

CS181 Practice Questions: Linear Regression, Continued

1. Posterior Weight Distribution Using Bayes' Rule for Linear Gaussian Systems

Some background: In section (2.3.3), Bishop derives the following facts about linear Gaussian systems: assuming we have a marginal distribution on \mathbf{x} and a conditional distribution on \mathbf{y} given by

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Lambda}^{-1}) \quad (2.113)$$

$$p(\mathbf{y} | \mathbf{x}) = \mathcal{N}(\mathbf{y} | \mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1}) \quad (2.114)$$

then

$$p(\mathbf{x} | \mathbf{y}) = \mathcal{N}(\mathbf{x} | \boldsymbol{\Sigma}(\mathbf{A}^\top \mathbf{L}(\mathbf{y} - \mathbf{b}) + \boldsymbol{\Lambda}\boldsymbol{\mu}), \boldsymbol{\Sigma}) \quad (2.116)$$

where

$$\boldsymbol{\Sigma} = (\boldsymbol{\Lambda} + \mathbf{A}^\top \mathbf{L} \mathbf{A})^{-1}. \quad (2.117)$$

Now, we know from (3.10) in Bishop that the regression likelihood can be written as

$$p(\mathbf{t} | \mathbf{w}) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^\top \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1}),$$

where $\beta = \frac{1}{\sigma^2}$. If the prior distribution on \mathbf{w} is given by $p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$, derive that the posterior distribution $p(\mathbf{w} | \mathbf{t})$ is given by

$$p(\mathbf{w} | \mathbf{t}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N)$$

where

$$\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\boldsymbol{\Phi}^\top \mathbf{t})$$

$$\mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta\boldsymbol{\Phi}^\top \boldsymbol{\Phi}$$

in the following way:

- Write down the likelihood in the form of a multivariate Gaussian.
- Explain what we need to substitute for $\mathbf{x}, \mathbf{y}, \boldsymbol{\mu}, \boldsymbol{\Lambda}, \mathbf{A}, \mathbf{b}, \mathbf{L}$ (respectively) in equations (2.113)-(2.117) to derive the posterior.

2. Posterior Weight Distribution By Completing the Square (Bishop 3.7)

We know from (3.10) in Bishop that the regression likelihood can be written as

$$\begin{aligned} p(\mathbf{t} | \mathbf{w}) &= \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^\top \phi(\mathbf{x}_n), \beta^{-1}) \\ &\propto \exp \left(-\frac{\beta}{2} (\mathbf{t} - \mathbf{\Phi} \mathbf{w})^\top (\mathbf{t} - \mathbf{\Phi} \mathbf{w}) \right), \end{aligned}$$

where $\beta = \frac{1}{\sigma^2}$ and in the second line above we have ignored the Gaussian normalizing constants. By completing the square, show that with a prior distribution on \mathbf{w} given by $p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$, the posterior distribution $p(\mathbf{w} | \mathbf{t})$ is given by

$$p(\mathbf{w} | \mathbf{t}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N)$$

where

$$\begin{aligned} \mathbf{m}_N &= \mathbf{S}_N (\mathbf{S}_0^{-1} \mathbf{m}_0 + \beta \mathbf{\Phi}^\top \mathbf{t}) \\ \mathbf{S}_N^{-1} &= \mathbf{S}_0^{-1} + \beta \mathbf{\Phi}^\top \mathbf{\Phi} \end{aligned}$$

3. Predictive Distribution

Bishop notes in section (3.2.2) that if our prior distribution on \boldsymbol{w} is

$$p(\boldsymbol{w}) = \mathcal{N}(\boldsymbol{w} \mid \mathbf{0}, \alpha^{-1} \boldsymbol{I}),$$

and if we assume again that our likelihood involves an inverse variance parameter β , then the predictive distribution of t for a new datapoint \boldsymbol{x} is given by

$$p(t \mid \boldsymbol{t}, \alpha, \beta) = \int p(t \mid \boldsymbol{w}, \beta) p(\boldsymbol{w} \mid \boldsymbol{t}, \alpha, \beta) d\boldsymbol{w} \quad (3.57)$$

Does the equation in (3.57) make any independence assumptions about the variables involved? If so, which?

4. Deriving Lasso Regularization with Lagrange Multipliers

Show that minimization of the unregularized sum-of-squares error function given by

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N (t_n - \mathbf{w}^\top \phi(\mathbf{x}_n))^2,$$

subject to the constraint

$$\sum_{j=1}^M |w_j| \leq \eta,$$

is equivalent to minimizing the regularized error function

$$\frac{1}{2} \sum_{n=1}^N (t_n - \mathbf{w}^\top \phi(\mathbf{x}_n))^2 + \frac{\lambda}{2} \sum_{j=1}^M |w_j|$$

5. Connection between Priors and Regularization

Consider the Bayesian linear regression model given in Bishop 3.3.1. The prior is given by

$$p(\mathbf{w}|\alpha) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I}),$$

where α is the precision parameter that controls the variance of the Gaussian prior. The likelihood can be written as

$$p(\mathbf{t}|\mathbf{w}) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^\top \phi(\mathbf{x}_n), \beta^{-1}),$$

Using the fact that the posterior is the product of the prior and the likelihood, show that maximizing the log posterior (i.e. $\ln p(\mathbf{w}|\mathbf{t}) = \ln p(\mathbf{w}|\alpha) + \ln p(\mathbf{t}|\mathbf{w})$) is equivalent to minimizing the regularized error term given by $E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$ with

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N (t_n - \mathbf{w}^\top \phi(\mathbf{x}_n))^2$$
$$E_W(\mathbf{w}) = \frac{\lambda}{2} \mathbf{w}^\top \mathbf{w}$$

Do this by writing $\ln p(\mathbf{w}|\mathbf{t})$ as a function of $E_D(\mathbf{w})$ and $E_W(\mathbf{w})$, dropping constant terms if necessary.

Conclude that maximizing this posterior is equivalent to minimizing the regularized error term given by $E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$.

(Hint: take $\lambda = \alpha/\beta$)

6. Bayesian Updates in Linear Regression, Bishop 3.8

Suppose we have the standard Bayesian linear regression model and we have already observed N data points, so the posterior distribution is the same as the one derived in problem 2,

$$p(\mathbf{w} \mid \mathbf{t}) = \mathcal{N}(\mathbf{w} \mid \mathbf{m}_N, \mathbf{S}_N)$$

where

$$\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\mathbf{\Phi}^\top \mathbf{t})$$

$$\mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta\mathbf{\Phi}^\top \mathbf{\Phi}$$

Suppose we observe a new data point (x_{N+1}, t_{N+1}) . Show that the resulting posterior distribution is of the same form with \mathbf{m}_{N+1} and \mathbf{S}_{N+1} .