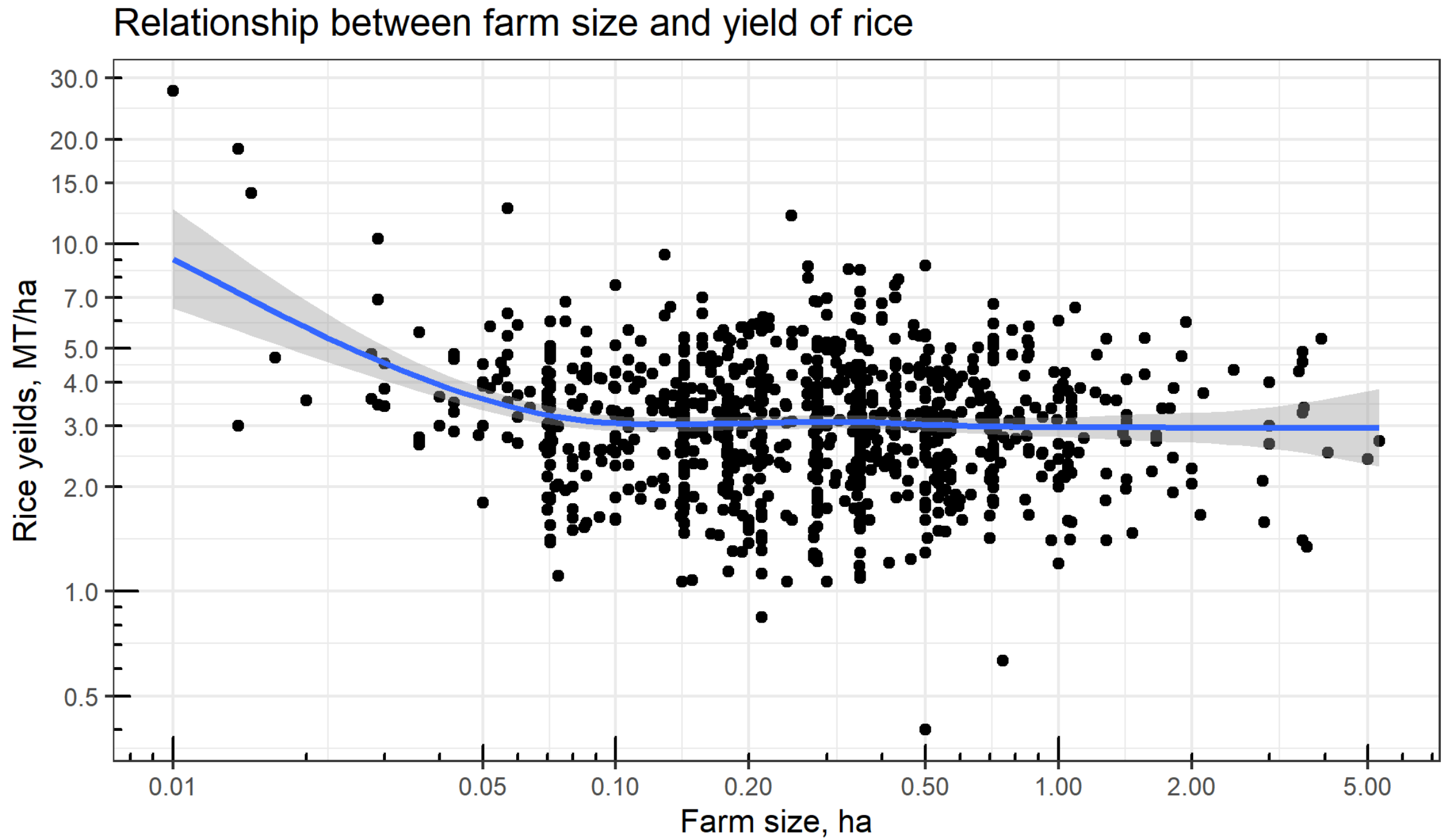
Farm size vs rice yields

```

1 yield_size <-
2   farm_dta_log %>%
3   ggplot() +
4   aes(x = land, y = yields_mt_ha) +
5   geom_point() +
6   geom_smooth() +
7   scale_x_log10("Farm size, ha",
8                 breaks = c(0.01, 0.05, 0.1, 0.2, 0.5, 1, 2, 5, 10)) +
9   scale_y_log10("Rice yeilds, MT/ha",
10                breaks = c(0.2, 0.5, 1, 2, 3, 4, 5, 7, 10, 15, 20, 30)) +
11   annotation_logticks() +
12   labs(title = "Relationship between farm size and yield of rice",
13         caption = "Loess non-parametric smoothing line highlights the trend")
14 yield_size

