

# Subquantile Minimization for Kernel Learning in the Huber $\epsilon$ -Contamination Model

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## Abstract

In this paper we propose Subquantile Minimization for learning with adversarial corruption in the training set. Superquantile objectives have been formed in the past in the context of fairness where one wants to learn an underrepresented distribution equally [17, 31]. Our intuition is to learn a more favorable representation of the *majority* class, thus we propose to optimize over the  $p$ -subquantile of the loss in the dataset. In particular, we study the Huber Contamination Problem for Kernel Learning where the distribution is formed as,  $\hat{\mathbb{P}} = (1 - \epsilon)\mathbb{P} + \epsilon\mathbb{Q}$ , and we want to find the function  $\inf_f \mathbb{E}_{\mathbf{x} \in \mathbb{P}} [\ell_f(\mathbf{x})]$ , from the noisy distribution,  $\hat{\mathbb{P}}$ . We assume the adversary has knowledge of the true distribution of  $\mathbb{P}$ , and is able to corrupt the covariates and the labels of  $\epsilon$  samples. To our knowledge, we are the first to study the problem of general kernel learning in the Huber Contamination Model. In our theoretical analysis, we analyze our non-convex concave objective function with the Moreau Envelope. We show (i) a stationary point with respect to the Moreau Envelope is a good point and (ii) we can reach a stationary point with gradient descent methods. Further, we analyze accelerated gradient methods for the non-convex concave minimax optimization problem. We empirically test Kernel Ridge Regression and Kernel Classification on various state of the art datasets and show Subquantile Minimization gives strong results. Furthermore, we run experiments on various datasets and compare with the state-of-the-art algorithms to show the superior performance of Subquantile Minimization.

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## A Concentration Inequalities

In this section we will give various concentration inequalities on the inlier data. We first restate our assumptions.

**Assumption 26.** (Sub-Gaussian Design of Covariates). We assume each covariate is drawn i.i.d from a zero-mean covariance  $\Sigma$  sub-Gaussian distribution with sub-Gaussian norm  $K_1 \in \mathbb{R}_{++}$ .

$$\mathbb{E}[\mathbf{x}_i] = \mathbf{0} \quad (46)$$

$$\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^\top] = \Sigma \quad (47)$$

For all  $i \in [n]$ ,

$$\mathbb{P}\{\mathbf{v}^\top \mathbf{x}_i \geq t\} \leq \exp\left(-\frac{t^2}{K_1^2}\right) \quad (48)$$

Is this equivalent to

$$\mathbb{P}\{\mathbf{v}^\top \mathbf{x}_i \geq t\} \leq \exp\left(-\frac{t^2}{K_1^2}\right) \quad (49)$$

**Assumption 27.** (Sub-Gaussian Design of Optimal Residuals). Recall the residual is defined as  $\eta_i \triangleq \mathbf{f}_{\mathbf{w}}^*(\mathbf{x}_i) - y_i$ . Then we assume for some  $K_2 \in \mathbb{R}_{++}$

$$\mathbb{E}[\eta_i] = 0 \quad (50)$$

$$\mathbb{P}\{|\eta_i| \geq t\} \leq 2 \exp\left(-\frac{t^2}{K_2^2}\right) \quad (51)$$

### A.1 Linear Kernel

We will first begin with the Linear Kernel Case.

