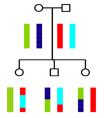
Inference of pairwise relatedness

Ida Moltke, Naples, May 2017



Outline

- 1. Introduction
 - Goal and motiviation
 - Definition of key concepts
 - Overview of current methods for inference
- 2. A simple maximum likelihood solution for called genotypes
 - Overview and intuition
 - Model
 - ML inference
- 3. Problems and more advanced methods
 - Problems with the standard methods
 - NGSrelate (for NGS data)
 - Methods for ancient DNA
 - RelateAdmix (for admixed individuals)
- 4. Exercises

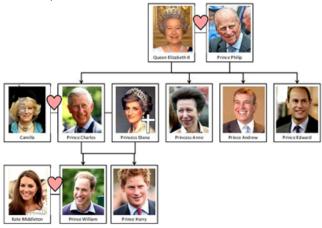
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Goal

- ▶ We want to infer how two individuals are related based on DNA
- ► For example:



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- ► Many methods assume individuals are unrelated.
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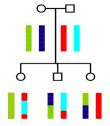
- ► Many methods assume individuals are unrelated.
- ► Violations can lead to wrong conclusions!
- ► Can be of interest in it self, e.g.:
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can reveal cultural practices in the past

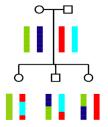


▶ One definition: DNA sequence identity due to recent common ancestry



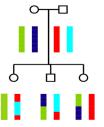
- ► Note with this definition:
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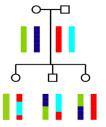
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 - ► Individuals can be identical in a locus without being IBD
 - ► IBD sharing cannot be observed, so needs to be inferred
 - ► Any non-inbred pair will in a locus share 0, 1 or 2 alleles IBD.
 - ► Importantly: the closer related, the more IBD sharing

IBD and relatedness coefficients (R)

▶ $R = (k_0, k_1, k_2)$: fractions of genome with 0, 1 and 2 alleles IBD

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- ▶ We expect different relationships to have different values of *R*, e.g.:

Relationship	k ₀	k_1	k ₂
Monozogotic twins	0	0	1
Parent-offspring	0	1	0
Siblings	0.25	0.5	0.25
Halfsiblings/Uncle-nephew/grandparent-child	0.5	0.5	0
First cousins	0.75	0.25	0
Second cousins	0.9375	0.0625	0
Unrelated	1	0	0

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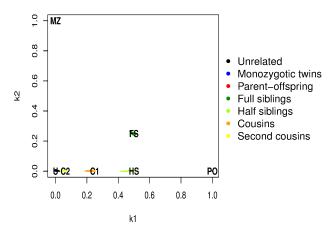
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- ▶ Hence we can (often) use R to infer how two individuals are related
- ► There are other measures like kinship
- ▶ But they are all functions of *R* so we will focus on that

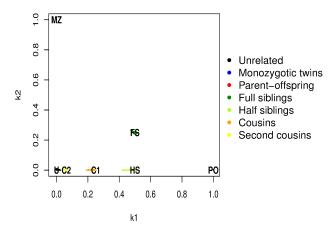
How?

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▶ However, we need to estimate R (because it can't be observed)!

There are a number of estimators:

- several method of moments estimators (e.g. PLINK)
- ▶ some maximum likelihood (ML) estimators (e.g. Thompson, 1975)
- ► most used program is probably PLINK

Importantly, all of them are based on several assumptions

▶ that the individuals are from a diploid species

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- ▶ that the individuals are from a diploid species
- ▶ that the individuals are not inbred
- ▶ that we have **called genotypes** for the two individuals in *L* loci
- ▶ that the individuals are from the same homogenous population
- that the population allele frequencies are known (or that you provide enough samples to estimate them)

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A simple ML approach (overview)

- 1. Define a model that allows us to calculate P(data|R) (likelihood)
- 2. Estimate R by finding the R value the maximises P(data|R) (MLE)

