# Multiple Linear Regression - Normal Equation

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Suppose the inputs are:

$$\mathbf{X} = \begin{bmatrix} x_1^{(1)} & x_2^{(1)} & \dots & x_n^{(1)} \\ x_1^{(2)} & x_2^{(2)} & \dots & x_n^{(2)} \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \vdots \\ x_1^{(m)} & x_2^{(m)} & \dots & x_n^{(m)} \end{bmatrix}, \mathbf{y} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ \vdots \\ y^{(m)} \end{bmatrix}$$
(1)

where each row in  $\mathbf{X}$  is the *i-th* sample. Each column in  $\mathbf{X}$  represents the feature (dependent variable) of the dataset.

The goal is to find a linear function **h** to approximate  $y^{(i)}$ , given  $\mathbf{x}^{(i)}$ 

 $x_0^{(i)}$  will be added into **X** where  $x_0^{(i)}=1$  to simplify the notations for the finding of the constant in the linear equation later. Hence,

$$\mathbf{X} = \begin{bmatrix} x_0^{(1)} & x_1^{(1)} & \dots & x_n^{(1)} \\ x_0^{(2)} & x_1^{(2)} & \dots & x_n^{(2)} \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \vdots \\ x_0^{(m)} & x_1^{(m)} & \dots & x_n^{(m)} \end{bmatrix}, \mathbf{y} = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ \vdots \\ y^{(m)} \end{bmatrix}$$
 (2)

The linear function  $\mathbf{h}$  is

$$\mathbf{h}_{\theta}(\mathbf{x}^{(i)}) = \theta_0 x_0^{(i)} + \theta_1 x_1^{(i)} + \dots + \theta_n x_n^{(i)}$$
(3)

or

$$\mathbf{h}_{\theta}(\mathbf{x}^{(i)}) = \sum_{j=0}^{n} \theta_{j} x_{j}^{(i)} \tag{4}$$

where  $\theta_i \in \mathbb{R}$  and i = 1, ..., m, such that

$$\mathbf{J}(\theta_0, ..., \theta_n) = \frac{1}{2m} \sum_{i=1}^{m} (\mathbf{h}_{\theta}(\mathbf{x}^{(i)}) - y^{(i)})^2$$
 (5)

is minimized.

Taking  $\theta$  as a vector,

$$\boldsymbol{\theta} = \begin{bmatrix} \theta_0 \\ \theta_1 \\ \vdots \\ \vdots \\ \theta_n \end{bmatrix} \tag{6}$$

we can rewrite (3) or (4) as

$$\mathbf{h}_{\theta}(\mathbf{x}^{(i)}) = \boldsymbol{\theta}^T \mathbf{x}^{(i)} \tag{7}$$

Since

$$(\mathbf{X}\boldsymbol{\theta} - \mathbf{y})^T (\mathbf{X}\boldsymbol{\theta} - \mathbf{y}) = \sum_{i=1}^m (\mathbf{h}_{\theta}(\mathbf{x}^{(i)}) - y^{(i)})^2 \quad , \tag{8}$$

(5) can be rewritten as (Note that  $X\theta$  is a vector.)

$$\mathbf{J}(\boldsymbol{\theta}) = \frac{1}{2m} (\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} - 2(\mathbf{X} \boldsymbol{\theta})^T \mathbf{y} + \mathbf{y}^T \mathbf{y}) \qquad See \ Appendix \ 1.1$$
 (9)

Since **J** is a polynomial function of degree 2, to find  $\boldsymbol{\theta}$  such that **J** is at minimum, we want to find

$$\frac{\partial \mathbf{J}}{\partial \boldsymbol{\theta}} = 0 \tag{10}$$

Since,

$$\frac{\partial \mathbf{J}}{\partial \boldsymbol{\theta}} = 2\mathbf{X}^T \mathbf{X} \boldsymbol{\theta} - 2\mathbf{X}^T \mathbf{y} \qquad See \ Appendix \ 1.2 \tag{11}$$

then,

$$2\mathbf{X}^{T}\mathbf{X}\boldsymbol{\theta} - 2\mathbf{X}^{T}\mathbf{y} = 0$$
$$\mathbf{X}^{T}\mathbf{X}\boldsymbol{\theta} = \mathbf{X}^{T}\mathbf{y}$$
 (12)

This system (12) is known as the **normal equations** for  $\theta$ . Furthermore, if  $\mathbf{X}^T\mathbf{X}$  is invertible, then there exist a unique solution for  $\theta$ , such that

$$\boldsymbol{\theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \tag{13}$$

