

# Large scale risks from pesticides in small streams

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## Abstract

Small streams are important refugia for biodiversity. In agricultural areas they may be at high risk from pesticide pollution. However, most related studies have been limited to a few streams on the regional level, hampering extrapolation to larger scales. We used data from German governmental water quality monitoring to quantify the drivers of pesticide risk and to assess pesticide risk in small streams on a large scale. The data set comprised of 1,766,104 measurements of 478 pesticides (including metabolites) related to 24,743 samples from 2,301 sampling sites. We investigated the influence of agricultural land use, catchment size, as well as precipitation and seasonal dynamics on pesticide risk using new statistical modelling techniques that explicitly consider the limit of quantification. Agricultural land use lead to a 3.7-fold increase in exceedance of risk thresholds when the proportion of agriculture in a catchment exceeded 28 percent. Precipitation increased pesticide risk by 36% and risk was the highest during summer months. Risk thresholds were exceeded in 26% of streams, with the highest risk related to neonicotinoid insecticides. We conclude that pesticides from agricultural land use

are a major threat to small streams and their biodiversity and that a realistic pesticide sampling would be driven by precipitation events.

## Introduction

More than 50% of the total land area in Germany is used by agriculture<sup>1</sup>. In the year 2014 more than 45,000 tonnes of 776 authorised plant protection products were sold for application on this area<sup>2</sup>. The applied pesticides may enter surface waters via spray-drift, edge-of-field run-off or drainage<sup>3–5</sup>. Once entered the surface waters they may have adverse effects on biota and ecosystem functioning<sup>6</sup>. Although it is known that pesticide pollution and its ecological effects increase with the fraction of agricultural land use in the catchment<sup>7</sup>, the shape of the relationship is unknown and studies on potential thresholds are lacking.

Two recent studies indicate that pesticides might threaten freshwater biodiversity in the European union. Malaj et al.<sup>8</sup> analysed data supplied to the European Union (EU) in the context of the Water Framework Directive (WFD) and showed that almost half of European water bodies are at risk from pesticides. Stehle and Schulz<sup>9</sup> compiled 1,566 measured concentrations of 23 insecticides in the EU from scientific publications. They found that many of these measurements exceed regulatory acceptable concentrations (RAC). However, these studies reflect only a small amount of potentially available data (173 sites in predominantly mid-sized and large rivers in Malaj et al.<sup>8</sup> and 138 measurements in Stehle and Schulz<sup>9</sup>), and it is unclear how representative they are for Germany. Much more comprehensive data on thousands of sites are available from national monitoring programs that are setup for the surveillance of water quality, which is done independently by the federal states in Germany in compliance with the WFD<sup>10</sup> and additional state-specific needs. Despite that these data are providing the opportunity to study pesticide risks and other research questions on a large scale with high spatial density, to date these data have not been compiled and related analyses are lacking.

Small streams comprise a major fraction of streams<sup>11</sup>, accommodate a higher proportion of biodiversity compared to larger freshwater systems<sup>12,13</sup> and play an important role in the recolonization of disturbed downstream reaches<sup>14,15</sup>. Nevertheless, a clear definition of small streams in terms of catchment or stream size is currently lacking<sup>16</sup>. For example, the WFD defines small streams with a catchment size between 10 and 100 km<sup>2</sup>, without further categorisation of streams <10km<sup>2</sup> and Lorenz et al.<sup>16</sup> defines small streams with catchment size <10km<sup>2</sup>. Moreover, small streams might particularly be at high risk of pesticide contamination in case of adjacent agricultural areas given their low dilution potential<sup>5,7</sup>. Indeed, meta-analyses using data from studies with a few sites reported higher pesticide pollution in smaller streams compared to bigger streams<sup>7,9</sup>. Despite their ecological relevance and potentially higher pesticide exposure, a recent analysis of pesticide studies showed that a disproportionately small fraction of studies was conducted in small water bodies, and these were largely limited to a few sites<sup>16</sup>. Consequently, knowledge on the pesticide pollution of small streams on larger scales is scant. In European law, the Directive 2009/128/EC<sup>17</sup> places an obligation on the EU Member States to adopt National Action Plans (NAP) for the Sustainable Use of Plant Protection Products and the German NAP also addresses the knowledge gap concerning pesticide impact on small streams, specifically including those with catchment size <10km<sup>2</sup>.

In this study, we compiled and analysed large-scale chemical monitoring data from small streams in Germany. First, we analysed the shape of the relationship between pesticide risk, agricultural land use, and catchment size and examined whether related thresholds for pesticide risks can be derived. Second, we investigated the influence of precipitation and seasonal dynamics on pesticide detections, given that precipitation proved an important driver of pesticide exposure in several small-scale studies<sup>7,18</sup>, but it is unknown whether a precipitation signal prevails on large scales. Finally, we quantified the current risks from pesticides in small streams in Germany and the compounds accountable for the risk.

