

# Large scale risks from agricultural pesticides in small streams

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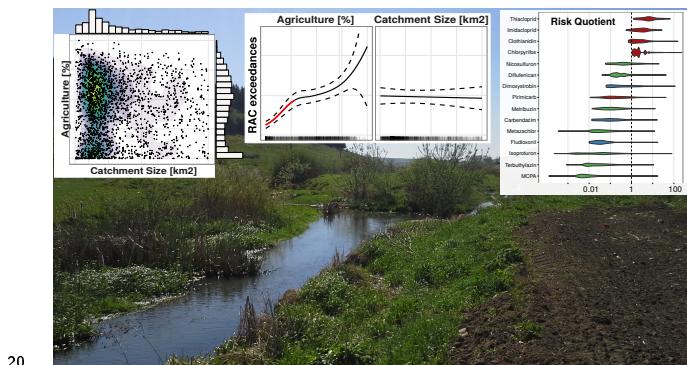
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## Abstract

Small streams are important refugia for biodiversity. In agricultural areas they may be at risk from pesticide pollution. However, most related studies have been limited to a few streams on the regional level, hampering extrapolation to larger scales. We quantified risks as exceedances of regulatory acceptable concentrations (RACs) and used German monitoring data to quantify the drivers thereof and to assess current risks in small streams on a large scale. The data set comprised of 1,766,104 measurements of 478 pesticides (including metabolites) related to 24,743 samples from 2,301 sampling sites. We investigated the influence of agricultural land use, catchment size, as well as precipitation and seasonal dynamics on pesticide risk taking also concentrations below the limit of quantification into account. The exceedances of risk thresholds dropped 3.7-fold at sites with no agriculture, indicating that agricultural land use is a major contributor of pesticides in streams. Precipitation increased detection probability by 43% and concentrations were the highest during summer months. RACs were exceeded in 26% of streams. We found the highest exceedances for neonicotinoid insecticides. We

16 conclude that pesticides from agricultural land use are a major threat to small streams  
17 and their biodiversity. To reflect peak concentrations, current pesticide monitoring  
18 needs to be refined.

19 **TOC Art**



21 **Introduction**

22 More than 50% of the total land area in Germany is used by agriculture<sup>1</sup>. In the year 2014  
23 more than 45,000 tonnes of 776 authorised plant protection products were sold for application  
24 on this area<sup>2</sup>. The applied pesticides may enter surface waters via spray-drift, edge-of-field  
25 run-off or drainage<sup>3–5</sup>. Once entered the surface waters they may have adverse effects on  
26 biota and ecosystem functioning<sup>6</sup>. Although it is known that pesticide pollution and its  
27 ecological effects increase with the fraction of agricultural land use in the catchment<sup>7</sup>, the  
28 shape of the relationship is unknown and studies on potential thresholds are lacking.

29 Two recent studies indicate that pesticides concentrations in streams might threaten  
30 freshwater biodiversity in the European union. Malaj et al.<sup>8</sup> analysed data supplied to  
31 the European Union (EU) in the context of the Water Framework Directive (WFD) and  
32 showed that almost half of European water bodies are at risk from pesticides. Stehle and  
33 Schulz<sup>9</sup> compiled 1,566 measured concentrations of 23 insecticides in the EU from scientific  
34 publications and considerable exceedances of regulatory acceptable concentrations (RAC).

35 However, these studies reflect only a small amount of potentially available data (173 sites in  
36 predominantly mid-sized and large rivers in Malaj et al.<sup>8</sup> and 138 measurements in Stehle  
37 and Schulz<sup>9</sup>), and it is unclear how representative they are for Germany. Much more com-  
38 prehensive data on thousands of sites are available from national monitoring programs that  
39 are setup for the surveillance of water quality, which is done independently by the federal  
40 states in Germany in compliance with the WFD<sup>10</sup>. Despite that these data are providing  
41 the opportunity to study pesticide risks and other research questions on a large scale with  
42 high spatial density, to date these data have not been compiled.

43 Small streams comprise a major fraction of streams<sup>11</sup>, accommodate a higher proportion  
44 of biodiversity compared to larger streams<sup>12,13</sup> and play an important role in the recoloniza-  
45 tion of disturbed downstream reaches<sup>14,15</sup>. Nevertheless, a clear definition of small streams  
46 in terms of catchment or stream size is currently lacking<sup>16</sup>. For example, the WFD defines  
47 small streams with a catchment size between 10 and 100 km<sup>2</sup>, without further categorisation  
48 of streams <10km<sup>2</sup> and Lorenz et al.<sup>16</sup> defines small streams with catchment size <10km<sup>2</sup>.  
49 Moreover, small streams might particularly be at high risk of pesticide contamination in case  
50 of adjacent agricultural areas and given their low dilution potential<sup>5,7</sup>. Indeed, meta-analyses  
51 using data from studies with a few sites reported higher pesticide pollution in smaller streams  
52 compared to bigger streams<sup>7,9</sup>. Despite their ecological relevance and potentially higher pes-  
53 ticide exposure, a recent review of pesticide studies showed that a disproportionately small  
54 fraction of studies was conducted in small water bodies, and these were largely limited to a  
55 few sites<sup>16</sup>. Consequently, knowledge on the pesticide pollution of small streams on larger  
56 scales is scant. In European law, the Directive 2009/128/EC<sup>17</sup> places an obligation on the  
57 EU Member States to adopt National Action Plans (NAP) for the Sustainable Use of Plant  
58 Protection Products and the German NAP also addresses the knowledge gap concerning  
59 pesticide impact on small streams, specifically including those with catchment size <10km<sup>2</sup>.

