## SVD and Least Squares Problems

Note Title

\* LS via SVD

Recall the LS solution via aR factorization:
(1) Compute reduced aR of A.
(2) Compute  $y = \hat{Q}^T b$ .

(3) Solve RX = Y - (\*)

If A: full rank, then Rii +0, 1≤i≤n, and the triangular system (\*) has a unique LS solution.

Now using the reduced SVD of A, i.e.,  $A = \hat{U} \Sigma V^T$ , we can also solve the normal egn:

 $A^{\mathsf{T}}A \times = A^{\mathsf{T}}B$   $\Leftrightarrow (\hat{U}\hat{\Sigma}V^{\mathsf{T}})^{\mathsf{T}}(\hat{U}\hat{\Sigma}V^{\mathsf{T}}) \times = (\hat{U}\hat{\Sigma}V^{\mathsf{T}})^{\mathsf{T}}B$ 

 $\Leftrightarrow$ 

 $V\hat{\Sigma}^{T}\hat{\Sigma}V^{T}X = V\hat{\Sigma}^{T}\hat{U}^{T}b$   $\hat{\Sigma}^{T}\hat{\Sigma}V^{T}X = \hat{\Sigma}^{T}\hat{U}^{T}b$  since V: ortho. $\Leftrightarrow$ 

 $\ddot{\Sigma} V^{\mathsf{T}} \times = \hat{\mathcal{O}}^{\mathsf{T}} \mathcal{B} \qquad \text{if } A : \text{full rank,}$ i.e., 0, >0, 15j5n

This can be solved easily.

(1) Compute reduced SVD of A.

(2) Compute y = ÛTb.

(3) Solve \(\hat{\Sigma}w = \mathfrak{y}. \tag{\pi\*}

(4) Set X = Vw

Note: (\*\*) is a diagonal system, easier to solve than (\*)!!

Pseudo inverse and SVD

Recall that if  $A \in \mathbb{R}^{m \times n}$  is full rank,  $m > n : A^{\dagger} = (A^{\dagger}A)^{-1}A^{\dagger}$   $m = n : A^{\dagger} = A^{-1}$   $m < n : A^{\dagger} = A^{\dagger}(AA^{\dagger})^{-1}$ 

However, we can define the pseudoinv. using SVD even if A is not full rank!

$$A = U \sum V^{\mathsf{T}}, \qquad \sum = \begin{bmatrix} \sigma_{1} & 0 & 0 \\ 0 & \sigma_{r} \end{bmatrix} r$$

$$A^{\dagger} := V \sum_{i=1}^{T} U_{i}^{T} \sum_{j=1}^{T} := \begin{bmatrix} x_{i,0} & 0 & 0 \\ 0 & x_{i,0} & 0 \end{bmatrix} r$$

as we discussed before, At salisfies the following Moore-Penrose conditions: (i)  $A \times A = A$ ; (ii)  $X A \times = X$ (iii)  $(A \times)^T = A \times$ ; (iv)  $(X A)^T = X A$ .

$$\triangle A \times A = A$$
;  $\triangle X + X = X$ 

$$(A \times)^T = A \times (A \times)^T = XA$$

Such X is uniquely determined and X = A<sup>+</sup>!!

\* Pseudo inverse & Orthogonal Projectors Thm AAt is an ortho. proj. onto range (A)
and AAt = Ur Ur  $A^{\dagger}A$  is an ortho. proj. onto range  $(A^{T})$  and  $A^{\dagger}A = V_{r}V_{r}^{T}$  where  $U_{r} \in \mathbb{R}^{m \times r}$ ,  $V_{r} \in \mathbb{R}^{m \times r}$  consist of the first r columns of U, V, respectively. r = rank (A). (Proof) Let PA := AAT, PAT := ATA. Now,  $P_A = U \Sigma V^T V \Sigma^+ U^T$  $= U \sum_{i=1}^{T} U^{T} = U \begin{bmatrix} I_{i} & 0 \\ 0 & 0 \end{bmatrix} U^{T}$ = UrUr / PA2 = Ur Ur Ur Ur = UrUr = PA  $P_{A}^{T} = (U_{r}U_{r}^{T})^{T} = (U_{r}^{T})^{T}U_{r}^{T} = U_{r}U_{r}^{T} = P_{A} \checkmark$ So it's an ortho. proj. ! Finally, it's also clear that PA maps onto range (A) since range (A) = < u1, ..., Ur>. You can do similarly for PAT ///

Note: Consider any X & Frange (A).