# Methods

## Data compilation

We queried pesticide monitoring data from sampling sites that can be classified as small streams (catchment sizes < 100 km<sup>2</sup> according to the WFD) from all 13 non-city federal states of Germany (see Supplemental Table S1 for the abbreviations of federal state names) for 2005 to 2015. We homogenised and unified all data provided by the federal states into a database and implemented a robust data-cleaning workflow (see Supplemental Figure S1 for details)<sup>19</sup>.

We identified precipitation at sampling sites by a spatio-temporal intersection of sampling events with gridded daily precipitation data (60×30 arcsec resolution) available from the German Meteorological Service (DWD). This data spatially interpolates daily precipitation values from local weather stations<sup>20</sup>. We performed the intersection for the actual sampling date and the day before and extracted precipitation during and up to 48 hours before sampling.

## Characterization of catchments

We compiled a total of 2,369 sampling sites in small streams with pesticide measurements. Alongside, we also queried catchment sizes and agricultural land use within the catchment for the sampling sites from the federal states. Catchment size was provided for 59% of sites. Additionally, we delineated upstream catchments for each of the sampling sites using (i) a digital elevation model (DEM)<sup>21</sup> and the multiple flow direction algorithm<sup>22</sup> as implemented in GRASS GIS 7<sup>23</sup> and (ii) from drainage basins provided by the Federal Institute of Hydrology (BfG). Delineated catchments were visually checked for accuracy by comparison with state stream networks and derived information amalgamated with existing data. Thus, catchment size information was available for 99% of all sites (59% from authorities, 24% from DEM and 16% from drainage basins).

For each derived catchment (either from DEM or drainage basins) we calculated the % agricultural land-use within the catchment based on the Authoritative Topographic-Cartographic Information System (ATKIS) of the land survey authorities<sup>24</sup>. Thus, agricultural land use information was available for 98% of all sites (24% from authorities, 52% from DEM and 22% from drainage basins). 68 sites (3%) that lacked catchment size or land use information were omitted from the analysis, resulting in 2301 sites used in the analyses outlined below.

## Characterization of pesticide pollution

We characterised pesticide pollution using regulatory acceptable concentrations (RAC)<sup>25</sup>. RACs are derived during pesticide authorisation as part of the ecological risk assessment. No unacceptable ecological effects are expected if the environmental concentration remains below this concentration. Stehle and Schulz<sup>9</sup> showed that RAC exceedances reflect a decrease in biodiversity and from this perspective are ecologically relevant indicators. The German Environment Agency (UBA) provided RACs for 107 compounds, including those with the highest detection rates (Supplemental Table S2). Based on these RACs, we calculated Risk Quotients (RQ):

$$RQ_i = \frac{C_i}{RAC_i} \quad (1)$$

where  $C_i$  is the concentration of a compound  $i$  in a sample and  $RAC_i$  the respective RAC.

## Statistical analyses

All data-processing and analyses were performed using R<sup>26</sup>. To display differences in the spectra of analysed compounds between federal states we used Multidimensional Scaling (MDS) based on Jaccard dissimilarity in conjunction with complete linkage hierarchical

clustering using the vegan package<sup>27</sup>. We determined the optimum number of clusters using the average silhouette width<sup>28</sup>.

We expected non-linear responses to agriculture and catchment size and therefore, used generalised additive models (GAM) to establish relationships<sup>29</sup>. We modelled the number of RAC exceedances ( $RQ > 1$ ) at a site as:

$$\begin{aligned} No(RQ > 1)_i &\sim NB(\mu_i, \kappa) \\ \log(\mu_i) &= \beta_0 + f_1(agri_i) + f_2(size_i) + \log(n_i) \end{aligned} \quad (2)$$

where  $No(RQ > 1)_i$  is the observed number of RAC exceedances at site  $i$ . We modelled  $No(RQ > 1)_i$  as resulting from a negative binomial distribution ( $NB$ ) with mean  $\mu_i$  and a quadratic mean-variance-relationship ( $Var(No(RQ > 1)_i) = \mu_i + \frac{\mu_i^2}{\kappa}$ ). The proportion of agriculture within the catchment ( $agri_i$ ) and the catchment size of the site ( $size_i$ ) were used as predictors of the number of RAC exceedances.  $\beta_0$  is the intercept and  $f_1$  and  $f_2$  are smoothing functions using penalized cubic regression splines<sup>30</sup>. The degree of smoothness was estimated using restricted maximum likelihood (REML) during the model fitting process<sup>31</sup>. The number of measurements per site ( $n_i$ ) was used as an offset to account for differences in sampling efforts (sampling interval and analysed compound spectrum) at a site and is equivalent to modelling the rate of exceedances. We used point-wise 95% Confidence Intervals (CI) of the first derivative of the fitted smooth to identify regions of statistically significant changes. GAMs were fitted using the mgcv package<sup>31</sup>.