Intuition behind (what information is used)

Although we cannot observe IBD, only genotypes, we note that

- ▶ If alleles are not identical they cannot be IBD
- ▶ If alleles are identical it could be due to either IBD or chance
- ▶ If an allele is frequent, identity will occur often simply by chance
- ► If an allele is rare, identity will occur rarely by chance
- So the rarer the allele, the more identity makes us believe the individuals are IBD

Model

- ▶ We want a model that allows us to calculate P(data|R)
- ▶ Notation: for a pair of non-inbred individuals genotyped in *L* loci we let
 - ▶ G_j be their genotypes in locus j, e.g. $G_j = (AA, aa)$
 - $ightharpoonup Z_j$ indicate how many alleles they share IBD in locus j (unobserved)

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- \blacktriangleright For a single locus, j, first we can write:

$$P(G_j|R) = P(Z_j = 0|R)P(G_j|Z_j = 0) +$$

$$P(Z_j = 1|R)P(G_j|Z_j = 1) +$$

$$P(Z_j = 2|R)P(G_j|Z_j = 2) = \sum_{i=0}^{2} P(Z_j = i|R)P(G_j|Z_j = i)$$

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▶ Note that $P(Z_j = i | R)$ is simply k_i for all $i \in \{0, 1, 2\}$, so we get

$$P(G_j|R) = \sum_{i=0}^2 k_i P(G_j|Z_j = i)$$

▶ Assuming Hardy-Weinberg Equilibrium we can derive $P(G_i|Z_i=i)$:

$Z_j=0$	$Z_j = 1$	$Z_j=2$	
f_{A}^{4}	f_A^3	f_A^2	$\forall A$
$2f_A^2f_a^2$	Ô	Ó	$A \neq a$
$4f_A^3f_a$	$2f_A^2f_a$	0	$A \neq a$
$2f_A^2f_a^2$	$f_A^2 f_a + f_A f_a^2$	$2f_A f_a$	$A \neq a$
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f_A^2	$\forall A$
0	$A \neq a$
0	$A \neq a$
$2f_A f_a$	$A \neq a$
	0

► E.g.

1.
$$P(AA, AA|Z_j = 0) = P(AA)P(AA|AA, Z_j = 0) = P(AA)P(AA) = f_A^2 \times f_A^2$$

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	f_A^4 $2f_A^2f_a^2$ $4f_A^3f_a$	$\begin{array}{ccc} f_A^4 & f_A^3 \\ 2f_A^2f_a^2 & 0 \\ 4f_A^3f_a & 2f_A^2f_a \end{array}$	f_A^4 f_A^3 f_A^2 $2f_A^2f_a^2$ 0 0 $4f_A^3f_a$ $2f_A^2f_a$ 0

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G_j	$Z_j=0$	$Z_j=1$	$Z_j=2$	
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Aa,Aa	$2f_A^2f_a^2$	$f_A^2 f_a + f_A f_a^2$	$2f_Af_a$	$A \neq a$

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3.
$$P(AA, Aa | Z_j = 2) = P(AA)P(Aa|AA, Z_j = 2) = P(AA) \times 0 = 0$$

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- ► E.g.
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 - 3. $P(AA, Aa | Z_j = 2) = P(AA)P(Aa|AA, Z_j = 2) = P(AA) \times 0 = 0$
- ► Captures connection between IBD, genotypes and allele frequencies:
 - ▶ 3) captures that alleles have to be identical to be IBD
 - ▶ 1) and 2) captures that the rarer an allele is, the more will observing identity make us believe the individuals are IBD

ML inference based on the model

► Assuming loci are independent we get the full likelihood:

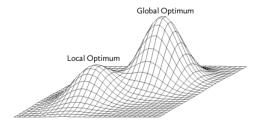
$$P(G_1, G_2, ..., G_L | R) = \prod_{j=1}^{L} P(G_j | R)$$

ML inference based on the model

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$$P(G_1, G_2, ..., G_L | R) = \prod_{j=1}^{L} P(G_j | R)$$

- ► This function is optimized for *R* and we get MLE of *R*
- Most often done using an EM algorithm
- ► NB does not always give you the MLE!