**Assumption 28.** Let the data have dimension  $d$  such that Assumption 26 holds. Then

$$n > \frac{1}{18} \left( \frac{d^2 K^2}{dK - 1} \right) \log(2d) \quad (52)$$

**Lemma 29.** Let  $\mathbf{x}_1, \dots, \mathbf{x}_n$  have sub-Gaussian design described in Assumption 26 and  $\eta_1, \dots, \eta_n$  have sub-Gaussian design described in Assumption 27. It then follows

$$\mathbb{E} \left| \sum_{i=1}^n \eta_i \|\mathbf{x}_i\|_2 \right| \leq \sqrt{\frac{nd}{K_1}} \left( \exp\left(\frac{C_1}{nd}\right) + \exp(C_2) \right) = \mathcal{O}(\sqrt{nd}) \quad (53)$$

**Proof.** Note  $\|\mathbf{x}\| = \sqrt{\sum_{i=1}^n x_i^2}$  is a  $\chi(n)$  random variable, thus  $\eta_i \|\mathbf{x}_i\|$  is symmetric around 0. Then for any  $t \geq 0$ ,

$$\mathbb{P} \left\{ \sum_{i=1}^n \eta_i \|\mathbf{x}_i\| \geq t \right\} = \mathbb{P} \left\{ \exp \left( \lambda \sum_{i=1}^n \eta_i \|\mathbf{x}_i\| \right) \geq \exp(\lambda t) \right\} \quad (54)$$

$$\stackrel{\text{Markov}}{\leq} e^{-\lambda t} \mathbb{E} \left[ \exp \left( \lambda \sum_{i=1}^n \eta_i \|\mathbf{x}_i\| \right) \right] \quad (55)$$

$$= e^{-\lambda t} \prod_{i=1}^n \underbrace{\mathbb{E}[\exp(\lambda \eta_i \|\mathbf{x}_i\|)]}_{\Upsilon} \quad (56)$$

Where  $\lambda > 0$  and will be chosen later. Now we will upper bound  $\Upsilon$ .

$$\Upsilon \triangleq \mathbb{E} [\lambda \exp(\eta_i \|\mathbf{x}_i\|)] \leq \mathbb{E} \left[ \exp \left( \lambda \left( \frac{\eta_i^2}{2} + \frac{\|\mathbf{x}_i\|^2}{2} \right) \right) \right] \quad (57)$$

$$= \mathbb{E} \left[ \exp \left( \lambda \frac{\eta_i^2}{2} \right) \exp \left( \lambda \frac{\|\mathbf{x}_i\|^2}{2} \right) \right] \quad (58)$$

$$\stackrel{\text{Young}}{\leq} \frac{1}{2} \mathbb{E} \left[ \exp(\lambda \eta_i^2) + \exp(\lambda \|\mathbf{x}_i\|^2) \right] \quad (59)$$

$$= \frac{1}{2} \mathbb{E} \left[ \exp(\lambda \eta_i^2) + \prod_{j=1}^d \exp(\lambda x_j^2) \right] \quad (60)$$

$$= \frac{1}{2} \left( \mathbb{E} [\exp(\lambda \eta_i^2)] + \prod_{j=1}^d \mathbb{E} [\exp(\lambda x_j^2)] \right) \quad (61)$$

$$\leq \frac{\zeta_1}{2} \exp \left( -\sqrt{\frac{K_1}{nd}} t \right) \left( \exp \left( \frac{C_1}{d} \right) + \exp(C_2) \right) \quad (62)$$

In  $\zeta_1$ , note  $\eta_i$  is sub-exponential and  $x_i^2$  for all  $i \in [n]$  are sub-exponential variables as  $x_i$  are sub-Gaussian as in Assumption 26. If we plug this back into Equation (56).

$$\mathbb{P} \left\{ \sum_{i=1}^n \eta_i \|\mathbf{x}_i\| \geq t \right\} \leq e^{-\lambda t} 2^{-n} (\exp(C_1 \lambda^2 K_2^2) + \exp(C_2 d \lambda^2 K_1^2))^n \quad (63)$$

$$\leq e^{-\lambda t} 2^{-n} (2^{n-1} \exp(C_1 n \lambda^2 K_2^2) + 2^{n-1} \exp(C_2 n d \lambda^2 K_1^2)) \quad (64)$$

$$= \frac{e^{-\lambda t}}{2} (\exp(C_1 n \lambda^2 K_2^2) + \exp(C_2 n d \lambda^2 K_1^2)) \quad (65)$$

