

# Large scale risks from agricultural pesticides in small streams

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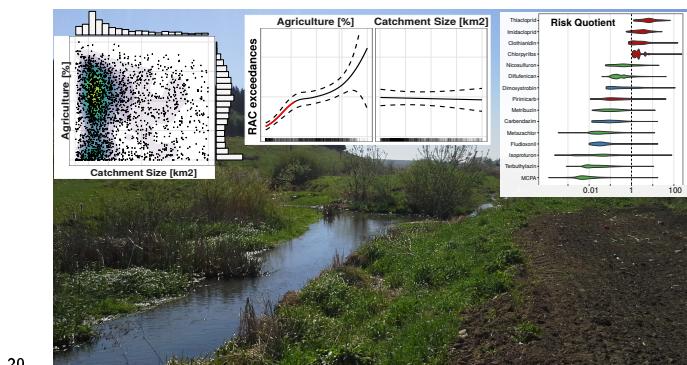
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## Abstract

Small streams are important refugia for biodiversity. In agricultural areas they may be at risk from pesticide pollution. However, most related studies have been limited to a few streams on the regional level, hampering extrapolation to larger scales. We quantified risks as exceedances of regulatory acceptable concentrations (RACs) and used German monitoring data to quantify the drivers thereof and to assess current risks in small streams on a large scale. The data set comprised of 1,766,104 measurements of 478 pesticides (including metabolites) related to 24,743 samples from 2,301 sampling sites. We investigated the influence of agricultural land use, catchment size, as well as precipitation and seasonal dynamics on pesticide risk taking also concentrations below the limit of quantification into account. The exceedances of risk thresholds dropped 3.7-fold at sites with no agriculture. Precipitation increased detection probability by 43% and concentrations were the highest from April to June. Overall, this indicates

14 that agricultural land use is a major contributor of pesticides in streams. RACs were  
15 exceeded in 26% of streams, with the highest exceedances found for neonicotinoid in-  
16 secticides. We conclude that pesticides from agricultural land use are a major threat to  
17 small streams and their biodiversity. To reflect peak concentrations, current pesticide  
18 monitoring needs refinement.

19 **TOC Art**



20

21 **Introduction**

22 More than 50% of the total land area in Germany is used by agriculture<sup>1</sup>. In the year 2014  
23 more than 45,000 tonnes of 776 authorised plant protection products were sold for application  
24 on this area<sup>2</sup>. The applied pesticides may enter surface waters via spray-drift, edge-of-field  
25 run-off or drainage<sup>3–5</sup>. Once having entered the surface waters they may have adverse effects  
26 on biota and ecosystem functioning<sup>6</sup>. Although it is known that pesticide pollution and its  
27 ecological effects increase with the fraction of agricultural land use in the catchment<sup>7</sup>, the  
28 shape of the relationship is unknown and studies on potential thresholds are lacking.

29 Two recent studies indicate that pesticide concentrations in streams might threaten fresh-  
30 water biodiversity in the European union. Malaj et al.<sup>8</sup> analysed data supplied to the Eu-  
31 ropean Union (EU) in the context of the Water Framework Directive (WFD) and showed  
32 that almost half of European water bodies are at risk from pesticides. Stehle and Schulz<sup>9</sup>

33 compiled 1,566 measured concentrations of 23 insecticides in the EU from scientific publi-  
34 cations and found considerable exceedances of regulatory acceptable concentrations (RAC).  
35 However, these studies reflect only a small amount of potentially available data (173 sites in  
36 predominantly mid-sized and large rivers in Malaj et al.<sup>8</sup> and 138 measurements in Stehle  
37 and Schulz<sup>9</sup>), and it is unclear how representative they are for Germany. Much more com-  
38 prehensive data on thousands of sites are available from national monitoring programs that  
39 are setup for the surveillance of water quality, which is done independently by the federal  
40 states in Germany in compliance with the WFD<sup>10</sup>. Although, these data are providing the  
41 opportunity to study pesticide risks and other research questions on a large scale with high  
42 spatial density, to date these data have not been compiled.

43 Small streams comprise a major fraction of streams<sup>11</sup>, accommodate a higher proportion  
44 of biodiversity compared to larger streams<sup>12,13</sup> and play an important role in the recoloniza-  
45 tion of disturbed downstream reaches<sup>14,15</sup>. Nevertheless, a clear definition of small streams  
46 in terms of catchment or stream size is currently lacking<sup>16</sup>. For example, the WFD defines  
47 small streams with a catchment size between 10 and 100 km<sup>2</sup>, without further categorisation  
48 of streams <10km<sup>2</sup> and Lorenz et al.<sup>16</sup> defines small streams with catchment size <10km<sup>2</sup>.  
49 Moreover, small streams might particularly be at high risk of pesticide contamination in case  
50 of adjacent agricultural areas and given their low dilution potential<sup>5,7</sup>. Indeed, meta-analyses  
51 using data from studies with a few sites reported higher pesticide pollution in smaller streams  
52 compared to bigger streams<sup>7,9</sup>. Despite their ecological relevance and potentially higher pes-  
53 ticide exposure, a recent review of pesticide studies showed that a disproportionately small  
54 fraction of studies was conducted in small water bodies, and these were largely limited to a  
55 few sites<sup>16</sup>. Consequently, knowledge on the pesticide pollution of small streams on larger  
56 scales is scant. In European law, the Directive 2009/128/EC<sup>17</sup> places an obligation on the  
57 EU Member States to adopt National Action Plans (NAP) for the Sustainable Use of Plant  
58 Protection Products and the German NAP also addresses the knowledge gap concerning  
59 pesticide impact on small streams, specifically including those with catchment size <10km<sup>2</sup>.

