

ENV797 - TIME SERIES ANALYSIS FOR ENERGY AND ENVIRONMENT APPLICATIONS

M8.1 – ARIMA Model Identification and Parameter Estimation

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Learning Goals

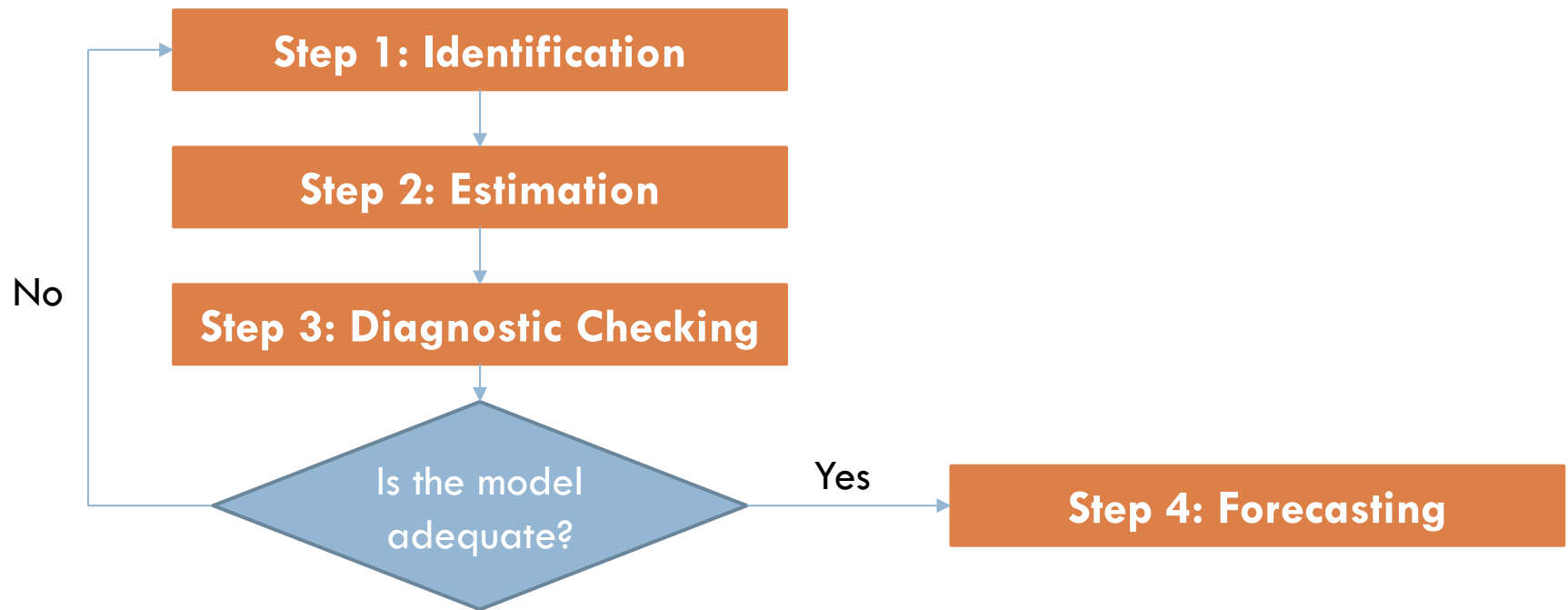


- Simple Example of ACF and PACF computation
- ARIMA Model Parameters Estimation
 - ▣ AR coefficients
 - ▣ MA coefficients
 - ▣ Variance of residuals



ARIMA Parameter Estimation

ARIMA Modeling - Process



Parameter Estimation

- We know that AR coefficients come from PACF
- How can we get PACF?
- Simple example: Excel spreadsheet
 - ▣ “*Temp_example_ACF_PACF_computation.xlsx*”

Simple Example: ACF Computation

<i>t</i>	<i>Y_t</i>						
1	14.2						
2	16.4						
3	11.9						
4	15.2						
5	18.5						
6	22.1						
7	19.4						
8	25.1						
9	23.4						
10	18.1						
11	22.6						
12	17.2						

