1. SIMPLE SEQUENTIAL TEST

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SINGLE-SAMPLING-PLAN(p_0, p_1, \alpha, \beta)
    n_{\min} \Leftarrow 1, \ n_{\max} \Leftarrow -1
    n \Leftarrow n_{\min}
    while n_{\max} < 0 \lor n_{\min} < n_{\max} do x_0 \Leftarrow \tilde{F}^{-1}(\alpha;n,p_0)
         x_1 \Leftarrow \tilde{F}^{-1}(1-\beta;n,p_1)
         if x_0 \ge x_1 \wedge x_0 \ge 0 then
             n_{\max} \Leftarrow n
         else
              n_{\min} \Leftarrow n + 1
         if n_{\rm max} < 0 then
             n \Leftarrow 2 \cdot n
         else
              n \Leftarrow \lfloor (n_{\min} + n_{\max})/2 \rfloor
    n \Leftarrow n_{\max} - 1
    repeat
         n \Leftarrow n+1
         c_0 \Leftarrow \lfloor \tilde{F}^{-1}(\alpha; n, p_0) \rfloor
c_1 \Leftarrow \lceil \tilde{F}^{-1}(1 - \beta; n, p_1) \rceil
    until c_0 \geq c_1
    return \langle n, \lfloor (c_0 + c_1)/2 \rfloor \rangle
```

Algorithm 2.1: Procedure for finding an optimal single sampling plan using binary search. $\tilde{F}^{-1}(y;n,p)$ can be computed by adding the terms of (2.3) until the sum equals or exceeds y.

```
\begin{array}{l} \text{SIMPLE-SEQUENTIAL-TEST} \left(p_0, p_1, \alpha, \beta\right) \\ & \langle n, c \rangle \Leftarrow \text{SINGLE-SAMPLING-PLAN} (p_0, p_1, \alpha, \beta) \\ & m \Leftarrow 0, \ d_m \Leftarrow 0 \\ & \text{while} \ d_m \leq c \wedge d_m + n - m > c \ \text{do} \\ & m \Leftarrow m + 1 \\ & d_m \Leftarrow d_{m-1} + x_m \\ & \text{if} \ d_m > c \ \text{then} \\ & \text{return} \ H_0 \\ & \text{else} \\ & \text{return} \ H_1 \end{array}
```

Algorithm 2.2: Sequential acceptance sampling procedure based on a single sampling plan.

2. SEQUENTIAL PROBABILITY RATIO TEST

```
\begin{aligned} & \text{SPRT}(p_0, p_1, \alpha, \beta) \\ & \text{if } p_0 = 1 \lor p_1 = 0 \text{ then} \\ & \text{return SIMPLE-SEQUENTIAL-TEST}(p_0, p_1, \alpha, \beta) \\ & \text{else} \\ & m \Leftarrow 0, \ f_m \Leftarrow 0 \\ & \text{while } \log \frac{\beta}{1-\alpha} < f_m < \log \frac{1-\beta}{\alpha} \text{ do} \\ & m \Leftarrow m+1 \\ & f_m \Leftarrow f_{m-1} + x_m \log \frac{p_1}{p_0} + (1-x_m) \log \frac{1-p_1}{1-p_0} \\ & \text{if } f_m \leq \log \frac{\beta}{1-\alpha} \text{ then} \\ & \text{return } H_0 \\ & \text{else} \\ & \text{return } H_1 \end{aligned}
```

Algorithm 2.3: Procedure implementing the sequential probability ratio test.

3. BAYESIAN HYPOTHESIS TESTING

```
Input: PBLTL property P_{\geqslant \theta}(\phi), acceptance threshold T \geqslant 1, prior density g for
            (unknown) probability p that the system satisfies \phi
   Output: "H_0: p \ge \theta accepted", or "H_1: p < \theta accepted"
                                               {number of traces drawn so far}
n := 0;
x := 0;
                                   {number of traces satisfying \phi so far}
3 loop
       \sigma := \text{draw a sample trace of the system (iid)};
                                                                                  {Sect. 2}
       n := n + 1;
       if \sigma \models \phi then
                                                                                  {Sect. 3}
           x := x + 1
       \mathcal{B} := \text{BayesFactor}(n, x);
                                                                                   {Sect. 5}
       if (B > T) then return "H_0 accepted";
       if (\mathcal{B} < \frac{1}{T}) then return "H_1 accepted"
12 end loop;
```

Algorithm 1: Statistical Model Checking by Bayesian Hypothesis Testing

Bayes Factor:

$$\mathcal{B}_n = \frac{\pi_1}{\pi_0} \cdot \left(\frac{1}{F_{(x+\alpha,n-x+\beta)}(\theta)} - 1 \right).$$

$$\forall u \in (0,1) \quad F_{(\alpha,\beta)}(u) = \int_0^u g(t,\alpha,\beta) \ dt = \frac{1}{B(\alpha,\beta)} \int_0^u t^{\alpha-1} (1-t)^{\beta-1} \ dt \quad (5)$$

$$B(\alpha,\beta) \stackrel{\frown}{=} \int_0^1 t^{\alpha-1} (1-t)^{\beta-1} dt .$$

4. BAYESIAN INTERVAL ESTIMATES

```
Input: BLTL Property \phi, half-interval size \delta \in (0, \frac{1}{2}), interval coverage coefficient
             c \in (\frac{1}{2}, 1), Prior Beta distribution with parameters \alpha, \beta for the (unknown)
             probability p that the system satisfies \phi
   Output: An interval (t_0, t_1) of width 2\delta with posterior probability at least c, estimate
             \hat{p} for the true probability p
                                                 {number of traces drawn so far}
 n := 0;
                                     {number of traces satisfying \phi so far}
 x := 0;
 3 repeat
       \sigma := \text{draw a sample trace of the system (iid)};
       n := n + 1;
       if \sigma \models \phi then x := x + 1;
       \hat{p} := (x + \alpha)/(n + \alpha + \beta);
                                                             {compute posterior mean}
       (t_0, t_1) := (\hat{p} - \delta, \hat{p} + \delta);
                                                       {compute interval estimate}
       if t_1 > 1 then (t_0, t_1) := (1 - 2 \cdot \delta, 1)
       else if t_0 < 0 then (t_0, t_1) := (0, 2 \cdot \delta);
       {compute posterior probability of p \in (t_0, t_1), by (9)}
       \gamma := Posterior Prob(t_0, t_1)
12 until (\gamma \geqslant c);
13 return (t_0, t_1), \hat{p}
```

Algorithm 2: Statistical Model Checking by Bayesian Interval Estimates

Posterior Probability:

$$\int_{t_0}^{t_1} f(u|x_1,\ldots,x_n)du = F_{(x+\alpha,n-x+\beta)}(t_1) - F_{(x+\alpha,n-x+\beta)}(t_0)$$