60 The aim of this study is to identify drivers and dynamics of pesticide concentrations  
61 in streams on large spatial scales. To achieve this, we compiled and analysed large-scale  
62 pesticide monitoring data from small streams in Germany and examined four hypotheses: 1)  
63 A major fraction of pesticides is applied to agricultural fields. Therefore, we hypothesised  
64 that the most frequent exceedances of RACs occur in streams with a high proportion of  
65 agricultural land use in the catchment. If agricultural land use was indeed the main source  
66 for pesticides in streams, the RAC exceedances should drop to negligible levels in the absence  
67 of agricultural land use in their catchments. Given this possible drop we expected a non-  
68 linear relationship between exceedances and agricultural land use. In case of a non-linear  
69 relationship, our analyses might guide the definition of reference streams without pesticide  
70 pollution in future monitoring. 2) Based on previous studies, we hypothesised that an  
71 increase in catchment size is associated with an decrease in RAC exceedances<sup>7,9</sup>. 3) However,  
72 also the timing of sampling may influence measured concentrations, as local and regional  
73 studies reported higher pesticide concentrations after precipitation events<sup>5,18</sup>. Therefore,  
74 we hypothesised the highest RAC exceedances to be found after precipitation events. 4)  
75 Pesticides are not applied throughout the whole year and highest RAC exceedances should  
76 be found during the main growing season. Finally, we quantified the current risks from  
77 pesticides in small streams in Germany and the compounds accountable for the risk.

## 78 Methods

### 79 Data compilation

80 We queried pesticide monitoring data from sampling sites that can be classified as small  
81 streams (catchment sizes < 100 km<sup>2</sup> according to the WFD) from all 13 non-city federal  
82 states of Germany (see Supporting Information Table S1 for the abbreviations of federal  
83 state names) for the years 2005 to 2015. We homogenised and unified all data provided  
84 by the federal states into a database and implemented a robust data-cleaning workflow (see

85 Supporting Information Figure S1 for details)<sup>19</sup>.

86 We identified precipitation at sampling sites by a spatio-temporal intersection of sam-  
87 pling events with gridded daily precipitation data ( $60 \times 30$  arcsec resolution) available from  
88 the German Meteorological Service (DWD). This data spatially interpolates daily precipi-  
89 tation values from local weather stations<sup>20</sup>. We performed the intersection for the actual  
90 sampling date and the day before and extracted precipitation during and up to 48 hours  
91 before sampling.

## 92 Characterization of catchments

93 We compiled a total of 2,369 sampling sites in small streams with pesticide measurements.  
94 Alongside, we also queried catchment sizes and agricultural land use within the catchment  
95 for the sampling sites from the federal states. Catchment size was provided for 59% of sites.  
96 Additionally, we delineated upstream catchments for each of the sampling sites using (i) a  
97 digital elevation model (DEM)<sup>21</sup> and the multiple flow direction algorithm<sup>22</sup> as implemented  
98 in GRASS GIS 7<sup>23</sup> and (ii) from drainage basins provided by the Federal Institute of Hy-  
99 drology (BfG). Delineated catchments were visually checked for accuracy by comparison of  
100 coverage with stream networks provided by the federal states. Thus, catchment size infor-  
101 mation was available for 99% of all sites (59% from authorities, 24% from DEM and 16%  
102 from drainage basins).

103 For each derived catchment (either from DEM or drainage basins) we calculated the  
104 % agricultural land-use within the catchment based on the Authoritative Topographic-  
105 Cartographic Information System (ATKIS) of the land survey authorities<sup>24</sup>. Thus, agri-  
106 cultural land use information was available for 98% of all sites (24% from authorities, 52%  
107 from DEM and 22% from drainage basins). 68 sites (3%) that lacked catchment size or land  
108 use information were omitted from the analysis, resulting in 2301 sites used in the analyses  
109 outlined below.

<sub>110</sub> **Characterization of pesticide pollution**

<sub>111</sub> We characterised pesticide pollution using regulatory acceptable concentrations (RAC)<sup>25</sup>.  
<sub>112</sub> RACs are derived during pesticide authorisation as part of the environmental risk assessment  
<sub>113</sub> (ERA). According to the goals of ERA, exceedances of RACs should not occur after pesticide  
<sub>114</sub> authorisation<sup>9</sup> and thus no unacceptable ecological effects are expected if the environmental  
<sub>115</sub> concentration remains below the RAC. Stehle and Schulz<sup>9</sup> showed that RAC exceedances  
<sub>116</sub> reflect a decrease in biodiversity and from this perspective are ecologically relevant indicators.  
<sub>117</sub> The German Environment Agency (UBA) provided RACs for 107 compounds, including  
<sub>118</sub> those with the highest detection rates (Supporting Information Table S2). Based on these  
<sub>119</sub> RACs, we calculated Risk Quotients (RQ):

$$RQ_i = \frac{C_i}{RAC_i} \quad (1)$$

<sub>120</sub> where  $C_i$  is the concentration of a compound  $i$  in a sample and  $RAC_i$  the respective  
<sub>121</sub> RAC.