To assess the influence of precipitation and seasonality, we modelled the RQ of individual compounds as the response variable. RQ and concentrations show a skewed distribution with an excess of zeros (no pesticides detected and quantified). Therefore, we modelled these as two processes (one generating values below the limit of quantification (LOQ) and one generating values above LOQ) using a Zero-Adjusted Gamma (ZAGA) distribution<sup>32,33</sup> (Equation 3). These two processes can be interpreted as changes in the mean value of RQ

(change in  $\mu$ ) and changes in the probability of exceeding LOQ and showing any risk (change in  $\nu$ ).

$$RQ_i \sim ZAGA(\mu_i, \sigma, \nu_i) = \begin{cases} (1 - \nu_i) & \text{if } y < LOQ \\ \nu_i \times f_{Gamma}(\mu_i, \sigma) & \text{if } y \geq LOQ \end{cases} \quad (3)$$

$\nu_i$  denotes the probability of a measurement  $i$  being above LOQ and  $f_{Gamma}$  denotes the gamma function and is used for values equal to or greater LOQ, with  $\mu$  being the mean and  $\sigma$  the standard deviation of RQ. We used the  $\log(x + 0.05)$  transformed precipitation at sampling date ( $\log prec_0$ ) and the day before ( $\log prec_{-1}$ ), as well as quarters of the year ( $Q1 - Q4$ ) as linear predictors for  $\mu$  and  $\nu$ . We used appropriate link functions for  $\mu$  and  $\nu$  and assumed  $\sigma$  to be constant. Equation 4 summarises the deterministic part of the model for a measurement  $i$ .

$$\begin{aligned} \log(\mu_i) &= \log prec_{0i} + \log prec_{-1i} + Q1_i + Q2_i + Q3_i + Q4_i \\ logit(\nu_i) &= \log prec_{0i} + \log prec_{-1i} + Q1_i + Q2_i + Q3_i + Q4_i \end{aligned} \quad (4)$$

To account for temporal autocorrelation and differences between federal states we used *site* nested within *state* as random intercepts. We implemented this model using the *gamlss* package.<sup>34</sup>

We fitted this model separately to each compound with a RAC, measured in at least 1000 samples and with more than 5% of values above LOQ ( $n = 22$  compounds, see Supplemental Table S3 for a list of compounds). To summarise the coefficients across the 22 modelled compounds we used a random effect meta-analysis for each model coefficient separately<sup>35</sup>, resulting in an averaged effect of the 22 compounds. The results of individual compounds are provided in the Supplemental Table S4 and Figure S7. The meta-analysis was performed

using the metafor package<sup>36</sup>.

## Results

### Overview of the compiled data

The compiled dataset used for analysis comprised 1,766,104 pesticide measurements in 24,743 samples from 2,301 sampling sites in small streams. These samples were all taken via grab sampling. We found large differences between federal states in the number of sampling sites and their spatial distribution (Figure 1 and Supplemental Table S1). The number of small stream sampling sites per state ranged from 1 (Lower Saxonia, NI) to 1139 (North Rhine-Westphalia, NW). No data were available from Brandenburg.

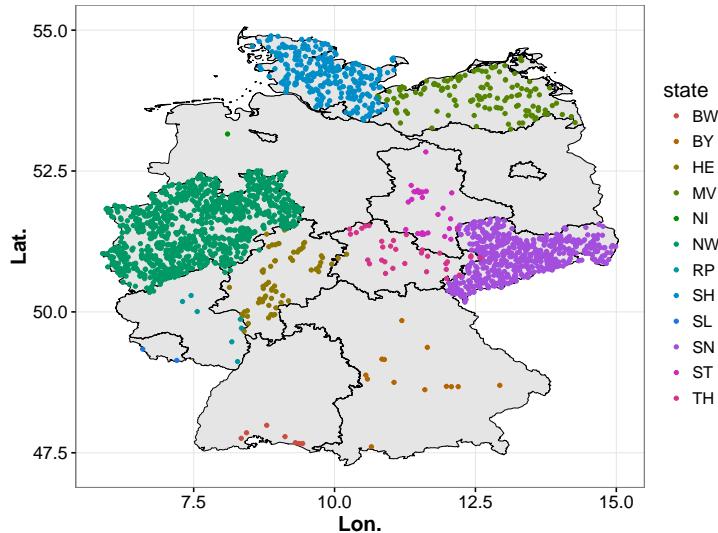


Figure 1: Spatial distribution of the 2,301 small stream sampling sites. Colour codes different federal states (see Supplemental Table S1 for abbreviations).

In total 478 different compounds used as pesticides and their metabolites were measured at least once (Supplemental Table S2). Most of the compounds were herbicides (179), followed by insecticides (117) and fungicides (109). Most samples were taken in the months

April till October, while fewer samples were taken during winter (see Supplemental Figure S2). We found substantial differences in the spectra of analysed pesticides between federal states (Figure 2). The number of different pesticides per state ranged from 57 (SL) to 236 (RP) (Supplemental Table S1). Hierarchical clustering revealed that RP and NI analysed distinct compound spectra compared to the cluster of other states. However, it has to be noted that both states surveyed these distinct spectra in special monitoring programs from only a few sites. Although there was high variability within the remaining cluster, this could not be further split (Figure 2, also Supplemental Figures S3 and S4). 4% (=71,113) of all measurements were concentrations above LOQ.