Here we can choose  $\lambda \triangleq \sqrt{\frac{K_1}{nd}}$ . If we assume  $K_1 = K_2$ , we then have

$$\mathbb{P} \left\{ \sum_{i=1}^n \eta_i \|\mathbf{x}_i\| \geq t \right\} = \frac{1}{2} \exp \left( -\sqrt{\frac{K_1}{nd}} t \right) \left( \exp \left( \frac{C_1}{d} \right) + \exp(C_2) \right) \quad (66)$$

We can now upper bound the expectation

$$\mathbb{E} \left[ \sum_{i=1}^n \eta_i \|\mathbf{x}_i\|_2 \right] = \frac{1}{2} \int_0^\infty \exp \left( -\sqrt{\frac{K_1}{nd}} t \right) \left( \exp \left( \frac{C_1}{nd} \right) + \exp(C_2) \right) dt \quad (67)$$

$$= \sqrt{\frac{nd}{K_1}} \left( \exp \left( \frac{C_1}{nd} \right) + \exp(C_2) \right) \quad (68)$$

Upper bound  $\Upsilon$  also follows from techniques in [34]. ■

**Lemma 30.** *Let  $\mathbf{x}_1, \dots, \mathbf{x}_n$  have sub-Gaussian design described in Assumption 26 such that  $n > d \log(d)$ . It then follows*

$$\mathbb{E} \left[ \sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) \right] \geq \sigma_{\min}(\Sigma) \left( n - \left( \sqrt{2(dK-1)n \log(2d)} + \frac{1}{3} dK \log(2d) \right) \right) = \Omega(\sigma_{\min}(\Sigma) n) \quad (69)$$

and with probability  $(1 - \delta)$

$$\sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) \geq \frac{1}{1 - \delta} \sigma_{\min}(\Sigma) \left( n - \left( \sqrt{2(dK-1)n \log(2d)} + \frac{1}{3} dK \log(2d) \right) \right) \quad (70)$$

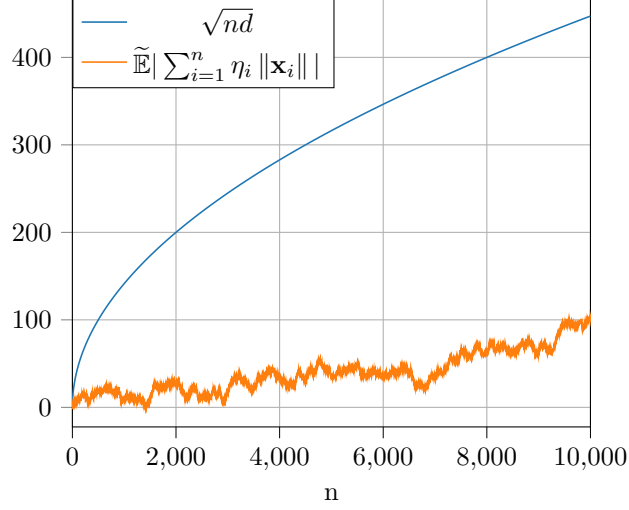


Figure 3: The bound for Lemma 29 compared to an average over 10 trials. We set  $\Sigma = \mathbf{I}$ ,  $K = 4$ ,  $d = 20$ .