60 The aim of this study is to identify drivers and dynamics of pesticide concentrations  
61 in streams. To achieve this, we compiled and analysed large-scale pesticide monitoring

62 data from small streams in Germany in order to identify drivers and dynamics of pesti-  
63 cide concentrations. We expect that the landscape is a determinant of measured pesticide  
64 concentrations. Because a major fraction of pesticides is applied to agricultural fields, we  
65 hypothesised highest concentrations and possible exceedances of RACs in streams with a  
66 high proportion of agriculture. Moreover, if agricultural land use is the main source for  
67 pesticides in streams, we expect that exceedances drop to zero if there is not agricultural  
68 land use in the catchment. Moreover, these relationships may show thresholds that could  
69 be to define reference streams without pollution for a future monitoring. Given their low  
70 dilution potential and direct adjacency to agricultural fields we expected that small streams  
71 show highest pesticide concentrations. However, also the timing of sampling may influence  
72 measured concentrations: A sampling directly after a precipitation might show higher con-  
73 centrations because of run-off. Furthermore, pesticides are not applied throughout the whole  
74 year and we expected highest concentrations during the main growing season. Finally, we  
75 quantified the current risks from pesticides in small streams in Germany and the compounds  
76 accountable for the risk.

## 77 Methods

### 78 Data compilation

79 We queried pesticide monitoring data from sampling sites that can be classified as small  
80 streams (catchment sizes < 100 km<sup>2</sup> according to the WFD) from all 13 non-city federal  
81 states of Germany (see Supplemental Table S1 for the abbreviations of federal state names)  
82 for 2005 to 2015. We homogenised and unified all data provided by the federal states into  
83 a database and implemented a robust data-cleaning workflow (see Supplemental Figure S1  
84 for details)<sup>18</sup>.

85 We identified precipitation at sampling sites by a spatio-temporal intersection of sam-  
86 pling events with gridded daily precipitation data (60×30 arcsec resolution) available from

87 the German Meteorological Service (DWD). This data spatially interpolates daily precipi-  
88 tation values from local weather stations<sup>19</sup>. We performed the intersection for the actual  
89 sampling date and the day before and extracted precipitation during and up to 48 hours  
90 before sampling.

## 91 Characterization of catchments

92 We compiled a total of 2,369 sampling sites in small streams with pesticide measurements.  
93 Alongside, we also queried catchment sizes and agricultural land use within the catchment  
94 for the sampling sites from the federal states. Catchment size was provided for 59% of sites.  
95 Additionally, we delineated upstream catchments for each of the sampling sites using (i) a  
96 digital elevation model (DEM)<sup>20</sup> and the multiple flow direction algorithm<sup>21</sup> as implemented  
97 in GRASS GIS 7<sup>22</sup> and (ii) from drainage basins provided by the Federal Institute of Hy-  
98 drology (BfG). Delineated catchments were visually checked for accuracy by comparison of  
99 coverage with stream networks provided by the federal states. Thus, catchment size infor-  
100 mation was available for 99% of all sites (59% from authorities, 24% from DEM and 16%  
101 from drainage basins).

102 For each derived catchment (either from DEM or drainage basins) we calculated the  
103 % agricultural land-use within the catchment based on the Authoritative Topographic-  
104 Cartographic Information System (ATKIS) of the land survey authorities<sup>23</sup>. Thus, agri-  
105 cultural land use information was available for 98% of all sites (24% from authorities, 52%  
106 from DEM and 22% from drainage basins). 68 sites (3%) that lacked catchment size or land  
107 use information were omitted from the analysis, resulting in 2301 sites used in the analyses  
108 outlined below.

## 109 Characterization of pesticide pollution

110 We characterised pesticide pollution using regulatory acceptable concentrations (RAC)<sup>24</sup>.  
111 RACs are derived during pesticide authorisation as part of the ecological risk assessment

112 (ERA). According to the goals of ERA, exceedances of RACs should not occur after pes-  
 113 ticide authorisation. No unacceptable ecological effects are expected if the environmental  
 114 concentration remains below the RAC. Stehle and Schulz<sup>9</sup> showed that RAC exceedances re-  
 115 flect a decrease in biodiversity and from this perspective are ecologically relevant indicators.  
 116 The German Environment Agency (UBA) provided RACs for 107 compounds, including  
 117 those with the highest detection rates (Supplemental Table S2). Based on these RACs, we  
 118 calculated Risk Quotients (RQ):

$$RQ_i = \frac{C_i}{RAC_i} \quad (1)$$

119 where  $C_i$  is the concentration of a compound  $i$  in a sample and  $RAC_i$  the respective  
 120 RAC.