```

Farm size vs rice yields



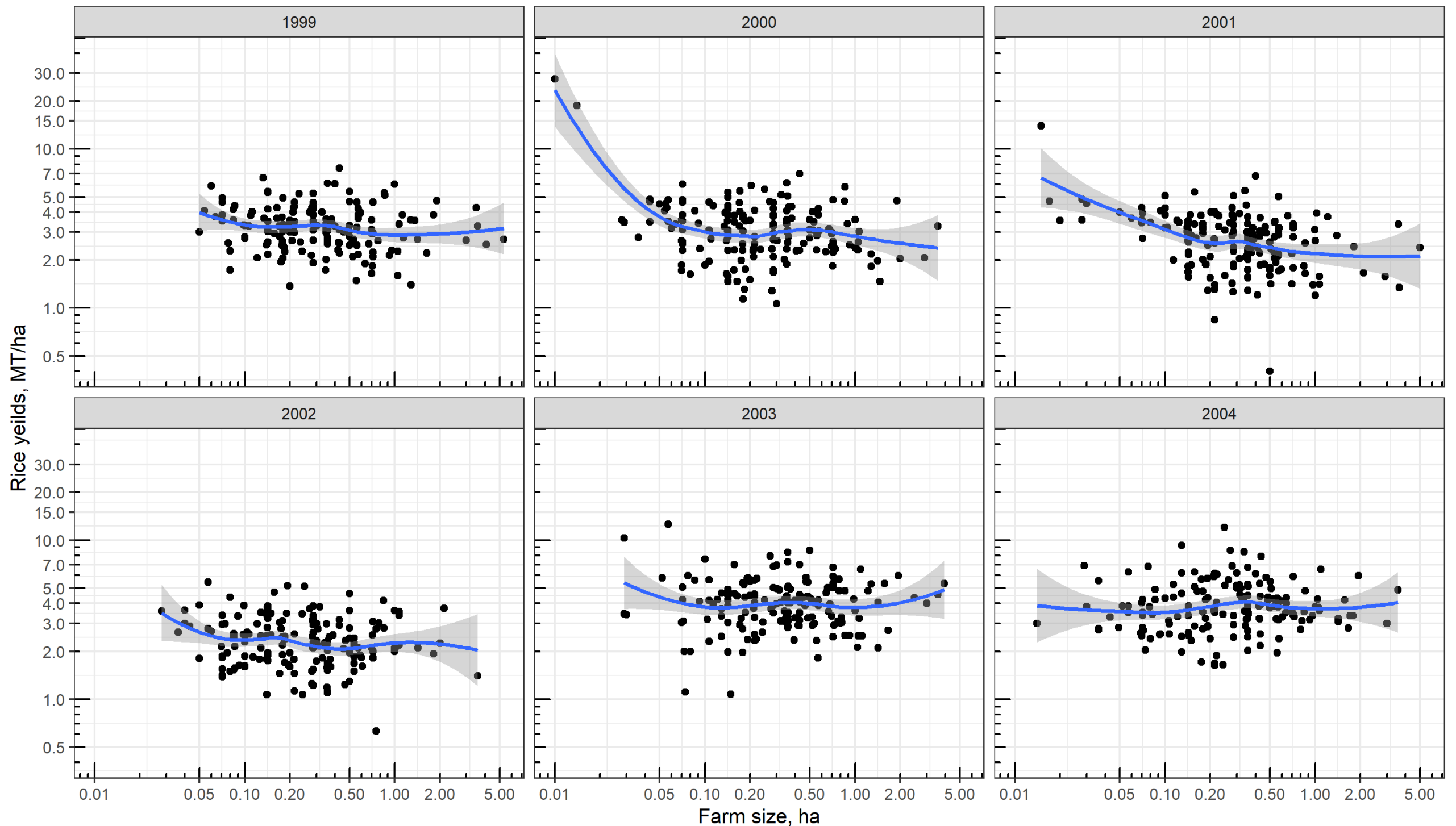
Loess non-parametric smoothing line highlights the trend

Farm size vs rice yields by year

```
1 yield_size + facet_wrap(. ~ time)
```



Relationship between farm size and yield of rice



Loess non-parametric smoothing line highlights the trend

Part 3. Estimating the models

Step 1. Pooled OLS (with `lm()` and `plm()` functions);

- Validate all assumptions (linearity, collinearity, homogeneity)

Step 2. Fixed Effect and Random Effect models (with `plm()`);

- Choose a consistent model (models) relying on: F-test (`pftest()`), Lagrange multiplier test (`plmtest()`), Hausman test (`phptest()`);

Step 3. Validate homogeneity assumption (cross-sectional dependency and autocorrelation)

- Wooldridge's test (`pwtest()`) and Lagrange-Multiplier tests (`pbsytest()`).

Step 4. Robust inference and results interpretation.