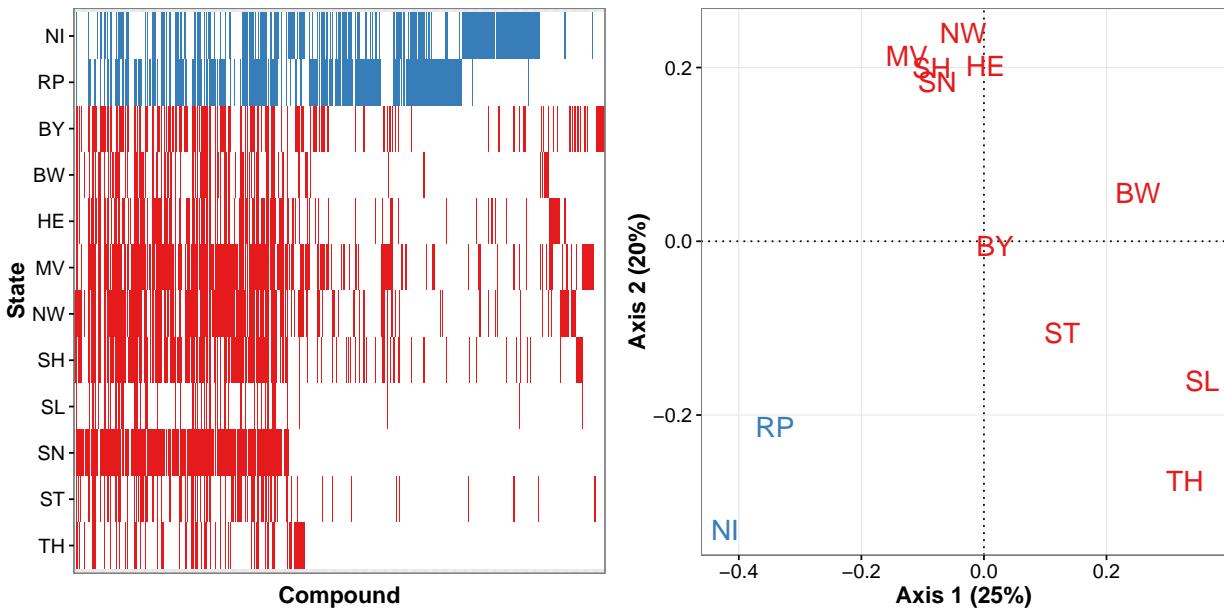


Figure 2: Compound spectra of the different federal states. Left: Barcode plot - each vertical line is an analysed compound. Right: MDS ordination. Colors according to two clusters determined by hierarchical clustering (see Supplemental Figure S3 and S4).

The distribution of sampling sites across catchment sizes indicated a disproportionately low number of sites with catchments below  $10 \text{ km}^2$ , with most sampling sites having catchment sizes between  $10$  and  $25 \text{ km}^2$  (Figure 3).