## 121 Statistical analyses

122 As outlined in the introduction, we expected non-linear responses to agricultural land use and  
 123 catchment size and searched for potential thresholds (defined as abrupt changes). Therefore,  
 124 we used generalised additive models (GAM) to establish relationships<sup>25</sup>. We modelled the  
 125 number of RAC exceedances ( $RQ > 1$ ) at a site as:

$$\begin{aligned} No(RQ > 1)_i &\sim NB(\mu_i, \kappa) \\ \log(\mu_i) &= \beta_0 + f_1(agri_i) + f_2(size_i) + \log(n_i) \end{aligned} \quad (2)$$

126 where  $No(RQ > 1)_i$  is the observed number of RAC exceedances at site  $i$ . Because of  
 127 overdispersion, we modelled  $No(RQ > 1)_i$  as resulting from a negative binomial distribution  
 128 ( $NB$ ) with mean  $\mu_i$  and a quadratic mean-variance-relationship ( $Var(No(RQ > 1)_i) =$   
 129  $\mu_i + \frac{\mu_i^2}{\kappa}$ ). The proportion of agricultural land use within the catchment ( $agri_i$ ) and the  
 130 catchment size of the site ( $size_i$ ) were used as predictors of the number of RAC exceedances.

131  $\beta_0$  is the intercept and  $f_1$  and  $f_2$  are smoothing functions using penalized cubic regression  
 132 splines<sup>26,27</sup>. The number of measurements per site ( $n_i$ ) was used as an offset to account  
 133 for differences in sampling efforts at a site (in terms of number of samples and analysed  
 134 compounds) and is equivalent to modelling the rate of exceedances. We used point-wise 95%  
 135 Confidence Intervals (CI) of the first derivative of the fitted smooth to identify regions of  
 136 statistically significant changes. All data-processing and analyses were performed using R<sup>28</sup>.  
 137 GAMs were fitted using the mgcv package<sup>27</sup>.

138 To assess the influence of precipitation and seasonality, we modelled the RQ of individual  
 139 compounds as the response variable. RQ and concentrations show a skewed distribution  
 140 with an excess of zeros (no pesticides detected and quantified). Therefore, we modelled  
 141 these as two processes (one generating values below the limit of quantification (LOQ) and  
 142 one generating values above LOQ) using a Zero-Adjusted Gamma (ZAGA) distribution<sup>29,30</sup>  
 143 (Equation 3). These two processes can be interpreted as changes in the mean value of RQ  
 144 (change in  $\mu$ ) and changes in the probability of exceeding LOQ and showing any risk (change  
 145 in  $\nu$ ).

$$RQ_i \sim ZAGA(\mu_i, \sigma, \nu_i) = \begin{cases} (1 - \nu_i) & \text{if } y < LOQ \\ \nu_i \times f_{Gamma}(\mu_i, \sigma) & \text{if } y \geq LOQ \end{cases} \quad (3)$$

146  $\nu_i$  denotes the probability of a measurement  $i$  being above LOQ and  $f_{Gamma}$  denotes the  
 147 gamma function and is used for values equal to or greater LOQ, with  $\mu$  being the mean  
 148 and  $\sigma$  the standard deviation of RQ. We used the  $\log(x + 0.05)$  transformed precipitation  
 149 at sampling date ( $\log prec_0$ ) and the day before ( $\log prec_{-1}$ ), as well as quarters of the year  
 150 ( $Q1 - Q4$ ) as linear predictors for  $\mu$  and  $\nu$ . We used appropriate link functions for  $\mu$  and  $\nu$   
 151 and assumed  $\sigma$  to be constant. Equation 4 summarises the deterministic part of the model  
 152 for a measurement  $i$ .

$$\begin{aligned}\log(\mu_i) &= \log(prec_{0i}) + \log(prec_{-1i}) + Q1_i + Q2_i + Q3_i + Q4_i \\ logit(\nu_i) &= \log(prec_{0i}) + \log(prec_{-1i}) + Q1_i + Q2_i + Q3_i + Q4_i\end{aligned}\tag{4}$$

153 To account for differences between federal states we used *site* nested within *state* as  
154 random intercepts. We implemented this model using the *gamlss* package.<sup>31</sup>

155 We fitted this model separately to each compound with a RAC, measured in at least 1000  
156 samples and with more than 5% of values above LOQ (n = 22 compounds, see Supplemental  
157 Table S3 for a list of compounds). To summarise the coefficients across the 22 modelled  
158 compounds we used a random effect meta-analysis for each model coefficient separately<sup>32</sup>,  
159 resulting in an averaged effect of the 22 compounds. The results of individual compounds  
160 are provided in the Supplemental Table S4 and Figures S6 and S7. The meta-analysis was  
161 performed using the *metafor* package<sup>33</sup>.