<sub>122</sub> **Statistical analyses**

<sub>123</sub> As outlined in the introduction, we expected non-linear responses to agricultural land use and  
<sub>124</sub> catchment size and searched for potential thresholds (defined as abrupt changes). Therefore,  
<sub>125</sub> we used generalised additive models (GAM) to establish relationships<sup>26</sup>. We modelled the  
<sub>126</sub> number of RAC exceedances ( $RQ > 1$ ) at a site as:

$$\begin{aligned} No(RQ > 1)_i &\sim NB(\mu_i, \kappa) \\ \log(\mu_i) &= \beta_0 + f_1(agri_i) + f_2(size_i) + \log(n_i) \end{aligned} \quad (2)$$

<sub>127</sub> where  $No(RQ > 1)_i$  is the observed number of RAC exceedances at site  $i$ . Because of  
<sub>128</sub> overdispersion, we modelled  $No(RQ > 1)_i$  as resulting from a negative binomial distribution

129 (*NB*) with mean  $\mu_i$  and a quadratic mean-variance-relationship ( $Var(No(RQ > 1)_i) =$   
 130  $\mu_i + \frac{\mu_i^2}{\kappa}$ ). The proportion of agricultural land use within the catchment ( $agri_i$ ) and the  
 131 catchment size of the site ( $size_i$ ) were used as predictors of the number of RAC exceedances.  
 132  $\beta_0$  is the intercept and  $f_1$  and  $f_2$  are smoothing functions using penalized cubic regression  
 133 splines<sup>27,28</sup>. The number of measurements per site ( $n_i$ ) was used as an offset to account  
 134 for differences in sampling efforts at a site (in terms of number of samples and analysed  
 135 compounds) and is equivalent to modelling the rate of exceedances. We used point-wise 95%  
 136 Confidence Intervals (CI) of the first derivative of the fitted smooth to identify regions of  
 137 statistically significant changes. All data-processing and analyses were performed using R<sup>29</sup>.  
 138 GAMs were fitted using the mgcv package<sup>28</sup>.

139 To assess the influence of precipitation and seasonality, we modelled the RQ of individual  
 140 compounds as the response variable. RQ and concentrations show a skewed distribution  
 141 with an excess of zeros (no pesticides detected and quantified). Therefore, we modelled  
 142 these as two processes (one generating values below the limit of quantification (LOQ) and  
 143 one generating values above LOQ) using a Zero-Adjusted Gamma (ZAGA) distribution<sup>30,31</sup>  
 144 (Equation 3). These two processes can be interpreted as changes in the mean value of RQ  
 145 (change in  $\mu$ ) and changes in the probability of exceeding LOQ and showing any risk (change  
 146 in  $\nu$ ).

$$RQ_i \sim ZAGA(\mu_i, \sigma, \nu_i) = \begin{cases} (1 - \nu_i) & \text{if } y < LOQ \\ \nu_i \times f_{Gamma}(\mu_i, \sigma) & \text{if } y \geq LOQ \end{cases} \quad (3)$$

147  $\nu_i$  denotes the probability of a measurement i being above LOQ and  $f_{Gamma}$  denotes the  
 148 gamma function and is used for values equal to or greater LOQ, with  $\mu$  being the mean  
 149 and  $\sigma$  the standard deviation of RQ. We used the  $\log(x + 0.05)$  transformed precipitation  
 150 at sampling date ( $\log prec_0$ ) and the day before ( $\log prec_{-1}$ ), as well as quarters of the year  
 151 (*Q1: Jan-Mar, Q2: Apr-Jun, Q3: Jul-Sep, Q4: Oct-Dec*) as linear predictors for  $\mu$  and  $\nu$ .

152 We used appropriate link functions for  $\mu$  and  $\nu$  and assumed  $\sigma$  to be constant. Equation 4  
153 summarises the deterministic part of the model for a measurement  $i$ .

$$\begin{aligned}\log(\mu_i) &= \log(prec_{0i}) + \log(prec_{-1i}) + Q1_i + Q2_i + Q3_i + Q4_i \\ logit(\nu_i) &= \log(prec_{0i}) + \log(prec_{-1i}) + Q1_i + Q2_i + Q3_i + Q4_i\end{aligned}\tag{4}$$

154 To account for differences between federal states we used *site* nested within *state* as  
155 random intercepts. We implemented this model using the *gamlss* package.<sup>32</sup>

156 We fitted this model separately to each compound with a RAC, measured in at least 1000  
157 samples and with more than 5% of values above LOQ ( $n = 22$  compounds, see Supporting  
158 Information Table S3 for a list of compounds). To summarise the coefficients across the  
159 22 modelled compounds we used a random effect meta-analysis for each model coefficient  
160 separately<sup>33</sup>, resulting in an averaged effect of the 22 compounds. The results of individual  
161 compounds are provided in the Supporting Information Table S4 and Figures S6 and S7.  
162 The meta-analysis was performed using the *metafor* package<sup>34</sup>.

## 163 Results

### 164 Overview of the compiled data

165 The compiled dataset used for analysis comprised 1,766,104 pesticide measurements in 24,743  
166 samples from 2,301 sampling sites in small streams. These samples were all taken via grab  
167 sampling. We found large differences between federal states in the number of sampling sites  
168 and their spatial distribution (Figure 1 and Supporting Information Table S1). The number  
169 of small stream sampling sites per state ranged from 1 (Lower Saxonia, NI) to 1139 (North  
170 Rhine-Westphalia, NW). No data were available from Brandenburg.