Step 1 Pooled OLS

```

1  rice_pooled <-
2    plm(l_output ~ l_land + l_labor + l_seed + l_urea +
3        l_pest_smart + pest_revdum,
4        data = farm_dta_log, model = "pooling", index = c("id", "time"))
5
6  rice_pooled_2 <-
7    lm(l_output ~ l_land + l_labor + l_seed + l_urea +
8        l_pest_smart + pest_revdum,
9        data = farm_dta_log)
10
11  rice_pooled

```

Model Formula: $l_output \sim l_land + l_labor + l_seed + l_urea + l_pest_smart + pest_revdum$

Coefficients:

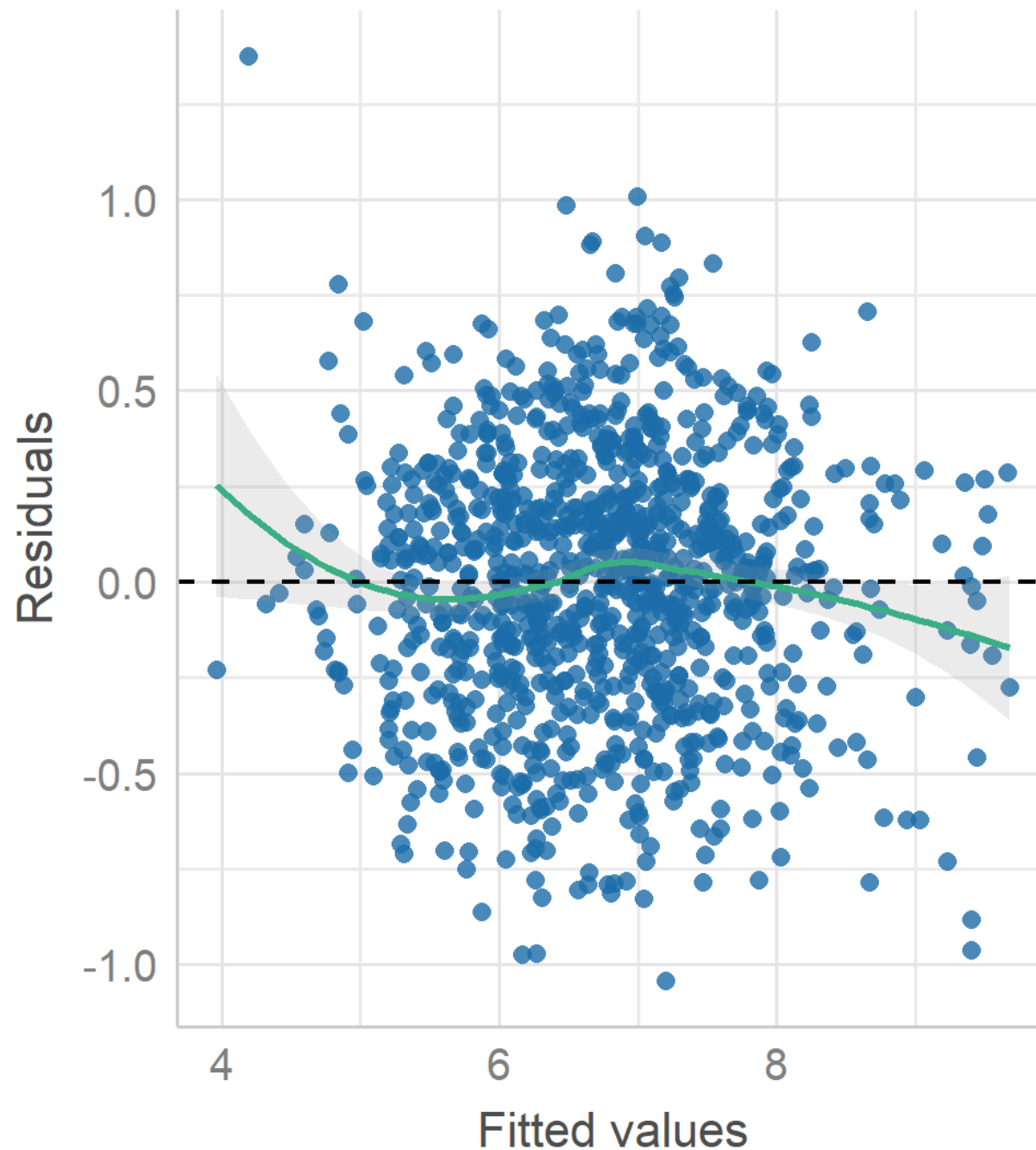
(Intercept)	l_land	l_labor	l_seed	l_urea	l_pest_smart
4.314437	0.434792	0.253482	0.148404	0.171475	0.091171
pest_revdum					
0.512910					

Step 1 Linearity and homoscedasticity

```
1 check_model(rice_pooled_2, check = c("linearity", "homogeneity"))
```