#### References

- [1] Nakos, G., and Joyner, D. (1998). Linear algebra with applications. PWS Publishing Company.
- [2] Derivation of the Normal Equation for linear regression Eli Bendersky's website, 2019

# 1 Appendix

### 1.1 Proof for (9)

Using equation (5) and (8),

$$\mathbf{J}(\boldsymbol{\theta}) = \frac{1}{2m} (\mathbf{X}\boldsymbol{\theta} - \mathbf{y})^T (\mathbf{X}\boldsymbol{\theta} - \mathbf{y})$$

 $Note: (X\pmb{\theta} - y)$  is a vector. To multiply a vector by its own transpose is equivalent to squaring the vector.

$$\begin{split} &= \frac{1}{2m} ((\mathbf{X}\boldsymbol{\theta})^T - \mathbf{y}^T) (\mathbf{X}\boldsymbol{\theta} - \mathbf{y}) \\ &= \frac{1}{2m} (\boldsymbol{\theta}^T \mathbf{X}^T - \mathbf{y}^T) (\mathbf{X}\boldsymbol{\theta} - \mathbf{y}) \\ &= \frac{1}{2m} (\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X}\boldsymbol{\theta} - (\mathbf{X}\boldsymbol{\theta})^T \mathbf{y} - \mathbf{y}^T (\mathbf{X}\boldsymbol{\theta}) + \mathbf{y}^T \mathbf{y}) \end{split}$$

Since  $(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y} = \mathbf{y}^T (\mathbf{X}\boldsymbol{\theta})$ , See Appendix 1.1.1

$$\mathbf{J}(\boldsymbol{\theta}) = \frac{1}{2m} (\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} - 2(\mathbf{X} \boldsymbol{\theta})^T \mathbf{y} + \mathbf{y}^T)$$

## 1.1.1 Proof for $(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y} = \mathbf{y}^T (\mathbf{X}\boldsymbol{\theta})$

$$(\mathbf{X}\boldsymbol{\theta})^{T}\mathbf{y} = \begin{pmatrix} \begin{bmatrix} x_{0}^{(1)} & \dots & x_{n}^{(1)} \\ \vdots & \ddots & \vdots \\ x_{0}^{(m)} & \dots & x_{n}^{(m)} \end{bmatrix} \begin{bmatrix} \theta_{0} \\ \vdots \\ \vdots \\ \theta_{n} \end{bmatrix} \end{pmatrix}^{T} \begin{bmatrix} y^{(1)} \\ \vdots \\ y^{(m)} \end{bmatrix}$$

$$= \begin{pmatrix} \begin{bmatrix} x_{0}^{(1)}\theta_{0} & + & \dots & + & x_{n}^{(1)}\theta_{n} \\ \vdots & \ddots & \ddots & \vdots \\ x_{0}^{(m)}\theta_{0} + & + & \dots & + & x_{n}^{(m)}\theta_{n} \end{bmatrix} \end{pmatrix}^{T} \begin{bmatrix} y^{(1)} \\ \vdots \\ y^{(m)} \end{bmatrix}$$

$$= \begin{bmatrix} (x_{0}^{(1)}\theta_{0})y^{(1)} & + & \dots & + & (x_{0}^{(m)}\theta_{0})y^{(m)} \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \ddots & \vdots \\ (x_{n}^{(1)}\theta_{n})y^{(1)} & + & \dots & + & (x_{n}^{(m)}\theta_{n})y^{(m)} \end{bmatrix}$$

$$= \begin{bmatrix} y^{(1)}(x_{0}^{(1)}\theta_{0}) & + & \dots & + & y^{(m)}(x_{0}^{(m)}\theta_{0}) \\ \vdots & \vdots & \ddots & \ddots & \vdots \\ y^{(1)}(x_{n}^{(1)}\theta_{n}) & + & \dots & + & y^{(m)}(x_{n}^{(m)}\theta_{n}) \end{bmatrix}$$

$$= \begin{pmatrix} \begin{bmatrix} y^{(1)} \\ \vdots \\ y^{(m)} \end{bmatrix} \end{pmatrix}^{T} \begin{pmatrix} \begin{bmatrix} x_{0}^{(1)}\theta_{0} & + & \dots & + & x_{n}^{(m)}\theta_{n} \\ \vdots & \ddots & \ddots & \vdots \\ x_{0}^{(m)}\theta_{0} + & + & \dots & + & x_{n}^{(m)}\theta_{n} \end{bmatrix} \end{pmatrix}$$