		Lag 1				
t	Y _t	Y1 _t	[Y _t - Mu(Y)]	[Y1 _t - Mu(Y)]	Mult	
1	14,2	-	-	-	-	
2	16,4	14,2	-2,275	-4,475	10,181	
3	11,9	16,4	-6,775	-2,275	15,413	
4	15,2	11,9	-3,475	-6,775	23,543	
5	18,5	15,2	-0,175	-3,475	0,608	
6	22,1	18,5	3,425	-0,175	-0,599	
7	19,4	22,1	0,725	3,425	2,483	
8	25,1	19,4	6,425	0,725	4,658	
9	23,4	25,1	4,725	6,425	30,358	
10	18,1	23,4	-0,575	4,725	-2,717	
11	22,6	18,1	3,925	-0,575	-2,257	
12	17,2	22,6	-1,475	3,925	-5,789	
Mean		18,68			SUM	75,882
Std. Dev.		3,84			Num. Obs.	12
					COVARIANCE(Y,Y1)	6,32
					CORRELATION(Y,Y1)	0,4288

		Lag 1				
t	Y _t	Y2 _t	[Y _t - Mu(Y)]	[Y2 _t - Mu(Y)]	Mult	
1	14,2	-	-	-	-	
2	16,4	-	-	-	-	
3	11,9	14,2	-6,775	-4,475	30,318	
4	15,2	16,4	-3,475	-2,275	7,906	
5	18,5	11,9	-0,175	-6,775	1,186	
6	22,1	15,2	3,425	-3,475	-11,902	
7	19,4	18,5	0,725	-0,175	-0,127	
8	25,1	22,1	6,425	3,425	22,006	
9	23,4	19,4	4,725	0,725	3,426	
10	18,1	25,1	-0,575	6,425	-3,694	
11	22,6	23,4	3,925	4,725	18,546	
12	17,2	18,1	-1,475	-0,575	0,848	
Mean		18,68			SUM	68,511
Std. Dev.		3,84			Num. Obs.	12
					CORRELATION(Y,Y2)	5,71
					CORRELATION(Y,Y2)	0,3871

$$Cov(Y_t Y_s) = E[(Y_t - \mu)(Y_s - \mu)]$$

$$Cov(Y_t Y_s) = \frac{1}{n} \sum (Y_t - \mu)(Y_s - \mu)$$

$$Corr(Y_t Y_s) = \frac{Cov(Y_t Y_s)}{\sigma_Y^2} = ACF(t-s)$$

Compare with R
function:

```
> print(acf_temp$acf)
, , 1
```

```

[1,] 1.00000000
[2,] 0.42875355
[3,] 0.38710748
[4,] 0.13060966
[5,] -0.24677581
[6,] -0.36416383
[7,] -0.30293249
[8,] -0.27678937
[9,] -0.23337053
[10,] 0.02054935
[11,] -0.08028336
```

Simple Example: ACF Computation

t	Y_t
1	14.2
2	16.4
3	11.9
4	15.2
5	18.5
6	22.1
7	19.4
8	25.1
9	23.4
10	18.1
11	22.6
12	17.2

ACF(3)			Lag 1			
	t	Y_t	Y3_t	[Y_t - Mu(Y)]	[Y3_t - Mu(Y)]	Mult
	1	14,2	-	-	-	-
	2	16,4	-	-	-	-
	3	11,9	-	-	-	-
	4	15,2	14,2	-3,475	-4,475	15,551
	5	18,5	16,4	-0,175	-2,275	0,398
	6	22,1	11,9	3,425	-6,775	-23,204
	7	19,4	15,2	0,725	-3,475	-2,519
	8	25,1	18,5	6,425	-0,175	-1,124
	9	23,4	22,1	4,725	3,425	16,183
	10	18,1	19,4	-0,575	0,725	-0,417
	11	22,6	25,1	3,925	6,425	25,218
	12	17,2	23,4	-1,475	4,725	-6,969
	Mean	18,68		SUM	23,116	
	Std. Dev.	3,84		Num. Obs.	12	
				CORRELATION(Y,Y2)	1,93	
				CORRELATION(Y,Y2)	0,1306	