**Proof.** In expectation we have

$$\mu_{\min} \triangleq \sigma_{\min} \left( \sum_{i=1}^n \mathbb{E} [\mathbf{x}_i \mathbf{x}_i^\top] \right) = \sigma_{\min} \left( \sum_{i=1}^n \Sigma \right) = n \sigma_{\min} (\Sigma) \quad (71)$$

Then we have by the sub-Gaussian design in Assumption 26.

$$\sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) = \sigma_{\min} \left( n \Sigma + \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top - n \Sigma \right) \quad (72)$$

$$\stackrel{\text{Weyl's}}{\geq} \sigma_{\min} (n \Sigma) - \sigma_{\max} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top - n \Sigma \right) \quad (73)$$

$$= \mu_{\min} - \underbrace{\left\| \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top - \Sigma \right\|_2}_A \quad (74)$$

We assume  $\Sigma \succcurlyeq \mathbf{0}$  and symmetric, therefore we can represent  $\Sigma = (\Sigma^{1/2})^\top \Sigma^{1/2} = \Sigma^{1/2} (\Sigma^{1/2})^\top$ . Next we will upper bound A.

$$A \triangleq \left\| \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top - \Sigma \right\|_2 \quad (75)$$

$$= \left\| \sum_{k \in [n]} \left( \Sigma^{1/2} \hat{\mathbf{x}}_k \right) \left( \Sigma^{1/2} \hat{\mathbf{x}}_k \right)^\top - \Sigma \right\|_2 \quad (76)$$

$$= \left\| \sum_{k \in [n]} \Sigma^{1/2} \hat{\mathbf{x}}_k \hat{\mathbf{x}}_k^\top \left( \Sigma^{1/2} \right)^\top - \Sigma \right\|_2 \quad (77)$$

$$= \left\| \Sigma^{1/2} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top - \mathbf{I} \right) \left( \Sigma^{1/2} \right)^\top \right\|_2 \quad (78)$$



$$\leq \underbrace{\left\| \sum_{k \in [n]} \hat{\mathbf{x}}_k \hat{\mathbf{x}}_k^\top - \mathbf{I} \right\|}_{\text{B}} \|\boldsymbol{\Sigma}\| \quad (79)$$

It then suffices to upper bound B. First, let  $\hat{\mathbf{x}} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$  and define  $K \geq 1$  s.t.  $\hat{\mathbf{x}} \leq dK$ , then

$$\mathbb{V}(\hat{\mathbf{x}}\hat{\mathbf{x}}^\top - \mathbf{I}) = \mathbb{E} \left[ (\hat{\mathbf{x}}\hat{\mathbf{x}}^\top - \mathbf{I})^\top (\hat{\mathbf{x}}\hat{\mathbf{x}}^\top - \mathbf{I}) \right] \quad (80)$$

$$= \mathbb{E} [\hat{\mathbf{x}}\hat{\mathbf{x}}^\top \hat{\mathbf{x}}\hat{\mathbf{x}}^\top] - 2\mathbb{E} [\hat{\mathbf{x}}\hat{\mathbf{x}}^\top] + \mathbf{I} \quad (81)$$

$$= \mathbb{E} [\hat{\mathbf{x}}\hat{\mathbf{x}}^\top \hat{\mathbf{x}}\hat{\mathbf{x}}^\top] - \mathbf{I} \quad (82)$$

$$= \mathbb{E} [\|\hat{\mathbf{x}}\|^2 \hat{\mathbf{x}}\hat{\mathbf{x}}^\top] - \mathbf{I} \quad (83)$$

$$\leq dK \mathbb{E} [\hat{\mathbf{x}}\hat{\mathbf{x}}^\top] - \mathbf{I} \quad (84)$$

$$= (dK - 1) \mathbf{I} \quad (85)$$

Then from Matrix Bernstein, we have

$$\mathbb{E}[\text{B}] \leq \sqrt{2(dK - 1)n \log(2d)} + \frac{1}{3}dK \log(2d) \quad (86)$$

Then we have in expectation

$$\mathbb{E} \left[ \sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) \right] \geq n \sigma_{\min}(\boldsymbol{\Sigma}) - \|\boldsymbol{\Sigma}\| \left( \sqrt{2(dK - 1)n \log(2d)} + \frac{1}{3}dK \log(2d) \right) \quad (87)$$

