Demand Estimation

MIXTAPE SESSION

Jeff Gortmaker and Ariel Pakes



$$\max_{j \in \mathcal{J}_t \cup \{0\}} u_{ijt} = \delta_{jt} + \mu_{ijt} + \varepsilon_{ijt}$$

• In each market $t \in \mathcal{T}$, individuals with types $i \in \mathcal{I}_t$ choose a $j \in \mathcal{J}_t \cup \{0\}$.

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- In each market $t \in \mathcal{T}$, individuals with types $i \in \mathcal{I}_t$ choose a $j \in \mathcal{J}_t \cup \{0\}$.
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- On day 1, we set $\mu_{ijt} = 0$ to get a conveniently linear estimating equation:

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Let's go over your first coding exercise.

Unrealistic Substitution Patterns

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- Last week we derived the own-price elasticity. What about the cross-price one?

$$\eta_{jkt} = \frac{\partial \log q_{jt}}{\partial \log p_{kt}} = \frac{\partial q_{jt}}{\partial p_{kt}} \frac{p_{kt}}{q_{jt}} = \frac{\partial s_{jt}}{\partial p_{kt}} \frac{p_{kt}}{s_{jt}} = -\alpha \cdot p_{kt} \cdot s_{kt}$$

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- Doesn't depend on the characteristics of j!
 - ightarrow Independence of Irrelevant Alternatives (IIA) property.

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- There are two options: buying a car or a blue bus. Each has a 50% market share.
- Introduce a second bus, but it's red. Pure logit (IIA) predicts 33% market shares.
 - \rightarrow In your exercise, consumers substituted *proportionally* from each cereal.
- In reality, we'd expect the car to still have 50% and each bus to have 25%.
 - $\,\rightarrow\,$ In your exercise, we'd hope for more substitution from more similar cereals.

Roadmap

Preference Heterogeneity

Mixed Logit Estimation

Numerical Best Practices

Differentiation Instruments

Coding Exercise :

Red Bus/Blue Bus Solution

• Our solution will be to re-introduce non-logit preference heterogeneity.

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- This will allow 50% of consumers to really like cars and 50% to really like buses.
 - → When a new bus is introduced, this doesn't really affect the car-lovers' choice.
- Want μ_{ijt} to dominate logit substitution from convenient but unrealistic ε_{ijt} .
 - ightarrow Want to add multiple dimensions of heterogeneity that really matter in our setting.

Random Coefficients

$$u_{ijt} = x'_{jt}\beta + \xi_{jt} + \varepsilon_{ijt}$$

- How to add preference heterogeneity to our pure logit model?
 - ightarrow For simplicity, I'll just let x_{jt} denote all characteristics, including prices p_{jt} .

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- Intuitively, we want to replace eta with random coefficients eta_{it} .
 - \rightarrow Random in that they're drawn from a distribution of consumer types $i \in \mathcal{I}_t$.
 - ightarrow For $x_{jt}=\mathrm{car}_{jt}$ and $\mathcal{I}_t=\{\mathrm{car-lovers},\mathrm{bus-lovers}\}$, want $eta_{it}\gg 0$ for car-lovers.

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- Most common specification is $\beta_{it} \sim N(\beta + \Pi y_{it}, \Sigma \Sigma')$.
 - ightarrow Π shifts preferences according to observed demographics $y_{it}\sim$ census.
 - ightarrow Σ shifts preferences according to unobserved preferences $u_{it} \sim N(0,I)$.

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- In your coding exercise, you'll just draw $|\mathcal{I}_t| = 100$ types per market.
 - \rightarrow Draw $\nu_{it} \sim N(0, I)$ from a random number generator.
 - \rightarrow Draw y_{it} from census data on demographics: income, etc.
 - ightarrow Each type is equally-likely, so use equal sampling weights $w_{it}=1/|\mathcal{I}_t|$.

