

Large scale risks from pesticides in small streams

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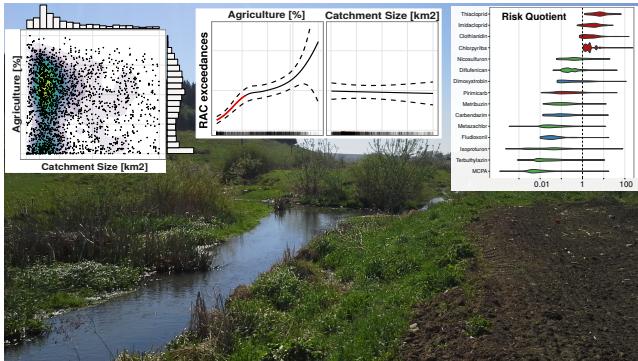
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Abstract

Small streams are important refugia for biodiversity. In agricultural areas they may be at high risk from pesticide pollution. However, most related studies have been limited to a few streams on the regional level, hampering extrapolation to larger scales. We used data from German governmental water quality monitoring to quantify the drivers of pesticide risk and to assess pesticide risk in small streams on a large scale. The data set comprised of 1,766,104 measurements of 478 pesticides (including metabolites) related to 24,743 samples from 2,301 sampling sites. We investigated the influence of agricultural land use, catchment size, as well as precipitation and seasonal dynamics on pesticide risk using new statistical modelling techniques that explicitly consider the limit of quantification. Agricultural land use lead to a 3.7-fold increase in exceedance of risk thresholds when the proportion of agriculture in a catchment exceeded 28 percent. Precipitation increased pesticide risk by 36% and risk was the highest during summer months. Risk thresholds were exceeded in 26% of streams, with the highest risk related to neonicotinoid insecticides. We conclude that pesticides from agricultural land use

16 are a major threat to small streams and their biodiversity and that a realistic pesticide
17 sampling would be driven by precipitation events.

18 **TOC Art**



19

20 **Introduction**

21 More than 50% of the total land area in Germany is used by agriculture¹. In the year 2014
22 more than 45,000 tonnes of 776 authorised plant protection products were sold for application
23 on this area². The applied pesticides may enter surface waters via spray-drift, edge-of-field
24 run-off or drainage^{3–5}. Once entered the surface waters they may have adverse effects on
25 biota and ecosystem functioning⁶. Although it is known that pesticide pollution and its
26 ecological effects increase with the fraction of agricultural land use in the catchment⁷, the
27 shape of the relationship is unknown and studies on potential thresholds are lacking.

28 Two recent studies indicate that pesticides might threaten freshwater biodiversity in the
29 European union. Malaj et al.⁸ analysed data supplied to the European Union (EU) in the
30 context of the Water Framework Directive (WFD) and showed that almost half of European
31 water bodies are at risk from pesticides. Stehle and Schulz⁹ compiled 1,566 measured con-
32 centrations of 23 insecticides in the EU from scientific publications. They found that many
33 of these measurements exceed regulatory acceptable concentrations (RAC). However, these
34 studies reflect only a small amount of potentially available data (173 sites in predominantly

35 mid-sized and large rivers in Malaj et al.⁸ and 138 measurements in Stehle and Schulz⁹),
36 and it is unclear how representative they are for Germany. Much more comprehensive data
37 on thousands of sites are available from national monitoring programs that are setup for the
38 surveillance of water quality, which is done independently by the federal states in Germany
39 in compliance with the WFD¹⁰ and additional state-specific needs. Despite that these data
40 are providing the opportunity to study pesticide risks and other research questions on a
41 large scale with high spatial density, to date these data have not been compiled and related
42 analyses are lacking.

43 Small streams comprise a major fraction of streams¹¹, accommodate a higher proportion
44 of biodiversity compared to larger freshwater systems^{12,13} and play an important role in the
45 recolonization of disturbed downstream reaches^{14,15}. Nevertheless, a clear definition of small
46 streams in terms of catchment or stream size is currently lacking¹⁶. For example, the WFD
47 defines small streams with a catchment size between 10 and 100 km², without further cate-
48 gorisation of streams <10km² and Lorenz et al.¹⁶ defines small streams with catchment size
49 <10km². Moreover, small streams might particularly be at high risk of pesticide contam-
50 ination in case of adjacent agricultural areas given their low dilution potential^{5,7}. Indeed,
51 meta-analyses using data from studies with a few sites reported higher pesticide pollution
52 in smaller streams compared to bigger streams^{7,9}. Despite their ecological relevance and
53 potentially higher pesticide exposure, a recent analysis of pesticide studies showed that a
54 disproportionately small fraction of studies was conducted in small water bodies, and these
55 were largely limited to a few sites¹⁶. Consequently, knowledge on the pesticide pollution
56 of small streams on larger scales is scant. In European law, the Directive 2009/128/EC¹⁷
57 places an obligation on the EU Member States to adopt National Action Plans (NAP) for
58 the Sustainable Use of Plant Protection Products and the German NAP also addresses the
59 knowledge gap concerning pesticide impact on small streams, specifically including those
60 with catchment size <10km².