ACF(4)			Lag 1			
	t	Y_t	Y4_t	[Y_t - Mu(Y)]	[Y4_t - Mu(Y)]	Mult
	1	14,2	-	-	-	-
	2	16,4	-	-	-	-
	3	11,9	-	-	-	-
	4	15,2	-	-	-	-
	5	18,5	14,2	-0,175	-4,475	0,783
	6	22,1	16,4	3,425	-2,275	-7,792
	7	19,4	11,9	0,725	-6,775	-4,912
	8	25,1	15,2	6,425	-3,475	-22,327
	9	23,4	18,5	4,725	-0,175	-0,827
	10	18,1	22,1	-0,575	3,425	-1,969
	11	22,6	19,4	3,925	0,725	2,846
	12	17,2	25,1	-1,475	6,425	-9,477
	Mean	18,68		SUM	-43,675	
	Std. Dev.	3,84		Num. Obs.	12	
				CORRELATION(Y,Y2)	-3,64	
				CORRELATION(Y,Y2)	-0,2468	

Compare with R values:

```
> print(acf_temp$acf)
, , 1
```

```

[1,] 1.00000000
[2,] 0.42875355
[3,] 0.38710748
[4,] 0.13060966
[5,] -0.24677581
[6,] -0.36416383
[7,] -0.30293249
[8,] -0.27678937
[9,] -0.23337053
[10,] 0.02054935
[11,] -0.08028336
```

PACF Concept

- Consider the two regression models

$$y_t = \beta_0 + \beta_1 y_{t-2}$$

$$y_t = \beta_0 + \beta_1 y_{t-1} + \beta_2 y_{t-2}$$

- What is the meaning of β_1 in the first model and β_2 in the second model?
- They both represent the linear dependence between observation y_t and y_{t-2}
- But what's difference between β_1 in the first model and β_2 in the second model?
- **The β_2 in the second model is the linear dependence between y_t and y_{t-2} WITH the dependency between y_t and y_{t-1} ALREADY accounted for**

ARIMA Parameter Estimation

- What do we need to estimate?
 - ▣ AR coefficients
 - ▣ MA coefficients
 - ▣ Variance of the residuals (error or innovation)
- Estimation Methods for ARIMA coefficients
 - ▣ Least squares method
 - ▣ Maximum Likelihood
 - ▣ Methods of Moments or Yule-Walker equations
- Other Methods: Bayesian estimation or Kalman Filtering

Estimating AR(p) parameters

- AR(p)

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + a_t \quad \text{for } t = 1, 2, \dots, n$$

$a_t \sim i.i.d. (0, \sigma^2)$

- Need to estimate $\phi = (\phi_1 \phi_2 \dots \phi_p)'$ and σ^2
- Estimation method: Method of moments (Yule-Walker equations)
- Yule Walker equations relate AR model coefficients to the autocovariance (ACF) of the random process
- Note: They do not address the model order, they are simply used to estimate AR parameters

Estimating AR coefficients

- Consider AR(p) model with a zero mean

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \cdots + \phi_p Y_{t-p} + a_t$$

$$Y_t Y_{t-h} = \phi_1 Y_{t-1} Y_{t-h} + \phi_2 Y_{t-2} Y_{t-h} + \cdots + \phi_p Y_{t-p} Y_{t-h} + a_t Y_{t-h} \quad (\text{with } h > 0)$$

$$E[Y_t Y_{t-h}] = E[\phi_1 Y_{t-1} Y_{t-h}] + E[\phi_2 Y_{t-2} Y_{t-h}] + \cdots + E[\phi_p Y_{t-p} Y_{t-h}] + E[a_t Y_{t-h}]$$

$$E[Y_t Y_{t-h}] = \phi_1 E[Y_{t-1} Y_{t-h}] + \cdots + \phi_p E[Y_{t-p} Y_{t-h}] + E[a_t Y_{t-h}] \quad \text{Eq. (1)}$$