## 162 Results

### 163 Overview of the compiled data

164 The compiled dataset used for analysis comprised 1,766,104 pesticide measurements in 24,743  
165 samples from 2,301 sampling sites in small streams. These samples were all taken via grab  
166 sampling. We found large differences between federal states in the number of sampling  
167 sites and their spatial distribution (Figure 1 and Supplemental Table S1). The number of  
168 small stream sampling sites per state ranged from 1 (Lower Saxonia, NI) to 1139 (North  
169 Rhine-Westphalia, NW). No data were available from Brandenburg.

170 In total 478 different compounds used as pesticides and their metabolites were measured  
171 at least once (Supplemental Table S2). Most of the compounds were herbicides (179), fol-  
172 lowed by insecticides (117) and fungicides (109). Most samples were taken in the months  
173 April till October, while fewer samples were taken during winter (see Supplemental Fig-

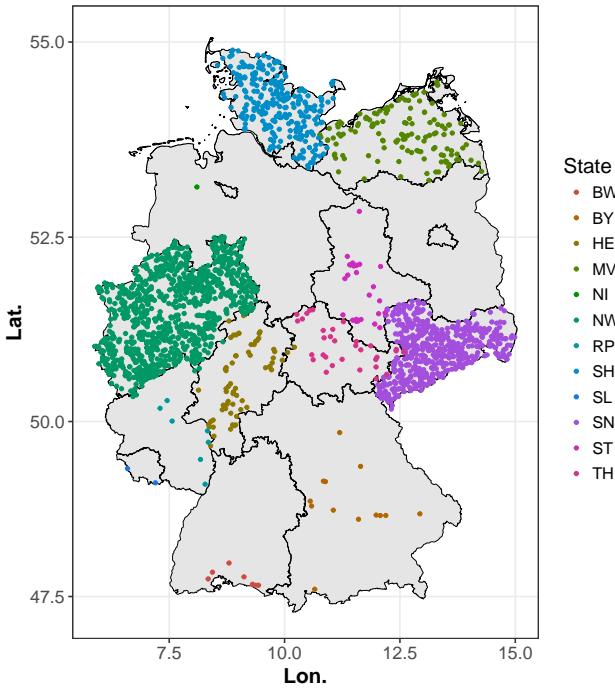


Figure 1: Spatial distribution of the 2,301 small stream sampling sites. Colour codes different federal states (see Supplemental Table S1 for abbreviations).

ure S2). We found substantial differences in the spectra of analysed pesticides between federal states (Figure 2). The number of analysed pesticides per state ranged from 57 (SL) to 236 (RP) (Supplemental Table S1). 4% (=71,113) of all measurements were concentrations above LOQ.

The distribution of sampling sites across catchment sizes indicated a disproportionately low number of sites with catchments below  $10 \text{ km}^2$ , with most sampling sites having catchment sizes between 10 and  $25 \text{ km}^2$  (Figure 3).

## **Influence of agricultural land use and catchment size**

We found a positive relationship between agricultural land use and the number of RAC exceedances. The non-linear model showed, that below 28% agriculture the mean number of RAC exceedances dropped statistically significant 3.7-fold from 0.39 (28% agriculture within

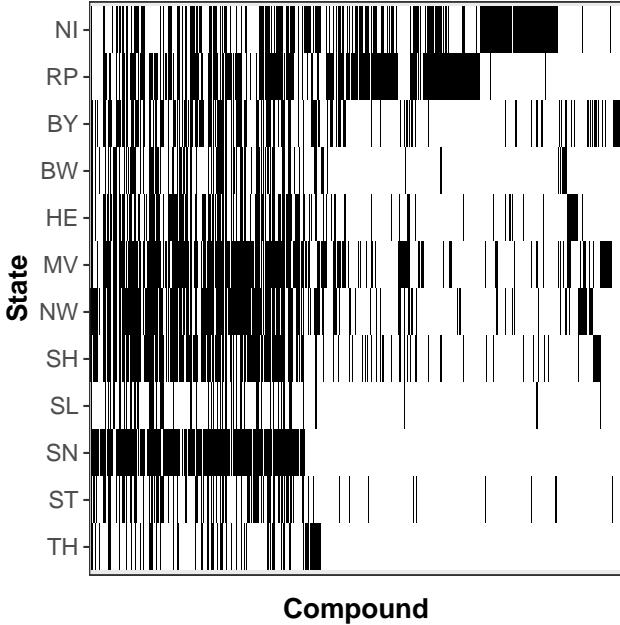


Figure 2: Barcode plot of compound spectra of the federal states. Each vertical line is an analysed compound.

the catchment) to 0.10 (no agriculture) (Figure 4, left). . Catchment size had no statistically significant effect on the number of RAC exceedances (Figure 4, right). We also could not detect a statistically significant interaction between catchment size and agriculture.