171 In total 478 different compounds used as pesticides and their metabolites were measured  
172 at least once (Supporting Information Table S2). Most of the compounds were herbicides

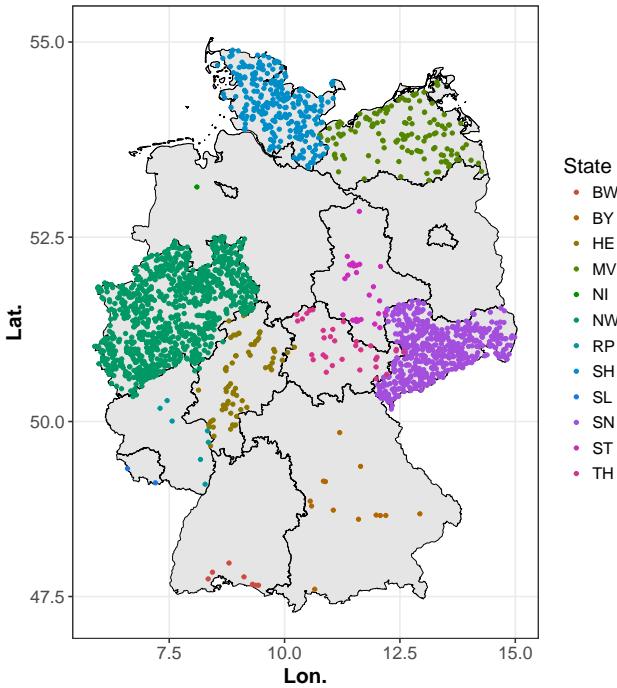


Figure 1: Spatial distribution of the 2,301 small stream sampling sites. Colour codes different federal states (see Supporting Information Table S1 for abbreviations).

173 (179), followed by insecticides (117) and fungicides (109). Most samples were taken in  
 174 the months April till October, while fewer samples were taken during winter (see Supporting  
 175 Information Figure S2). We found substantial differences in the spectra of analysed pesticides  
 176 between federal states (Figure 2). The number of analysed pesticides per state ranged from  
 177 57 (SL) to 236 (RP) (Supporting Information Table S1). 4% (=71,113) of all measurements  
 178 were concentrations above LOQ.

179 The distribution of sampling sites across catchment sizes indicated a disproportionately low  
 180 number of sites with catchments below  $10 \text{ km}^2$ , with most sampling sites having catchment  
 181 sizes between 10 and  $25 \text{ km}^2$  (Figure 3).

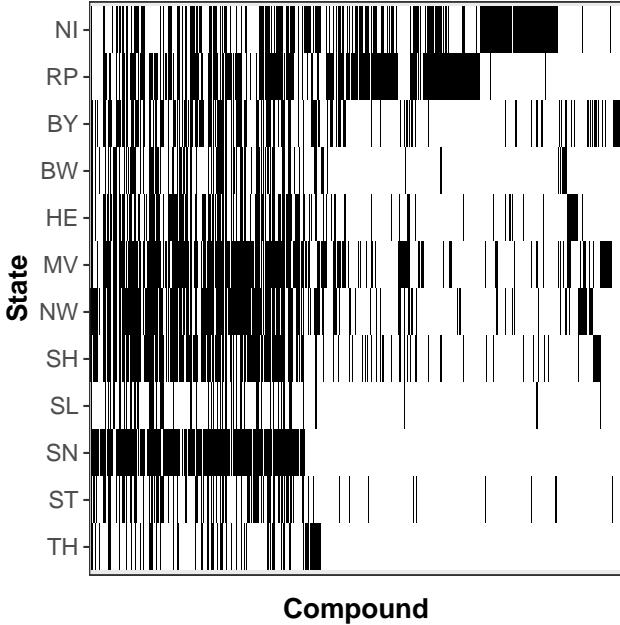


Figure 2: Barcode plot of compound spectra of the federal states. Each vertical line is an analysed compound.

## **182 Influence of agricultural land use and catchment size**

**183** We found a positive relationship between agricultural land use and the number of RAC  
**184** exceedances. The non-linear model revealed, that below 28% agriculture the mean number of  
**185** RAC exceedances show a 3.7-fold decrease from 0.39 (28% agriculture within the catchment)  
**186** to 0.10 (no agriculture) (Figure 4, left). Catchment size had no statistically significant effect  
**187** on the number of RAC exceedances (Figure 4, right). We also could not detect a statistically  
**188** significant interaction between catchment size and agriculture.

## **189 Effect of precipitation on pesticide risk**

**190**  $prec_0$  and  $prec_{-1}$  increased the probability of exceeding LOQ and RQ. In Q2 an increase  
**191** from 0.1 mm to 15 mm of precipitation before sampling ( $prec_{-1}$ ) lead on average to a 43%  
**192** higher mean RQ of 0.05 (Supporting Information Figure S7). The probability to exceed  
**193** LOQ increases in Q2 1.6-fold from 8.7% to 13.5% (Figure 5). Precipitation before sampling  
**194** ( $prec_{-1}$ ) had a stronger effect than precipitation during sampling ( $prec_0$ ) on the probability