61 In this study, we compiled and analysed large-scale chemical monitoring data from small

62 streams in Germany. First, we analysed the shape of the relationship between pesticide
63 risk, agricultural land use, and catchment size and examined whether related thresholds
64 for pesticide risks can be derived. Second, we investigated the influence of precipitation
65 and seasonal dynamics on pesticide detections, given that precipitation proved an important
66 driver of pesticide exposure in several small-scale studies^{7,18}, but it is unknown whether a
67 precipitation signal prevails on large scales. Finally, we quantified the current risks from
68 pesticides in small streams in Germany and the compounds accountable for the risk.

69 Methods

70 Data compilation

71 We queried pesticide monitoring data from sampling sites that can be classified as small
72 streams (catchment sizes < 100 km² according to the WFD) from all 13 non-city federal
73 states of Germany (see Supplemental Table S1 for the abbreviations of federal state names)
74 for 2005 to 2015. We homogenised and unified all data provided by the federal states into
75 a database and implemented a robust data-cleaning workflow (see Supplemental Figure S1
76 for details)¹⁹.

77 We identified precipitation at sampling sites by a spatio-temporal intersection of sam-
78 pling events with gridded daily precipitation data (60×30 arcsec resolution) available from
79 the German Meteorological Service (DWD). This data spatially interpolates daily precipi-
80 tation values from local weather stations²⁰. We performed the intersection for the actual
81 sampling date and the day before and extracted precipitation during and up to 48 hours
82 before sampling.

83 Characterization of catchments

84 We compiled a total of 2,369 sampling sites in small streams with pesticide measurements.
85 Alongside, we also queried catchment sizes and agricultural land use within the catchment

86 for the sampling sites from the federal states. Catchment size was provided for 59% of sites.
87 Additionally, we delineated upstream catchments for each of the sampling sites using (i)
88 a digital elevation model (DEM)²¹ and the multiple flow direction algorithm²² as imple-
89 mented in GRASS GIS 7²³ and (ii) from drainage basins provided by the Federal Institute of
90 Hydrology (BfG). Delineated catchments were visually checked for accuracy by comparison
91 with state stream networks and derived information amalgamated with existing data. Thus,
92 catchment size information was available for 99% of all sites (59% from authorities, 24%
93 from DEM and 16% from drainage basins).

94 For each derived catchment (either from DEM or drainage basins) we calculated the
95 % agricultural land-use within the catchment based on the Authoritative Topographic-
96 Cartographic Information System (ATKIS) of the land survey authorities²⁴. Thus, agri-
97 cultural land use information was available for 98% of all sites (24% from authorities, 52%
98 from DEM and 22% from drainage basins). 68 sites (3%) that lacked catchment size or land
99 use information were omitted from the analysis, resulting in 2301 sites used in the analyses
100 outlined below.

101 Characterization of pesticide pollution

102 We characterised pesticide pollution using regulatory acceptable concentrations (RAC)²⁵.
103 RACs are derived during pesticide authorisation as part of the ecological risk assessment.
104 No unacceptable ecological effects are expected if the environmental concentration remains
105 below this concentration. Stehle and Schulz⁹ showed that RAC exceedances reflect a decrease
106 in biodiversity and from this perspective are ecologically relevant indicators. The German
107 Environment Agency (UBA) provided RACs for 107 compounds, including those with the
108 highest detection rates (Supplemental Table S2). Based on these RACs, we calculated Risk
109 Quotients (RQ):

$$RQ_i = \frac{C_i}{RAC_i} \quad (1)$$

110 where C_i is the concentration of a compound i in a sample and RAC_i the respective
111 RAC.

112 Statistical analyses

113 All data-processing and analyses were performed using R²⁶. To display differences in the
114 spectra of analysed compounds between federal states we used Multidimensional Scaling
115 (MDS) based on Jaccard dissimilarity in conjunction with complete linkage hierarchical
116 clustering using the vegan package²⁷. We determined the optimum number of clusters using
117 the average silhouette width²⁸.

118 We expected non-linear responses to agriculture and catchment size and therefore, used
119 generalised additive models (GAM) to establish relationships²⁹. We modelled the number
120 of RAC exceedances ($RQ > 1$) at a site as:

$$\begin{aligned} No(RQ > 1)_i &\sim NB(\mu_i, \kappa) \\ \log(\mu_i) &= \beta_0 + f_1(agri_i) + f_2(size_i) + \log(n_i) \end{aligned} \tag{2}$$

121 where $No(RQ > 1)_i$ is the observed number of RAC exceedances at site i . We modelled
122 $No(RQ > 1)_i$ as resulting from a negative binomial distribution (NB) with mean μ_i and
123 a quadratic mean-variance-relationship ($Var(No(RQ > 1)_i) = \mu_i + \frac{\mu_i^2}{\kappa}$). The proportion
124 of agriculture within the catchment ($agri_i$) and the catchment size of the site ($size_i$) were
125 used as predictors of the number of RAC exceedances. β_0 is the intercept and f_1 and f_2 are
126 smoothing functions using penalized cubic regression splines³⁰. The degree of smoothness was
127 estimated using restricted maximum likelihood (REML) during the model fitting process³¹.
128 The number of measurements per site (n_i) was used as an offset to account for differences
129 in sampling efforts (sampling interval and analysed compound spectrum) at a site and is
130 equivalent to modelling the rate of exceedances. We used point-wise 95% Confidence Intervals
131 (CI) of the first derivative of the fitted smooth to identify regions of statistically significant

132 changes. GAMs were fitted using the mgcv package³¹.

