

Data Driven Adaptive Sampling for Low-Rank Matrix Approximation

Arvind Rathnashyam*

Nicolas Boullé†

Alex Townsend‡

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Abstract

We consider the problem of low-rank matrix approximation in case the when the matrix \mathbf{A} is accessible only via matrix-vector products and we are given a budget of $k + p$ matrix-vector products. This situation arises in practice when the cost of data acquisition is high, despite the Numerical Linear Algebra (NLA) costs being low. We create an adaptive sampling algorithm to optimally choose vectors to sample. The Randomized Singular Value Decomposition (rSVD) is an effective algorithm for obtaining the low rank representation of a matrix developed by [17]. Recently, [2] generalized the rSVD to Hilbert-Schmidt Operators where functions are sampled from non-standard Covariance Matrices when there is already prior information on the right singular vectors within the column space of the target matrix, \mathbf{A} . In this work, we develop an adaptive sampling framework for the Matrix-Vector Product Model which does not need prior information on the matrix \mathbf{A} . We provide a novel theoretic analysis of our algorithm with subspace perturbation theory. We extend the analysis of [38] for right singular vector approximations from the randomized SVD in the context of non-symmetric rectangular matrices. We also test our algorithm on various synthetic, real-world application, and image matrices. Furthermore, we show our theory bounds on matrices are stronger than state-of-the-art methods with the same number of matrix-vector product queries.

*Math and CS, Rensselaer Polytechnic Institute, rathna@rpi.edu

†Math, Cambridge, nb690@cam.ac.uk

‡Math, Cornell, ajt453@cornell.edu

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1 Introduction

In many real-world applications, it is often not possible to run experiments in parallel. Consider the following setting, there are a set of n inputs and m outputs, and there exists a PDE such it maps any set of inputs in $\mathbb{C}^m \rightarrow \mathbb{C}^n$. However, to run experiments, it takes hours for set up, execution, or it is expensive, e.g. aerodynamics [13], fluid dynamics [25]. Thus, after each experimental run, we want to sample a function such that in expectation, we will be exploring an area of the PDE which we have the least knowledge of. For Low-Rank Approximation the Randomized SVD, [17], has been theoretically analyzed and used in various applications. Even more recently, [4] discovered if we have prior information on the right singular vectors of \mathbf{A} , we can modify the Covariance Matrix such that the sampled vectors are within the column space of \mathbf{A} . They extended the theory for Randomized SVD where the covariance matrix is now a general PSD matrix. The basis of our analysis is the idea of sampling vectors in the Null-Space of the Low-Rank Approximation. This idea has been introduced recently in Machine Learning in [40] for training neural networks for sequential tasks. In a Bayesian sense, we want to maximize the expected information gain of the PDE in each iteration by sampling in the space where we have no information. This leads to the formulation of our iterative algorithm for sampling vectors for the Low-Rank Approximation. The current state of the art algorithms for low-rank matrix approximation in the matrix-vector product model used a fixed covariance matrix structure. Algorithms for adaptive sampling for CUR low-rank matrix approximation have been explored in [34]. To our knowledge, we are the first paper to give an algorithm for low-rank approximation in the non-symmetric matrix low-rank approximation in the matrix-vector product model.

Contributions.

1. We develop a novel adaptive sampling algorithm for Low-Rank Matrix Approximation problem in the matrix-vector product model which does not utilize prior information of \mathbf{A} .
2. We provide a novel theoretical analysis which utilizes subspace perturbation theory.
3. We perform extensive experiments on matrices with various spectrums and compare with the state of the art methods.

2 Notation, Background Materials, and Relevant Work

In this section we will introduce the notation we use throughout the paper, perturbations of singular spaces, as well as relevant work in the Low-Rank Matrix Approximation Literature.

2.1 Notation

Let $\mathbf{A} \in \mathbb{R}^{m \times n}$ represent the target matrix. $\|\cdot\|$ represents the spectral norm, which is equivalent to the max eigenvalue of the argument, $\sigma_{\max}(\cdot)$. Quasimatrices (matrices with infinite rows and finite columns) will be denoted as a variation of the symbol, $\mathbf{\Omega}$. The pseudoinverse is represented by $(\cdot)^\dagger$ s.t. $\mathbf{X}^\dagger = (\mathbf{X}^* \mathbf{X})^{-1} \mathbf{X}^*$. The Projection Matrix is defined as $\Pi_{\mathbf{Y}} = \mathbf{Y} \mathbf{Y}^\dagger = \mathbf{Y} (\mathbf{Y}^* \mathbf{Y})^{-1} \mathbf{Y}^*$ as the projection on to the column space of \mathbf{Y} . If \mathbf{Y} has orthogonal columns, then $\Pi_{\mathbf{Y}}$ is the Orthogonal Projection defined as $\Pi_{\mathbf{Y}} = \mathbf{Y} \mathbf{Y}^*$. Let $a \wedge b = \min(a, b)$ and $a \vee b = \max(a, b)$. Let $\mathbb{O}_{n,k}$ be the set of all $n \times k$ matrices with orthogonal columns, i.e. $\{\mathbf{V} : \mathbf{V}^* \mathbf{V} = \mathbf{I}_{k \times k}\}$. We also denote $\mathcal{MN}(\mathbf{0}, \mathbf{I}_{n \times n}, \mathbf{I}_{m \times m})$, denote the distribution of $m \times n$ standard gaussian matrices. The Frobenius norm for a matrix is defined as,

$$\|\mathbf{A}\|_F = \left(\sum_{i \in [m]} \sum_{j \in [d]} A_{i,j}^2 \right)^{1/2} = \sqrt{\text{Tr}(\mathbf{A}^* \mathbf{A})} = \sqrt{\text{Tr}(\mathbf{A} \mathbf{A}^*)} \quad (1)$$

We define $\llbracket \mathbf{A} \rrbracket_r$ as the best rank- r approximation to \mathbf{A} w.r.t the Frobenius norm. We use Big-O notation, $y \leq \mathcal{O}(x)$, to denote $y \leq Cx$ for some positive constant, C . We define \mathbb{E} as expectation, \mathbb{P} as probability, and \mathbb{V} as variance.

2.2 Singular Subspace Perturbations

To represent the distance between subspaces we utilize the $\sin \Theta$ norm. Let \mathcal{X}, \mathcal{Y} be subspaces, then we denote the principal angles between subspaces (PABS) \mathcal{X} and \mathcal{Y} as $\frac{\pi}{2} \geq \Theta_1(\mathcal{X}, \mathcal{Y}) \geq \dots \geq \Theta_{m \wedge n}(\mathcal{X}, \mathcal{Y})$. Typically, the norm for distance between subspaces \mathcal{X} and \mathcal{Y} is defined as,

$$\|\sin \Theta(\mathcal{X}, \mathcal{Y})\|_F = \|\Pi_{\mathcal{X}} - \Pi_{\mathcal{Y}}\|_F \quad (2)$$

In a landmark paper by [9], they introduced upper bounds for $\|\sin \Theta(\mathcal{X}, \mathcal{Y})\|$ and $\|\tan \Theta(\mathcal{X}, \mathcal{Y})\|$. A generalized version of the $\sin \Theta$ theorem for rectangular matrices is given in [42].

Theorem 1. [42]. Let $\mathbf{A}, \hat{\mathbf{A}} \in \mathbb{R}^{m \times n}$ have singular values $\sigma_1 \geq \dots \geq \sigma_{m \vee n}$ and $\hat{\sigma}_1 \geq \dots \geq \hat{\sigma}_{m \vee n}$, respectively. Given $j \in 1, \dots, m \vee n$, it follows

$$\|\sin \Theta(\hat{\mathbf{v}}_j, \mathbf{v}_j)\|_F \leq \frac{2 \left(2\sigma_1 + \|\hat{\mathbf{A}} - \mathbf{A}\| \right) \|\hat{\mathbf{A}} - \mathbf{A}\|_F}{\sigma_j^2 - \sigma_{j+1}^2} \wedge 1 \quad (3)$$

However, we would like to note this theorem tends to not be sharp enough for theoretical use. Instead, we introduce Wedin's Theorem,

Theorem 2. [41]. Let $\mathbf{A}, \hat{\mathbf{A}} \in \mathbb{R}^{m \times n}$ have singular values $\sigma_1 \geq \dots \geq \sigma_{m \vee n}$ and $\hat{\sigma}_1 \geq \dots \geq \hat{\sigma}_{m \vee n}$, respectively. Given $j \in 1, \dots, m \vee n$,

$$\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \frac{\|\mathbf{A} - \hat{\mathbf{A}}\|}{\sigma_1 - \hat{\sigma}_2} \quad (4)$$

Theorem 3. [32]. Let $\mathbf{A}, \hat{\mathbf{A}} \in \mathbb{R}^{m \times n}$ have singular values $\sigma_1 \geq \dots \geq \sigma_{m \vee n}$ and $\hat{\sigma}_1 \geq \dots \geq \hat{\sigma}_{m \vee n}$, respectively. Then,

$$\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \frac{2 \|\mathbf{A} - \hat{\mathbf{A}}\|}{\sigma_1 - \sigma_2} \quad (5)$$

Now we give some introduction to singular vector perturbation theory. Given two vectors, $\mathbf{v}, \tilde{\mathbf{v}} \in \mathbb{R}^n$ s.t. $\|\mathbf{v}\| = \|\tilde{\mathbf{v}}\| = 1$, it follows $\cos \Theta(\mathbf{v}, \tilde{\mathbf{v}}) = \mathbf{v}^* \tilde{\mathbf{v}}$. Let \mathbf{V} be the matrix representing an orthonormal basis of vectors in \mathbb{R}^n : $\{\mathbf{v}_1, \dots, \mathbf{v}_n\}$. Using basic ideas in trigonometry we find that,

$$\begin{aligned} \sin^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}) &= 1 - \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}) = \|\mathbf{V}^* \tilde{\mathbf{v}}\|^2 - \|\mathbf{v}_j^* \tilde{\mathbf{v}}\|^2 \\ &= \sum_{i=1}^n \mathbb{I}\{i \neq j\} \|\mathbf{v}_i^* \tilde{\mathbf{v}}\|^2 \triangleq \|\mathbf{V}_{\perp, j}^* \tilde{\mathbf{v}}\|^2 \end{aligned} \quad (6)$$

2.3 Relevant Works

The Randomized Singular Value Decomposition was developed and analyzed thoroughly in [17]; throughout this paper we will refer to this algorithm as HMT. The review work by [27] gives significant theory on the Randomized SVD. [2] proposed learning the Hilbert-Schmidt Operators associated with the Green's Functions with the randomized SVD algorithm. One of their key findings is they can better approximate the HS Operator when they use functions drawn from $\mathcal{GP}(\mathbf{0}, \mathbf{K})$ where \mathbf{K} is not the identity. [3] extended upon previous work on generalizing the Randomized SVD to learning HS Operators. [20] empirically look at Entropy Search for Probabilistic Optimization. [33] analyzed faster algorithms for the approximation of the null-space.