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## 188 Effect of precipitation on pesticide risk

189  $prec_0$  and  $prec_{-1}$  increased the probability of exceeding LOQ and RQ. In  $Q2$  an increase from  
 190 0.1 mm to 15 mm of precipitation before sampling ( $prec_{-1}$ ) lead on average to a 43% higher  
 191 mean RQ of 0.05 (Supplemental Figure S7). The probability to exceed LOQ increases in  $Q2$   
 192 1.6-fold from 8.7% to 13.5% (Figure 5). Precipitation before sampling ( $prec_{-1}$ ) had a stronger  
 193 effect than precipitation during sampling ( $prec_0$ ) on the probability of exceeding LOQ. This  
 194 difference was less pronounced for the mean value of RQ (Supplemental Figure S7, top left).  
 195 Moreover, effects differed between individual compounds (see Supplemental Table S4).

196 The first quarter showed the lowest RQ and probability of exceeding LOQ. Both increased  
 197 during summer months and decreased towards the end of the year. There was a 2.5-fold

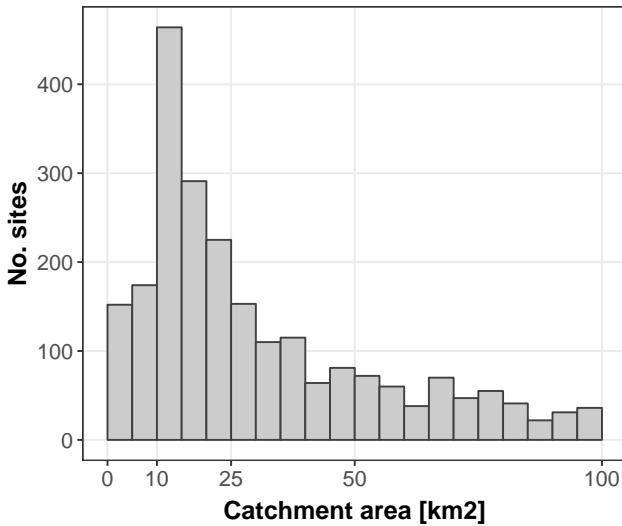


Figure 3: Distribution of catchment area across the sampling sites.

higher probability of exceeding LOQ in  $Q_2$  (10.6%) than in  $Q_1$  (4.6%) (Figure 5). The differences were less pronounced for the mean value of RQ and with less precision (see Supplemental Figure S7, left). Individual compounds showed different temporal patterns (see Supplemental Table S4).

## Current pesticide risks in small streams

We found RAC exceedances in 25.5% of sampling sites and  $RQ > 0.1$  in 54% of sites. In 23% of sites none of the chemicals, for which RACs were available, were detected (see also Supplemental Figure S8). Neonicotinoid insecticides and Chlorpyrifos showed the highest RQ (Figure 6). For Thiacloprid and Chlorpyrifos the RAC was equal or less than LOQ, therefore, all detections have a  $RQ \geq 1$ . The herbicides Nicosulfuron and Diflufenican, as well as the fungicide Dimoxystrobin also showed high exceedances of RQ (26.7, 14.1 and 21.1 % of measurements  $>$  LOQ), see also Supplemental Table S5). RAC exceedances were found in 14% of samples with concentrations  $>$  LOQ (and 7.3% of all samples).

The highest RQs were observed for Chlorpyrifos ( $\max(RQ) = 220$ ), Clothianidin ( $\max(RQ) = 157$ ), Dimoxystrobin ( $\max(RQ) = 117$ ) and Isoproturon ( $\max(RQ) = 80$ ). Where analysed, metabolites exhibited the highest detection rates (for example, Metazachlor sulfonic