133 To assess the influence of precipitation and seasonality, we modelled the RQ of individual
134 compounds as the response variable. RQ and concentrations show a skewed distribution
135 with an excess of zeros (no pesticides detected and quantified). Therefore, we modelled
136 these as two processes (one generating values below the limit of quantification (LOQ) and
137 one generating values above LOQ) using a Zero-Adjusted Gamma (ZAGA) distribution^{32,33}
138 (Equation 3). These two processes can be interpreted as changes in the mean value of RQ
139 (change in μ) and changes in the probability of exceeding LOQ and showing any risk (change
140 in ν).

$$RQ_i \sim ZAGA(\mu_i, \sigma, \nu_i) = \begin{cases} (1 - \nu_i) & \text{if } y < LOQ \\ \nu_i \times f_{Gamma}(\mu_i, \sigma) & \text{if } y \geq LOQ \end{cases} \quad (3)$$

141 ν_i denotes the probability of a measurement i being above LOQ and f_{Gamma} denotes the
142 gamma function and is used for values equal to or greater LOQ, with μ being the mean
143 and σ the standard deviation of RQ. We used the $\log(x + 0.05)$ transformed precipitation
144 at sampling date ($\log prec_0$) and the day before ($\log prec_{-1}$), as well as quarters of the year
145 ($Q1 - Q4$) as linear predictors for μ and ν . We used appropriate link functions for μ and ν
146 and assumed σ to be constant. Equation 4 summarises the deterministic part of the model
147 for a measurement i .

$$\begin{aligned} \log(\mu_i) &= \log prec_{0i} + \log prec_{-1i} + Q1_i + Q2_i + Q3_i + Q4_i \\ logit(\nu_i) &= \log prec_{0i} + \log prec_{-1i} + Q1_i + Q2_i + Q3_i + Q4_i \end{aligned} \quad (4)$$

148 To account for temporal autocorrelation and differences between federal states we used
149 $site$ nested within $state$ as random intercepts. We implemented this model using the gamlss

¹⁵⁰ package.³⁴

¹⁵¹ We fitted this model separately to each compound with a RAC, measured in at least 1000
¹⁵² samples and with more than 5% of values above LOQ ($n = 22$ compounds, see Supplemental
¹⁵³ Table S3 for a list of compounds). To summarise the coefficients across the 22 modelled
¹⁵⁴ compounds we used a random effect meta-analysis for each model coefficient separately³⁵,
¹⁵⁵ resulting in an averaged effect of the 22 compounds. The results of individual compounds
¹⁵⁶ are provided in the Supplemental Table S4 and Figure S7. The meta-analysis was performed
¹⁵⁷ using the metafor package³⁶.

¹⁵⁸ Results

¹⁵⁹ Overview of the compiled data

¹⁶⁰ The compiled dataset used for analysis comprised 1,766,104 pesticide measurements in 24,743
¹⁶¹ samples from 2,301 sampling sites in small streams. These samples were all taken via grab
¹⁶² sampling. We found large differences between federal states in the number of sampling
¹⁶³ sites and their spatial distribution (Figure 1 and Supplemental Table S1). The number of
¹⁶⁴ small stream sampling sites per state ranged from 1 (Lower Saxonia, NI) to 1139 (North
¹⁶⁵ Rhine-Westphalia, NW). No data were available from Brandenburg.

¹⁶⁶ In total 478 different compounds used as pesticides and their metabolites were measured
¹⁶⁷ at least once (Supplemental Table S2). Most of the compounds were herbicides (179), fol-
¹⁶⁸ lowed by insecticides (117) and fungicides (109). Most samples were taken in the months
¹⁶⁹ April till October, while fewer samples were taken during winter (see Supplemental Fig-
¹⁷⁰ ure S2). We found substantial differences in the spectra of analysed pesticides between
¹⁷¹ federal states (Figure 2). The number of different pesticides per state ranged from 57 (SL)
¹⁷² to 236 (RP) (Supplemental Table S1). Hierarchical clustering revealed that RP and NI anal-
¹⁷³ ysed distinct compound spectra compared to the cluster of other states. However, it has
¹⁷⁴ to be noted that both states surveyed these distinct spectra in special monitoring programs

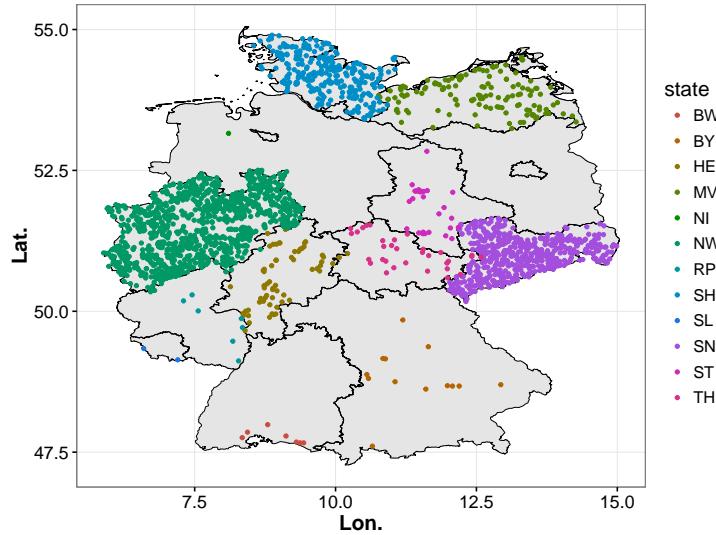


Figure 1: Spatial distribution of the 2,301 small stream sampling sites. Colour codes different federal states (see Supplemental Table S1 for abbreviations).

from only a few sites. Although there was high variability within the remaining cluster, this could not be further split (Figure 2, also Supplemental Figures S3 and S4). 4% (=71,113) of all measurements were concentrations above LOQ.