Block iterative methods have also been studied extensively in [18]. The study of block Krylov subspaces have also seen increased attention in the last few years, [37]. Bounds for the $(1 + \varepsilon) \|\mathbf{A} - \mathbf{A}_k\|$ approximation error with randomized block Krylov Subspace methods have been explored in [30, 1].

The most relevant work to ours is likely [12]. The measure of accuracy in the Krylov Subspace is measured by the $\sin \Theta$ norm. We would like to note the Krylov Subspace method takes q times more matrix-vector

products and thus is not a suitable method for our problem. Work similar to ours with regards to upper bounding the sine of the principal angles between subspaces in the context of the Randomized SVD is explored in [11] and [35].

Upper bounds on the tangent of principal angles are visited in [31] and improved in [28]. The highly studied $\sin \Theta$ norms are studied in depth in [8, 32].

Learning algorithms for Low-Rank Matrix Approximations have also been explored. In [23] and [22], the sketching matrix is learned.

A similar analysis of a power method is explored in [19] utilizing subspace perturbation theory. In this work, they consider the Matrix-Vector products have noise. In this work, similarly to [12], it takes d times more matrix-vector products to recover the right singular space. Furthermore, a similar projection-based analysis based on the sines of the singular vector perturbations is done in [26].

The theoretical analysis of greedy algorithms for low-rank matrix approximation tend to be difficult as noted in the discussion of [16]. Such algorithms have been developed for Nyström Approximation in [14, 24].

3 Data Driven Sampling

In this section, we will go over the covariance matrices proposed papers and we consider choosing the optimal covariance matrix adaptively for sampling vectors. In the seminal paper by [17], the covariance matrix is given as:

$$\mathbf{C} = \mathbf{I} \quad (7)$$

In the generalization of the Randomized SVD, when given some prior information of the matrix, the covariance matrix is given as:

$$\mathbf{C} = \mathbf{K} \quad (8)$$

where \mathbf{K} should be close to the right singular vectors of \mathbf{A} . Let $\tilde{\mathbf{V}}$ be the right singular vectors of the SVD of the low-rank approximation at iteration $k - 1$, then the update for the covariance matrix is given as follows:

$$\mathbf{C}^{(k+1)} = \tilde{\mathbf{V}}_{(:,k)} \tilde{\mathbf{V}}_{(:,k)}^* \quad (9)$$

Throughout this paper we will only consider $\mathbf{C}^{(0)} = \mathbf{I}$ due to simplified analysis, however using theory from [2], this can be extended to $\mathbf{C}^{(0)} = \mathbf{K}$ if one has some knowledge of the right singular vectors. A similar algorithm can be found in [40]. Naturally, want to continuously sample in the null space of the the matrix approximation we have already obtained. This ensures we are learning new information in each iteration as we don't want to 'waste' samples which do not learn any new information about the matrix. To further motivate our covariance update, we will introduce the following remark.

Remark 4. Let $\mathbf{U}\Sigma\mathbf{V}^*$ be the SVD of \mathbf{A} , then the Covariance update described in Equation (9) is the optimal covariance update is the optimal covariance matrix for sampling vectors at iteration k .

Remark 4 is an intuitive result, in that when we are learning a matrix \mathbf{A} , we would optimally want to sample the right singular vectors, so the resultant matrix product is the left singular vectors.

3.1 Algorithm

The Pseudo Code for the optimal function sampling is given in Algorithm 1. For efficient updates, we frame all operations as rank-1 updates.

In Algorithm 1, we first sample a standard normal gaussian matrix which can be considered as the oversampling vectors. These oversampling vectors are used to approximate the first singular vector. This is the first vector which is *adaptively* sampled. Next, we form the low-rank approximation $\mathbf{Q}\mathbf{Q}^*\mathbf{A}$ with the adaptive matrix vector query. From here, we adaptively query with the k th right singular value of the SVD of the the low rank approximation at iteration $k - 1$. We believe this algorithm to be the closest to replicate the idea given in Remark 4.

Algorithm 1 Optimal Function Sampling

1: **Input:** HS Operator: \mathcal{F} , Rank: r , Initial Covariance: \mathbf{C} , Oversampling Parameter: p
2: **Output:** Rank- r Approximation, $\hat{\mathbf{A}}_r$
3: $\mathbf{\Omega} \leftarrow \underbrace{\left[\mathcal{N}(\mathbf{0}, \mathbf{C}) \quad \overset{\text{i.i.d.}}{\vdots} \quad \mathcal{N}(\mathbf{0}, \mathbf{C}) \right]}_p$ \triangleright Sample Oversampling Vectors from Standard Normal Matrix
4: $\mathbf{Y} \leftarrow \mathbf{A}\mathbf{\Omega}$ \triangleright Matrix Vector Products
5: $[\mathbf{Q}_0, \sim] \leftarrow \text{QR}(\mathbf{Y})$ \triangleright Find Orthonormal Basis
6: $\tilde{\mathbf{A}} \leftarrow \mathbf{0}_{m \times n}$ \triangleright Initial Low-Rank Approximation
7: **for** $k \in 1, 2, \dots, r$ **do**
8: $\tilde{\mathbf{A}}_k \leftarrow \tilde{\mathbf{A}}_{k-1} + \mathbf{Q}_{(k-1)} \mathbf{Q}_{(k-1)}^* \mathbf{A}$ \triangleright Rank-1 update to the low-rank approximation
9: $[\tilde{\mathbf{U}}, \tilde{\mathbf{\Sigma}}, \tilde{\mathbf{V}}] \leftarrow \text{SVD}(\tilde{\mathbf{A}}_k)$ \triangleright SVD of current low-rank approximation
10: $\mathbf{C}^{(k+1)} \leftarrow \tilde{\mathbf{V}}_{(:,k)} \tilde{\mathbf{V}}_{(:,k)}^*$ \triangleright Form new Covariance Matrix
11: $\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{C}^{(k+1)})$ \triangleright Adaptive Sampling of Vector
12: $\mathbf{Y} \leftarrow [\mathbf{Y} \quad \mathbf{A}\mathbf{x}]$ \triangleright Matrix-Vector Product
13: $[\mathbf{Q}_k, \sim] \leftarrow \text{QR}(\mathbf{Y})$ \triangleright Find Orthonormal Basis
14: **end for**
15: $\tilde{\mathbf{A}}_r \leftarrow \mathbf{Q}_r \mathbf{Q}_r^* \mathbf{A}$ \triangleright Final Low-Rank Approximation
16: **Return:** $\tilde{\mathbf{A}}_r$

4 Theory

In this section we will give the mathematical setup for the theoretical analysis. We will then represent theorems from relevant works on the error bounds for their low-rank approximation methods. We will then give our error bounds and general theory of Algorithm 1 with the proofs in the appendix.

4.1 Setup.

We follow a similar setup as previous literature. Let $\rho \triangleq \text{rank}(\mathbf{A}) \leq m \wedge n$, we will factorize \mathbf{A} as

$$\mathbf{A} = \begin{bmatrix} \mathbf{U}_k & \mathbf{U}_{\rho-k} \end{bmatrix} \begin{bmatrix} \mathbf{\Sigma}_k & \mathbf{0} \\ \mathbf{0} & \mathbf{\Sigma}_{\rho-k} \end{bmatrix} \begin{bmatrix} \mathbf{V}_k^* \\ \mathbf{V}_{\rho-k}^* \end{bmatrix} = \sum_{i=1}^{\rho} \sigma_i \mathbf{u}_i \mathbf{v}_i^* = \sum_{i=1}^{\rho} \mathbf{U}_{(i)} \mathbf{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \quad (10)$$

Furthermore, we let $\mathbf{A}_{(k)} \triangleq \sigma_k \mathbf{u}_k \mathbf{v}_k^*$. Let $\mathbf{\Omega} \in \mathbb{R}^{n \times \ell}$ be a test matrix where $\ell = k + p$ denotes the number of samples and p is the oversampling parameter.

4.2 Previous Literature

We first restate the expected Frobenius error in the Low-Rank Approximation obtained by the Randomized SVD with data sampling from a central and uncorrelated Normal Distribution.

Theorem 5. [17][Theorem 10.5] *Let $\mathbf{A} \in \mathbb{R}^{m \times n}$, $k \geq 2$, oversampling parameter $p \geq 2$, where $k + p \leq m \wedge n$. Let $\mathbf{\Omega} \sim \mathcal{MN}(\mathbf{0}, \mathbf{I}_{n \times n}, \mathbf{I}_{k+p \times k+p})$, and $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\mathbf{\Omega})$. Then,*

$$\mathbb{E} \|\mathbf{A} - (\mathbf{Q}\mathbf{Q}^* \mathbf{A})_k\|_F \leq \left(1 + \frac{k}{p-1}\right)^{1/2} \sqrt{\sum_{j=k+1}^n \sigma_j^2} \quad (11)$$

We will now restate the expected Frobenius norm error in the Low-Rank Approximation obtained by the Randomized SVD with vector sampling from a central and correlated Normal Distribution.

Theorem 6. [2][Theorem 2] Under the same conditions as Theorem 5, except assume the columns of $\mathbf{\Omega}$ are sampled from $\mathcal{N}(\mathbf{0}, \mathbf{K})$.

$$\mathbb{E} \|\mathbf{A} - (\mathbf{Q}\mathbf{Q}^* \mathbf{A})_k\|_F \leq \left(1 + \sqrt{\frac{\beta_k (k+p)}{\gamma_k (p-1)}}\right) \sqrt{\sum_{j=k+1}^n \sigma_j^2} \quad (12)$$

where $\gamma_k = \frac{k}{\lambda_1 \text{Tr}((\mathbf{V}_1^* \mathbf{K} \mathbf{V}_1)^{-1})}$ and $\beta_k = \frac{\text{Tr}(\mathbf{\Sigma}_2^2 \mathbf{V}_2^* \mathbf{K} \mathbf{V}_2)}{\lambda_1 \|\mathbf{\Sigma}_2\|_F^2}$.