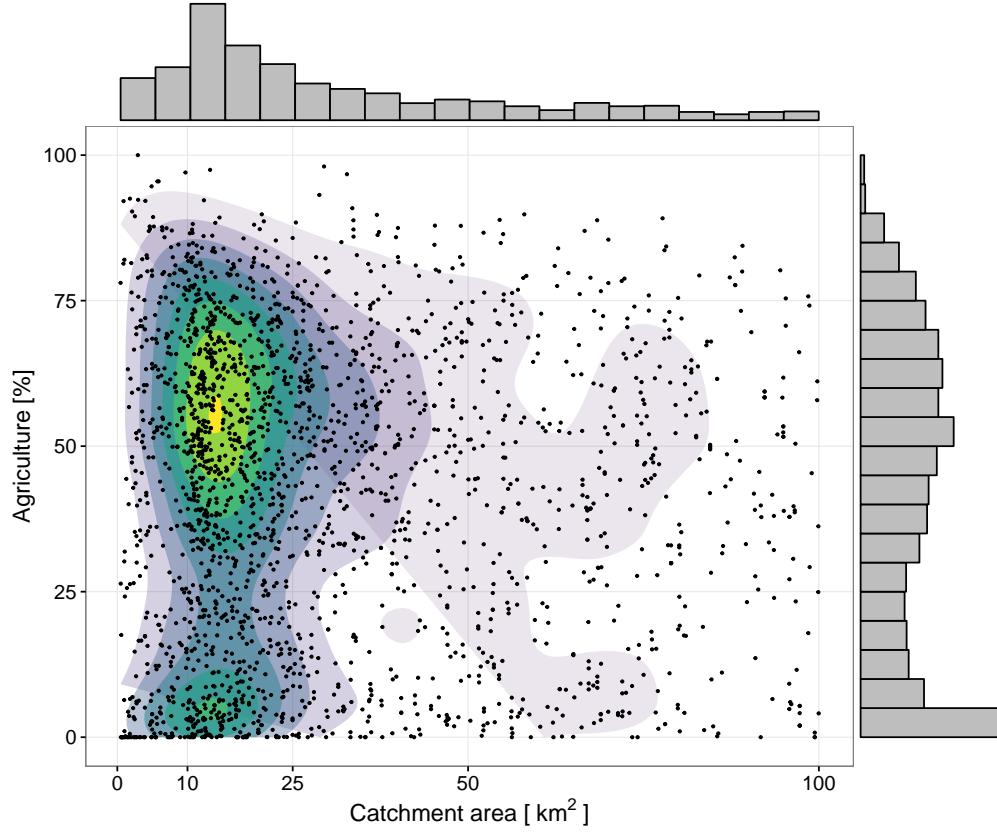


Figure 3: Distribution of catchment area and agriculture within the catchment area across the sampling sites. Colour codes the 2-dimensional density of points.

### Influence of agricultural land use and catchment size

The number of RAC exceedances increased strongly and statistically significant up to 28% agriculture within the catchment. The mean number of RAC exceedances per site increased 3.7-fold from 0.10 (no agriculture) to 0.39 (28% agriculture within the catchment). Above this threshold the exceedances levelled. Above 75% agriculture within the catchment the number of exceedances further increased, but the increase was not statistically significant (Figure 4, left). Catchment size had no statistically significant effect on the number of RAC exceedances (Figure 4, right). We also could not detect a statistically significant interaction between catchment size and agriculture.

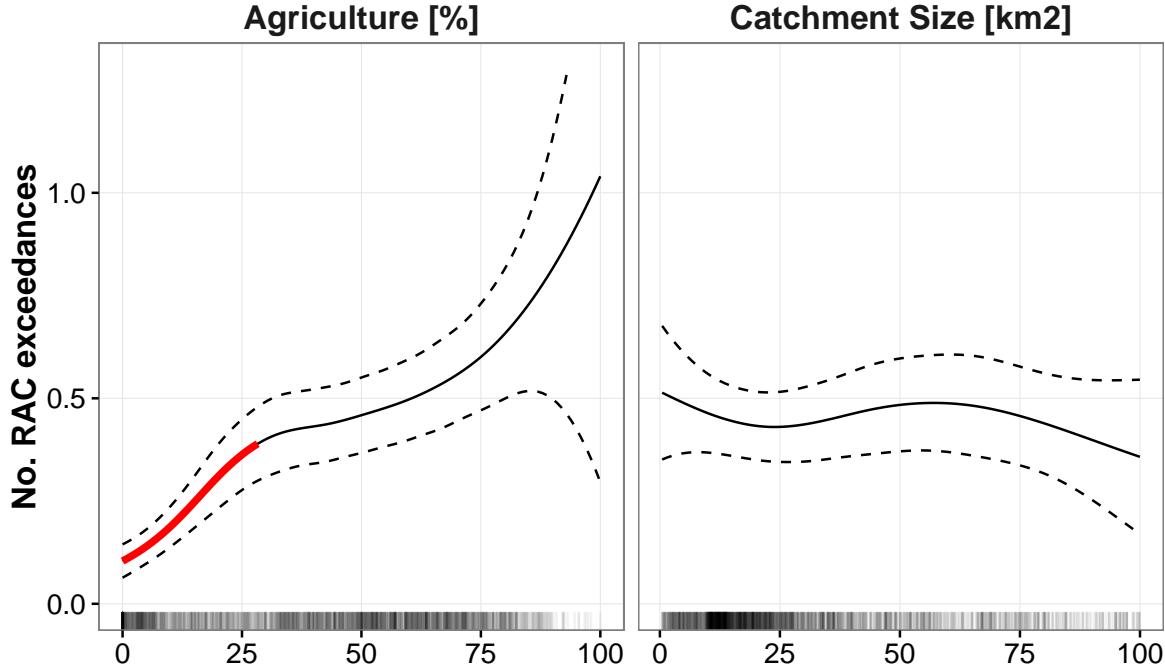


Figure 4: Effect of percent agriculture within the catchment (left) and catchment size (right) on the number of RAC exceedances. Red line marks statistically significant changes. Dashed lines denote 95% point-wise Confidence Intervals.

## Effect of precipitation on pesticide risk

The spatio-temporal intersection revealed that most samples were taken during periods of low precipitation. For example, only 5% of the samples were taken at or after days with rainfall events greater than 10mm / day that may lead to run-off (Supplemental Figure S6).

$prec_0$  and  $prec_{-1}$  increased the probability of exceeding LOQ and RQ. In Q2 an increase from 0.1 mm to 10 mm of precipitation before sampling ( $prec_{-1}$ ) lead on average to a 36% higher mean RQ of 0.05. The probability to exceed LOQ increases 1.6-fold from 8.7% to 13.5% (Figure 5, top). Effects differed between individual compounds and are provided in the Supplemental Table S4. Precipitation before sampling ( $prec_{-1}$ ) had a stronger effect than precipitation during sampling ( $prec_0$ ) on the probability of exceeding LOQ. This difference was less pronounced for the mean value of RQ (Figure 5, top).

