

Design and Analysis of Experiments

06 - Simple Comparisons

Version 2.11

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"Science is simply common sense at its best, that is, rigidly accurate in observation, and merciless to fallacy in logic."

> Thomas H. Huxley 1825-1895 English biologist



Simple Comparative Experiments

Statistical inference for two samples

The concepts of comparison between two populations based on information obtained from their samples follow the same principles used for testing hypotheses about a single population;

Inferences for two samples frequently arise when comparing the effect of a technique (treatment) against a *control group*: placebo, classical technique, random search, etc;

Usual questions involve:

- Comparison of means;
- Comparison of variances;
- Comparison of proportions;
- etc.

Example: Length of steel rods



One of the critical aspects of manufacturing steel rods is cutting the bars with a precise length, which is expected by the customers.

This process is prone to errors, which result in additional costs for standardizing and reprocessing the rods.

An engineer is interested in comparing the current controller of the cutting scissors with a new method that could potentially improve the performance of the process.

Example: Length of steel rods

A possible statistical model for this kind of data would be:

$$y_{ij} = \mu_i + \epsilon_{ij} \begin{cases} i = 1, 2 \\ j = 1, \dots, n_i \end{cases}$$

Lets initially assume that the residuals ϵ_{ij} are iid $\mathcal{N}\left(0, \sigma_i^2\right)$, which implies:

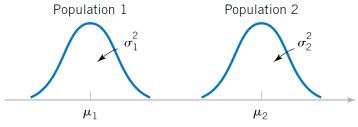


Image: D.C.Montgomery,G.C. Runger, Applied Statistics and Probability for Engineers, Wiley 2003.

Definitions

What we wish is to perform an inference about the difference in the mean values of constructive deviations for the two controllers. In this case, a reasonable response variable would be the *absolute error*, from which we could test our hypotheses.

The statistical hypotheses can be stated as:

$$\begin{cases} H_0: \mu_1 - \mu_2 = 0 \\ H_1: \mu_1 - \mu_2 \neq 0 \end{cases} \quad \text{or, equivalently,} \quad \begin{cases} H_0: \mu_1 = \mu_2 \\ H_1: \mu_1 \neq \mu_2 \end{cases}$$

Suppose a desired significance level $\alpha=0.05$, and that the engineer is interested in detecting any difference larger than 15mm in the mean absolute error with a power $(1-\beta)=0.8$.

Also, lets assume that the variance of the process is unknown but similar for both controllers.

Definitions

Since the variance is unknown, it will have to be estimated from the data. As we are assuming $\sigma_1^2 \approx \sigma_2^2$, we can use the pooled variance estimator:

$$S_p^2 = \frac{(n_1 - 1) S_1^2 + (n_2 - 1) S_2^2}{n_1 + n_2 - 2} = wS_1^2 + (1 - w) S_2^2$$

Based on this estimator and the stated assumptions, we have that:

$$T = rac{(ar{y_1} - ar{y_2}) - (\mu_1 - \mu_2)}{S_{
ho}\sqrt{rac{1}{n_1} + rac{1}{n_2}}} \sim t_{(n_1 + n_2 - 2)}$$

Rejection threshold

If we recall our working hypotheses:

$$\begin{cases} H_0: \mu_1 - \mu_2 = 0 \\ H_1: \mu_1 - \mu_2 \neq 0 \end{cases}$$

we have that, under H_0 :

$$t_0 = \frac{(\bar{y_1} - \bar{y_2}) - (\mu_1 - \bar{\mu_2})}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{(\bar{y_1} - \bar{y_2})}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{(n_1 + n_2 - 2)}$$

We'll reject H_0 at the $(1-\alpha)$ confidence level if $|t_0| \geq t_{\alpha/2;(n_1+n_2-2)}$

Sample sizes

Now recall that the process engineer was interested in some very specific characteristics for his test:

• Significance: $\alpha = 0.05$;

• Power: $(1 - \beta) = 0.8$;

• Minimally interesting effect: $\delta^* = 15mm$

From these specifications, we can obtain the required sample sizes. The derivation of the sample size formulas is not particularly difficult, but we'll concentrate only on the results. More details can be easily found in the literature^a.

^aCheck, for instance, Paul Mathews' Sample Size Calculations, MMB, 2010.

Sample sizes

For the general case of unequal sample sizes, we have:

$$n_1 = \left(1 + \frac{n_1}{n_2}\right) \left(\frac{s_p}{\delta^*}\right)^2 \left(t_{\alpha/2} + t_\beta\right)^2$$

where s_p is the estimated common standard deviation, and $t_{\alpha/2}$ and t_{β} are the $\alpha/2$ and β quantiles of the $t_{(n1+n2-2)}$ distribution. The sample size n_2 can be calculated by simply substituting (n_1/n_2) by (n_2/n_1) .

For equal sample sizes $(n_1 = n_2 = n)$ the expression is simplified to:

$$n=2\left(rac{s_{
ho}}{\delta^*}
ight)^2\left(t_{lpha/2}+t_{eta}
ight)^2$$

Comparison of two means Sample sizes

These formulas are very convenient, but leave us with a riddle: we need variance estimate in order to calculate the sample size, but we need observations to be able to estimate the variance.

There are a few ways to proceed in this case. The most practical are:

- Use process knowledge or historical data to obtain an (initial) estimate of the variance;
- Perform a pilot study and collect samples to estimate the variance.

The first method is almost always preferable since it does not imply additional costs for the experiment.

Sample sizes

If no information is available to estimate the variance, a pilot study must be performed to obtain this value. The sample size required for this pilot study is given by:

$$n_{pilot} \approx 2 \left(\frac{z_{\alpha_n/2}}{e_n}\right)^2$$

where $(1 - \alpha_n)$ is the desired confidence level for the sample size estimate of the main study, and e_n is the maximum relative error allowed for the sample size.

This calculation can yield some scarily large sample sizes for a pilot study (much larger than would be actually required for the main study itself), so use this with caution.

Sample sizes

For the steel rods experiment, suppose that the engineer uses data available from the controller manuals, as well as historical measurements, to estimate the common standard deviation for the cutting process as $\hat{\sigma} \approx 15 mm$.

Assuming that equal sample sizes are desired, we can simply use the formula:

$$n=2\left(rac{\hat{\sigma}}{\delta^*}
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Easy, right?

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Sample sizes

The last problem we have to solve is that the values of $t_{\alpha/2}$ and t_{β} are also dependent of n, which makes the equation

$$n=2\left(rac{\hat{\sigma}}{\delta^*}
ight)^2\left(t_{lpha/2}+t_{eta}
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transcendental in n. We'll have to iterate until we find the smallest n that satisfies:

$$n \geq 2 \left(\frac{\hat{\sigma}}{\delta^*} \right)^2 \left(t_{lpha/2} + t_eta
ight)^2$$

Usually $t_{\alpha/2} \approx z_{\alpha/2}$ is used for the first iteration. Easy, right?

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Bibliography

Required reading



Recommended reading



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