In the literature, approximation error bounds on $\|\mathbf{A} - \mathbf{Q}\mathbf{Q}^* \mathbf{A}\|$ typically are of the form

$$\left(1 + \underbrace{\left\| \mathbf{\Sigma}_{\rho-k} (\mathbf{V}_{\rho-k}^* \mathbf{\Omega}) (\mathbf{V}_k^* \mathbf{\Omega})^\dagger \right\|}_{\psi}\right)^{1/2} \|\mathbf{\Sigma}_{\rho-k}\|_F \quad (13)$$

See Theorems 5 and 6 and [6, 15]. However, we find working with $\|(\mathbf{V}_k^* \mathbf{\Omega})^\dagger\|$ is difficult since this norm can be extremely large. The ' ψ ' term in Equation (13) has been studied w.r.t to Krylov sSubspaces in [12]. In [12][Theorem 2.2], Drineas et al. find

$$\psi \leq \left\| \tan \mathbf{\Theta}(\tilde{\mathbf{v}}, \mathbf{v}_k) \right\| \quad (14)$$

Since the tan function is unbounded, upper bounding ψ is difficult and may not lead to strong bounds. First we will introduce a lemma for the resultant vector of sampling from $\mathbf{C}^{(k)}$. Since our general proof technique will be an induction. We first want to understand how well we are able to approximate the first right singular vector. To do this, we must know the singular vector perturbation from the error of the low-rank matrix approximation. With Theorem 1, we can now put bounds on the top right singular vector approximation by the Randomized SVD.

Lemma 7. Let $\mathbf{A} \in \mathbb{C}^{m \times n}$ and \mathbf{Q} be an orthogonal matrix representing the basis of the subspace of $\mathbf{Y} \in \mathbb{C}^{m \times k}$. Let $\mathbf{v} \sim \mathcal{N}_n(\mathbf{0}, \mathbf{I})$ and $\tilde{\mathbf{v}}$ represent the right singular vector of $\mathbf{Q}\mathbf{Q}^* \mathbf{A}$. Denote $\hat{\mathbf{Q}} \triangleq \text{orth}([\mathbf{Y} \ \mathbf{v}])$ and $\tilde{\hat{\mathbf{Q}}} \triangleq \text{orth}([\mathbf{Y} \ \tilde{\mathbf{v}}])$, then we have with probability $1 - \delta$

Claim (i):

$$\left\| \mathbf{A} - \tilde{\hat{\mathbf{Q}}} \tilde{\hat{\mathbf{Q}}}^* \mathbf{A} \right\|_F \leq \Xi \quad (15)$$

Claim (ii):

$$\left\| \mathbf{A} - \hat{\mathbf{Q}} \hat{\mathbf{Q}}^* \mathbf{A} \right\|_F \leq \left\| \mathbf{A} - \tilde{\hat{\mathbf{Q}}} \tilde{\hat{\mathbf{Q}}}^* \mathbf{A} \right\|_F - \left\| (\mathbf{I} - \hat{\mathbf{Q}} \hat{\mathbf{Q}}^*) \mathbf{A} \right\|_2 \left(\mathcal{O}\left(\frac{1}{n^2}\right) \left(1 - \sqrt{\frac{1}{c} \log\left(\frac{2}{\delta}\right)}\right) \right)^{1/2} \quad (16)$$

4.3 Deterministic Results

First, we introduce a necessary result.

Lemma 8. Let $\mathbf{V}_{\rho-k}$ be the last $\rho - k$ right singular vectors of \mathbf{A} and let $\tilde{\mathbf{V}}_k$ be the k orthonormal adaptively sampled vectors, then

$$\left\| \mathbf{V}_{\rho-k}^* \tilde{\mathbf{V}}_k \right\|_F \leq \left(\sum_{i=1}^k \sin \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \right)^{1/2} \quad (17)$$

Now, we will come to our main results.

Lemma 9. Let \mathbf{A} have singular values $\sigma_1 \geq \dots \geq \sigma_{m \vee n}$ and $\tilde{\mathbf{A}}_k$ be the rank- k approximation from Algorithm 1 with oversampling parameter p . Let $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\mathbf{X}) = \text{orth}(\mathbf{A}[\tilde{\mathbf{v}}_1 \ \dots \ \tilde{\mathbf{v}}_k])$. Then we have the following bounds ,

$$\left\| \mathbf{A} - \tilde{\mathbf{Q}} \tilde{\mathbf{Q}}^* \mathbf{A} \right\|_F \leq \left(\left\| \sin \mathbf{\Theta}(\mathbf{U}_k, \tilde{\mathbf{U}}) \mathbf{\Sigma}_k \right\|_F^2 - (k-1) \|\mathbf{\Sigma}_k\|_F^2 \right)^{1/2} \quad (18)$$

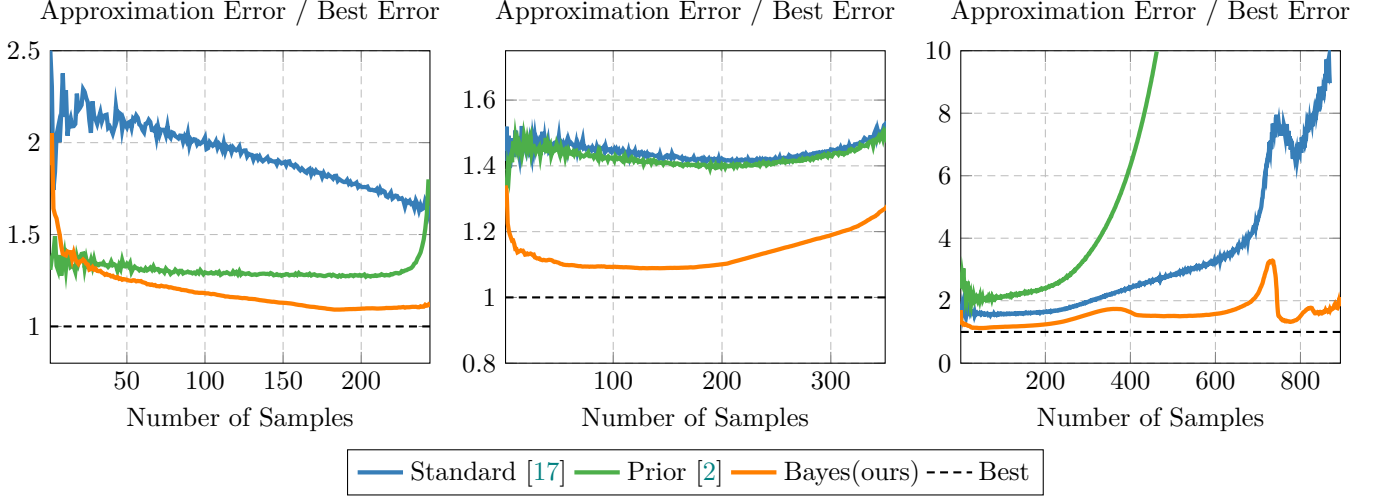


Figure 1: Low Rank Approximation for the Inverse Differential Operator given in Equation (31) (Left), Differential Operator Matrix Poisson2D [10] (Center), and Differential Operator Matrix DK01R [10] (Right). The experiment on the left is from [2, Figure 2].

and

$$\begin{aligned} \left\| \mathbf{A} - \tilde{\mathbf{Q}}\tilde{\mathbf{Q}}^* \mathbf{A} \right\|_{\text{F}} &\leq \left(\sum_{i=1}^k \sigma_k^2 \left(\sum_{j=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) - (k-1) \right) \right)^{1/2} + \left(\sum_{i=k+1}^{\rho} \sigma_i^2 \sum_{j=1}^k \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) \right)^{1/2} \\ &\quad + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \end{aligned} \quad (19)$$

Remark 10. If we have a completely flat spectrum, then for any algorithm which produces orthonormal vectors of $\tilde{\mathbf{V}}$ will give an optimal approximation.

Now we are interested in the sines of the angles between the sampled vectors and the right singular vectors.

4.4 Expected Bounds

Now we will analyze the expected approximation error bounds. First, we want to know the effect of our choices of adaptively sampled vectors in $\tilde{\mathbf{V}}$ on the orthonormal basis \mathbf{U} .

Lemma 11. Let p be our predetermined oversampling parameter, then let $\mathbf{\Omega} \in \mathbb{R}^{n \times p}$ be a standard gaussian matrix. Define $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\mathbf{\Omega})$, let $\tilde{\mathbf{v}}_1$ be the first singular vector of the approximation $\mathbf{Q}\mathbf{Q}^* \mathbf{A}$, then

$$\sin \Theta(\mathbf{u}_1, \tilde{\mathbf{u}}_1) = \left(1 + \left(\frac{\sigma_1^2 \cos^2 \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1)}{\sum_{i=2}^n \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1)} \right) \right)^{-1/2} \leq (1 + \cot^2 \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1))^{-1/2} \quad (20)$$

More generally,

$$\sin \Theta(\mathbf{u}_j, \tilde{\mathbf{u}}_1) = \left(1 + \left(\frac{\sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_1)}{\sum_{i=1}^n \mathbb{I}_{i \neq j} \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1)} \right) \right)^{-1/2} \leq \left(1 + \left(\frac{\sigma_j}{\sigma_1} \cot \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_1) \right)^2 \right)^{-1/2} \quad (21)$$

The proof is deferred to § Appendix A.4. Note the upper bound in Lemma 11 only sharp when the spectrum is completely flat from σ_1 to σ_j , so if there exists spectral decay then this bound tends to be very loose and it is better to look at the equality, despite it being more complicated. We now have an upper bound that

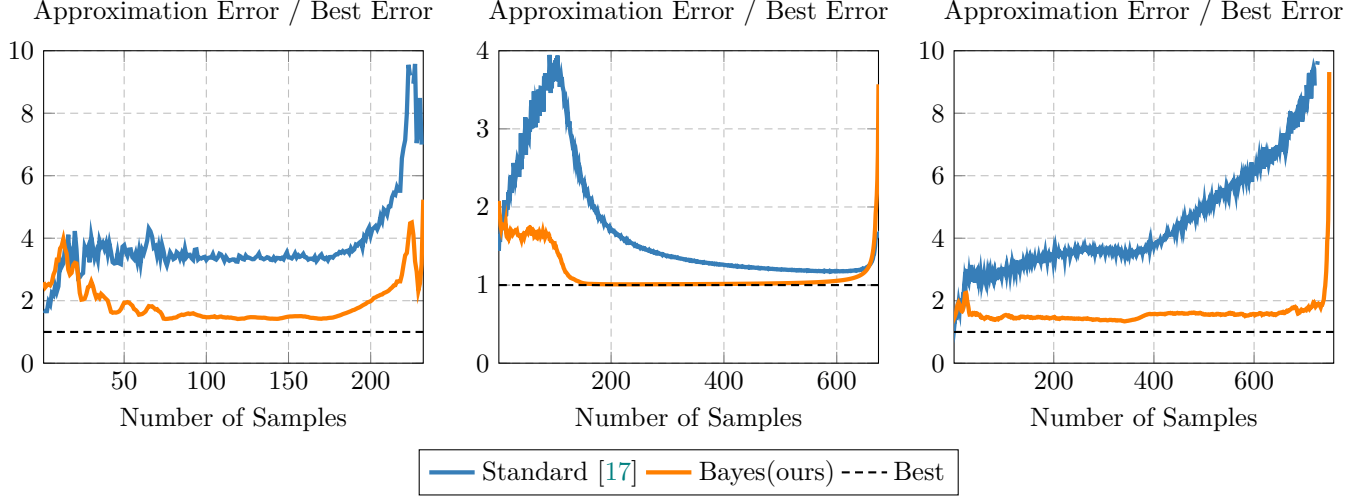


Figure 2: Low Rank Approximation for a matrix for a Computational Fluid Dynamics Problem, **saylr1** (Left) from [10]. Subsequent 2D/3D Problem **fs-680-2** (Center) from [10]. Astrophysics 2D/3D Problem **msfe** (Right) from [10].

depends upon the spectral gap for the dominant left singular vector. This results motivates our analysis in Lemma 12. This bound is dependent on both the spectral gap and the spectral decay. When we have a flat spectrum, the denominator is minimized and our approximation of the left singular vector will be at it's worst. Now we will extend this lemma to the the j -th right singular vector.