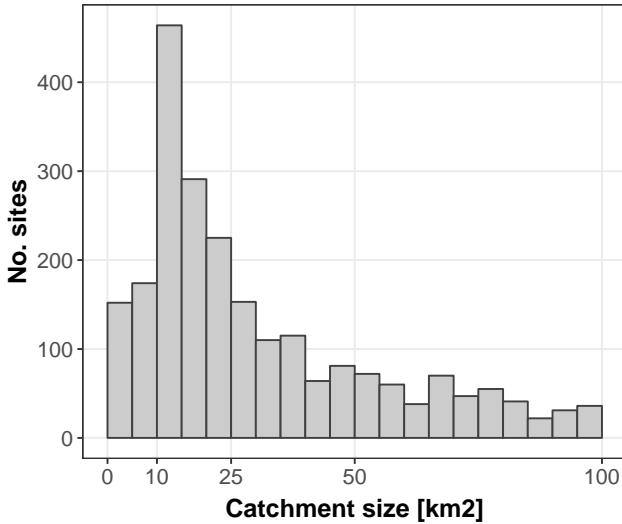


Figure 3: Distribution of catchment sizes across the sampling sites.

of exceeding LOQ. This difference was less pronounced for the mean value of RQ (Supporting Information Figure S7, top left). Moreover, effects differed between individual compounds (see Supporting Information Table S4).

The first quarter showed the lowest RQ and probability of exceeding LOQ. Both increased in Q2 and decreased towards the end of the year. There was a 2.5-fold higher probability of exceeding LOQ in Q2 (10.6%) than in Q1 (4.6%) (Figure 5). The differences were less pronounced for the mean value of RQ and with less precision (see Supporting Information Figure S7, left). Individual compounds showed different temporal patterns (see Supporting Information Table S4).

## Current pesticide risks in small streams

We found RAC exceedances in 25.5% of sampling sites and RQ > 0.1 in 54% of sites. In 23% of sites none of the chemicals for which RACs were available were detected (see also Supporting Information Figure S8). Neonicotinoid insecticides and Chlorpyrifos showed the highest RQ (Figure 6). For Thiacloprid and Chlorpyrifos the RAC was equal or less than LOQ, therefore, all detections have a  $RQ \geq 1$ . The herbicides Nicosulfuron and Diflufenican, as well as the fungicide Dimoxystrobin also showed high exceedances of RACs (26.7, 14.1

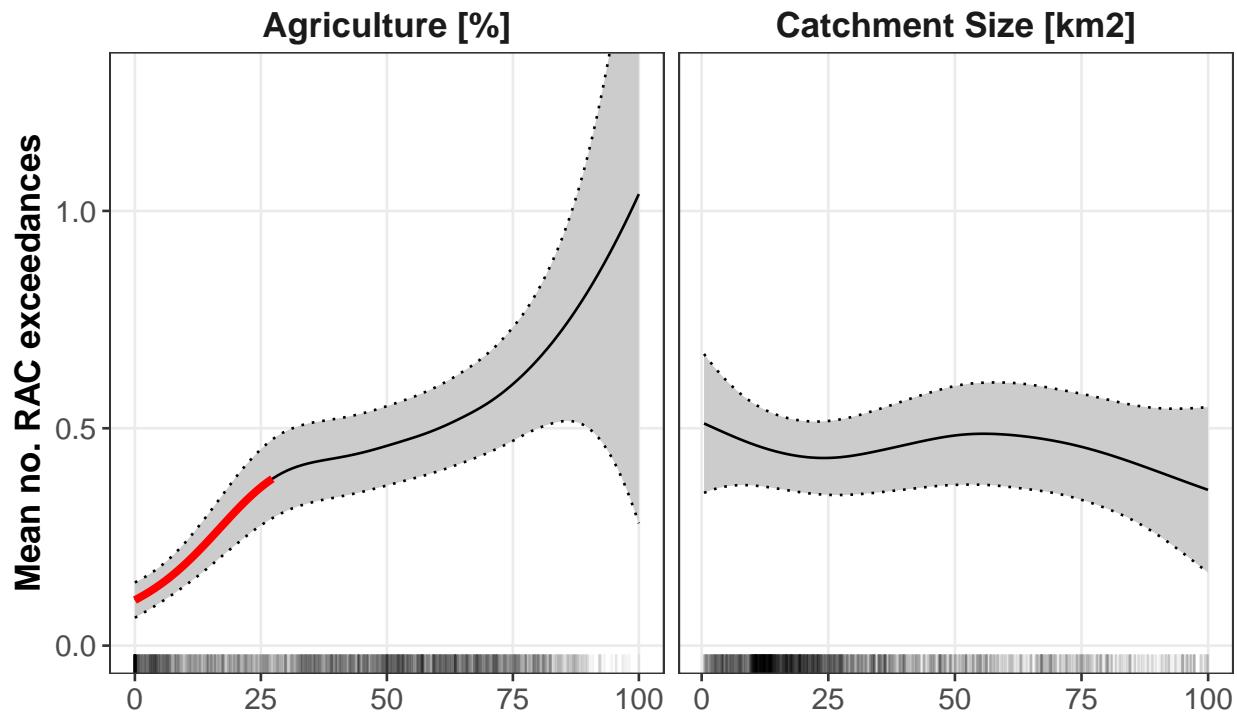


Figure 4: Effect of percent agriculture within the catchment (left) and of catchment size (right) on the mean number of RAC exceedances per site. Red line marks statistically significant changes. Dashed lines denote 95% point-wise Confidence Intervals.

and 21.1 % of measurements > LOQ), see also Supporting Information Table S5). RAC exceedances were found in 14% of samples with concentrations >LOQ (and 7.3% of all samples).