The distribution of sampling sites across catchment sizes indicated a disproportionately low number of sites with catchments below 10 km^2 , with most sampling sites having catchment sizes between 10 and 25 km^2 (Figure 3).

Influence of agricultural land use and catchment size

The number of RAC exceedances increased strongly and statistically significant up to 28% agriculture within the catchment. The mean number of RAC exceedances per site increased 3.7-fold from 0.10 (no agriculture) to 0.39 (28% agriculture within the catchment). Above this threshold the exceedances levelled. Above 75% agriculture within the catchment the number of exceedances further increased, but the increase was not statistically significant (Figure 4, left). Catchment size had no statistically significant effect on the number of RAC

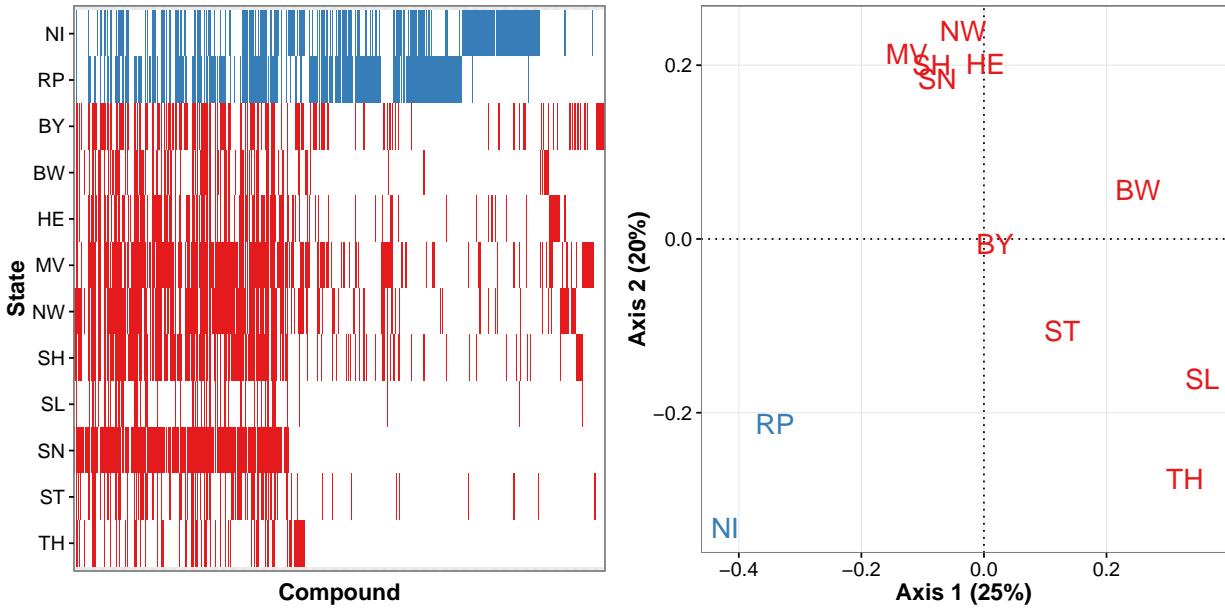


Figure 2: Compound spectra of the different federal states. Left: Barcode plot - each vertical line is an analysed compound. Right: MDS ordination. Colors according to two clusters determined by hierarchical clustering (see Supplemental Figure S3 and S4).

¹⁸⁸ exceedances (Figure 4, right). We also could not detect a statistically significant interaction
¹⁸⁹ between catchment size and agriculture.

¹⁹⁰ Effect of precipitation on pesticide risk

¹⁹¹ The spatio-temporal intersection revealed that most samples were taken during periods of
¹⁹² low precipitation. For example, only 5% of the samples were taken at or after days with
¹⁹³ rainfall events greater than 10mm / day that may lead to run-off (Supplemental Figure S6).
¹⁹⁴ $prec_0$ and $prec_{-1}$ increased the probability of exceeding LOQ and RQ. In Q2 an increase
¹⁹⁵ from 0.1 mm to 10 mm of precipitation before sampling ($prec_{-1}$) lead on average to a 36%
¹⁹⁶ higher mean RQ of 0.05. The probability to exceed LOQ increases 1.6-fold from 8.7% to
¹⁹⁷ 13.5% (Figure 5, top). Effects differed between individual compounds and are provided in
¹⁹⁸ the Supplemental Table S4. Precipitation before sampling ($prec_{-1}$) had a stronger effect than
¹⁹⁹ precipitation during sampling ($prec_0$) on the probability of exceeding LOQ. This difference
²⁰⁰ was less pronounced for the mean value of RQ (Figure 5, top).