- Closer look at last term $E[a_t Y_{t-h}]$

- ▣ a_t is uncorrelated with Y_{t-h} , therefore $E[a_t Y_{t-h}] = a_t E[Y_{t-h}]$

- ▣ $E[Y_t] = 0$ since this is a zero mean process, therefore

$$E[a_t Y_{t-h}] = a_t E[Y_{t-h}] = a_t * 0 = 0$$

- Back in Eq. (1)

$$E[Y_t Y_{t-h}] = \phi_1 E[Y_{t-1} Y_{t-h}] + \cdots + \phi_p E[Y_{t-p} Y_{t-h}] + E[a_t Y_{t-h}]$$

$$E[Y_t Y_{t-h}] = \phi_1 E[Y_{t-1} Y_{t-h}] + \cdots + \phi_p E[Y_{t-p} Y_{t-h}] \quad \text{Eq. (2)}$$

Estimating AR coefficients (cont'd)

□ Recall two definitions from Lect. 3

1. $\gamma_{t,s} = \text{Cov}(Y_t Y_s) = E[Y_t Y_s] - \mu_t \mu_s$

2. For a stationary process $\gamma_{t,s} = \gamma_{0,|t-s|}$

□ Since we are considering a zero mean series the above relationships leads to

$$\gamma_{t,s} = E[Y_t Y_s] - \cancel{\mu_t \mu_s}^0 \because E[Y_t Y_s] = \gamma_{t,s}$$

□ Rewriting in terms of t and $t - h$ we get

$$E[Y_t Y_{t-h}] = \gamma_{t,t-h} = \gamma_{0,h} \text{ or simply } E[Y_t Y_{t-h}] = \gamma_h$$

□ Substituting in Eq. (2)

$$E[Y_t Y_{t-h}] = \phi_1 E[Y_{t-1} Y_{t-h}] + \cdots + \phi_p E[Y_{t-p} Y_{t-h}]$$

$$\gamma_h = \phi_1 \gamma_{h-1} + \cdots + \phi_p \gamma_{h-p}$$

Estimating AR coefficients (cont'd)

- For a zero mean process autocovariance divided by variance equal autocorrelation, therefore

$$\rho_h = \phi_1 \rho_{h-1} + \cdots + \phi_p \rho_{h-p}$$

- Writing this equation for $h = 1, 2, \dots, p$ we get

$$h = 1 \quad \rho_1 = \phi_1 \rho_0 + \phi_2 \rho_1 + \cdots + \phi_p \rho_{p-1}$$

$$h = 2 \quad \rho_2 = \phi_1 \rho_1 + \phi_2 \rho_0 + \cdots + \phi_p \rho_{p-2}$$

$$\vdots \quad \quad \quad \vdots$$

$$h = p \quad \rho_p = \phi_1 \rho_{p-1} + \phi_2 \rho_{p-2} + \cdots + \phi_p \rho_0$$

Note that for $h=1$:

$$\rho_{h-p} = \rho_{1-p} = \rho_{p-1}$$

**Yule-Walker
equations**

Estimating AR coefficients (cont'd)

- In matrix form

$$\underbrace{\begin{bmatrix} \rho_1 \\ \rho_2 \\ \rho_3 \\ \vdots \\ \rho_p \end{bmatrix}}_{\boldsymbol{\rho}} = \underbrace{\begin{bmatrix} \rho_0 & \rho_1 & \rho_2 & \cdots & \rho_{p-1} \\ \rho_1 & \rho_0 & \rho_1 & \cdots & \rho_{p-2} \\ \rho_2 & \rho_1 & \rho_0 & \cdots & \rho_{p-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \rho_{p-1} & \rho_{p-2} & \rho_{p-3} & \cdots & \rho_0 \end{bmatrix}}_{\mathbf{R}} \cdot \underbrace{\begin{bmatrix} \phi_1 \\ \phi_2 \\ \phi_3 \\ \vdots \\ \phi_p \end{bmatrix}}_{\boldsymbol{\phi}}$$