Lemma 12. *Consider the same setup in Lemma 11, with $j - 1$ vectors sampled as described in Algorithm 1, then*

$$\sin \Theta(\mathbf{u}_j, \tilde{\mathbf{u}}_k) \leq \left(1 + \left(\frac{\sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_k)}{\sum_{j=1}^{\rho} \sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i) + 2\sigma_i(i-1) \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \sum_{j=1}^{\rho} \sigma_j \cos \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i)} \right) \right)^{-1/2} \quad (22)$$

The proof is deferred to § Appendix A.5. Here we see we lose a little accuracy from Lemma 11. Note in Lemmas 11 and 12, as $\cos \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_k) \rightarrow 1$, then $\sin \Theta(\mathbf{u}_j, \tilde{\mathbf{u}}_k) \rightarrow 0$. Now we will introduce the most important theorem of our work. From the result in Lemma 12, we actually see our results will be best when there is a flatter spectrum. From Lemma 11, we now have an idea on the effect of sampling the eigenvector approximations.

Lemma 13. *Let p be our predetermined oversampling parameter, then let $\mathbf{\Omega} \in \mathbb{R}^{n \times p}$ be a standard gaussian matrix. Define $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\mathbf{\Omega})$, let $\tilde{\mathbf{v}}_1$ be the first singular vector of the approximation $\mathbf{Q}\mathbf{Q}^*\mathbf{A}$, then*

$$\mathbb{E} \sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \left(\frac{\sigma_2}{\sigma_1} \right) \left(1 + \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \right) \quad (23)$$

The proof is deferred to § Appendix A.6.

Lemma 14. *Consider the same setup in Lemma 13, with $j - 1$ vectors sampled as described in Algorithm 1, then*

$$\sin \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_j) \leq \Xi \quad (24)$$

We will connect this together with the error bounds of sampling k right singular vector approximations. With Lemma 9 and Lemma 11, we have the following theorem.

Theorem 15. Let $\mathbf{A} \in \mathbb{R}^{m \times n}$, $k \geq 2$, oversampling parameter $p \geq 2$, where $k + p \leq m \wedge n$. Let $\tilde{\mathbf{A}}_k$ be the matrix returned by Algorithm 1. Then,

$$\mathbb{E} \left\| \mathbf{A} - \tilde{\mathbf{A}}_k \right\|_F \leq \sum_{i=1}^n \sigma_i^2 \left(\sum_{j=1}^k \left(1 + \left(\frac{\sigma_1^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_k)}{\sum_{i=1}^n \mathbb{I}_{i \neq j} \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_k)} \right) \right)^{-1/2} - (k-1) \right) \quad (25)$$

Proof. This theorem follows from the induction formed in Lemma 11 and Lemma 12 and plugging this in to Lemma 9. \blacksquare

4.5 Probabilistic Bounds

Lemma 16. Let p be our predetermined oversampling parameter, then let $\mathbf{\Omega} \in \mathbb{R}^{n \times p}$ be a standard gaussian matrix. Define $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\mathbf{\Omega})$, let $\tilde{\mathbf{v}}_1$ be the first singular vector of the approximation $\mathbf{Q}\mathbf{Q}^* \mathbf{A}$, with Markov's Inequality, we have

$$\mathbb{P} \left\{ \sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \geq t \cdot \left(\frac{\sigma_2}{\sigma_1} \right) \left(1 + \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \right) \right\} \leq \frac{1}{t} \quad (26)$$

$$\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \left(\frac{\sigma_2}{\sigma_1} \right) \left(1 + \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \right) + \sqrt{\frac{1}{2} \ln \left(\frac{2}{\delta} \right)} \quad (27)$$

Proof. Apply Hoeffding's Inequality [21] to Lemma 13. Note since we consider all principal angles in range $[0, \pi/2]$, it follows $0 \leq \sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq 1$. Then we have with probability $(1 - \delta)$,

$$\mathbb{P} \{ |\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) - \mathbb{E}[\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1)]| \geq t \} \leq 2e^{-2t^2} \quad (28)$$

from here we find for $\delta \in (0, 1)$, with probability $1 - \delta$

$$\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \left(\frac{\sigma_2}{\sigma_1} \right) \left(1 + \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \right) + \sqrt{\frac{1}{2} \ln \left(\frac{2}{\delta} \right)} \quad (29)$$

This concludes the proof. \blacksquare

Corollary 17. Let $\mathbf{A} \in \mathbb{R}^{m \times n}$, $k \geq 2$, oversampling parameter $p \geq 2$, where $k + p \leq m \wedge n$. Let $\tilde{\mathbf{A}}_k$ be the matrix returned by Algorithm 1. Then,

$$\left\| \mathbf{A} - \tilde{\mathbf{A}}_k \right\|_F \leq \quad (30)$$

Proof. This follows from a simple upper bound on the multiplicative term on σ_{k+1} in Lemma 9 by using the fact $\sin \Theta(\mathbf{x}, \mathbf{y}) \leq 1$ for all \mathbf{x}, \mathbf{y} . \blacksquare

5 Numerical Experiments

In this section we will test various Synthetic Matrices, Differential Operators, Images, and real-world applications, with our framework compared to fixed covariance matrices. In our first experiment we attempt to learn the discretized 250×250 matrix of the inverse of the following differential operator:

$$\mathcal{L}u = \frac{\partial^2 u}{\partial x^2} - 100 \sin(5\pi x) u, \quad x \in [0, 1] \quad (31)$$

Learning the inverse operator of a PDE is equivalent to learning the Green's Function of a PDE. This has been theoretically proven for certain classes of PDEs (Linear Parabolic [5, 3]) as the inverse Differential operator is compact and there are nice theoretical properties, such as data efficiency.

In Figure 1 (*Right*), note if the Covariance Matrix has eigenvectors orthogonal to the left singular vectors of \mathbf{A} , then the randomized SVD will not perform well. Furthermore, in Figure 1 (*Left*), we can note even without knowledge of the Green’s Function, our method achieves lower error than with the Prior Covariance. We also test our algorithm against various Sparse Matrices in the Texas A& M Sparse Matrix Suite, [10]. In Figure 2 (*Left*), we choose a fluid dynamics problem due to its relevance in low-rank approximation [7]. The synthetic matrix is developed in the following scheme:

$$\mathbf{A} = \sum_{i=1}^{\rho} \frac{100i^{\ell}}{n} \mathbf{U}_{(:,i)} \mathbf{V}_{(i,:)}^*, \quad \mathbf{U} \in \mathbb{O}_{m,k}, \mathbf{V} \in \mathbb{O}_{n,k} \quad (32)$$

We find our theoretical bounds stronger than Theorem 5.

6 Conclusions

We have theoretically and empirically analyzed a novel Covariance Update to iteratively construct the sampling matrix, $\mathbf{\Omega}$ in the Randomized SVD algorithm. Our covariance update for generating sampling vectors and functions can find use various PDE learning applications, [4, 7]. Numerical Experiments indicate without prior knowledge of the matrix, we are able to obtain superior performance to the Randomized SVD and generalized Randomized SVD with covariance matrix utilizing prior information of the PDE. Theoretically, we provide an analysis of our update extended to k -steps and show in expectation, under certain singular value decay conditions, we obtain better performance expectation.

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A Deferred Proofs of Main Results

In this section we give proofs for results we deferred from the main text.

A.1 Proof of Lemma 7

We have \mathbf{Q} is an orthonormal basis of $\mathbf{Y} \in \mathbb{C}^{n \times k}$. Then let us denote $\widehat{\mathbf{Q}} \triangleq \text{orth}([\mathbf{Y} \quad \mathbf{A}\mathbf{v}])$ where $\mathbf{v} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ and $\widetilde{\mathbf{Q}} \triangleq \text{orth}([\mathbf{Y} \quad \mathbf{A}\widetilde{\mathbf{v}}])$ where $\widetilde{\mathbf{v}}$ is the k th right singular vector of $\mathbf{Q}\mathbf{Q}^*\mathbf{A}$. Note $\widehat{\mathbf{q}} \in \text{Span}(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*\mathbf{A})$.

$$\left\| \mathbf{A} - \widehat{\mathbf{Q}}\widehat{\mathbf{Q}}^*\mathbf{A} \right\|_{\text{F}}^2 = \left\| \mathbf{A} - \mathbf{Q}\mathbf{Q}^*\mathbf{A} - \widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} \right\| \quad (33)$$

$$= \text{Tr} \left(\mathbf{A}^*\mathbf{A} - \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} - \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} - \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} + \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} \right. \\ \left. + \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} - \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} + \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} + \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} \right) \quad (34)$$

$$= \text{Tr} \left(\mathbf{A}^*\mathbf{A} - \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} - \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} - \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} + \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} \right. \\ \left. - \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} + \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A} \right) \quad (35)$$

$$= \text{Tr} (\mathbf{A}^*\mathbf{A} - \mathbf{A}^*\mathbf{Q}\mathbf{Q}^*\mathbf{A} - \mathbf{A}^*\widehat{\mathbf{q}}\widehat{\mathbf{q}}^*\mathbf{A}) \quad (36)$$

$$= \left\| \mathbf{A} \right\|_{\text{F}}^2 - \left\| \mathbf{Q}^*\mathbf{A} \right\|_{\text{F}}^2 - \underbrace{\left\| \widehat{\mathbf{q}}^*\mathbf{A} \right\|_{\text{F}}^2}_{\text{C}_1} \quad (37)$$

$$= \left\| \mathbf{A} - \mathbf{Q}\mathbf{Q}^*\mathbf{A} \right\|_{\text{F}}^2 - \left\| \widehat{\mathbf{q}}^*\mathbf{A} \right\|_{\text{F}}^2 \quad (38)$$