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- The goal is to have a dataset that reflects the *distribution* of individuals.

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Coding Exercise 2

From Linear Regression to GMM

$$\log \frac{s_{jt}}{s_{0t}} = \delta_{jt} = x'_{jt}\beta + \xi_{jt}$$

- In your exercise, you estimated $\boldsymbol{\beta}$ by running the above regression.
 - \rightarrow Again, let x_{jt} include price, a constant, any other characteristics.
 - ightarrow Let z_{jt} include our price IV and exogenous characteristics in x_{jt} .

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- Our exclusion restriction implies the moment condition $\mathbb{E}[\xi_{jt} \cdot z_{jt}] = 0$.
- We'd get the exact same $\hat{\beta}$ by optimizing the following GMM objective:

$$\hat{\beta} = \operatorname*{argmin}_{\beta} g(\beta) W g(\beta)' \quad \text{where} \quad g(\beta) = \frac{1}{N} \sum_{t \in \mathcal{T}} \sum_{j \in \mathcal{J}_t} (\delta_{jt} - x'_{jt}\beta) \cdot z_{jt}$$

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- Many ways to solve and speed up BLP's (1995) contraction.
 - → PyBLP will take care of this, but see Conlon and Gortmaker (2020) if interested.

The BLP Estimator

$$\hat{\theta} = \operatorname*{argmin}_{\theta} g(\theta) W g(\theta)' \quad \text{where} \quad g(\theta) = \frac{1}{N} \sum_{t \in \mathcal{T}} \sum_{j \in \mathcal{J}_t} (\delta_{jt}(\Sigma, \Pi) - x'_{jt}\beta) \cdot z_{jt}$$

- BLP estimation consists of two nested loops.
 - 1. In the "outer" loop, we optimize over $\theta = (\beta, \Sigma, \Pi)$.
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- Actually, since $g(\theta)$ is linear in x_{jt} , we can "concentrate out" β and optimize (Σ, Π) .
 - $\rightarrow \operatorname{Get} \hat{\beta}$ by running an IV regression of $\delta_{jt}(\Sigma,\Pi)$ on x_{jt} , like in the pure logit exercise.
- What about the GMM weighting matrix W?
 - \rightarrow If you're just-identified (dim $z_{jt}=\dim \theta$), it doesn't matter. You'll get a zero objective.
 - ightarrow Otherwise, you may want to repeat optimization with optimal the two-step GMM \hat{W} .

Roadmap

Preference Heterogeneity

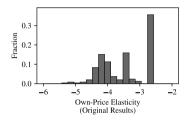
Mixed Logit Estimation

Numerical Best Practices

Differentiation Instruments

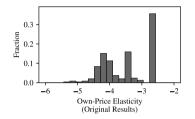
Coding Exercise 2

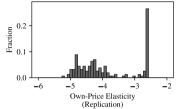
Motivation for Numerical Best Practices



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Motivation for Numerical Best Practices







- Variation in BLP estimates across different optimization algorithms and starting values has disillusioned some researchers (Knittel and Metaxoglou, 2014).
- But there are some numerical best practices that you can follow to avoid these kinds of issues (Conlon and Gortmaker, 2020).
 - ightarrow They're likely to be useful for most structural estimation, not just BLP!

Nonlinear Optimization

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- Set box constraints $\theta \in [\underline{\theta}, \overline{\theta}]$ to preclude unrealistic and unstable guesses of θ .
 - \rightarrow E.g. huge Σ values can make the inner loop unstable.
 - → Economic intuition and initial estimates will give a sense for reasonable bounds.

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- Check that 3-5 different starting values $\theta \sim U(\underline{\theta}, \overline{\theta})$ give the same $\hat{\theta}$.
 - ightarrow For 2-step GMM, do this twice, once for each step (6-10 jobs total).
 - \rightarrow If you have access to a cluster, each can be a separate job, run in parallel.