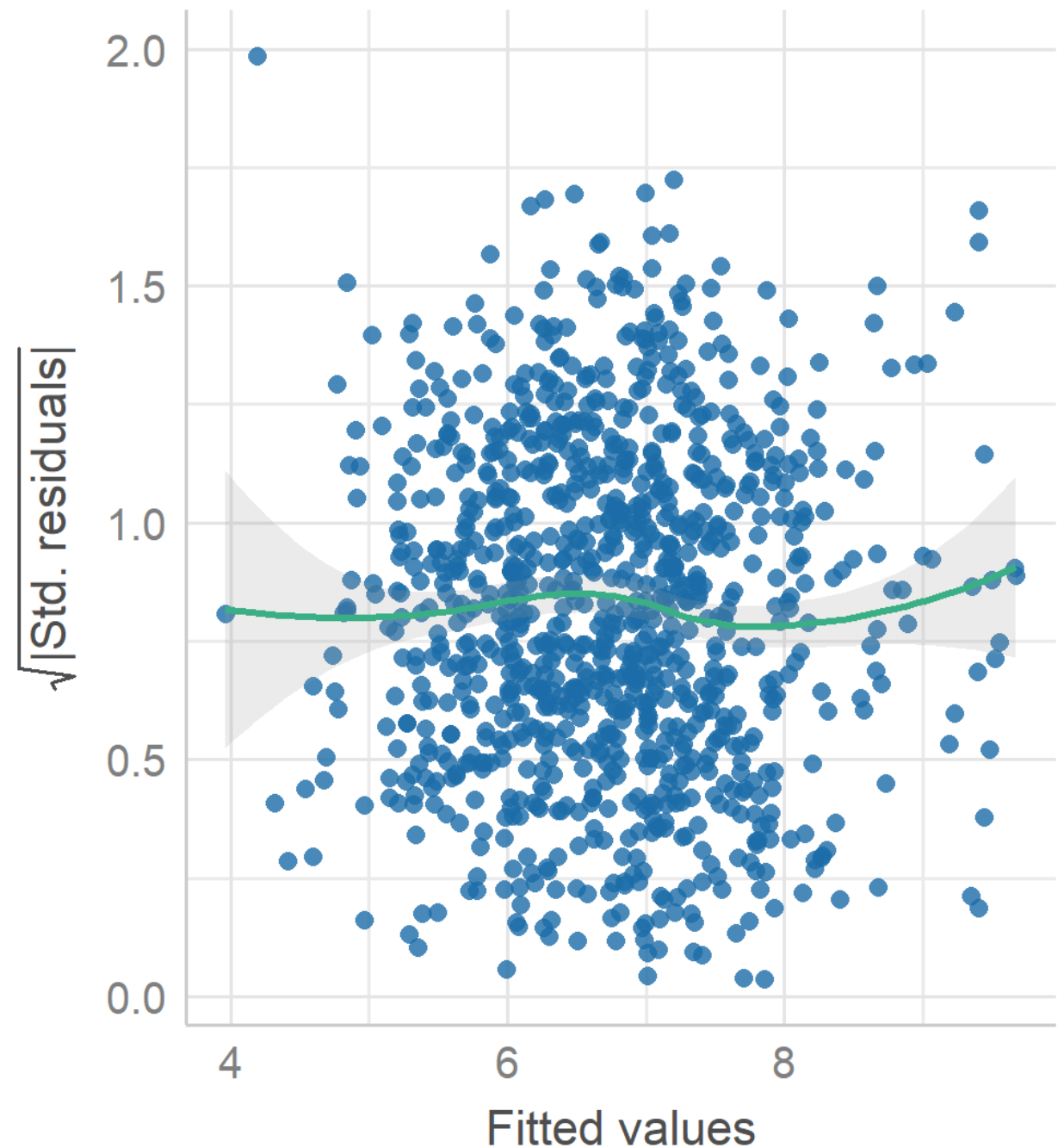
Linearity

Reference line should be flat and horizontal



Homogeneity of Variance

Reference line should be flat and horizontal



Step 1 Collinearity

```
1 check_collinearity(rice_pooled_2)
```



```
# Check for Multicollinearity
```

Low Correlation

Term	VIF	VIF 95% CI	Increased SE	Tolerance	Tolerance 95% CI
l_labor	4.77	[4.28, 5.32]	2.18	0.21	[0.19, 0.23]
l_urea	2.67	[2.42, 2.95]	1.63	0.37	[0.34, 0.41]

Moderate Correlation

Term	VIF	VIF 95% CI	Increased SE	Tolerance	Tolerance 95% CI
l_land	6.94	[6.21, 7.78]	2.64	0.14	[0.13, 0.16]
l_seed	5.07	[4.55, 5.66]	2.25	0.20	[0.18, 0.22]

High Correlation

Step 2 Fixed Effect

```

1 rice_fe <-
2   plm(l_output ~ l_land + l_labor + l_seed + l_urea +
3       l_pest_smart + pest_revdum,
4       data = farm_dta_log,
5       model = "within",
6       effect = "individual",
7       index = c("id", "time"))
8 rice_fe

```

Model Formula: `l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + pest_revdum`

Coefficients:

<code>l_land</code>	<code>l_labor</code>	<code>l_seed</code>	<code>l_urea</code>	<code>l_pest_smart</code>	<code>pest_revdum</code>
0.41984	0.27084	0.12381	0.16291	0.11029	0.63498

Step 2 FE with lazy log(0)

```

1 rice_fe_lazy <-
2   plm(l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_lazy ,
3       data = farm_dta_log,
4       model = "within",
5       effect = "individual",
6       index = c("id", "time"))
7 rice_fe_lazy

```

Model Formula: `l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_lazy`

Coefficients:

l_land	l_labor	l_seed	l_urea	l_pest_lazy
0.421042	0.263876	0.134373	0.174145	0.007703