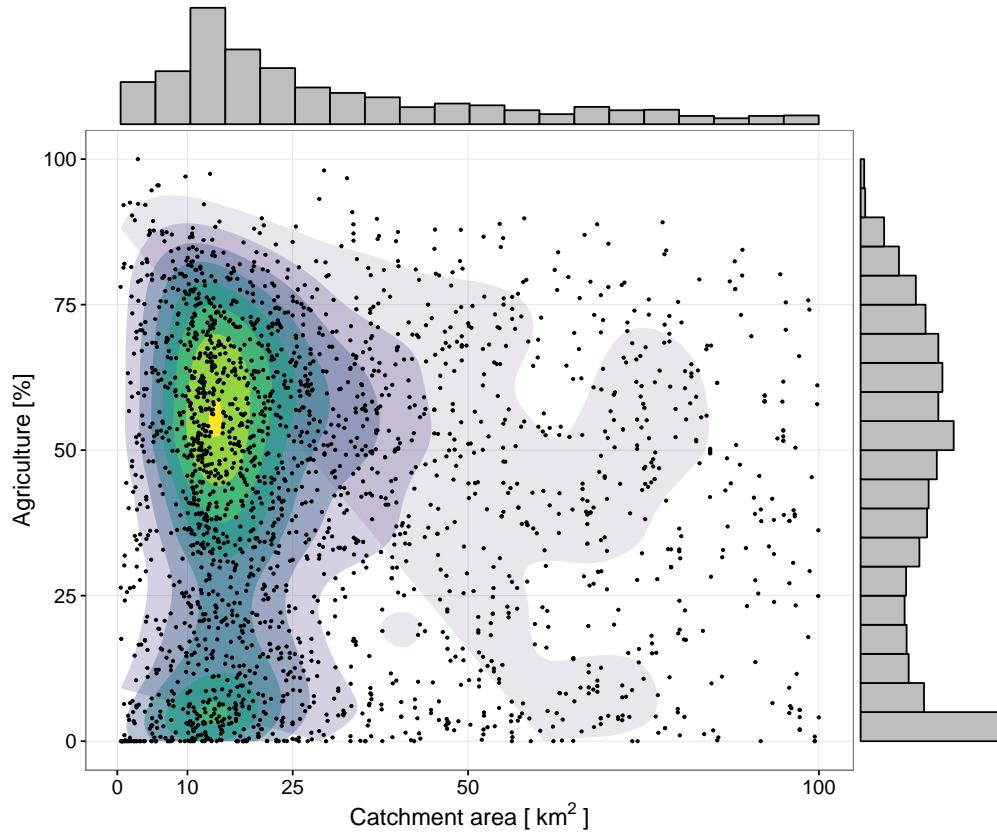


Figure 3: Distribution of catchment area and agriculture within the catchment area across the sampling sites. Colour codes the 2-dimensional density of points.

201 The first quarter showed the lowest RQ and probability of exceeding LOQ. Both increased
 202 during summer months and decreased towards the end of the year. There was a 2.5-fold
 203 higher probability of exceeding LOQ in Q2 (10.6%) than in Q1 (4.6%). The differences
 204 were less pronounced for the mean value of RQ and with less precision (Figure 5, bottom).
 205 Individual compounds showed different temporal patterns (see Supplemental Table S4).

206 **Pesticide risk in small streams**

207 We found RAC exceedances in 25.5% of sampling sites and RQ > 0.1 in 54% of sites. In
 208 23% of sites none of the chemicals, for which RACs were available, were detected (see also
 209 Supplemental Figure S8). Neonicotinoid insecticides and Chlorpyrifos showed the highest
 210 RQ (Figure 6). For Thiacloprid and Chlorpyrifos the RAC was equal or less than LOQ,

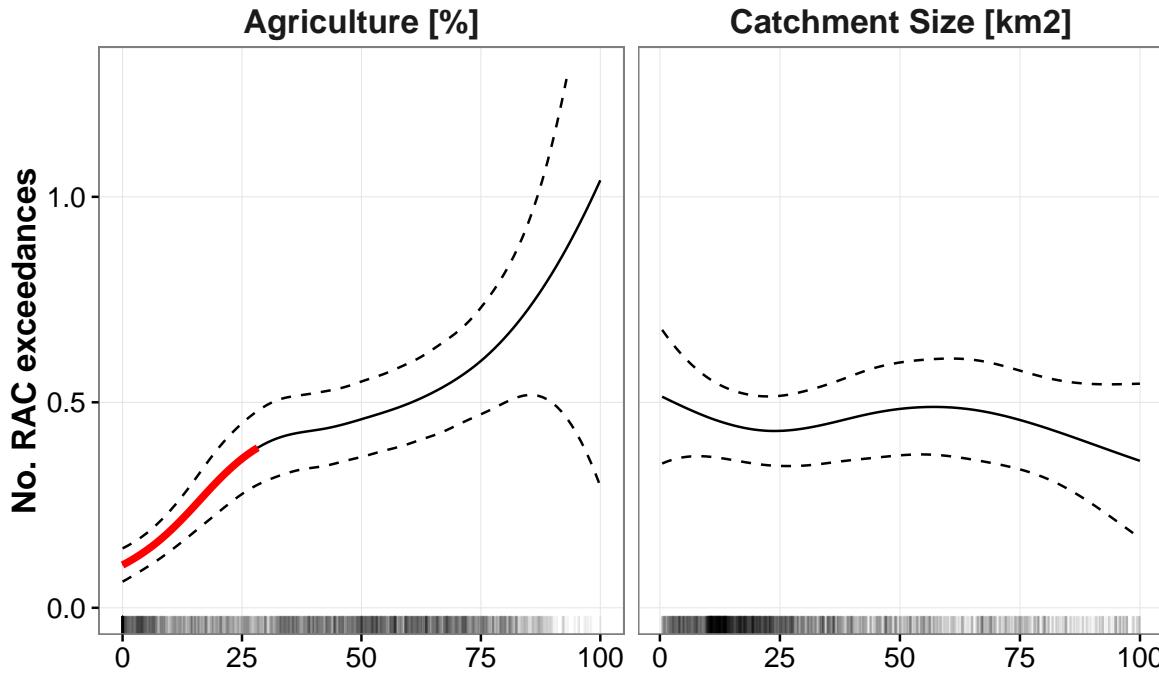


Figure 4: Effect of percent agriculture within the catchment (left) and catchment size (right) on the number of RAC exceedances. Red line marks statistically significant changes. Dashed lines denote 95% point-wise Confidence Intervals.