Similarly from Equation (37), we have

$$\left\| \mathbf{A} - \widetilde{\mathbf{Q}}\widetilde{\mathbf{Q}}^*\mathbf{A} \right\|_{\text{F}}^2 = \left\| \mathbf{A} \right\|_{\text{F}}^2 - \left\| \mathbf{Q}^*\mathbf{A} \right\|_{\text{F}}^2 - \underbrace{\left\| \widetilde{\mathbf{q}}^*\mathbf{A} \right\|_{\text{F}}^2}_{\text{C}_2} \quad (39)$$

Since we have $\widehat{\mathbf{q}}, \widetilde{\mathbf{q}} \in \text{Span}(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*\mathbf{A})$, then from our formulation in Equations (37) and (39), we want to our sampled vector to be in the dominant singular space of the span of the singular vectors of $\mathbf{I} - \mathbf{Q}\mathbf{Q}^*\mathbf{A}$.

$$\mathbf{q}_{\text{OPT}} = \max_{\mathbf{q} \in \text{Span}(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) : \|\mathbf{q}\|=1} \left\| \mathbf{q}^*\mathbf{A} \right\| \quad (40)$$

Let us note for any column $\mathbf{q} \in \mathbf{Q}$, we have

$$((\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{v})^*\mathbf{q} = (\mathbf{v}^* - \mathbf{v}^*\mathbf{Q}\mathbf{Q}^*)\mathbf{q} = \mathbf{v}^*\mathbf{q} - \mathbf{v}^*\mathbf{q} = 0 \quad (41)$$

Claim (ii):

Let us define $\widetilde{\mathbf{U}}\widetilde{\Sigma}\widetilde{\mathbf{V}}^* = \text{SVD}(\mathbf{Q}\mathbf{Q}^*\mathbf{A})$.

$$\text{C}_1 = \left\| \widehat{\mathbf{q}}^*\mathbf{A} \right\|_{\text{F}}^2 = \left\| \widehat{\mathbf{q}}^*\mathbf{U}\Sigma\mathbf{V}^* \right\|_2^2 = \left\| \widehat{\mathbf{q}}^*\mathbf{U}\Sigma \right\|_2^2 \stackrel{\text{eqn. (41)}}{=} \frac{\left\| ((\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A}\widehat{\mathbf{v}})^*\mathbf{U}\Sigma \right\|_2^2}{\left\| ((\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A}\widehat{\mathbf{v}}) \right\|_2^2} \quad (42)$$

$$= \frac{\left\| \widehat{\mathbf{v}}^*\mathbf{A}^*(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A} \right\|_2^2}{\left\| ((\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A}\widehat{\mathbf{v}}) \right\|_2^2} \geq \frac{\left\| \widehat{\mathbf{v}}^*\mathbf{A}^*(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A} \right\|_2^2}{\left\| (\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A} \right\|_2^2 \left\| \widehat{\mathbf{v}} \right\|_2^2} \quad (43)$$

For shorthand, let $\widetilde{\mathbf{A}} \triangleq \mathbf{A}^*(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A}$. Note, we have $\left\| (\mathbf{I} - \mathbf{Q}\mathbf{Q}^*)\mathbf{A} \right\|$ is given. Let us note since $\widehat{\mathbf{v}} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, we have

$$\mathbb{E}[\widehat{\mathbf{x}}] = \widetilde{\mathbf{A}}\mathbb{E}[\widehat{\mathbf{v}}] = \mathbf{0} \quad (44)$$

$$\mathbb{V}(\widehat{\mathbf{x}}) = \widetilde{\mathbf{A}}\mathbb{E}[\widehat{\mathbf{v}}\widehat{\mathbf{v}}^\top]\widetilde{\mathbf{A}}^* = \widetilde{\mathbf{A}}\widetilde{\mathbf{A}}^* \quad (45)$$

We will then upper bound $\|\hat{\mathbf{x}}\|$ when $\hat{\mathbf{x}} \sim \mathcal{N}(\mathbf{0}, \tilde{\mathbf{A}}\tilde{\mathbf{A}}^*)$. It then follows $\hat{\mathbf{x}} = \tilde{\mathbf{A}}\mathbf{x}$ where $\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$. From Lemma 21, we have

$$\mathbb{P}\left\{\left|\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}} \mathbf{x} - \mathbb{E}\left[\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}} \mathbf{x}\right]\right| \geq t\right\} \leq 2 \exp\left(-c \frac{t^2}{\left\|\tilde{\mathbf{A}}^* \tilde{\mathbf{A}}\right\|_{\text{F}}^2}\right) \quad (46)$$

It then follows with probability $1 - \delta$

$$\left\|\tilde{\mathbf{A}}\mathbf{x}\right\|^2 \geq \mathbb{E}\left[\left\|\tilde{\mathbf{A}}\mathbf{x}\right\|^2\right] - \left\|\tilde{\mathbf{A}}\right\|_{\text{F}}^2 \sqrt{\frac{1}{c} \log\left(\frac{2}{\delta}\right)} \quad (47)$$

We further simplify this with the fact

$$\mathbb{E}\left\|\tilde{\mathbf{A}}\mathbf{x}\right\|^2 = \mathbb{E}\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}} \mathbf{x} = \mathbb{E}\text{Tr}\left(\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}} \mathbf{x}\right) = \mathbb{E}\text{Tr}\left(\mathbf{x}\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}}\right) = \text{Tr}\left(\mathbb{E}\left[\mathbf{x}\mathbf{x}^* \tilde{\mathbf{A}}^* \tilde{\mathbf{A}}\right]\right) \quad (48)$$

$$= \text{Tr}\left(\mathbb{E}\left[\mathbf{x}\mathbf{x}^*\right] \tilde{\mathbf{A}}^* \tilde{\mathbf{A}}\right) = \text{Tr}\left(\tilde{\mathbf{A}}^* \tilde{\mathbf{A}}\right) = \left\|\tilde{\mathbf{A}}\right\|_{\text{F}}^2 \quad (49)$$

We will also note the following fact for a general matrix \mathbf{M} and an orthonormal rectangular matrix \mathbf{X} ,

$$\left\|(\mathbf{I} - \mathbf{X}\mathbf{X}^*)\mathbf{M}\right\|^2 = \left\|\mathbf{M}^*(\mathbf{I} - \mathbf{X}\mathbf{X}^*)(\mathbf{I} - \mathbf{X}\mathbf{X}^*)\mathbf{M}\right\| = \left\|\mathbf{M}^*(\mathbf{I} - \mathbf{X}\mathbf{X}^*)\mathbf{M}\right\| \quad (50)$$

Now we can calculate the expectation. Assume $\|\mathbf{v}\|_2^2 \leq K^2$, for some $K \in \mathbb{R}_+$, then

$$\mathbb{E}C_1 \geq \frac{\mathbb{E}\left[\left\|\hat{\mathbf{v}}^* \mathbf{A}^* (\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2\right]}{\left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 K^2} = \frac{\left\|\mathbf{A}^* (\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_{\text{F}}^2}{\left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 K^2} \quad (51)$$

$$\geq \frac{\left\|\mathbf{A}^* (\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2}{\left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 K^2} \stackrel{\text{eqn. (50)}}{=} \frac{\left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^4}{\left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 K^2} \quad (52)$$

$$= K^{-2} \left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 \quad (53)$$

where $K = \mathcal{O}(n^2)$. We thus have with probability $1 - \delta$

$$C_1 \geq \left\|(\mathbf{I} - \mathbf{Q}\mathbf{Q}^*) \mathbf{A}\right\|_2^2 \left(\mathcal{O}\left(\frac{1}{n^2}\right) \left(1 - \sqrt{\frac{1}{c} \log\left(\frac{2}{\delta}\right)}\right)\right) \quad (54)$$

This completes our proof of Claim (ii). ■

A.2 Proof of Lemma 8

Let $\tilde{\mathbf{V}}_k$ be the orthonormal basis for the adaptively sampled vectors.

$$\zeta_1 \triangleq \left\|\mathbf{V}_{\rho-k}^* \tilde{\mathbf{V}}\right\|_{\text{F}} = \left(\sum_{i=1}^k \sum_{j=k+1}^{\rho} \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2\right)^{1/2} = \left(\sum_{i=1}^k \sum_{j=1}^n \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2 - \sum_{i=1}^k \sum_{j=1}^k \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2\right)^{1/2} \quad (55)$$

$$= \left(\sum_{i=1}^k \sum_{j=1}^n \mathbb{I}_{i \neq j} \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2 - \sum_{i=1}^k \sum_{j=1}^k \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2 + \sum_{i=1}^k \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2\right)^{1/2} \quad (56)$$

$$= \left(\left(\sum_{i=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) + \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i)\right) - \sum_{i=1}^k \sum_{j=1}^k \left\|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\right\|^2\right)^{1/2} \quad (57)$$

$$= \left(k - \sum_{i=1}^k \sum_{j=1}^k \|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\|^2 \right)^{1/2} = \left(k - \sum_{i=1}^k \|\tilde{\mathbf{v}}_i^* \mathbf{v}_i\|^2 - \underbrace{\sum_{i=1}^k \sum_{j=1}^k \mathbb{I}\{i \neq j\} \|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\|^2}_{\beta} \right)^{1/2} \quad (58)$$

$$= \left(\sum_{i=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) - \underbrace{\sum_{i=1}^k \sum_{j=1}^k \mathbb{I}\{i \neq j\} \|\tilde{\mathbf{v}}_i^* \mathbf{v}_j\|^2}_{\beta} \right)^{1/2} \leq \left(\sum_{i=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \right)^{1/2} \quad (59)$$

Here we have the desired result. ■

In Equation (59), when we use the approximation below, we lose the β term, is there any way to lower bound it?