The first quarter showed the lowest RQ and probability of exceeding LOQ. Both increased

during summer months and decreased towards the end of the year. There was a 2.5-fold higher probability of exceeding LOQ in  $Q_2$  (10.6%) than in  $Q_1$  (4.6%). The differences were less pronounced for the mean value of RQ and with less precision (Figure 5, bottom). Individual compounds showed different temporal patterns (see Supplemental Table S4).

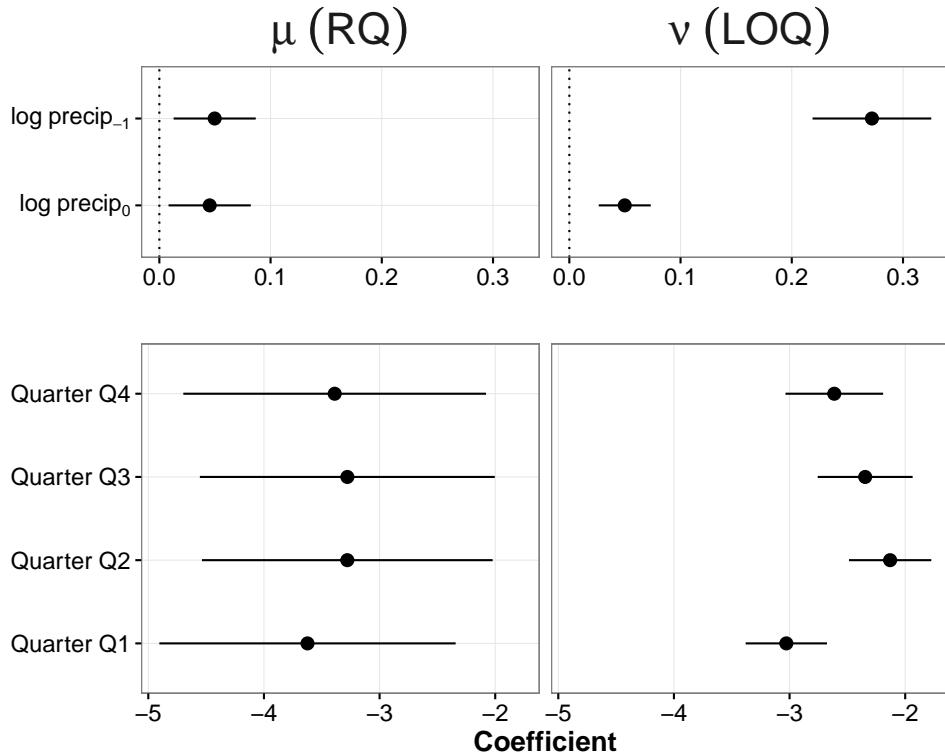


Figure 5: Summarised coefficients (and their 95% CI) for precipitation (top row) and quarter (bottom row) from a meta-analysis of the 22 modelled compounds. Left: coefficients for mean RQ ( $\mu$ ), right: coefficients for probability of exceeding LOQ ( $\nu$ ). Coefficients are shown on the link scale (see Eq. 4). Single compound coefficients are provided in Supplemental Table S4 and Figure S7).

## Pesticide risk in small streams

We found RAC exceedances in 25.5% of sampling sites and  $RQ > 0.1$  in 54% of sites. In 23% of sites none of the chemicals, for which RACs were available, were detected (see also Supplemental Figure S8). Neonicotinoid insecticides and Chlorpyrifos showed the highest RQ (Figure 6). For Thiacloprid and Chlorpyrifos the RAC was equal or less than LOQ,

therefore, all detections have a  $RQ \geq 1$ . The herbicides Nicosulfuron and Diflufenican, as well as the fungicide Dimoxystrobin also showed high exceedances of RQ (26.7, 14.1 and 21.1 % of measurements  $>$  LOQ), see also Supplemental Table S5). RAC exceedances were found in 14% of samples with concentrations  $>$  LOQ (and 7.3% of all samples).

The highest RQs were observed for Chlorpyrifos (max(RQ) = 220), Clothianidin (max(RQ) = 157), Dimoxystrobin(max(RQ) = 117) and Isoproturon (max(RQ) = 80). Where analysed, metabolites exhibited the highest detection rates (for example, Metazachlor sulfonic acid was detected in 84% of all samples where it was analysed ( $n = 3038$ , see also Supplemental Figure S9)). Glyphosate was the compound with the highest detection rates (41%,  $n = 3557$  samples), followed by Boscalid (23%,  $n = 9886$ ) and Isoproturon (22%,  $n = 19112$ ). However, only the latter showed RAC exceedances (Figure 6). In 45.9% of samples more than one compound was quantified, with a maximum of 54 different compounds in one sample (Supplemental Figure S10).