Then = y \in IR^n s.t. X = Ay.

Now PA X = AA<sup>t</sup> X = AA<sup>t</sup> Ay

= Ay = X. "A via

Moore-Pourose (i)

\* Principal Component Analysis (PCA)

(a.k.a. Karhunen-Loève Transform)

is a data analysis technique that

uses an orthogonal transformation to

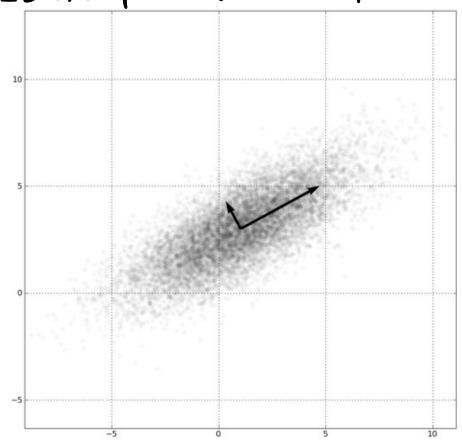
convert a set of observations of possibly

correlated variables into a set of

linearly uncorrelated variables called

"principal components."





One can understand PCA using SVD! But before doing so, we need a bit of Statistics.

Suppose we are given a set of vectors (observations) often  $X_1, X_2, \cdots, X_n$ there are viewed and each Xj EIRd. d: could be huge as n ex. a face image database). Let X:= [X1 ×2···×n] ∈ IR realizations of some stochastic You know the mean (or average) process. of this data set  $\overline{\mathbb{X}} := \frac{1}{n} \sum_{j=1}^{\infty} \mathbb{X}_{j}$ and define the centered data matrix  $X := \begin{bmatrix} X_1 - \overline{X} & X_2 - \overline{X} & \cdots & X_n - \overline{X} \end{bmatrix}$ Note:  $\hat{X} = X \left( I_n - \frac{1}{n} I_n I_n^T \right)$ Good exercise! Now the sample covariance matrix S is defined as  $S := \frac{1}{n} \widetilde{X} \widetilde{X}^{T} \in \mathbb{R}^{d \times d}$ 

Sij indicates the covariance or mutual correlation between the ith and jth entries of data vectors.

PCA is nothing but an eigenvalue decomposition of S, i.e.,  $S = \Phi \Lambda \Phi^T$ ,  $\Lambda = diag(\lambda_1, ..., \lambda_d)$ 

Let's sort  $\lambda_i$ 's as  $\lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_d$ Because  $S^T = S$ , and  $S = \frac{1}{N} \times X^T$ ,

We can show that  $\lambda_i \geq 0$ .  $1 \leq i \leq d$   $\Phi = [\Phi_1 \cdots \Phi_d] \in \mathbb{R}^{d \times d}$ is a matrix containing the eigenvectors. Also thanks to  $S^T = S$ ,  $\Phi$  is an orthogonal matrix whose columns form an ONB of  $\mathbb{R}^d$ .

The change of the bases from  $[\Phi_1 \cdots \Phi_d]$  is achieved simply by  $\Phi^T \times X$ .  $\Phi_i^T \times Y$  is called the ith principal components of X.

PCA was known for a long time, e.g., since the time of Pearson (1901) and Hotelling (1933).

Those days, the measurement dimension I was much smaller than the number of samples n, i.e. I << n

This is called the "classical" setting.

Ex. 5 exam scores of 2000 students

I = 5, n = 2000.

Due to the advent of computers and pensor technology, now we often have

d >> n , the "neo-classical" netting.

Ex. The face database: d=1282, n=143.