A.3 Proof of Lemma 9

Proof. We will utilize the fact that when working with projection we do not have to account for the orthogonalization of the sampled vectors from the previous eigenvector approximations. First, we define $\tilde{\mathbf{U}} \triangleq \text{orth}(\mathbf{A}\tilde{\mathbf{V}})$.

$$\left\| \mathbf{A} - \tilde{\mathbf{A}}_k \right\|_{\text{F}} \triangleq \underbrace{\left\| \mathbf{A} - \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} \right\|_{\text{F}}}_{\xi_1} \quad (60)$$

We will upper bound the square.

$$\left\| \mathbf{A} - \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} \right\|_{\text{F}}^2 = \text{Tr} \left(\mathbf{A}^* \mathbf{A} - \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} - \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} + \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} \right) \quad (61)$$

$$= \text{Tr} \left(\mathbf{A}^* \mathbf{A} - \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} - \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} + \mathbf{A}^* \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} \right) \quad (62)$$

$$= \left\| \mathbf{A} \right\|_{\text{F}}^2 - \left\| \tilde{\mathbf{U}}^* \mathbf{A} \right\|_{\text{F}}^2 = \left\| \boldsymbol{\Sigma} \right\|_{\text{F}}^2 - \text{Tr} \left(\tilde{\mathbf{U}}^* \mathbf{U} \boldsymbol{\Sigma} \mathbf{V}^* \mathbf{V} \boldsymbol{\Sigma} \mathbf{U}^* \tilde{\mathbf{U}} \right) \quad (63)$$

$$= \left\| \boldsymbol{\Sigma} \right\|_{\text{F}}^2 - \left\| \tilde{\mathbf{U}}^* \mathbf{U} \boldsymbol{\Sigma} \right\|_{\text{F}}^2 = \sum_{i=1}^n \sigma_i^2 - \sum_{i=1}^n \sigma_i^2 \sum_{j=1}^k \cos^2 \Theta(\mathbf{u}_i, \tilde{\mathbf{u}}_j) \quad (64)$$

$$= \sum_{i=1}^n \sigma_i^2 \left(1 - \sum_{j=1}^k \cos^2 \Theta(\mathbf{u}_i, \tilde{\mathbf{u}}_j) \right) = \sum_{i=1}^n \sigma_i^2 \left(1 - \left\| \cos \boldsymbol{\Theta}(\mathbf{U}_{(i)}, \tilde{\mathbf{U}}_k) \right\|_{\text{F}}^2 \right) \quad (65)$$

$$\stackrel{\text{lem. 19}}{=} \sum_{i=1}^n \sigma_i^2 \left(\left\| \sin \boldsymbol{\Theta}(\mathbf{U}_{(i)}, \tilde{\mathbf{U}}_k) \right\|_{\text{F}}^2 - \frac{k-1}{2} \right) \quad (66)$$

$$= \sum_{i=1}^n \frac{\sigma_i^2}{2} \left(\underbrace{\left\| \tilde{\mathbf{U}}_k \tilde{\mathbf{U}}_k^* - \mathbf{U}_{(i)} \mathbf{U}_{(i)}^* \right\|_{\text{F}}^2}_{\xi_1} - (k-1) \right) \quad (67)$$

We will also prove a similar result which is more interpretable.

$$\left\| \mathbf{A} - \tilde{\mathbf{U}}\tilde{\mathbf{U}}^* \mathbf{A} \right\|_{\text{F}} = \left\| \mathbf{A} - \mathbf{A} \tilde{\mathbf{V}} (\mathbf{A} \tilde{\mathbf{V}})^{\dagger} \mathbf{A} \right\|_{\text{F}} \stackrel{??}{\leq} \left\| \mathbf{A} - \mathbf{A} \tilde{\mathbf{V}} \mathbf{V}_k^* \right\|_{\text{F}} \quad (68)$$

$$= \left\| \mathbf{A}_k (\mathbf{I} - \tilde{\mathbf{V}} \mathbf{V}_k^*) \right\|_{\text{F}} + \left\| \mathbf{A}_{\rho-k} \tilde{\mathbf{V}} \mathbf{V}_k^* \right\|_{\text{F}} + \left\| \mathbf{A}_{\rho-k} \right\|_{\text{F}} \quad (69)$$

$$\leq \left\| \boldsymbol{\Sigma}_k (\mathbf{I} - \tilde{\mathbf{V}}^* \mathbf{V}_k) \right\|_{\text{F}} + \left\| \boldsymbol{\Sigma}_{\rho-k} \mathbf{V}_{\rho-k}^* \tilde{\mathbf{V}} \right\|_{\text{F}} + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \quad (70)$$

$$= \left\| \boldsymbol{\Sigma}_k (\mathbf{I} - \tilde{\mathbf{V}}^* \mathbf{V}_k) \right\|_{\text{F}} + \left(\sum_{i=k+1}^{\rho} \sigma_i^2 \sum_{j=1}^k \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) \right)^{1/2} + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \quad (71)$$

$$= \left(\sum_{i=1}^k \sigma_k^2 \left(1 - \sum_{j=1}^k \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) \right) \right)^{1/2} + \left(\sum_{i=k+1}^{\rho} \sigma_i^2 \sum_{j=1}^k \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) \right)^{1/2} + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \quad (72)$$

$$\stackrel{\text{lem. 8}}{\leq} \left(\sum_{i=1}^k \sigma_k^2 \left(1 - \sum_{j=1}^k \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) \right) \right)^{1/2} + \sigma_{k+1} \left(\sum_{i=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \right) + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \quad (73)$$

$$= \left(\sum_{i=1}^k \sigma_k^2 \left(\sum_{j=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_j) - (k-1) \right) \right)^{1/2} + \sigma_{k+1} \left(\sum_{i=1}^k \sin^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \right) + \left(\sum_{i=1}^{\rho} \sigma_i^2 \right)^{1/2} \quad (74)$$

The last line gives us a formulation so it is sufficient to only upper bound the sines. This completes the proof. \blacksquare

A.4 Proof of Lemma 11

Proof. Let $\hat{\mathbf{A}}_1$ be the approximation obtained by $(\mathbf{Q}\mathbf{Q}^*\mathbf{A})_1$ where \mathbf{Q} is obtained by orth $(\mathbf{A}\mathbf{\Omega})$ and $\mathbf{\Omega} \in \mathbb{R}^{n \times p}$ is a standard gaussian matrix. It is important all our sine bounds are less than 1 or they will not be useful. We will first convert this bound into a bound w.r.t. \mathbf{v}_1 and $\tilde{\mathbf{v}}_1$.

$$\sin^2 \Theta(\mathbf{u}_1, \tilde{\mathbf{u}}_1) \stackrel{\text{eqn. (6)}}{=} \frac{\|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2}{\|\mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2} = \frac{\|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2}{\|\mathbf{U}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2} = \frac{\|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2}{\|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2 + \|\mathbf{U}_1^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}}^2} \quad (75)$$

$$= \left(1 + \left(\frac{\|\mathbf{U}_1^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}} \mu_1}{\|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\|_{\text{F}} \mu_2} \right)^2 \right)^{-1} \quad (76)$$

We now lower bound μ_1 and upper bound μ_2 .

$$\mu_1 \triangleq \|\mathbf{U}_1^* \mathbf{A} \tilde{\mathbf{v}}_1\| = \|\mathbf{\Sigma}_1 \mathbf{V}_1^* \tilde{\mathbf{v}}_1\| = \sigma_1 \cos \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \quad (77)$$

$$\mu_2 \triangleq \|\mathbf{U}_{\perp,1}^* \mathbf{A} \tilde{\mathbf{v}}_1\| = \|\mathbf{\Sigma}_{\perp,1} \mathbf{V}_{\perp,1}^* \tilde{\mathbf{v}}_1\| = \left(\sum_{i=2}^n \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1) \right)^{1/2} \quad (78)$$

The denominator in equation 75 is a result of the Matrix Pythagoras theorem.

Plugging this back into Equation (75), we have

$$\sin \Theta(\mathbf{u}_1, \tilde{\mathbf{u}}_1) = \left(1 + \left(\frac{\sigma_1^2 \cos^2 \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1)}{\sum_{i=2}^n \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1)} \right) \right)^{-1/2} \quad (79)$$

More generally,

$$\sin \Theta(\mathbf{u}_j, \tilde{\mathbf{u}}_1) = \left(1 + \left(\frac{\sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_1)}{\sum_{i=1}^n \mathbb{I}\{i \neq j\} \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1)} \right) \right)^{-1/2} \quad (80)$$

We now have a relation dependent only on the angles between the adaptively sampled vectors and the right singular vectors. This completes the proof. \blacksquare

A.5 Proof of Lemma 12

Proof. We must now handle the orthogonalization of $\tilde{\mathbf{U}}$.

$$\sin^2 \Theta(\mathbf{u}_i, \tilde{\mathbf{u}}_i) \stackrel{\text{eqn. (6)}}{=} \frac{\|\mathbf{U}_{\perp,i}^* \text{orth}(\mathbf{A} \tilde{\mathbf{v}}_i)\|_{\text{F}}^2}{\|\mathbf{U}_{\perp,i}^* \text{orth}(\mathbf{A} \tilde{\mathbf{v}}_i)\|_{\text{F}}^2 + \|\mathbf{U}_{(i)}^* \text{orth}(\mathbf{A} \tilde{\mathbf{v}}_i)\|_{\text{F}}^2} \quad (81)$$

$$= \left(1 + \frac{\left\| \mathbf{U}_{(i)}^* \text{orth}(\mathbf{A}\tilde{\mathbf{v}}_i) \right\|_{\text{F}}^2}{\left\| \mathbf{U}_{\perp,i}^* \text{orth}(\mathbf{A}\tilde{\mathbf{v}}_i) \right\|_{\text{F}}^2} \mu_1 \right)^{-1} \quad (82)$$

First, we will lower bound μ_1 .

$$\sqrt{\mu_1} \triangleq \left\| \mathbf{U}_{(i)}^* \text{orth}(\mathbf{A}\tilde{\mathbf{v}}_i) \right\|_{\text{F}} = \left\| \mathbf{U}_{(i)}^* \left(\mathbf{A}\tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \tilde{\mathbf{u}}_j \right) \right\|_{\text{F}} \quad (83)$$

$$= \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* \tilde{\mathbf{u}}_j}{\|\tilde{\mathbf{u}}_j\|^2} (\mathbf{U}_{(i)}^* \tilde{\mathbf{u}}_j) \right\|_{\text{F}} \quad (84)$$

$$\geq \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i \right\|_{\text{F}} - \sum_{j=1}^{i-1} \left\| \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* \tilde{\mathbf{u}}_j}{\|\tilde{\mathbf{u}}_j\|^2} (\mathbf{U}_{(i)}^* \tilde{\mathbf{u}}_j) \right\|_{\text{F}} \quad (85)$$

$$\geq \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i \right\|_{\text{F}} - \sum_{j=1}^{i-1} \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}} \quad (86)$$

Now let us consider the square.

$$\mu_1 \geq \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i \right\|_{\text{F}}^2 - 2(i-1) \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i \right\|_{\text{F}} \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}} + (i-1)^2 \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}}^2 \quad (87)$$

$$= \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) - 2(i-1) \sigma_i \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}} + (i-1)^2 \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}}^2 \quad (88)$$

$$\stackrel{(a)}{\geq} \sigma_i^2 \cos^2 \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \quad (89)$$