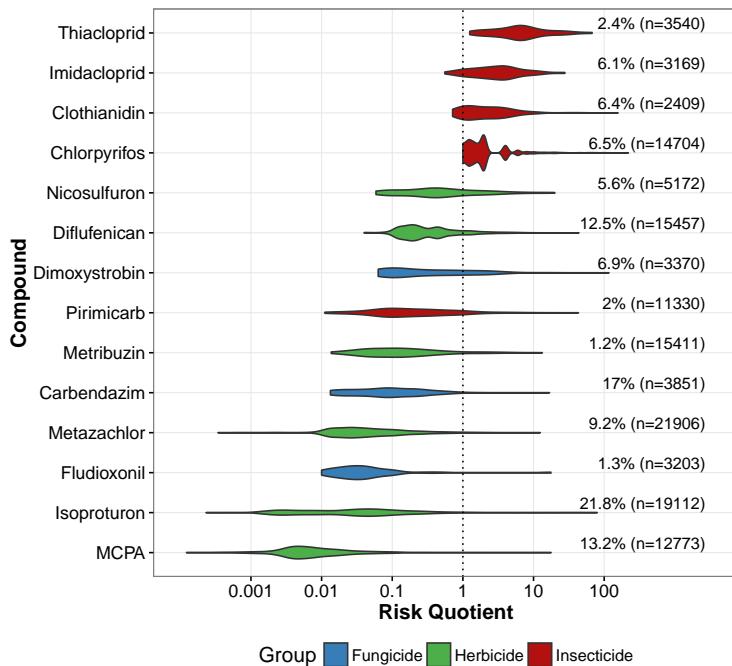


Figure 6: 15 compounds with the highest risk quotients in small streams. Non-detects are not shown due to the logarithmic axis. Numbers on the right give the percentage of values  $>$  LOQ and the total number of samples were the compound was analysed.

# Discussion

## Overview on the compiled dataset

The compiled dataset of governmental monitoring data, with a particular focus on small streams, represents currently the most comprehensive available for Germany. Similar nationwide datasets have been compiled for the Netherlands<sup>37</sup>, Switzerland<sup>38</sup> and the United States<sup>39</sup>. While the compilations from Europe are of similar quantity and quality to the data compiled and analysed here, the compilation used in Stone et al.<sup>39</sup> is much smaller, though these data may be complemented by more data in future analyses.

A nationwide assessment of pesticide pollution is hampered by inhomogeneous data across federal states: Beside large differences in the spatial distribution and quantity of sampling sites (Figure 1), the spectrum of analysed compounds (Figure 2) and the quality of chemical analyses differed between states. Despite the outlined differences between states, all ecoregions occurring in Germany<sup>40,41</sup> were covered by the presented dataset and thus it might nonetheless represent a sample covering all types of small streams in Germany. For Thiaclorpid and Chlorpyrifos the LOQs were above the RAC, which means that exceedances are likely underestimated. For these compounds a lowering of LOQ through an improvement of chemical analysis is essential for reliable assessment. Moreover, a nationwide assessment would benefit from a harmonised spectrum of analysed compounds between federal states.

Given their high abundance in the landscape<sup>11</sup> small streams below 10 km<sup>2</sup> are disproportionately less sampled in current monitoring (Figure 3), which may be attributed to the missing categorisation in the WFD. Clearly, there is currently a lack of knowledge on stressor effects on small streams. We analysed only data from small streams, however, for lentic small water bodies this lack might be even greater<sup>16</sup>.

## Influence of agricultural land use and catchment size

We found a strong influence of agriculture on the pollution of streams. Above 25% agriculture within a catchment, it is likely that a RAC will be exceeded, with a further increase in entirely agricultural catchments (above 75 % agriculture). To our knowledge, this is the first study investigating such thresholds of pesticide risk. Previous studies examined thresholds for the percentage of agricultural land use with respect to the response of biological communities, integrating different agricultural stressors. Feld<sup>42</sup> found change points of biological community metrics at 40% agricultural land use in lowland streams in Europe. Similarly, Waite<sup>43</sup> found a threshold for aquatic diatoms at 40% agricultural land use in wadeable streams in the United States. Our results coincide with these thresholds and suggest that pesticides might contribute to the observed biological changes.

We did not find a statistically significant relationship between pesticide pollution and catchment size. However, previous studies showed that small streams are more polluted than bigger streams<sup>7,9,44</sup>. This can be explained by the relatively short gradient of catchment sizes in our dataset, with most of the streams with catchments above 10 km<sup>2</sup> and below 100 km<sup>2</sup> (Figure 3, top). For example, the gradient of Schulz<sup>7</sup> covered 6 orders of magnitude.

## Effect of precipitation on pesticide risk

Our results revealed that pesticide sampling for chemical monitoring in Germany is mainly performed when no precipitation occurs. Nevertheless, we found a 36% higher RQ if samples were taken after rainfall events. Samples taken on the day of a rainfall event showed less risk than samples taken one day after a rainfall event. This could be explained by the sampling preceding the rainfall event and the delay between the start of a rain event and the peak in discharge or runoff. The effects of precipitation were more pronounced for the probability to exceed LOQ, with smaller effect sizes for the absolute value of RQ. This may be explained by a higher variability of absolute concentrations. Overall, our results indicate that current pesticide monitoring relying on grab sampling, largely disconnected from precipitation

events, underestimates pesticide risks. Automatic event-driven samplers<sup>3</sup> and passive samplers<sup>45,46</sup> may help overcome these shortcomings and provide a better representation of risks, especially for small water bodies<sup>16</sup>.