$$= \begin{bmatrix} y^{(1)} \end{bmatrix}^{T} \begin{pmatrix} \begin{bmatrix} x_{0}^{(1)}\theta_{0} & + & \dots & + & x_{n}^{(m)}\theta_{n} \\ \vdots & \ddots & \ddots & \vdots \\ x_{0}^{(m)}\theta_{0} + & + & \dots & + & x_{n}^{(m)}\theta_{n} \end{bmatrix} \end{pmatrix}$$

$$= \mathbf{y}^{T}(\mathbf{X}\boldsymbol{\theta})$$

$$= \mathbf{y}^{T}(\mathbf{X}\boldsymbol{\theta})$$

#### 1.2 Proof for (11)

$$\mathbf{J}(\boldsymbol{\theta}) = \frac{1}{2m} (\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} - 2(\mathbf{X} \boldsymbol{\theta})^T \mathbf{y} + \mathbf{y}^T \mathbf{y})$$

The equation above (9) be broken down to three expressions,

$$\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} \tag{14}$$

 $2(\mathbf{X}\boldsymbol{\theta})^T\mathbf{y} \tag{15}$ 

$$\mathbf{y}^T \mathbf{y} \tag{16}$$

Expression (16) does not depend on  $\theta$ , hence the partial derivative of the expression with respect to  $\theta$  results in 0.

As for expression (14),

Since,

$$\begin{split} \frac{\partial}{\partial \theta_0}(\pmb{\theta}^T\mathbf{X}^T\mathbf{X}\pmb{\theta}) &= [2(x_0^{(1)}x_0^{(1)} + \ldots + x_0^{(m)}x_0^{(m)})\theta_0 + \ldots + (x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_n] \\ &+ \ldots + [(x_0^{(1)}x_n^{(1)} + \ldots + x_0^{(m)}x_n^{(m)})\theta_n] \\ &= [2(x_0^{(1)}x_0^{(1)} + \ldots + x_0^{(m)}x_0^{(m)})\theta_0 + \ldots + 2(x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_n] \\ &= 2[(x_0^{(1)}x_0^{(1)} + \ldots + x_0^{(m)}x_0^{(m)})\theta_0 + \ldots + (x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_n] \\ &\cdot \\ &\cdot \\ &\cdot \end{split}$$

$$\begin{split} \frac{\partial}{\partial \theta_n}(\pmb{\theta}^T\mathbf{X}^T\mathbf{X}\pmb{\theta}) &= [(x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_0] + \ldots + [(x_0^{(1)}x_n^{(1)} + \ldots + x_0^{(m)}x_n^{(m)})\theta_0 \\ &+ \ldots + 2(x_n^{(1)}x_n^{(1)} + \ldots + x_n^{(m)}x_n^m)\theta_n] \\ &= [2(x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_0] + \ldots + 2(x_n^{(1)}x_n^{(1)} + \ldots + x_n^{(m)}x_n^m)\theta_n] \\ &= 2[(x_n^{(1)}x_0^{(1)} + \ldots + x_n^{(m)}x_0^{(m)})\theta_0] + \ldots + (x_n^{(1)}x_n^{(1)} + \ldots + x_n^{(m)}x_n^m)\theta_n] \end{split}$$

then,

$$\frac{\partial}{\partial \boldsymbol{\theta}} (\boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta}) = 2 \mathbf{X}^T \mathbf{X} \boldsymbol{\theta}$$

As for expression (15),

$$\frac{\partial}{\partial \boldsymbol{\theta}} (2(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y}) = \frac{\partial}{\partial \boldsymbol{\theta}} \left( 2 \begin{bmatrix} (x_0^{(1)} \theta_0) y^{(0)} & + & \dots & + & (x_0^{(m)} \theta_0) y^{(n)} \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ (x_n^{(1)} \theta_n) y^{(0)} & + & \dots & + & (x_n^{(m)} \theta_n) y^{(n)} \end{bmatrix} \right)$$

Since

$$\frac{\partial}{\partial \theta_0} (2(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y}) = 2((x_0^{(1)})y^{(0)} + \dots + (x_0^{(m)})y^{(n)})$$

$$\vdots$$

$$\vdots$$

$$\frac{\partial}{\partial \theta_n} (2(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y}) = 2((x_n^{(1)})y^{(0)} + \dots + (x_n^{(m)})y^{(n)})$$

then,

$$\frac{\partial}{\partial \boldsymbol{\theta}} (2(\mathbf{X}\boldsymbol{\theta})^T \mathbf{y}) = 2\mathbf{X}^T \mathbf{y}$$