(a) follows if $(i-1) \|\mathbf{A}\tilde{\mathbf{v}}_i\| \geq 2\sigma_i \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i)$, which is true. Now we will analyze μ_2 .

$$\mu_2 \triangleq \left\| \mathbf{U}_{\perp,i}^* \text{orth}(\mathbf{A}\tilde{\mathbf{v}}_i) \right\|_{\text{F}}^2 \quad (90)$$

$$= \left\| \mathbf{U}^* \left(\mathbf{A}\tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \tilde{\mathbf{u}}_j \right) \right\|_{\text{F}}^2 - \left\| \mathbf{U}_{(i)}^* \left(\mathbf{A}\tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \tilde{\mathbf{u}}_j \right) \right\|_{\text{F}}^2 \quad (91)$$

$$= \left\| \mathbf{A}\tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \tilde{\mathbf{u}}_j \right\|_{\text{F}}^2 - \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i - \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \mathbf{U}_{(i)}^* \tilde{\mathbf{u}}_j \right\|_{\text{F}}^2 \quad (92)$$

$$\stackrel{(a)}{\leq} \|\mathbf{A}\tilde{\mathbf{v}}_i\|_{\text{F}}^2 + 2 \left\| \boldsymbol{\Sigma}_{(i)} \mathbf{V}_{(i)}^* \tilde{\mathbf{v}}_i \right\|_{\text{F}} \left\| \sum_{j=1}^{i-1} \frac{(\mathbf{A}\tilde{\mathbf{v}}_i)^* (\tilde{\mathbf{u}}_j)}{\|\tilde{\mathbf{u}}_j\|^2} \mathbf{U}_{(i)}^* \tilde{\mathbf{u}}_j \right\|_{\text{F}} \quad (93)$$

$$= \sum_{j=1}^{\rho} \sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i) + 2\sigma_i(i-1) \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \sum_{j=1}^{\rho} \sigma_j \cos \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i) \quad (94)$$

In our analysis of μ_2 , we make the realization in (a) the Gram-Schmidt Orthogonalization [36], the norm of the orthogonalized vector will be less than the original vector due to the subtraction of the projections and for all j we have $\|\tilde{\mathbf{v}}_i\| = \|\tilde{\mathbf{v}}_j\| = 1$. We thus have,

$$\sin \Theta(\mathbf{u}_j, \tilde{\mathbf{u}}_k) \leq \left(1 + \left(\frac{\sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_k)}{\sum_{j=1}^{\rho} \sigma_j^2 \cos^2 \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i) + 2\sigma_i(i-1) \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_i) \sum_{j=1}^{\rho} \sigma_j \cos \Theta(\mathbf{v}_j, \tilde{\mathbf{v}}_i)} \right) \right)^{-1/2} \quad (95)$$

This completes the proof. ■

A.6 Proof of Lemma 13

Proof.

$$\sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \stackrel{\text{thm. 3}}{\leq} \frac{\|\mathbf{A}_1 - \llbracket \mathbf{Q}\mathbf{Q}^* \mathbf{A} \rrbracket_1\|}{\sigma_1} \quad (96)$$

$$= \frac{\|\mathbf{A}_1 - \mathbf{Q}\mathbf{Q}^* \mathbf{A}_1 + \mathbf{Q}\mathbf{Q}^* \mathbf{A}_1 - \llbracket \mathbf{Q}\mathbf{Q}^* \mathbf{A} \rrbracket_1\|}{\sigma_1} \quad (97)$$

$$\leq \frac{\|\mathbf{A}_1 - \mathbf{Q}\mathbf{Q}^* \mathbf{A}_1\| + \|\mathbf{Q}\mathbf{Q}^* \mathbf{A}_1 - \llbracket \mathbf{Q}\mathbf{Q}^* \mathbf{A} \rrbracket_1\|}{\sigma_1} \quad (98)$$

$$\leq \frac{\overbrace{\|(\mathbf{I} - \Pi_{\mathbf{Y}}) \mathbf{A}_1\|}^{C_1} + \|\llbracket \mathbf{Q}\mathbf{Q}^* \mathbf{A} \rrbracket_{\perp, 1}\|}{\sigma_1} \quad (99)$$

Now we will upper bound C_1 with standard theory.

$$C_1 = \left\| \mathbf{A}_1 - \mathbf{A}\mathbf{\Omega}(\mathbf{A}\mathbf{\Omega})^\dagger \mathbf{A}_1 \right\| \stackrel{??}{\leq} \left\| \mathbf{A}_1 - \mathbf{A}\mathbf{\Omega}(\mathbf{V}_1^* \mathbf{\Omega})^\dagger \mathbf{V}_1^* \right\| \quad (100)$$

$$= \left\| \mathbf{A}_{\rho-1} \mathbf{\Omega}(\mathbf{V}_1^* \mathbf{\Omega})^\dagger \mathbf{V}_1^* \right\| \quad (101)$$

In expectation,

$$\mathbb{E} \left\| \mathbf{A}_{\rho-1} \mathbf{\Omega}(\mathbf{V}_1^* \mathbf{\Omega})^\dagger \mathbf{V}_1^* \right\| \leq \|\Sigma_{\rho-1}\| \mathbb{E} \left\| \mathbf{V}_{\rho-1}^* \mathbf{\Omega}(\mathbf{V}_1^* \mathbf{\Omega})^\dagger \right\| \quad (102)$$

$$\stackrel{\text{Cauchy-Schwarz}}{\leq} \sigma_2 \left(\mathbb{E} \left\| \mathbf{V}_{\rho-1}^* \mathbf{\Omega} \right\|^2 \mathbb{E} \left\| (\mathbf{V}_1^* \mathbf{\Omega})^\dagger \right\|^2 \right)^{1/2} \quad (103)$$

$$\leq \sigma_2 \left(\mathbb{E} \left\| \mathbf{V}_{\rho-1}^* \mathbf{\Omega} \right\|^2 \left(\frac{1}{p-1} \right) \right)^{1/2} \quad (104)$$

$$\stackrel{??}{\leq} \sigma_2 \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \quad (105)$$

We thus have

$$\mathbb{E} \sin \Theta(\mathbf{v}_1, \tilde{\mathbf{v}}_1) \leq \sigma_2 \left(1 + \left(\frac{\sqrt{n} + \sqrt{p}}{p-1} \right)^{1/2} \right) \quad (106)$$

This completes the proof. ■

A.7 Proof of Lemma 14

Proof. We will use Wedin's Theorem [41]. Here define $\mathbf{Q} \triangleq \text{orth}(\mathbf{A}\tilde{\mathbf{V}})$ where $\tilde{\mathbf{V}}$ contains $i-1$ sampled vectors. Let us first consider $i=2$, we have the $\sin \Theta$ from Lemma 13.

$$\sin \Theta(\mathbf{v}_2, \tilde{\mathbf{v}}_2) \leq \frac{\|\mathbf{A}_{(2)} - \llbracket \mathbf{Q}\mathbf{Q}^* \mathbf{A} \rrbracket_{(1)}\|}{\sigma_2} = \frac{\|\mathbf{A}_{(2)} - \mathbf{Q}\mathbf{Q}^* \mathbf{A}\|}{\sigma_2} = \frac{\left\| \mathbf{A}_{(2)} - \mathbf{A}\tilde{\mathbf{V}}_1 \left(\mathbf{A}\tilde{\mathbf{V}}_1 \right)^\dagger \mathbf{A} \right\|}{\sigma_2} \quad (107)$$

$$\stackrel{??}{\leq} \frac{\left\| \mathbf{A}_{(2)} - \mathbf{A}\tilde{\mathbf{V}}_1 \left(\mathbf{V}_{(2)}^* \tilde{\mathbf{V}}_1 \right)^\dagger \mathbf{V}_{(2)}^* \right\|}{\sigma_2} = \frac{\left\| \mathbf{A}_{\perp, 2} \tilde{\mathbf{V}}_1 \left(\mathbf{V}_{(2)}^* \tilde{\mathbf{V}}_1 \right)^\dagger \mathbf{V}_{(2)}^* \right\|}{\sigma_2} \quad (108)$$

$$= \frac{\sum_{i=1}^p \mathbb{I}\{i \neq 2\} \sigma_i \cos \Theta(\mathbf{v}_i, \tilde{\mathbf{v}}_1)}{\cos \Theta(\mathbf{v}_2, \tilde{\mathbf{v}}_1) \sigma_2} \leq \left(\frac{\sigma_1}{\sigma_2} \right) \tan \Theta(\mathbf{v}_2, \tilde{\mathbf{v}}_1) \quad (109)$$

where $\tilde{\mathbf{v}}_1$ has been determined by the randomized SVD. ■

B Singular Subspace Perturbation Lemmas

Lemma 18. *Let \mathbf{v} and $\tilde{\mathbf{v}}$ be vectors s.t. $\|\mathbf{v}\| = \|\tilde{\mathbf{v}}\| = 1$ and $\mathbf{v}^* \tilde{\mathbf{v}} \geq 0$. Then,*

$$\|\mathbf{v} - \tilde{\mathbf{v}}\| \leq \sqrt{2} \sin \Theta(\mathbf{v}, \tilde{\mathbf{v}}) \quad (110)$$

Proof.

$$\sin^2 \Theta(\mathbf{v}, \tilde{\mathbf{v}}) = 1 - (\mathbf{v}^* \tilde{\mathbf{v}})^2 \stackrel{(a)}{\geq} 1 - \mathbf{v}^* \tilde{\mathbf{v}} = 1 + \frac{1}{2} \|\mathbf{v} - \tilde{\mathbf{v}}\|^2 - \frac{1}{2} \|\mathbf{v}\|^2 - \frac{1}{2} \|\tilde{\mathbf{v}}\|^2 = \frac{1}{2} \|\mathbf{v} - \tilde{\mathbf{v}}\|^2 \quad (111)$$

(a) follows from $0 \leq \mathbf{v}^* \tilde{\mathbf{v}} \leq 1$, therefore $\mathbf{v}^* \tilde{\mathbf{v}} \geq (\mathbf{v}^* \tilde{\mathbf{v}})^2$.