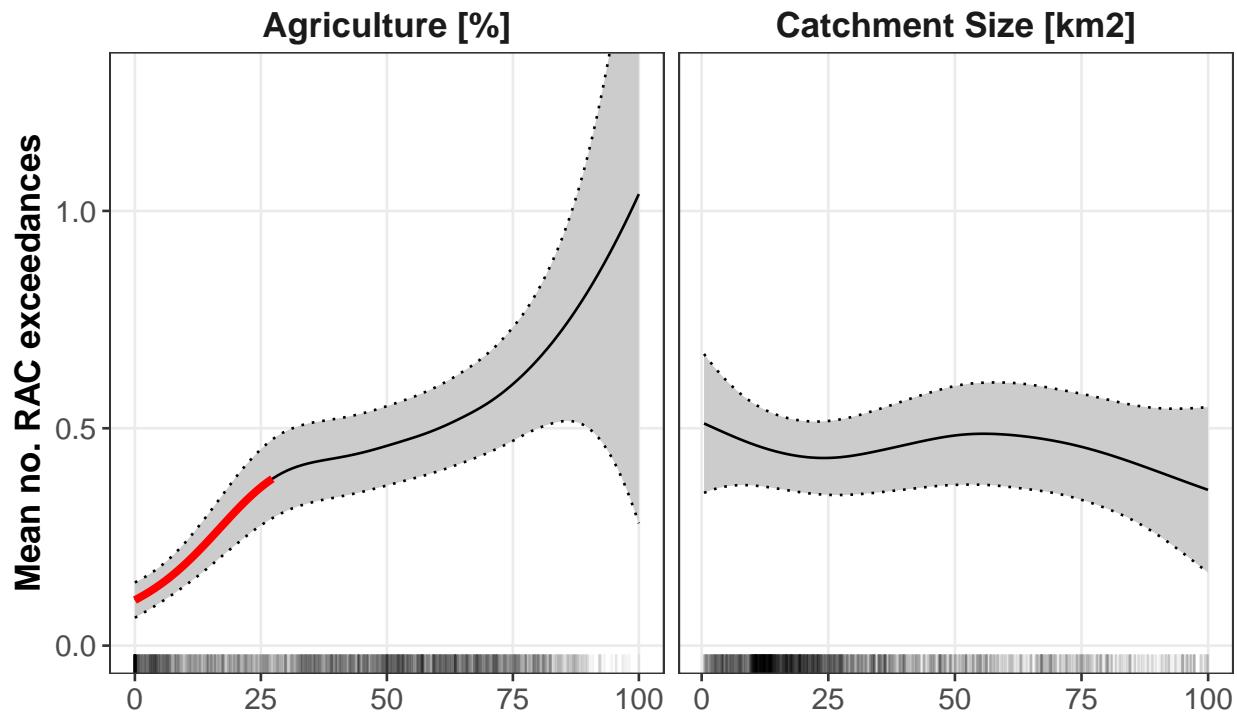


Figure 4: Effect of percent agriculture within the catchment (left) and catchment size (right) on the mean number of RAC exceedances per site. Red line marks statistically significant changes. Dashed lines denote 95% point-wise Confidence Intervals.

acid was detected in 84% of all samples where it was analysed ( $n = 3038$ , see also Supplemental Figure S9). Glyphosate was the compound with the highest detection rates (41%,  $n = 3557$  samples), followed by Boscalid (23%,  $n = 9886$ ) and Isoproturon (22%,  $n = 19112$ ). However, only the latter showed RAC exceedances (Figure 6). In 45.9% of samples more than one compound was quantified, with a maximum of 54 different compounds in one sample (Supplemental Figure S10).

## Discussion

### Overview on the compiled dataset

The compiled dataset of governmental monitoring data, with a particular focus on small streams, represents currently the most comprehensive available for Germany. Similar na-

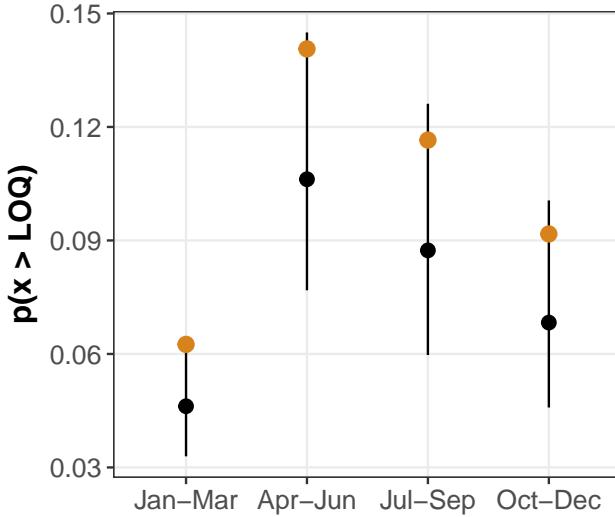


Figure 5: Summarised model predictions for the probability to exceed LOQ throughout the year. Black points indicate the probabilities at 0.1 mm precipitation (and their 95% CI). Orange points indicated the probabilities at 15 mm precipitation. Probabilities have been summaries from a meta-analysis of the 22 modelled compounds. Single compound coefficients are provided in Supplemental Table S4 and Figure S7.

224 nationwide datasets have been compiled for the Netherlands<sup>34</sup>, Switzerland<sup>35</sup> and the United  
 225 States<sup>36</sup>. While the compilations from Europe are of similar quantity and quality to the data  
 226 compiled and analysed here, the compilation used in Stone et al.<sup>36</sup> is much smaller, though  
 227 these data may be complemented by more data in future analyses.

228 A nationwide assessment of pesticide pollution is hampered by inhomogeneous data across  
 229 federal states: Beside large differences in the spatial distribution and quantity of sampling  
 230 sites (Figure 1), the spectrum of analysed compounds (Figure 2) and the quality of chemical  
 231 analyses differed between states. Despite the outlined differences between states, all ecore-  
 232 gions occurring in Germany<sup>37,38</sup> were covered by the presented dataset (see Supplemental  
 233 Figure S3) and thus it might nonetheless represent a major fraction of small stream types  
 234 in Germany.