Problems with the standard methods NGSrelate (for NGS data) Methods for ancient DNA RelateAdmix (for admixed individuals

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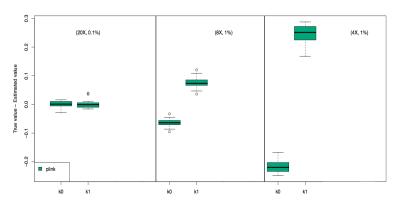
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Problems

Most current methods

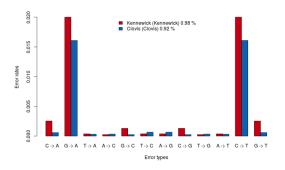
- ▶ work on called genotypes and perform poorly on low depth NGS data
- ► E.g. PLINK on simulated data from 20 1st cousins (k0 is overestimated):



Problems (cont.)

For ancient DNA there is a range of additional issues:

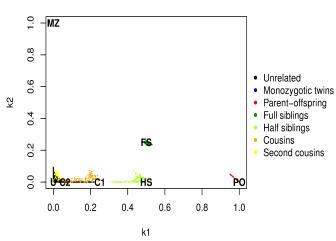
- ▶ very low coverage
- ► increased error rates (especially transitions)



▶ not enough samples available for proper allele frequency estimation

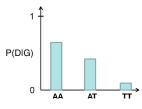
Problems (cont.)

Finally, most current methods have problems with admixed individuals. E.g. PLINK:



► A solution is to use genotype likelihoods (GLs) instead of called genotypes

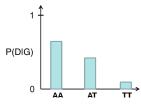




▶ Quantifies the uncertainty in what the genotype is

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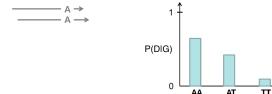




- Quantifies the uncertainty in what the genotype is
- ► We can write a new likelihood for a locus *j* which takes this uncertainty into account using these:

$$P(D_j|R) = \sum_{G_j \in \{\text{all possible genotype pairs}\}} P(D_{j,i1}|G_{j,i1}) P(D_{j,i2}|G_{j,i2}) P(G_j|R)$$

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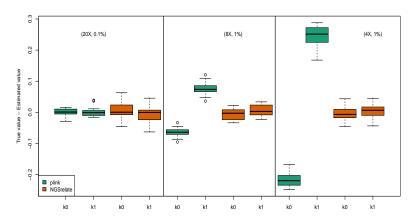


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▶ NB when genotypes are known the GLs are 0 for all but the true genotype, so the likelihood becomes the same as before!

► NGSrelate vs PLINK:

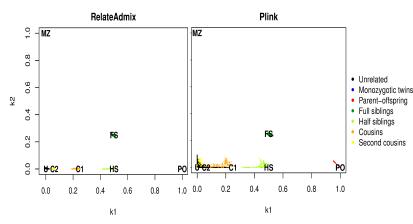


Methods for ancient DNA

- ► In some cases methods like NGSrelate can be used (if frequency info is available and one removes transitions)
- ▶ In other cases you need special methods (e.g. due to lack of allele frequencies)
- Still highly active research field, this year at least 2 new methods came out (Kuhn et al. BioRXiv and Theunert et al. Genetics)

Handling admixture

- ▶ the likelihood can also be adjusted to take admixture into account
- ▶ this is done in the software relateAdmix
- ► comparison of relateAdmix and PLINK on admixed individuals:



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Exercise

Go to http://popgen.dk/ida/EMBONaples2017/web/ and solve exercise 1 & 2

(run the exercises on the server logged in with ssh -Y)