Plugging this back into the first inequality and taking the square root gives us the desired result. \blacksquare

Lemma 19. *Let \mathbf{P} and \mathbf{Q} be projection matrices s.t. $\mathbf{P} \triangleq \mathbf{U}\mathbf{U}^\dagger$ and $\mathbf{Q} \triangleq \mathbf{V}\mathbf{V}^\dagger$, then*

$$\|\cos \Theta(\mathbf{U}, \mathbf{V})\|_{\text{F}}^2 = \frac{\text{rank}(\mathbf{U}) + \text{rank}(\mathbf{V})}{2} - \|\sin \Theta(\mathbf{U}, \mathbf{V})\|_{\text{F}}^2 \quad (112)$$

Proof.

$$\|\sin \Theta(\mathbf{U}, \mathbf{V})\|_{\text{F}}^2 = \frac{1}{2} \|\Pi_{\mathbf{U}} - \Pi_{\mathbf{V}}\|_{\text{F}}^2 = \frac{1}{2} \text{Tr}(\Pi_{\mathbf{U}}^* \Pi_{\mathbf{U}} - \Pi_{\mathbf{U}}^* \Pi_{\mathbf{V}} - \Pi_{\mathbf{V}}^* \Pi_{\mathbf{U}} + \Pi_{\mathbf{V}}^* \Pi_{\mathbf{V}}) \quad (113)$$

$$= \frac{1}{2} (\text{rank}(\mathbf{U}) + \text{rank}(\mathbf{V}) - 2 \text{Tr}(\Pi_{\mathbf{U}}^* \Pi_{\mathbf{V}})) \quad (114)$$

$$= \frac{1}{2} (\text{rank}(\mathbf{U}) + \text{rank}(\mathbf{V})) - \|\cos \Theta(\mathbf{U}, \mathbf{V})\|_{\text{F}}^2 \quad (115)$$

Subtract the sine term from both sides and add the cosine terms to both sides and the proof is complete. \blacksquare

C Concentration Inequalities

Lemma 20 ([39], Exercise 6.3.5). *Let $\mathbf{B} \in \mathbb{C}^{m \times n}$ and $\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, then we have*

$$\mathbb{P}\{\|\mathbf{B}\mathbf{x}\|_2 \geq CK\|\mathbf{B}\|_{\text{F}} + t\} \leq \exp\left(-\frac{ct^2}{K\|\mathbf{B}\|^2}\right) \quad (116)$$

Lemma 21 (Hanson-Wright Inequality). *Let $\mathbf{x} \in \mathbb{R}^n$ be a vector with sub-Gaussian random vector symmetric about $\mathbf{0}$. Let \mathbf{B} be a symmetric $n \times n$ matrix, then $\forall t \geq 0$,*

$$\mathbb{P}\{|\mathbf{x}^\top \mathbf{B} \mathbf{x} - \mathbb{E}[\mathbf{x}^\top \mathbf{B} \mathbf{x}]| \geq t\} \leq 2 \exp\left(-c \min\left\{\frac{t^2}{K^4 \|\mathbf{B}\|_{\text{F}}^2}, \frac{t}{K^2 \|\mathbf{B}\|}\right\}\right) \quad (117)$$

Lemma 22. [29]. *For any matrices \mathbf{A} , \mathbf{B} , and \mathbf{C} , then for any unitarily invariant norm $\|\cdot\|$, it follows*

$$\|\mathbf{ABC}\| \leq \min\{\|\mathbf{A}\| \|\mathbf{B}\|_2 \|\mathbf{C}\|_2, \|\mathbf{A}\|_2 \|\mathbf{B}\| \|\mathbf{C}\|_2, \|\mathbf{A}\|_2 \|\mathbf{B}\|_2 \|\mathbf{C}\|\} \quad (118)$$

D Additional Experiments

In this section we perform more experiments on learning the inverse operator for PDE matrices with State of the Art Matrix Experiments.

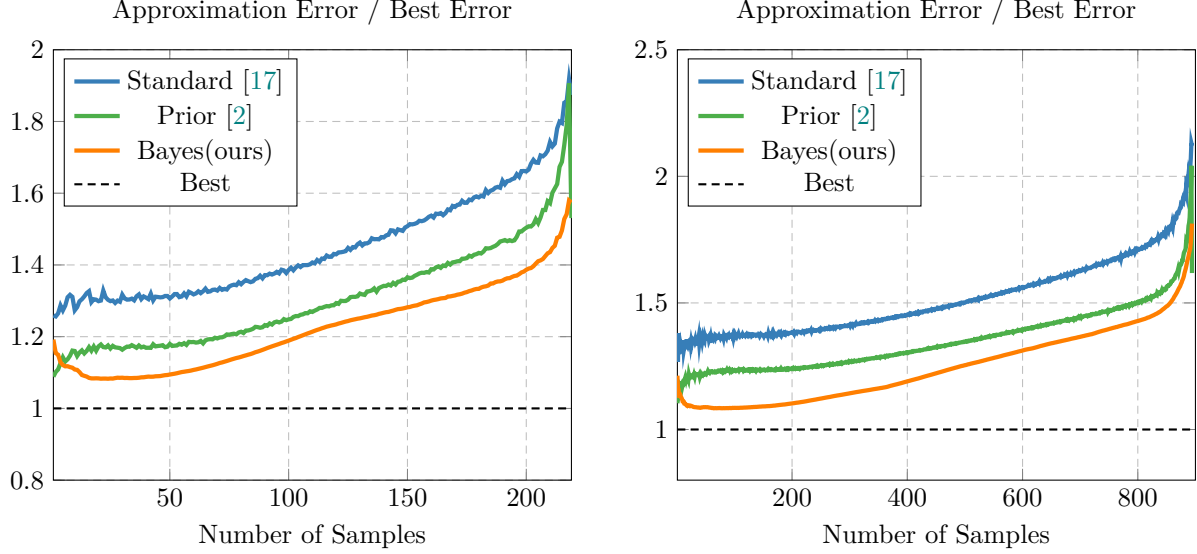


Figure 3: In (*Left*), Matrix from TAMU Sparse Matrix Suite `pde 225`. In (*Right*), Matrix from TAMU Sparse Matrix Suite `pde 900`. With the prior, we use the covariance matrix associated with the discrete Green’s Function for the Laplacian as in Equation (31).

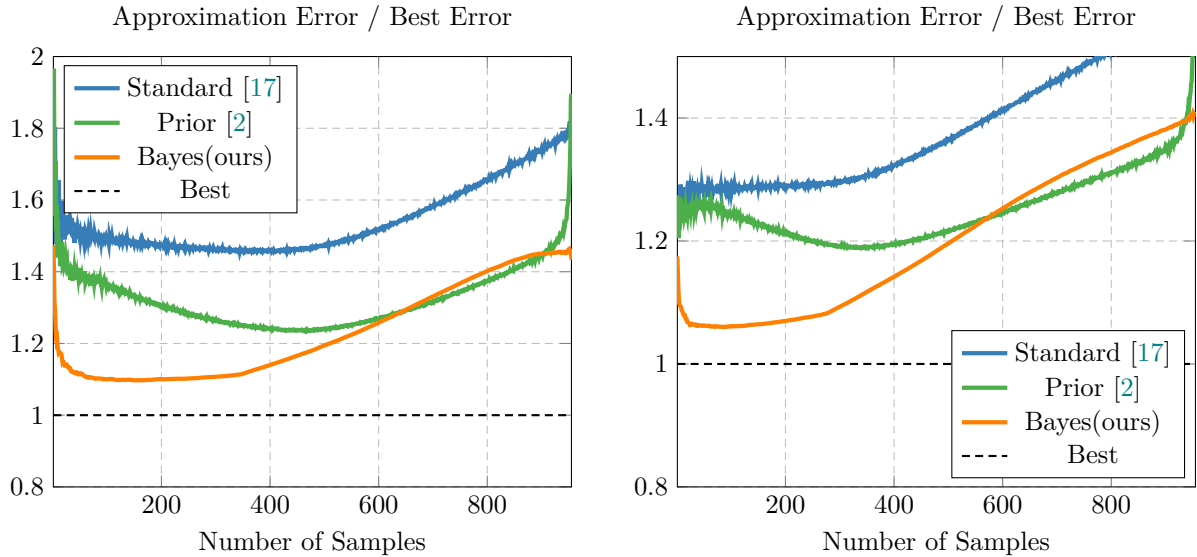


Figure 4: In these figures we look at matrices from Computational Fluid Dynamics. In (*Left*), Matrix from TAMU Sparse Matrix Suite `cdde1`. In (*Right*), Matrix from TAMU Sparse Matrix Suite `cdde1`. With the prior, we use the covariance matrix associated with the discrete Green’s Function for the Laplacian as in Equation (31).

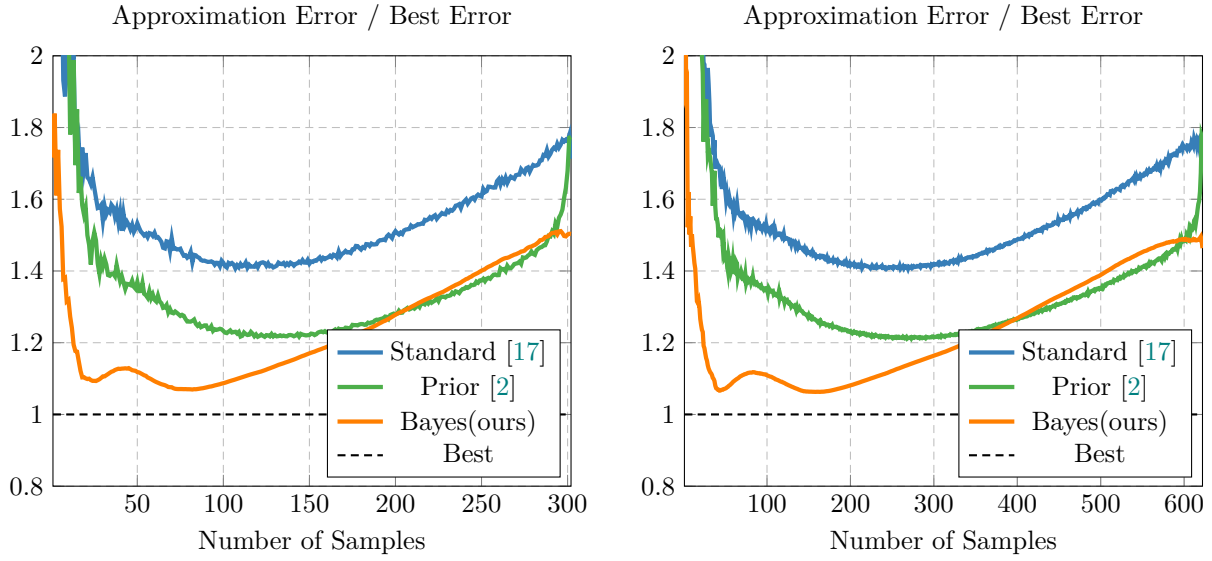


Figure 5: In these figures we look at matrices from the PDE for the Poisson Differential Operator. In (*Left*), Matrix from TAMU Sparse Matrix Suite cz308. In (*Right*), Matrix from TAMU Sparse Matrix Suite cz628. With the prior, we use the covariance matrix associated with the discrete Green's Function for the Laplacian as in Equation (31).