We found the highest the probability of exceeding LOQ during summer (10% for Q2) and lowest in the first quarter of the year (4%, Figure 5, bottom right). This annual pattern coincides with the main application season for pesticides in Central Europe. Nevertheless, there are compound-specific differences in the annual pattern, which explains the wide CI for the absolute RQ (Figure 5, bottom left). For example, the herbicide Diflufenican showed the highest RQ and the highest probability of exceeding LOQ during the winter quarters Q1 and Q4 (Supplemental Table S4), which coincides with the application period it is registered for in Germany<sup>47</sup>. Our study suggests that pesticide risks display compound specific spatio-temporal dynamics. Currently, little is known about these and further research on those might provide useful information for future ecological risk assessment. For example, the sensitivity of organisms is often life stage dependent<sup>48</sup> and knowledge on temporal dynamics could inform on concurrent exposure to multiple pesticides, as well as assist to parameterise toxicokinetic and toxicodynamic models<sup>49</sup>. Moreover, our results show that analysing absolute concentrations and probabilities of LOQ together might deliver valuable insights into risk dynamics.

## Pesticides in small streams

Our results suggest that small streams are frequently exposed to ecologically relevant pesticide concentrations. In one-quarter of small streams RACs were exceeded at least once. Stehle and Schulz<sup>9</sup> found the highest percentage of RAC exceedances for organophosphate insecticides. By contrast, we found that neonicotinoid insecticides have highest exceedances of RACs, followed by the organophosphate chlorpyrifos. This difference can be attributed to the low sample size for neonicotinoid insecticides in their study ( $n = 33$ ) compared to the dataset presented here (for example 3,540 samples of Thiacloprid, Figure 6). Overall, our

results suggest that neonicotinoids may currently pose a high risk to freshwater ecosystems. Moreover, our results add further evidence to the growing literature on the risks arising from neonicotinoids for aquatic<sup>50</sup> and terrestrial<sup>51</sup> ecosystems.

Compared to Stehle and Schulz<sup>9</sup> we found higher rates of RAC exceedances for insecticides. They found exceedances in 37.1% of insecticide measurements >LOQ (n = 1352, 23 insecticides), whereas, we found exceedances in 67% of insecticide measurements with RACs >LOQ (n = 1855, 22 insecticides). This could be attributed to different insecticides considered and different underlying RACs. Our study has only 7 insecticides with RACs in common with the insecticides investigated by Stehle and Schulz<sup>9</sup>. Moreover, all RACs were lower in our study (average difference = -0.71 µg/L, range = [-2.757; -0.005]). Nevertheless, it must be noted that the dataset compiled here comprised only samples from grab sampling, which may considerably underestimate pesticide exposure<sup>3,52</sup>.

By contrast, Knauer<sup>44</sup> found exceedances from monitoring data mainly for herbicides and fungicides and only one insecticide Chlorpyrifos-methyl. Moreover, RAC exceedances in Switzerland were generally lower and less abundant (for example 6 exceedances (=0.2%) for Isoproturon with a maximum RQ of 2) compared to our results for Germany. This might reflect differences in pesticide use between countries, ecoregions and RACs used. From the definition of RAC it follows that if the concentration of a compound exceeds its RAC ecological effects are expected. Indeed, Stehle and Schulz<sup>53</sup> found that the biological diversity of stream invertebrates was significantly reduced by 30% at RQ = 1.12 and by 10% at 1/10 of RAC. We found RQ values greater than 1.12 in 25% of small streams and RQ at 1/10 of RAC in 54% of small streams. Consequently, we conclude that agricultural pesticides are on a large scale a major threat to small streams, the biodiversity they host and the services they provide. This threat may exacerbate because pesticides often occur in mixtures<sup>54</sup> and may co-occur with other stressors<sup>55</sup>.

Monitoring data, despite the outlined limitations, provides an opportunity to study large-scale environmental occurrence patterns of pesticides. Furthermore, such nationwide com-

pilations, may not only be used for governmental surveillance, but also to answer other questions, like validation of exposure modelling,<sup>56</sup> retrospective evaluation of regulatory risk assessment<sup>9,44</sup> or occurrences of pesticide mixtures.<sup>54</sup> However, the sampling design needs to account for precipitation events to provide robust data. Our results suggest that exceedances of RACs are landscape dependent and therefore, pesticide regulation should account for landscape features. Moreover, the high exceedances of RACs indicate that greater efforts are needed to describe causal links, which may lead to further developments of the current authorisation procedure.

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## Supporting Information Available

The following files are available free of charge.

- Supplemental\_Materials.pdf : Supplemental Materials (Figures, Tables, Models).

This material is available free of charge via the Internet at <http://pubs.acs.org/>.

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# Graphical TOC Entry