Step 2.2 F test for individual effects

- Compares FE model to OLS. OLS is always consistent, when Gauss-Markov assumptions are satisfied.
 - H0: One model is inconsistent.
 - H1: Both models are equally consistent.

```
1 pFtest(rice_fe, rice_pooled)
```

F test for individual effects

```
data:  l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + ...  
F = 1.4988, df1 = 170, df2 = 849, p-value = 0.0001704  
alternative hypothesis: significant effects
```

Step 2.3 Lagrange Multiplier Tests

- Compares FE model to OLS. OLS is always consistent, when Gauss-Markov assumptions are satisfied.
 - H0: One model is inconsistent.
 - H1: Both models are equally consistent.

```
1 plmtest(rice_pooled, effect = "individual", type = "honda")
```

Lagrange Multiplier Test - (Honda)

```
data:  l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + ...
normal = 3.7129, p-value = 0.0001025
alternative hypothesis: significant effects
```

```
1 plmtest(rice_pooled, effect = "individual", type = "bp")
```

Lagrange Multiplier Test - (Breusch-Pagan)

```
data:  l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + ...
chisq = 13.785, df = 1, p-value = 0.0002049
alternative hypothesis: significant effects
```


Step 2 Random Effect

```

1 rice_re <-
2   plm(l_output ~ l_land + l_labor + l_seed + l_urea +
3       l_pest_smart + pest_revdm,
4       data = farm_dta_log,
5       model = "random",
6       effect = "individual",
7       index = c("id", "time"))
8 rice_re

```

Model Formula: `l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + pest_revdm`

Coefficients:

(Intercept)	<code>l_land</code>	<code>l_labor</code>	<code>l_seed</code>	<code>l_urea</code>	<code>l_pest_smart</code>
4.276601	0.433744	0.257738	0.142589	0.169423	0.096608
<code>pest_revdm</code>					
0.546997					

Step 2 Hausman Test for Panel Models

- Compares RE to FE model. FE is assumed to be consistent
 - H0: One model is inconsistent.
 - H1: Both models are equally consistent.

```
1 phtest(rice_fe, rice_re)
```

Hausman Test

```
data:  l_output ~ l_land + l_labor + l_seed + l_urea + l_pest_smart + ...  
chisq = 7.285, df = 6, p-value = 0.2953  
alternative hypothesis: one model is inconsistent
```

- Fixed Effect model is recommended

Step 3 Serial correlation and cross-sectional dependence

- Wooldridge's test for unobserved individual effects
 - H0: no unobserved effects
 - H1: some effects also dues to serial correlation

```
1 pwtest(rice_pooled, effect = "individual")
```

Wooldridge's test for unobserved individual effects

```
data: formula
z = 2.1603, p-value = 0.03075
alternative hypothesis: unobserved effect
```

```
1 pwtest(rice_pooled, effect = "time")
```

Wooldridge's test for unobserved time effects

```
data: formula
z = 1.6899, p-value = 0.09105
alternative hypothesis: unobserved effect
```

Step 3 lm tests for random effects and/or serial correlation

- H0: serial correlation is zero
- H1: some serial correlation

```
1 pbsytest(rice_pooled)
```

Bera, Sosa-Escudero and Yoon locally robust test

```
data: formula  
chisq = 20.988, df = 1, p-value = 4.622e-06  
alternative hypothesis: AR(1) errors sub random effects
```

Step 4. Robust inference

```
1 library(lmtest)
2
3 rice_pooled_robust <- coeftest(
4   rice_pooled,
5   vcovHC(rice_pooled, method = "arellano", type = "HC3", cluster = "group")
6 )
7
8 rice_fe_robust <- coeftest(
9   rice_fe,
10  vcovHC(rice_fe, method = "arellano", type = "HC3", cluster = "group")
11 )
12
13 rice_felazy_robust <- coeftest(
14   rice_fe_lazy,
15   vcovHC(rice_fe_lazy, method = "arellano", type = "HC3", cluster = "group")
16 )
```

Step 4. Robust inference

```

1 modelsummary(
2   list(
3     `Pooled` = rice_pooled_robust,
4     `FE (rev. dum.)` = rice_fe_robust,
5     `FE (lazy)` = rice_felazy_robust
6   ),
7   fmt = 4, statistic = NULL,
8   estimate = "{estimate}{stars} ({std.error})",
9   notes = "Robust standard errors clustered at the group level are reported in the brackets.")

```

	Pooled	FE (rev. dum.)	FE (lazy)
(Intercept)	4.3144*** (0.2691)		
l_land	0.4348*** (0.0418)	0.4198*** (0.0479)	0.4210*** (0.0486)
l_labor	0.2535*** (0.0333)	0.2708*** (0.0347)	0.2639*** (0.0348)
l_seed	0.1484*** (0.0344)	0.1238*** (0.0370)	0.1344*** (0.0386)
l_urea	0.1715*** (0.0228)	0.1629*** (0.0262)	0.1741*** (0.0269)
l_pest_smart	0.0912*** (0.0172)	0.1103*** (0.0185)	
pest_rev dum	0.5129*** (0.1141)	0.6350*** (0.1290)	
l_pest_lazy			0.0077*** (0.0020)
Num.Obs.	1026	1026	1026
AIC	2796.9	2187.6	2222.6
BIC	7824.1	6376.1	6416.0
Robust standard errors clustered at the group level are reported in the brackets.			