235 Nevertheless, the differences in monitoring efforts among states hamper a spatial interpo-  
 236 lation. We accounted for differences in sampling efforts per site, by taking the total number of  
 237 measurement into account. However, we acknowledge that this takes only partly differences

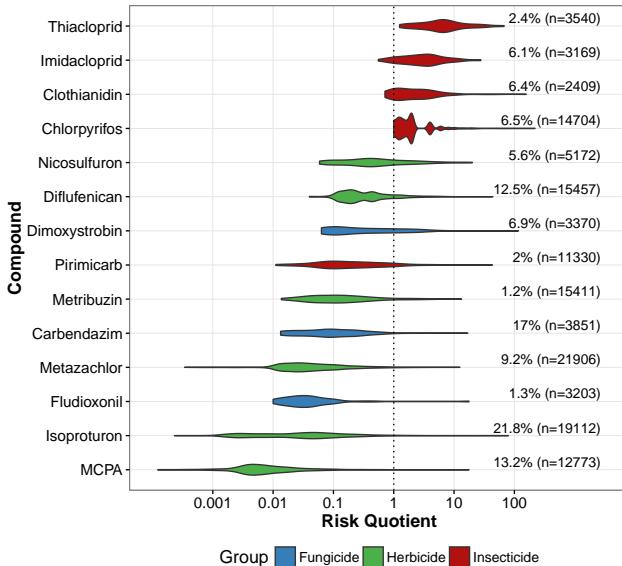


Figure 6: 15 compounds with the highest observed risk quotients in small streams. Non-detects are not shown due to the logarithmic axis. Numbers on the right give the percentage of values >LOQ and the total number of samples were the compound was analysed.

238 between sampling intervals into account which might influence our results<sup>3,39</sup>. Moreover, it  
 239 is known that differences in analytical quality might influence estimated effects<sup>7</sup>. However,  
 240 the used model (Equation 3) takes explicitly LOQs and differences therein into account.

241 For Thiacloprid and Chlorpyrifos the LOQs were above the RAC, which means that  
 242 exceedances are likely underestimated. For compounds with low RACs a lowering of LOQ  
 243 through an improvement of chemical analysis is essential for reliable assessment. Moreover,  
 244 a nationwide assessment would benefit from a harmonised spectrum of analysed compounds  
 245 between federal states.

246 Given their high abundance in the landscape<sup>11</sup> small streams below 10 km<sup>2</sup> are dispro-  
 247 portionally less sampled in current monitoring (Figure 3), which may be attributed to the  
 248 missing categorisation in the WFD. Clearly, there is currently a lack of knowledge on stres-  
 249 sor effects on small streams. We analysed only data from small streams, however, for lentic  
 250 small water bodies this lack might be even greater<sup>16</sup>.

**251 Influence of agricultural land use and catchment size**

**252** As hypothesised, we found a positive relationship between agricultural land use and the  
**253** number of RAC exceedances. Especially, we found a statistically significant drop below 28%  
**254** of agricultural land use (Figure 4, left). This drop indicates that agricultural land use might  
**255** be a major contributor to the observed exceedances. The absence of such a drop would have  
**256** indicated that there are exceedances that cannot be explained by agricultural land use, like  
**257** inputs from urban land uses.

**258** We did not find a statistically significant relationship between pesticide pollution and  
**259** catchment size. However, previous studies showed that small streams are more polluted than  
**260** bigger streams<sup>7,9,40</sup>. This can be explained by the relatively short gradient of catchment sizes  
**261** in our dataset, with most of the streams with catchments above 10 km<sup>2</sup> and below 100 km<sup>2</sup>  
**262** (Figure 3, top). For example, the gradient of Schulz<sup>7</sup> covered 6 orders of magnitude.

**263 Effect of precipitation on pesticide risk**

**264** We found a 36% higher RQ if samples were taken after rainfall events, that correspond to  
**265** the hypothesis that pesticides enter streams via run-off. However, samples taken on the  
**266** day of a rainfall event showed less risk than samples taken one day after a rainfall event.  
**267** This discrepancy could be explained by a sampling preceding the rainfall event because the  
**268** temporal resolution of our dataset was 1 day. Additionally, this might be also explained by  
**269** a delay between the start of a rain event and the peak in discharge or runoff.

**270** The effects of precipitation were more pronounced for the probability to exceed LOQ,  
**271** with smaller effect sizes for the absolute value of RQ. This may be explained by a higher  
**272** variability of absolute concentrations. Overall, our results indicate that current pesticide  
**273** monitoring relying on grab sampling, largely disconnected from precipitation events, under-  
**274** estimates pesticide risks. Automatic event-driven samplers<sup>3</sup> and passive samplers<sup>41,42</sup> may  
**275** help overcome these shortcomings and provide a better representation of risks, especially for  
**276** small water bodies<sup>16</sup>.