Assume the spectrum of  $\boldsymbol{\Sigma}$  is flat, then we have

$$\mathbb{E} \left[ \sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) \right] \geq \sigma_{\min}(\boldsymbol{\Sigma}) \left( n - \left( \sqrt{2(dK - 1)n \log(2d)} + \frac{1}{3}dK \log(2d) \right) \right) \quad (88)$$

Similarly from the Matrix Bernstein Inequality in [33], we have

$$\mathbb{P} \left\{ \sigma_{\min} \left( \sum_{k \in [n]} \mathbf{x}_k \mathbf{x}_k^\top \right) > t \right\} \leq (n + d) \exp \left( \frac{-t^2}{2n(dK - 1) + Kt/3} \right) \quad (89)$$

■

**Lemma 31.** Let  $\eta_i \in P \cap S$  be defined as in Assumption 27, then it follows

$$\mathbb{E} \left[ \sum_{i \in P \cap S} \eta_i^2 \right] \leq \Xi \quad (90)$$

**Proof.** We have

$$\sum_{i \in P \cap S} \eta_i^2 \leq \sum_{i \in P \cap S} (\eta_i^2 - \mathbb{V}(\eta_i)) + \mathbb{V}(\eta_i) \quad (91)$$

■

## A.2 Polynomial Kernel

The polynomial kernel is given by

$$K(\mathbf{x}_1, \mathbf{x}_2) = (\mathbf{x}_1^\top \mathbf{x}_2 + C)^p \quad (92)$$

The feature map for the polynomial kernel is given as

$$\phi_{\text{poly}}(\mathbf{x}) = [x_1, \dots, x_d, x_1^2, \dots, x_d^2, \dots, x_1^p, \dots, x_d^p, x_1 x_2, \dots, x_{d-1} x_d] \in \mathbb{R}^{d^p} \quad (93)$$

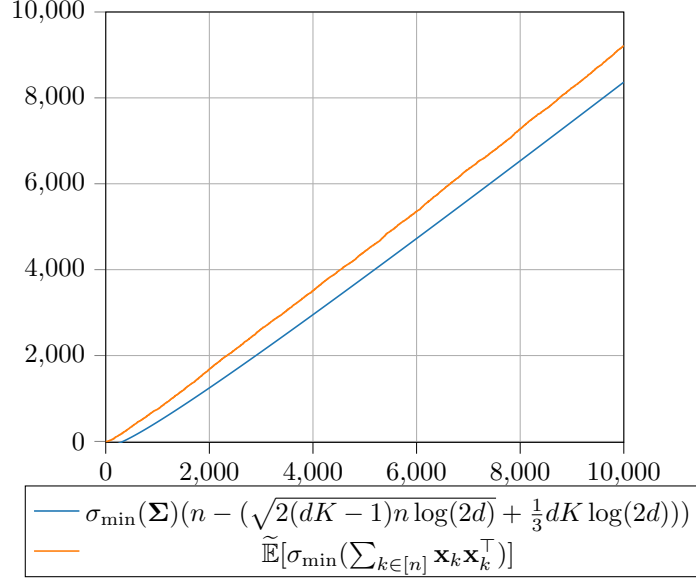


Figure 4: The bound for Lemma 30 compared to an average over 10 trials. We set  $\Sigma = \mathbf{I}$ ,  $K = 4$ ,  $d = 20$ .

**Lemma 32.** Let  $\mathbf{x}_1, \dots, \mathbf{x}_n \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ , and let  $\eta_1, \dots, \eta_n \sim \mathcal{N}(0, \sigma^2)$  be sub-Gaussian. Given  $C \in \mathbb{R}_+$ , it then follows

$$\mathbb{E} \left| \sum_{i=1}^n \eta_i \left( \|\mathbf{x}_i\|^2 + C \right)^{p/2} \right| \leq \Xi \quad (94)$$