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- Prefer using gradient-based algorithms for "smooth" problems like BLP.
 - → Avoid derivative-free methods like Nelder-Mead/simplex, which tend to work worse.
 - ightarrow I prefer trust-region algorithms, e.g. SciPy's trust-constr or Knitro if you have it.

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- Try to terminate on strict first-order conditions, e.g. $\|gradient\|_{\infty} < 1e-8$.
 - → Inner loop should be tighter to prevent error "bubbling up." PyBLP default is very tight.
 - ightarrow Can also check second-order conditions, i.e. Hessian eigenvalues are positive.

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- Prefer using gradient-based algorithms for "smooth" problems like BLP.
- Try to terminate on strict first-order conditions, e.g. $\|gradient\|_{\infty} < 1e-8$.
- Configure your optimizer! Defaults may not work for your setting.

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- Individual types i are typically an approximation to a population distribution.
- Sometimes there are only a few types that we can integrate exactly.
 - \rightarrow E.g. high- and low-income types $i \in \{1,2\}$ with known shares w_{1t} and $w_{2t} = 1 w_{1t}$.

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- But usually we approximate the distribution with Monte Carlo integration.
 - \rightarrow Use a random number generator (RNG) to draw $|\mathcal{I}_t| \approx 1,000$ of (ν_{it}, y_{it}) 's per market.
 - → Even better than your default RNG are quasi-Monte Carlo sequences.
 - → I recommend scrambled Halton sequences. R: Owen (2017). Python: SciPy or PyBLP.

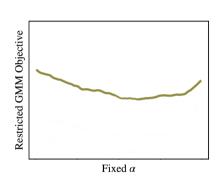
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- But usually we approximate the distribution with Monte Carlo integration.
- If you just need a few $\nu_{it} \sim N(0, I)$'s, try out Gauss-Hermite quadrature.
 - \rightarrow 10-100× fewer carefully-chosen (w_{it}, ν_{it}) 's that do just as well as Monte Carlo.
 - $\,\rightarrow\,$ Chosen to exactly integrate a polynomial expansion of the integrand.

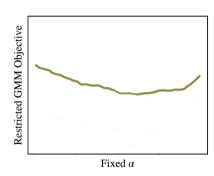
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- But usually we approximate the distribution with Monte Carlo integration.
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- Keep increasing $|\mathcal{I}_t|$ until your estimates stabilize across draws/starting values.

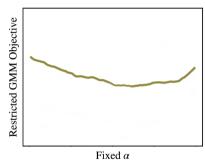
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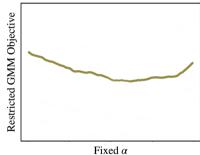
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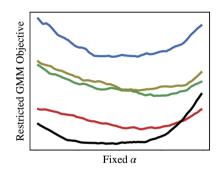
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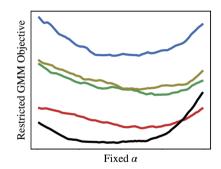
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 - → Poorly-configured optimizers can stop too early.
- Different instruments give different objectives.
 - \rightarrow Even if they're all valid, some may be weaker.
 - \rightarrow Weaker means flatter and harder to optimize.



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- This makes your estimation strategy clear, and makes optimization easier.
 - ightarrow Just-identified models give $Q(\hat{\theta}) \approx 0$ at the optimum.
 - ightarrow This is regardless of your weighting matrix W, so you don't need 2-step GMM.

- For each new parameter in (Σ, Π) , we need another instrument in z_{jt} .
 - → If you have fewer moments than parameters, you're under-identified.
- In general, I recommend starting with one instrument per parameter.
 - \rightarrow Try to choose an instrument that "targets" that parameter.
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- Later, adding more can help with weakness and testing exclusion restrictions.

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• Let's use our stronger intuition about linear regression to think about instruments!