211 therefore, all detections have a $RQ \geq 1$. The herbicides Nicosulfuron and Diflufenican, as
 212 well as the fungicide Dimoxystrobin also showed high exceedances of RQ (26.7, 14.1 and
 213 21.1 % of measurements $>$ LOQ), see also Supplemental Table S5). RAC exceedances were
 214 found in 14% of samples with concentrations $>$ LOQ (and 7.3% of all samples).

215 The highest RQ s were observed for Chlorpyrifos ($\text{max}(RQ) = 220$), Clothianidin ($\text{max}(RQ)$
 216 $= 157$), Dimoxystrobin($\text{max}(RQ) = 117$) and Isoproturon ($\text{max}(RQ) = 80$). Where anal-
 217 ysed, metabolites exhibited the highest detection rates (for example, Metazachlor sulfonic
 218 acid was detected in 84% of all samples where it was analysed ($n = 3038$, see also Supple-
 219 mental Figure S9). Glyphosate was the compound with the highest detection rates (41%, n
 220 $= 3557$ samples), followed by Boscalid (23%, $n = 9886$) and Isoproturon (22%, $n = 19112$).
 221 However, only the latter showed RAC exceedances (Figure 6). In 45.9% of samples more than
 222 one compound was quantified, with a maximum of 54 different compounds in one sample

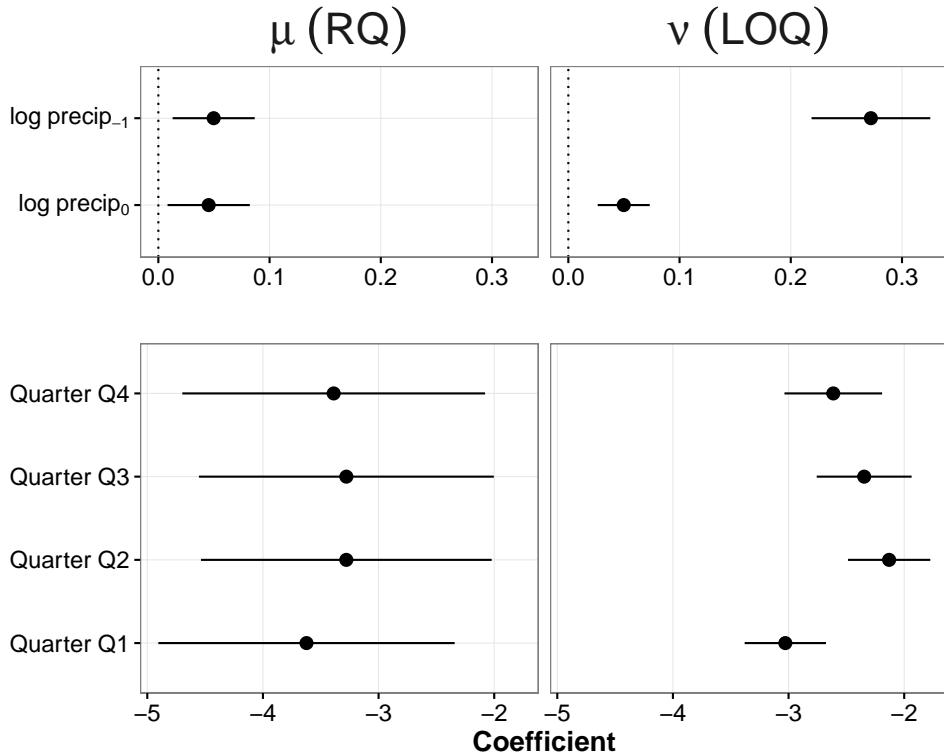


Figure 5: Summarised coefficients (and their 95% CI) for precipitation (top row) and quarter (bottom row) from a meta-analysis of the 22 modelled compounds. Left: coefficients for mean RQ (μ), right: coefficients for probability of exceeding LOQ (ν). Coefficients are shown on the link scale (see Eq. 4). Single compound coefficients are provided in Supplemental Table S4 and Figure S7).

²²³ (Supplemental Figure S10).