The highest RQs were observed for Chlorpyrifos ( $\text{max}(\text{RQ}) = 220$ ), Clothianidin ( $\text{max}(\text{RQ}) = 157$ ), Dimoxystrobin( $\text{max}(\text{RQ}) = 117$ ) and Isoproturon ( $\text{max}(\text{RQ}) = 80$ ). Where analysed, metabolites exhibited the highest detection rates (for example, Metazachlor sulfonic acid was detected in 84% of all samples where it was analysed ( $n = 3038$ , see also Supporting Information Figure S9)). Glyphosate was the compound with the highest detection rates (41%,  $n = 3557$  samples), followed by Boscalid (23%,  $n = 9886$ ) and Isoproturon (22%,  $n = 19112$ ). However, only the latter showed RAC exceedances (Figure 6). In 45.9% of samples more than one compound was quantified, with a maximum of 54 different compounds in one sample (Supporting Information Figure S10).

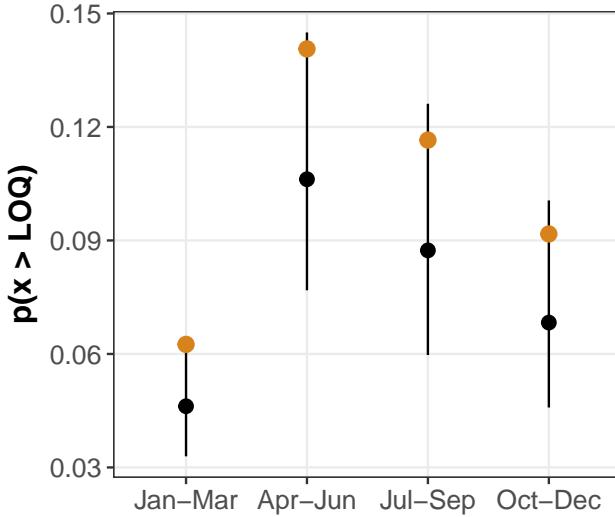


Figure 5: Summarised model predictions for the probability to exceed LOQ throughout the year. Black dots indicate the probabilities at 0.1 mm precipitation (and their 95% CI). Orange dots indicate the probabilities at 15 mm precipitation. Probabilities were summarised from a meta-analysis of the 22 modelled compounds. Single compound coefficients are provided in Supporting Information Table S4 and Figure S7.

## 223 Discussion

### 224 Overview of the compiled dataset

225 The compiled dataset of governmental monitoring data, with a particular focus on small  
 226 streams, represents currently the most comprehensive for Germany. Similar nationwide  
 227 datasets have been compiled for the Netherlands<sup>35</sup>, Switzerland<sup>36</sup> and the United States<sup>37</sup>.  
 228 While the compilations from Europe are of similar quantity and quality to the data compiled  
 229 and analysed here, the compilation used in Stone et al.<sup>37</sup> is much smaller, though these data  
 230 may be complemented by more data in future analyses.

231 A nationwide assessment of pesticide pollution is hampered by inhomogeneous data across  
 232 federal states: Beside large differences in the spatial distribution and quantity of sampling  
 233 sites (Figure 1), the spectrum of analysed compounds (Figure 2) and the quality of chem-  
 234 ical analyses differed between states. Despite the outlined differences between states, all  
 235 ecoregions occurring in Germany<sup>38,39</sup> and all major stream types were covered by the data

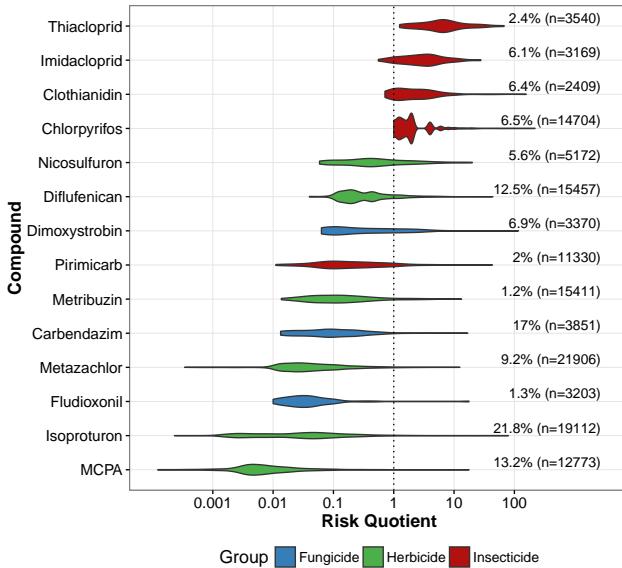


Figure 6: 15 compounds with the highest observed risk quotients in small streams. Non-detects are not shown due to the logarithmic axis. Numbers on the right give the percentage of values >LOQ and the total number of samples where the compound was analysed.

236 set. The unequal distribution of sampling sites and the different sampling strategies hamper  
 237 inference on the total population of small streams in Germany. We accounted for differences  
 238 in sampling efforts per site by including the total number of measurement into the statistical  
 239 models. However, we acknowledge that additional differences such as sampling frequency and  
 240 temporal distribution of the samplings might incur bias between states<sup>3,18</sup>. Consequently,  
 241 we did not compare the results between states. Moreover, it is known that differences in  
 242 analytical quality can influence estimated effects<sup>40,41</sup>. However, the used model (Equation 3)  
 243 explicitly accounted for LOQs and differences therein.