**Lemma 33.** Let  $\mathbf{x}_1, \dots, \mathbf{x}_n \sim \mathcal{N}_d(\mathbf{0}, \mathbf{I})$  such that  $n > d^p$ . It then follows

$$\sigma_{\min} \left( \sum_{i=1}^n \phi_{\text{poly}}(\mathbf{x}_i) \phi_{\text{poly}}(\mathbf{x}_i)^\top \right) \geq \Xi \quad (95)$$

### A.3 Gaussian Kernel

The gaussian kernel is given for  $\gamma > 0$  by

$$K(\mathbf{x}_1, \mathbf{x}_2) = \exp \left( -\gamma \|\mathbf{x}_1 - \mathbf{x}_2\|_2^2 \right) \quad (96)$$

**Lemma 34.** Let  $\eta_1, \dots, \eta_n \sim \mathcal{N}(0, \sigma^2)$  be sub-Gaussian. Then it follows

$$\mathbb{E} \left| \sum_{i=1}^n \eta_i \right| \leq \Xi \quad (97)$$

We will utilize the Random Fourier Features (RFF) representation [27]. Let  $\mathbf{w}_1, \dots, \mathbf{w}_d \sim \mathcal{N}_d(\mathbf{0}, \frac{2}{\gamma} \mathbf{I})$ , then the RFF representation is given by

$$\phi_{\text{RFF}}(\mathbf{x}) = \frac{1}{\sqrt{d}} [\cos(\mathbf{w}_1^\top \mathbf{x}), \sin(\mathbf{w}_1^\top \mathbf{x}), \dots, \cos(\mathbf{w}_d^\top \mathbf{x}), \sin(\mathbf{w}_d^\top \mathbf{x})] \quad (98)$$

The key idea behind the Randomized Fourier Features is

$$\mathbb{E} [\phi_{\text{RFF}}(\mathbf{x}_1)^\top \phi_{\text{RFF}}(\mathbf{x}_2)] = \exp \left( -\gamma \|\mathbf{x}_1 - \mathbf{x}_2\|_2^2 \right) \quad (99)$$

We will first derive a relation between  $\mathbb{E} [\phi_{\text{RFF}}(\mathbf{x}_1)^\top \phi_{\text{RFF}}(\mathbf{x}_2)]$  and  $\mathbb{E} [\phi_{\text{RFF}}(\mathbf{x}_1)]^\top \mathbb{E} [\phi_{\text{RFF}}(\mathbf{x}_2)]$ .

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## D Necessary Lemmas

**Lemma 33** (MGF of Sub-Exponential Random Variable). *Let  $x$  be a Sub-exponential variable, then we have*

$$\mathbb{E}[\exp(\lambda x)] \leq \exp(C\lambda^2) \quad (135)$$

**Theorem 34** (Matrix Chernoff, [?]). *Let  $\mathbf{X}_k$  be a sequence of independent, random, self-adjoint matrices with dimension  $d$  s.t.*

$$\mathbf{X}_k \succcurlyeq \mathbf{0} \quad \text{and} \quad \lambda_{\max}(\mathbf{X}_k) \leq R \quad \text{almost surely} \quad (136)$$

*Define*

$$\mu_{\min} \triangleq \lambda_{\min}\left(\sum_k \mathbb{E}\mathbf{X}_k\right) \quad (137)$$

*Then for  $\delta \in [0, 1]$*

$$\mathbb{P}\left\{\lambda_{\min}\left(\sum_k \mathbf{X}_k\right) \geq (1-\delta)\mu_{\min}\right\} \leq d \cdot \left[\frac{e^{-\delta}}{(1-\delta)^{1-\delta}}\right]^{\mu_{\min}/R} \quad (138)$$

**Theorem 35** (Matrix Bernstein, [33]). *Let  $\mathbf{X}_k$  be a sequence of independent, random, self-adjoint matrices with dimension  $d$  s.t.*

$$\mathbb{E}\mathbf{X}_k = \mathbf{0} \quad (139)$$