²²⁴ Discussion

²²⁵ Overview on the compiled dataset

²²⁶ The compiled dataset of governmental monitoring data, with a particular focus on small
²²⁷ streams, represents currently the most comprehensive available for Germany. Similar na-
²²⁸ tionwide datasets have been compiled for the Netherlands³⁷, Switzerland³⁸ and the United
²²⁹ States³⁹. While the compilations from Europe are of similar quantity and quality to the data
²³⁰ compiled and analysed here, the compilation used in Stone et al.³⁹ is much smaller, though

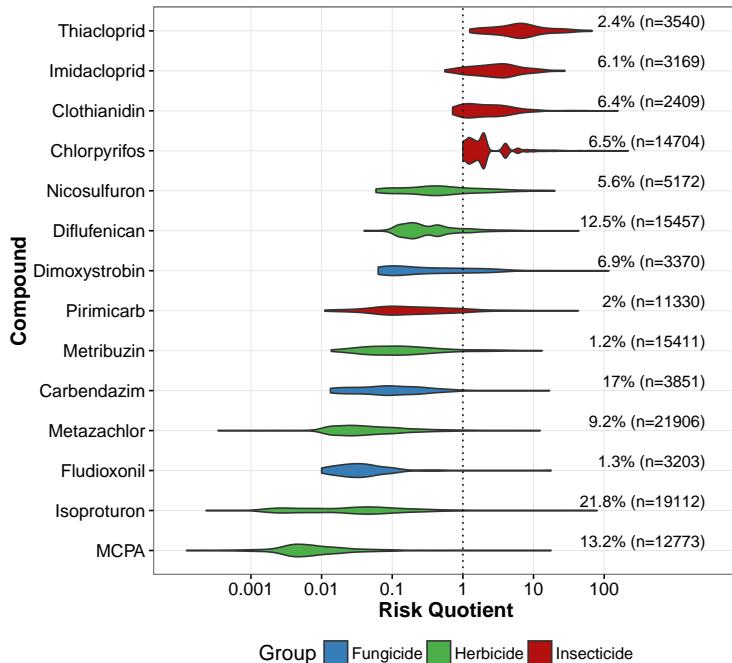


Figure 6: 15 compounds with the highest risk quotients in small streams. Non-detects are not shown due to the logarithmic axis. Numbers on the right give the percentage of values $>\text{LOQ}$ and the total number of samples were the compound was analysed.

231 these data may be complemented by more data in future analyses.

232 A nationwide assessment of pesticide pollution is hampered by inhomogeneous data across
 233 federal states: Beside large differences in the spatial distribution and quantity of sampling
 234 sites (Figure 1), the spectrum of analysed compounds (Figure 2) and the quality of chemical
 235 analyses differed between states. Despite the outlined differences between states, all ecore-
 236 gions occurring in Germany^{40,41} were covered by the presented dataset and thus it might
 237 nonetheless represent a sample covering all types of small streams in Germany. For Thia-
 238 cloprid and Chlorpyrifos the LOQs were above the RAC, which means that exceedances are
 239 likely underestimated. For these compounds a lowering of LOQ through an improvement
 240 of chemical analysis is essential for reliable assessment. Moreover, a nationwide assessment
 241 would benefit from a harmonised spectrum of analysed compounds between federal states.

242 Given their high abundance in the landscape¹¹ small streams below 10 km² are dispro-
 243portionally less sampled in current monitoring (Figure 3), which may be attributed to the

²⁴⁴ missing categorisation in the WFD. Clearly, there is currently a lack of knowledge on stress-
²⁴⁵ sor effects on small streams. We analysed only data from small streams, however, for lentic
²⁴⁶ small water bodies this lack might be even greater¹⁶.

²⁴⁷ **Influence of agricultural land use and catchment size**

²⁴⁸ We found a strong influence of agriculture on the pollution of streams. Above 25% agricultur-
²⁴⁹ e within a catchment, it is likely that a RAC will be exceeded, with a further increase in
²⁵⁰ entirely agricultural catchments (above 75 % agriculture). To our knowledge, this is the first
²⁵¹ study investigating such thresholds of pesticide risk. Previous studies examined thresholds
²⁵² for the percentage of agricultural land use with respect to the response of biological commu-
²⁵³ nities, integrating different agricultural stressors. Feld⁴² found change points of biological
²⁵⁴ community metrics at 40% agricultural land use in lowland streams in Europe. Similarly,
²⁵⁵ Waite⁴³ found a threshold for aquatic diatoms at 40% agricultural land use in wadeable
²⁵⁶ streams in the United States. Our results coincide with these thresholds and suggest that
²⁵⁷ pesticides might contribute to the observed biological changes.

²⁵⁸ We did not find a statistically significant relationship between pesticide pollution and
²⁵⁹ catchment size. However, previous studies showed that small streams are more polluted than
²⁶⁰ bigger streams^{7,9,44}. This can be explained by the relatively short gradient of catchment sizes
²⁶¹ in our dataset, with most of the streams with catchments above 10 km² and below 100 km²
²⁶² (Figure 3, top). For example, the gradient of Schulz⁷ covered 6 orders of magnitude.

²⁶³ **Effect of precipitation on pesticide risk**

²⁶⁴ Our results revealed that pesticide sampling for chemical monitoring in Germany is mainly
²⁶⁵ performed when no precipitation occurs. Nevertheless, we found a 36% higher RQ if samples
²⁶⁶ were taken after rainfall events. Samples taken on the day of a rainfall event showed less risk
²⁶⁷ than samples taken one day after a rainfall event. This could be explained by the sampling
²⁶⁸ preceding the rainfall event and the delay between the start of a rain event and the peak in

269 discharge or runoff. The effects of precipitation were more pronounced for the probability to
270 exceed LOQ, with smaller effect sizes for the absolute value of RQ. This may be explained
271 by a higher variability of absolute concentrations. Overall, our results indicate that cur-
272 rent pesticide monitoring relying on grab sampling, largely disconnected from precipitation
273 events, underestimates pesticide risks. Automatic event-driven samplers³ and passive sam-
274 plers^{45,46} may help overcome these shortcomings and provide a better representation of risks,
275 especially for small water bodies¹⁶.