244 For Thiacloprid and Chlorpyrifos the LOQs were above the RAC, which means that  
 245 exceedances are likely underestimated. For compounds with low RACs a lowering of LOQ  
 246 through an improvement of chemical analysis is essential for reliable assessment. Moreover,  
 247 a nationwide assessment would benefit from a harmonised spectrum of analysed compounds  
 248 between federal states.

249 Given their high abundance in the landscape<sup>11</sup> small streams below 10 km<sup>2</sup> are dispro-  
 250portionally less sampled in current monitoring (Figure 3). This may be attributed to the

<sup>251</sup> missing categorisation in the WFD, but also to technical (e.g. sampling at low flow) and  
<sup>252</sup> natural reasons (e.g. ephemeral streams).

<sup>253</sup> Clearly, there is currently a lack of knowledge on stressor effects on small streams. We  
<sup>254</sup> analysed only data from small streams, however, for lentic small water bodies this lack might  
<sup>255</sup> be even greater<sup>16</sup>.

## <sup>256</sup> **Influence of agricultural land use and catchment size**

<sup>257</sup> As hypothesised, we found a positive relationship between agricultural land use and the  
<sup>258</sup> number of RAC exceedances. Especially, we found a statistically significant drop below 28%  
<sup>259</sup> of agricultural land use (Figure 4, left). This drop indicates that agricultural land use is a  
<sup>260</sup> major contributor to the observed RAC exceedances. We note that this drop would have  
<sup>261</sup> been missed by linear modeling (Supporting Information Figure S5). The absence of such  
<sup>262</sup> a drop would have suggested that other inputs such as from urban gardening are relevant  
<sup>263</sup> contributors to RAC exceedances.

<sup>264</sup> We did not find a statistically significant relationship between pesticide pollution and  
<sup>265</sup> catchment size (Hypothesis 2). However, previous studies showed that small streams are  
<sup>266</sup> more polluted than bigger streams<sup>7,9,42</sup>. This can be explained by the relatively short gra-  
<sup>267</sup> dient of catchment sizes in our dataset, with most of the streams with catchments above  
<sup>268</sup>  $10 \text{ km}^2$  and below  $100 \text{ km}^2$  (Figure 3, top). For example, the gradient of Schulz<sup>7</sup> covered  
<sup>269</sup> 6 orders of magnitude. Although catchment size and stream size were strongly correlated  
<sup>270</sup> (Supporting Information Figure S11), other factors such as geology, hydrology and precip-  
<sup>271</sup> itation regime also determine the stream size. Therefore, translating our results to stream  
<sup>272</sup> size bears uncertainties.

## <sup>273</sup> **Effect of precipitation on pesticide risk**

<sup>274</sup> We found a 43% higher mean RQ if samples were taken after rainfall events, which conforms  
<sup>275</sup> to the hypothesis that run-off is a major entry path way for pesticides into streams on the

276 large scale. However, samples taken on the day of a rainfall event showed less risk than  
277 samples taken one day after a rainfall event. This discrepancy could be explained by a  
278 sampling preceding the rainfall event because the temporal resolution of our dataset was 1  
279 day. Additionally, this might be explained by a delay between the start of a rain event and  
280 the peak in discharge or runoff.

281 The effects of precipitation were more pronounced for the probability to exceed LOQ,  
282 with smaller effect sizes for the absolute value of RQ. This may be explained by a higher  
283 variability of absolute concentrations. Overall, our results indicate that current pesticide  
284 monitoring relying on grab sampling, largely disconnected from precipitation events, under-  
285 estimates pesticide risks. Automatic event-driven samplers<sup>3</sup> and passive samplers<sup>43,44</sup> may  
286 help overcome these shortcomings and provide a better representation of risks. Our results  
287 demonstrate that future monitoring of small water bodies should also capture precipitation  
288 events, which is in agreement with other studies, such as Lorenz et al.<sup>16</sup>.

289 We found the highest the probability of exceeding LOQ from April to June (10% for  
290 Q2) and lowest in the first quarter of the year (4%, Figure 5, bottom right). This annual  
291 pattern coincides, as hypothesised, with the main application season for pesticides in Central  
292 Europe. Nevertheless, there are compound-specific differences in the annual pattern, which  
293 explains the wide CI for the absolute RQ (Figure 5, bottom left). For example, the herbicide  
294 Diflufenican showed the highest RQ and the highest probability of exceeding LOQ during  
295 the winter quarters Q1 and Q4 (Supporting Information Table S4), which coincides with  
296 the application period it is registered for in Germany<sup>45</sup>. Moreover, compound properties,  
297 like half-life or water solubility, might influence compound dynamics. Our study suggests  
298 that pesticide risks display compound specific spatio-temporal dynamics. Currently, little is  
299 known about these and further research on those might provide useful information for future  
300 environmental risk assessment. For example, the sensitivity of organisms is often life stage  
301 dependent<sup>46</sup> and knowledge on temporal dynamics could inform on concurrent exposure to  
302 multiple pesticides, as well as assist to parameterise toxicokinetic and toxicodynamic mod-

<sup>303</sup> els<sup>47</sup>. Moreover, our results show that analysing absolute concentrations and probabilities  
<sup>304</sup> of LOQ together might deliver valuable insights into risk dynamics. The influence of agri-  
<sup>305</sup> cultural land use within the catchment area and the coincidence with the growing season  
<sup>306</sup> indicates that agricultural land use a major contributor of pesticides in streams.