277 We found the highest the probability of exceeding LOQ during summer (10% for Q2)  
278 and lowest in the first quarter of the year (4%, Figure 5, bottom right). This annual pattern  
279 coincides, as expected, with the main application season for pesticides in Central Europe.  
280 Nevertheless, there are compound-specific differences in the annual pattern, which explains  
281 the wide CI for the absolute RQ (Figure 5, bottom left). For example, the herbicide Di-  
282 flufenican showed the highest RQ and the highest probability of exceeding LOQ during the  
283 winter quarters Q1 and Q4 (Supplemental Table S4), which coincides with the application  
284 period it is registered for in Germany<sup>43</sup>. Moreover, compound properties, like half-life or  
285 water solubility, might influence compound dynamics. Our study suggests that pesticide  
286 risks display compound specific spatio-temporal dynamics. Currently, little is known about  
287 these and further research on those might provide useful information for future ecological  
288 risk assessment. For example, the sensitivity of organisms is often life stage dependent<sup>44</sup> and  
289 knowledge on temporal dynamics could inform on concurrent exposure to multiple pesticides,  
290 as well as assist to parameterise toxicokinetic and toxicodynamic models<sup>45</sup>. Moreover, our  
291 results show that analysing absolute concentrations and probabilities of LOQ together might  
292 deliver valuable insights into risk dynamics.

## 293 Pesticides in small streams

294 Our results suggest that small streams are frequently exposed to ecologically relevant pes-  
295 ticide concentrations. In one-quarter of small streams RACs were exceeded at least once.  
296 Stehle and Schulz<sup>9</sup> found the highest percentage of RAC exceedances for organophosphate  
297 insecticides. By contrast, we found that neonicotinoid insecticides have highest exceedances  
298 of RACs, followed by the organophosphate chlorpyrifos. This difference can be attributed to  
299 the low sample size for neonicotinoid insecticides in their study ( $n = 33$ ) compared to the  
300 dataset presented here (for example 3,540 samples of Thiacloprid, Figure 6). Overall, our  
301 results suggest that neonicotinoids may currently pose a high risk to freshwater ecosystems.  
302 Moreover, our results add further evidence to the growing literature on the risks arising from

303 neonicotinoids for aquatic<sup>46</sup> and terrestrial<sup>47</sup> ecosystems.

304 Compared to Stehle and Schulz<sup>9</sup> we found higher rates of RAC exceedances for insec-  
305 ticides. They found exceedances in 37.1% of insecticide measurements >LOQ (n = 1352,  
306 23 insecticides), whereas, we found exceedances in 67% of insecticide measurements with  
307 RACs >LOQ (n = 1855, 22 insecticides). This could be attributed to different insecticides  
308 considered and different underlying RACs. Our study has only 7 insecticides with RACs in  
309 common with the insecticides investigated by Stehle and Schulz<sup>9</sup>. Moreover, all RACs were  
310 lower in our study (average difference = -0.71 µg/L, range = [-2.757; -0.005]). Nevertheless,  
311 it must be noted that the dataset compiled here comprised only samples from grab sampling,  
312 which may considerably underestimate pesticide exposure<sup>3,39</sup>.

313 By contrast, Knauer<sup>40</sup> found exceedances from monitoring data mainly for herbicides  
314 and fungicides and only one insecticide Chlorpyrifos-methyl. Moreover, RAC exceedances in  
315 Switzerland were generally lower and less abundant (for example 6 exceedances (=0.2%) for  
316 Isoproturon with a maximum RQ of 2) compared to our results for Germany. This might  
317 reflect differences in pesticide use between countries, ecoregions and RACs used. From  
318 the definition of RAC it follows that if the concentration of a compound exceeds its RAC  
319 ecological effects are expected. Indeed, Stehle and Schulz<sup>48</sup> found that the biological diversity  
320 of stream invertebrates was significantly reduced by 30% at RQ = 1.12 and by 10% at 1/10  
321 of RAC. We found RQ values greater than 1.12 in 25% of small streams and RQ at 1/10 of  
322 RAC in 54% of small streams. Consequently, we conclude that agricultural pesticides are  
323 on a large scale a major threat to small streams, the biodiversity they host and the services  
324 they provide. This threat may exacerbate because pesticides often occur in mixtures<sup>49</sup> and  
325 may co-occur with other stressors<sup>50</sup>.

326 Monitoring data, despite the outlined limitations, provides an opportunity to study large-  
327 scale environmental occurrence patterns of pesticides. Furthermore, such nationwide com-  
328 pilations, may not only be used for governmental surveillance, but also to answer other  
329 questions, like validation of exposure modelling,<sup>51</sup> retrospective evaluation of regulatory risk

330 assessment<sup>9,40</sup> or occurrences of pesticide mixtures.<sup>49</sup> However, the sampling design needs to  
331 account for precipitation events to provide robust data. Our results suggest that exceedances  
332 of RACs are landscape dependent and therefore, pesticide regulation should account for  
333 landscape features. Moreover, the high exceedances of RACs indicate that greater efforts  
334 are needed to describe causal links, which may lead to further developments of the current  
335 authorisation procedure.

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## 344 Supporting Information Available

345 The following files are available free of charge.

- 346 • Supplemental\_Materials.pdf : Supplemental Materials (Figures, Tables, Models).
- 347 This material is available free of charge via the Internet at <http://pubs.acs.org/>.

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