Step 4. Robust inference

```
1 compare_performance(
2   list(
3     `Pooled` = rice_pooled,
4     `FE with rev. dum` = rice_fe,
5     `FE lazy` = rice_fe_lazy
6   )
7 )
```

Comparison of Model Performance Indices

Name	Model	AIC (weights)	AICc (weights)	BIC (weights)	R2	R2 (adj.)
RMSE Sigma						

Pooled	plm	774.9 (<.001)	775.1 (<.001)	814.4 (<.001)	0.876	0.875
0.350 0.351						
FE with rev. dum	plm	503.6 (>.999)	503.7 (>.999)	538.2 (>.999)	0.732	0.676
0.307 0.308						
FE lazy	plm	534.6 (<.001)	534.6 (<.001)	564.2 (<.001)	0.723	0.666
0.312 0.313						

Part 4. Return to scale (1)

```
1 rice_fe
```

Model Formula: $\text{l_output} \sim \text{l_land} + \text{l_labor} + \text{l_seed} + \text{l_urea} + \text{l_pest_smart} + \text{pest_revdum}$

Coefficients:

l_land	l_labor	l_seed	l_urea	l_pest_smart	pest_revdum
0.41984	0.27084	0.12381	0.16291	0.11029	0.63498

Computing sum of the coefficients and robust SE:

```
1 library(car)
2 deltaMethod(
3   rice_fe,
4   "l_land + l_labor + l_seed + l_urea + l_pest_smart",
5   vcov = vcovHC(rice_fe, method = "arellano", type = "HC3", cluster = "group")
6 )
```

	Estimate	SE	2.5 %	97.5 %
l_land + l_labor + l_seed + l_urea + l_pest_smart	1.087685	0.031039	1.026850	
l_land + l_labor + l_seed + l_urea + l_pest_smart	1.1485			

Part 4. Return to scale (2)

HT about the sum of the coefficients:

- H_0 : return to scale = 1
- H_0 : return to scale $\neq 1$

```
1 linearHypothesis(rice_fe,
2   "l_land + l_labor + l_seed + l_urea + l_pest_smart = 1",
3   vcov = vcovHC(rice_fe, method = "arellano", type = "HC3", cluster = "group"))
```

Linear hypothesis test

Hypothesis:

$l_land + l_labor + l_seed + l_urea + l_pest_smart = 1$

Model 1: restricted model

Model 2: $l_output \sim l_land + l_labor + l_seed + l_urea + l_pest_smart +$
 $pest_revdum$

Note: Coefficient covariance matrix supplied.

```
      Res.Df Df    Chisq Pr(>Chisq)
1         850
2         849  1  7.9806   0.004728 **
---
adj. R-sq: 0.99999 0.99999 0.99999 0.99999 0.99999 0.99999 0.99999 0.99999 0.99999 0.99999
```

Conclusions

Regression results

	Pooled	FE (rev. dum.)	FE (lazy)
(Intercept)	4.3144*** (0.2691)		
l_land	0.4348*** (0.0418)	0.4198*** (0.0479)	0.4210*** (0.0486)
l_labor	0.2535*** (0.0333)	0.2708*** (0.0347)	0.2639*** (0.0348)
l_seed	0.1484*** (0.0344)	0.1238*** (0.0370)	0.1344*** (0.0386)
l_urea	0.1715*** (0.0228)	0.1629*** (0.0262)	0.1741*** (0.0269)
l_pest_smart	0.0912*** (0.0172)	0.1103*** (0.0185)	
pest_revstum	0.5129*** (0.1141)	0.6350*** (0.1290)	
l_pest_lazy			0.0077*** (0.0020)
Num.Obs.	1026	1026	1026

Return to scale

```
1 library(car)
2 deltaMethod(rice_fe,
3             "l_land + l_labor + l_seed + l_urea + l_pest_smart",
4             vcov = vcovHC(rice_fe, method = "arellano", type =
```

SE 2.5 % Estimate

l_land + l_labor + l_seed + l_urea + l_pest_smart 1.087685

0.031039 1.026850

97.5 %

l_land + l_labor + l_seed + l_urea + l_pest_smart 1.1485

```
1 linearHypothesis(rice_fe,
2                   "l_land + l_labor + l_seed + l_urea + l_pest_s
3                   vcov = vcovHC(rice_fe, method = "arellano", ty
```

Linear hypothesis test

Hypothesis:

l_land + l_labor + l_seed + l_urea + l_pest_smart = 1

Model 1: restricted model

Model 2: l_output ~ l_land + l_labor + l_seed + l_urea +
l_pest_smart +
pest_revstum