276 We found the highest the probability of exceeding LOQ during summer (10% for Q2)
277 and lowest in the first quarter of the year (4%, Figure 5, bottom right). This annual pattern
278 coincides with the main application season for pesticides in Central Europe. Nevertheless,
279 there are compound-specific differences in the annual pattern, which explains the wide CI
280 for the absolute RQ (Figure 5, bottom left). For example, the herbicide Diflufenican showed
281 the highest RQ and the highest probability of exceeding LOQ during the winter quarters Q1
282 and Q4 (Supplemental Table S4), which coincides with the application period it is registered
283 for in Germany⁴⁷. Our study suggests that pesticide risks display compound specific spatio-
284 temporal dynamics. Currently, little is known about these and further research on those
285 might provide useful information for future ecological risk assessment. For example, the
286 sensitivity of organisms is often life stage dependent⁴⁸ and knowledge on temporal dynamics
287 could inform on concurrent exposure to multiple pesticides, as well as assist to parameterise
288 toxicokinetic and toxicodynamic models⁴⁹. Moreover, our results show that analysing abso-
289 lute concentrations and probabilities of LOQ together might deliver valuable insights into
290 risk dynamics.

291 Pesticides in small streams

292 Our results suggest that small streams are frequently exposed to ecologically relevant pes-
293 ticide concentrations. In one-quarter of small streams RACs were exceeded at least once.
294 Stehle and Schulz⁹ found the highest percentage of RAC exceedances for organophosphate

295 insecticides. By contrast, we found that neonicotinoid insecticides have highest exceedances
296 of RACs, followed by the organophosphate chlorpyrifos. This difference can be attributed to
297 the low sample size for neonicotinoid insecticides in their study ($n = 33$) compared to the
298 dataset presented here (for example 3,540 samples of Thiacloprid, Figure 6). Overall, our
299 results suggest that neonicotinoids may currently pose a high risk to freshwater ecosystems.
300 Moreover, our results add further evidence to the growing literature on the risks arising from
301 neonicotinoids for aquatic⁵⁰ and terrestrial⁵¹ ecosystems.

302 Compared to Stehle and Schulz⁹ we found higher rates of RAC exceedances for insec-
303 ticides. They found exceedances in 37.1% of insecticide measurements $>\text{LOQ}$ ($n = 1352$,
304 23 insecticides), whereas, we found exceedances in 67% of insecticide measurements with
305 RACs $>\text{LOQ}$ ($n = 1855$, 22 insecticides). This could be attributed to different insecticides
306 considered and different underlying RACs. Our study has only 7 insecticides with RACs in
307 common with the insecticides investigated by Stehle and Schulz⁹. Moreover, all RACs were
308 lower in our study (average difference = $-0.71 \mu\text{g}/\text{L}$, range = [-2.757; -0.005]). Nevertheless,
309 it must be noted that the dataset compiled here comprised only samples from grab sampling,
310 which may considerably underestimate pesticide exposure^{3,52}.

311 By contrast, Knauer⁴⁴ found exceedances from monitoring data mainly for herbicides
312 and fungicides and only one insecticide Chlorpyrifos-methyl. Moreover, RAC exceedances in
313 Switzerland were generally lower and less abundant (for example 6 exceedances (=0.2%) for
314 Isoproturon with a maximum RQ of 2) compared to our results for Germany. This might
315 reflect differences in pesticide use between countries, ecoregions and RACs used. From
316 the definition of RAC it follows that if the concentration of a compound exceeds its RAC
317 ecological effects are expected. Indeed, Stehle and Schulz⁵³ found that the biological diversity
318 of stream invertebrates was significantly reduced by 30% at RQ = 1.12 and by 10% at 1/10
319 of RAC. We found RQ values greater than 1.12 in 25% of small streams and RQ at 1/10 of
320 RAC in 54% of small streams. Consequently, we conclude that agricultural pesticides are
321 on a large scale a major threat to small streams, the biodiversity they host and the services

322 they provide. This threat may exacerbate because pesticides often occur in mixtures⁵⁴ and
323 may co-occur with other stressors⁵⁵.

324 Monitoring data, despite the outlined limitations, provides an opportunity to study large-
325 scale environmental occurrence patterns of pesticides. Furthermore, such nationwide com-
326 pilations, may not only be used for governmental surveillance, but also to answer other
327 questions, like validation of exposure modelling,⁵⁶ retrospective evaluation of regulatory risk
328 assessment^{9,44} or occurrences of pesticide mixtures.⁵⁴ However, the sampling design needs to
329 account for precipitation events to provide robust data. Our results suggest that exceedances
330 of RACs are landscape dependent and therefore, pesticide regulation should account for
331 landscape features. Moreover, the high exceedances of RACs indicate that greater efforts
332 are needed to describe causal links, which may lead to further developments of the current
333 authorisation procedure.

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342 Supporting Information Available

343 The following files are available free of charge.

- 344 • Supplemental _ Materials.pdf : Supplemental Materials (Figures, Tables, Models).

³⁴⁵ This material is available free of charge via the Internet at <http://pubs.acs.org/>.

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