### <sup>307</sup> Pesticides in small streams

<sup>308</sup> Our results suggest that small streams are frequently exposed to ecologically relevant pes-  
<sup>309</sup> ticide concentrations. In one-quarter of small streams RACs were exceeded at least once.  
<sup>310</sup> Stehle and Schulz<sup>9</sup> found the highest percentage of RAC exceedances for organophosphate  
<sup>311</sup> insecticides. By contrast, we found that neonicotinoid insecticides have highest exceedances  
<sup>312</sup> of RACs, followed by the organophosphate chlorpyrifos. This difference can be attributed to  
<sup>313</sup> the low sample size for neonicotinoid insecticides in their study ( $n = 33$ ) compared to the  
<sup>314</sup> dataset presented here (for example 3,540 samples of Thiacloprid, Figure 6). Overall, our  
<sup>315</sup> results suggest that neonicotinoids may currently pose a high risk to freshwater ecosystems.  
<sup>316</sup> Moreover, our results add further evidence to the growing literature on the risks arising from  
<sup>317</sup> neonicotinoids for aquatic<sup>48</sup> and terrestrial<sup>49</sup> ecosystems.

<sup>318</sup> Compared to Stehle and Schulz<sup>9</sup> we found higher rates of RAC exceedances for insec-  
<sup>319</sup> ticides. They found exceedances in 37.1% of insecticide measurements >LOQ ( $n = 1352$ ,  
<sup>320</sup> 23 insecticides), whereas, we found exceedances in 67% of insecticide measurements with  
<sup>321</sup> RACs >LOQ ( $n = 1855$ , 22 insecticides). This could be attributed to different insecticides  
<sup>322</sup> considered and different underlying RACs. Our study has only 7 insecticides with RACs in  
<sup>323</sup> common with the insecticides investigated by Stehle and Schulz<sup>9</sup>. Moreover, all RACs were  
<sup>324</sup> lower in our study (average difference =  $-0.71 \mu\text{g/L}$ , range = [-2.757; -0.005]). Nevertheless,  
<sup>325</sup> it must be noted that the dataset compiled here comprised only samples from grab sampling,  
<sup>326</sup> which may considerably underestimate pesticide exposure<sup>3,18</sup>.

<sup>327</sup> By contrast, Knauer<sup>42</sup> found exceedances from monitoring data mainly for herbicides  
<sup>328</sup> and fungicides and only one insecticide Chlorpyrifos-methyl. Moreover, RAC exceedances in

329 Switzerland were generally lower and less abundant (for example 6 exceedances (=0.2%) for  
330 Isoproturon with a maximum RQ of 2) compared to our results for Germany. This might  
331 reflect differences in pesticide use between countries, ecoregions and RACs used. From  
332 the definition of RAC it follows that if the concentration of a compound exceeds its RAC  
333 ecological effects are expected. Indeed, Stehle and Schulz<sup>50</sup> found that the biological diversity  
334 of stream invertebrates was significantly reduced by 30% at RQ = 1.12 and by 10% at 1/10  
335 of RAC. We found RQ values greater than 1.12 in 25% of small streams and RQ at 1/10 of  
336 RAC in 54% of small streams. Consequently, we conclude that agricultural pesticides are  
337 on a large scale a major threat to small streams, the biodiversity they host and the services  
338 they provide. This threat may exacerbate because pesticides often occur in mixtures<sup>51</sup> and  
339 may co-occur with other stressors<sup>52</sup>.

340 Monitoring data, despite the outlined limitations, provide an opportunity to study large-  
341 scale environmental occurrence patterns of pesticides. Furthermore, such nationwide com-  
342 pilations, may not only be used for governmental surveillance, but also to answer other  
343 questions, like validation of exposure modelling,<sup>53</sup> retrospective evaluation of regulatory risk  
344 assessment<sup>9,42</sup> or occurrences of pesticide mixtures.<sup>51</sup> However, the sampling design needs  
345 to account for precipitation events to provide robust data. Therefore, non-linear modeling  
346 can provide additional insights to risk assessment compared to linear modeling<sup>26</sup>. Our re-  
347 sults suggest that exceedances of RACs are landscape dependent and therefore, pesticide  
348 regulation should account for landscape features. Moreover, the high exceedances of RACs  
349 indicate that greater efforts are needed to describe causal links, which may lead to further  
350 developments of the current authorisation procedure.

## 351 Acknowledgement

352 The authors thank the federal state authorities and the German Working Group on water  
353 issues of the Federal States (LAWA) for providing chemical monitoring data and the Ger-

<sup>354</sup> man Environment Agency (UBA) for funding a related project (FKZ 3714 67 4040 / 1).  
<sup>355</sup> We thank Alexandra Müller, Wolfram König and Volker Mohaupt (German Environment  
<sup>356</sup> Agency (UBA)), Martin Keller and Beate Bänsch-Baltruschat (German Federal Institute  
<sup>357</sup> of Hydrology (BfG)), Matthias Liess and Kaarina Foit (Center for Environmental Research  
<sup>358</sup> (UFZ)) for their contributions to this project.

## <sup>359</sup> Supporting Information Available

<sup>360</sup> The following files are available free of charge.

- <sup>361</sup> • Supplemental \_ Materials.pdf : Supporting Information (Figures, Tables, Models).

<sup>362</sup> This material is available free of charge via the Internet at <http://pubs.acs.org/>.

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