Note: Coefficient covariance matrix supplied.

```
Res.Df Df  Chisq Pr(>Chisq)
1      850
2      849  1  7.9806  0.004728 **
---
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Individual fixed effects

From the regression equation below,

$$\begin{aligned} \ln \text{output}_{it} = & A_0 + \beta_1 \cdot \ln \text{land}_{it} + \beta_2 \cdot \ln \text{labor}_{it} \\ & + \beta_3 \cdot \ln \text{seed}_{it} + \beta_4 \cdot \ln \text{urea}_{it} \\ & + \beta_5 \cdot \ln \text{pesticide}_{it} \\ & + \alpha_i + e_{it} \end{aligned}$$

we know that α_i are the individual fixed effects.

- R calculates them and we can explore them.
- In fact, those individual fixed effects are the simplest possible measured of farms efficiency!

Extra on fixed effects

Individual fixed effects (extraction 1)

In the model:

```
1 rice_fe
```

Model Formula: $l_output \sim l_land + l_labor + l_seed + l_urea + l_pest_smart + pest_revdum$

Coefficients:

l_land	l_labor	l_seed	l_urea	l_pest_smart	$pest_revdum$
0.41984	0.27084	0.12381	0.16291	0.11029	0.63498

Individual fixed effects can be extracted as:

```
1 fixef(rice_fe)
```

101001	101017	101026	101035	101056	101057	101067	101068	101069	101073	101089
3.8470	4.1040	4.2375	4.2631	4.5649	4.2051	4.2591	4.3471	4.3598	4.0599	4.0611
101094	102111	102113	102119	102126	102157	102194	102220	201001	201002	201003
4.0503	4.1591	4.1073	4.2289	3.9991	4.2420	3.9798	4.2032	4.3270	4.3719	4.2906
201009	202039	202061	202066	203079	203080	204096	204114	204116	204124	205132
4.1036	4.2036	4.1584	4.1318	4.2522	4.0977	4.2539	3.9338	4.3095	4.0783	3.9765
205136	205151	205153	206158	206169	207209	208225	209232	209241	209250	301004
3.9870	4.2319	4.2504	4.4004	4.0218	4.2235	4.2522	4.1496	4.0292	4.0267	4.0658
301010	301023	301038	301055	301058	301067	301070	301075	301084	301105	301110
3.7676	4.1062	4.1388	4.1066	4.1385	4.1110	4.0471	4.3944	3.8461	3.9620	3.9876
302116	302120	302131	302134	302137	302142	302143	302144	302146	302147	302151
4.1645	3.8708	4.0910	4.2774	4.1428	4.0946	4.0225	4.1853	4.1072	3.8265	4.1892
302153	302161	302163	302169	302182	302189	302192	302194	302195	302197	302199
3.9304	4.1563	4.1978	4.2176	4.1114	4.2771	4.3379	4.1691	4.3303	4.2350	4.0927
302205	302207	302209	401002	401006	401032	401034	401036	401037	401041	401043
4.1504	3.8822	4.1140	4.0507	3.8850	4.0725	4.0764	3.8721	3.8727	4.0275	4.1460

Individual fixed effects (extraction 3)

Or, we can extract individual fixed effects with effect-specific standard errors:

```
1 fef_dta <-
2   fixef(rice_fe) %>%
3   summary() %>%
4   as.data.frame() %>% rownames_to_column("id") %>%
5   as_tibble() %>%
6   mutate(id = as.double(id))
7 fef_dta
```

```
# A tibble: 171 × 5
      id Estimate `Std. Error` `t-value` `Pr(>|t|)`
  <dbl>    <dbl>      <dbl>    <dbl>    <dbl>
1 101001      3.85      0.317     12.2 1.96e-31
2 101017      4.10      0.282     14.6 5.20e-43
3 101026      4.24      0.285     14.9 1.09e-44
4 101035      4.26      0.279     15.3 1.11e-46
5 101056      4.56      0.285     16.0 9.02e-51
6 101057      4.21      0.292     14.4 3.46e-42
7 101067      4.26      0.296     14.4 2.89e-42
8 101068      4.35      0.283     15.4 3.87e-47
9 101069      4.36      0.279     15.6 1.62e-48
10 101073      4.06      0.294     13.8 3.35e-39
# i 161 more rows
```

Farm size and efficiency

Let us compute average farm size and plot fixed effects versus farm size:

```
1 farm_mean_size <- farm_dta_log %>% group_by(id) %>% summarise(mean_size = mean(land))
2 glimpse(farm_mean_size)
```

Rows: 171

Columns: 2

```
$ id      <dbl> 101001, 101017, 101026, 101035, 101056, 101057, 101067, 1010...
$ mean_size <dbl> 2.5240000, 0.7230000, 0.1558333, 0.3211667, 0.3360000, 0.189...
```