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- If we set $\sigma = \pi = 0$ like on day 1, we get our familiar pure logit regression.
 - \rightarrow Use the same IV as before to target β : if $x_{jt}=p_{jt}$, a price IV; if exogenous, x_{jt} itself.

$$\log \frac{s_{jt}}{s_{0t}} \approx \beta x_{jt} + \sigma^2 d_{jt}^x + \xi_{jt} \quad \text{where} \quad d_{jt}^x = \left(\frac{x_{jt}}{2} - \sum_{k \in \mathcal{J}_t} s_{kt} x_{kt}\right) x_{jt}$$

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 - ightarrow Can technically identify π from higher-order variation, e.g. in variance v_t^y

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- To target π , we can interact x_{jt} with mean within-market income m_t^y .
- In your exercise, you'll target (β, σ, π) with $z_{jt} = (x_{jt}, \sum_{k \neq j} (x_{jt} x_{kt})^2, m_t^y x_{jt})$.
 - \rightarrow If $x_{jt} = p_{jt}$, can replace x_{jt} with fitted values \hat{p}_{jt} from the price IV's first stage.

- There are many valid instruments that satisfy exclusion restrictions $\mathbb{E}[\xi_{it} \mid z_{it}] = 0$.
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- But adding a ton of instruments will bias your estimator.
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$$f^*(z_{jt}) = \mathbb{E}\left[\frac{\partial \xi_{jt}}{\partial \theta'} \middle| z_{jt}\right]$$

- Can be a bit tricky to compute, but with PyBLP it's just one line of code.
 - ightarrow In practice, can update your IVs along with your weighting matrix for a second GMM step.

Roadmap

Preference Heterogeneity

Mixed Logit Estimation

Numerical Best Practices

Differentiation Instruments

- Try to do the second exercise before day 3's class, when I'll do it live.
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 - 2. Mixed logit estimation.
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 - ightarrow What dimensions of preference heterogeneity are missing?
- If you have time, try the supplemental exercises.
 - → Quadrature.
 - \rightarrow Supply-side restrictions.
 - $\rightarrow \ \, \text{Optimal IVs.}$

References I

- **Angrist, Joshua D, Guido W Imbens, and Alan B Krueger**, "Jackknife instrumental variables estimation," *Journal of Applied Econometrics*, 1999, 14 (1), 57–67.
- **Berry, Steven, James Levinsohn, and Ariel Pakes**, "Automobile prices in market equilibrium," *Econometrica*, 1995, *63* (4), 841–890.
- **Berry, Steven T and Philip A Haile**, "Identification in differentiated products markets using market level data," *Econometrica*, 2014, 82 (5), 1749–1797.
- and __ , "Nonparametric identification of differentiated products demand using micro data," 2023.
- **Chamberlain, Gary**, "Asymptotic efficiency in estimation with conditional moment restrictions," *Journal of Econometrics*, 1987, 34 (3), 305–334.

References II

- **Conlon, Christopher and Jeff Gortmaker**, "Best practices for differentiated products demand estimation with PyBLP," *The RAND Journal of Economics*, 2020, *51* (4), 1108–1161.
- **Gandhi, Amit and Jean-François Houde**, "Measuring substitution patterns in differentiated-products industries," 2020.
- **Han, Chirok and Peter CB Phillips**, "GMM with many moment conditions," *Econometrica*, 2006, 74 (1), 147–192.
- **Knittel, Christopher R and Konstantinos Metaxoglou**, "Estimation of random-coefficient demand models: Two empiricists' perspective," *Review of Economics and Statistics*, 2014, 96 (1), 34–59.

References III

Newey, Whitney K and Frank Windmeijer, "Generalized method of moments with many weak moment conditions," *Econometrica*, 2009, 77 (3), 687–719.

Owen, Art B, "A randomized Halton algorithm in R," 2017.

Salanié, Bernard and Frank A Wolak, "Fast, detail-free, and approximately correct: Estimating mixed demand systems," 2022.