```
1 plot_dta <- farm_mean_size %>% left_join(fef_dta)
2 glimpse(plot_dta)
```

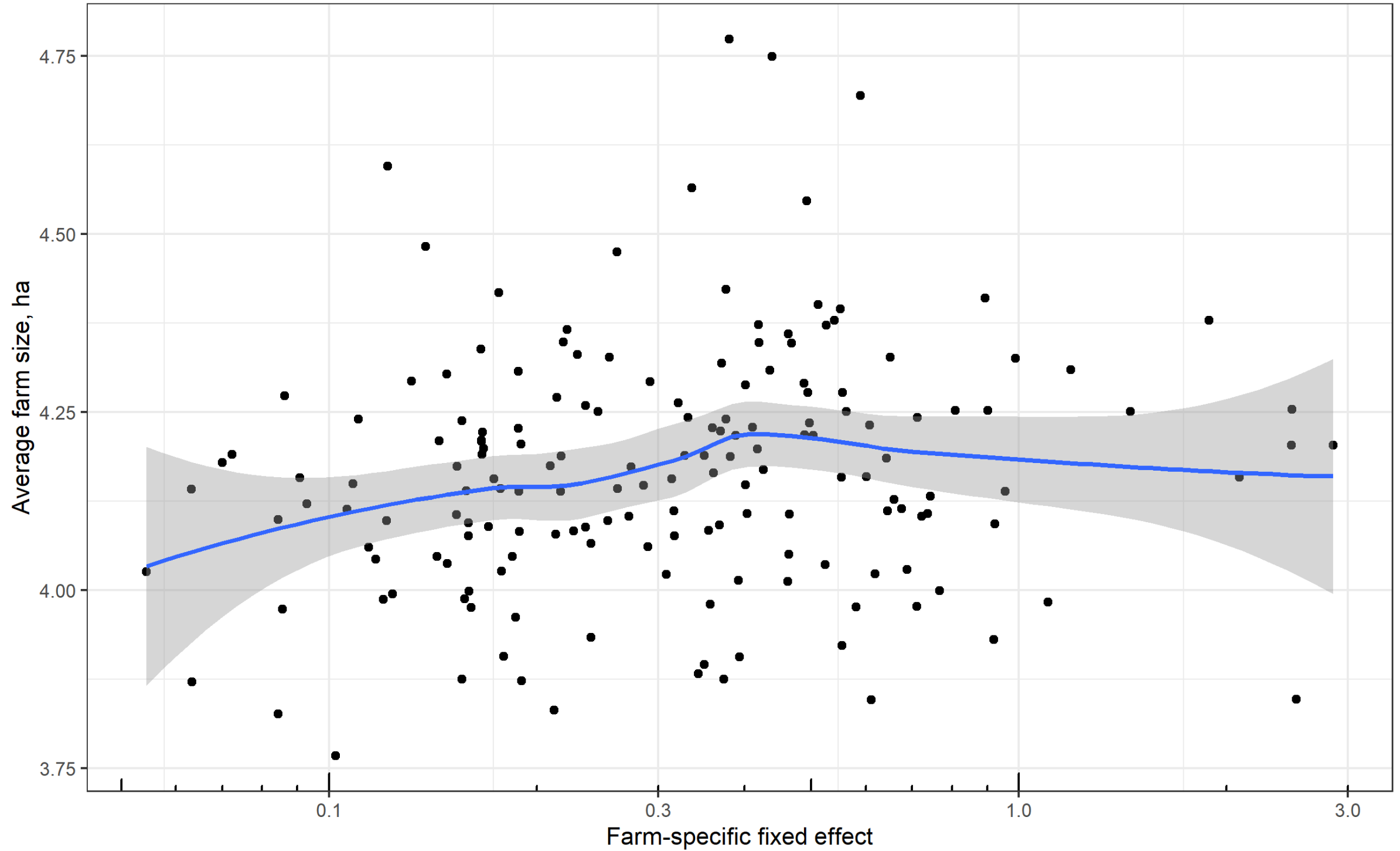
Rows: 171

Columns: 6

```
$ id      <dbl> 101001, 101017, 101026, 101035, 101056, 101057, 101067, 1...
$ mean_size <dbl> 2.5240000, 0.7230000, 0.1558333, 0.3211667, 0.3360000, 0...
$ Estimate <dbl> 3.847036, 4.103960, 4.237548, 4.263063, 4.564856, 4.20514...
$ `Std. Error` <dbl> 0.3166160, 0.2819603, 0.2847204, 0.2792276, 0.2846006, 0...
$ `t-value` <dbl> 12.15048, 14.55510, 14.88319, 15.26734, 16.03951, 14.3927...
$ `Pr(>|t|)` <dbl> 1.958956e-31, 5.202280e-43, 1.090622e-44, 1.109427e-46, 9...
```

```
1 plot_dta %>% ggplot() + aes(x = mean_size, y = Estimate) +
2   geom_point() + geom_smooth() + scale_x_log10() +
3   xlab("Farm-specific fixed effect") +
4   ylab("Average farm size, ha") +
5   annotation_logticks(side = "b")
```


Farm size and efficiency



References

Battese, G. E. (1997). A note on the estimation of cobb-douglas production functions when some explanatory variables have zero values. *Journal of Agricultural Economics*, 48(1-3), 250–252.
<http://doi.org/10.1111/j.1477-9552.1997.tb01149.x>