- This a linear system with p equations and p variables

$$\boldsymbol{\rho}_p = \mathbf{R}_p \cdot \boldsymbol{\phi}_p \quad \longrightarrow \quad \boldsymbol{\phi}_p = \mathbf{R}_p^{-1} \cdot \boldsymbol{\rho}_p$$

Simple Example: PACF Computation

PACF(2)

$$\begin{array}{c} \text{ACFs } (\backslash \rho) \\ \begin{array}{|c|} \hline 0,4288 \\ \hline 0,3871 \\ \hline \end{array} \end{array} = \begin{array}{c} \text{R (ACF matrix)} \\ \begin{array}{|c|c|} \hline 1 & 0,4288 \\ \hline 0,4288 & 1 \\ \hline \end{array} \end{array} * \begin{array}{c} \backslash \phi_1 \\ \backslash \phi_2 \end{array}$$

$$\begin{array}{c} \backslash \phi_1 \\ \backslash \phi_2 \end{array} = \begin{array}{c} \text{R inverse} \\ \begin{array}{|c|c|} \hline 1,2252 & -0,5253 \\ \hline -0,5253 & 1,2252 \\ \hline \end{array} \end{array} * \begin{array}{c} \text{ACFs } (\backslash \rho) \\ \begin{array}{|c|} \hline 0,4288 \\ \hline 0,3871 \\ \hline \end{array} \end{array} = \begin{array}{c} \begin{array}{|c|} \hline 0,322 \\ \hline \end{array} \\ \begin{array}{|c|} \hline 0,249 \\ \hline \end{array} \end{array}$$

Compare with R values:

```
> print(pacf_temp$acf)
, , 1
```

```

      [,1]
[1,]  0.4287535
[2,]  0.2490630
[3,] -0.1316882
[4,] -0.4645661
[5,] -0.2730998
[6,]  0.2045947
[7,]  0.1920660
[8,] -0.2915612
[9,] -0.1206727
[10,] -0.1351969

```

Estimating the variance of AR(p)

- To get an estimate for the variance use same approach

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \cdots + \phi_p Y_{t-p} + a_t$$

- Multiply by Y_{t-h} and take expectations

$$E[Y_t Y_{t-h}] = \phi_1 E[Y_{t-1} Y_{t-h}] + \cdots + \phi_p E[Y_{t-p} Y_{t-h}] + E[a_t Y_{t-h}]$$

- But now we consider the case where $h = 0$, so the last term is no equal to zero, instead

$$E[a_t Y_t] = \sigma^2$$

- Therefore

$$\rho_0 = \phi_1 \rho_1 + \cdots + \phi_p \rho_p + \sigma^2 \quad \Rightarrow \quad \sigma^2 = \rho_0 - \phi_1 \rho_1 - \cdots - \phi_p \rho_p$$

- From estimates of ϕ_1, \dots, ϕ_p we can get an estimate for σ^2

Estimating MA coefficients

- Difficult because regressors are unknown residuals
- Assuming **invertibility** we can write a MA process as an AR and then use the Yule-Walker equations
- Example: MA(1) to AR(∞)
 - ▣ MA(1): $Y_t = \epsilon_t - \theta\epsilon_{t-1}$
 - ▣ Define operator L such that: $L\epsilon_t = \epsilon_{t-1}$ (also denoted by B and known as back shift operator)
 - Therefore $LL\epsilon_t = L\epsilon_{t-1} = \epsilon_{t-2}$
 - ▣ In terms of back shift operator: $Y_t = (1 - \theta L)\epsilon_t$
 - ▣ Rewriting: $\frac{Y_t}{(1-\theta L)} = \epsilon_t$

Sum of Geometric Progression: $S_\infty = a + ra + r^2a + \dots = \frac{a}{1-r}$ if $|r| < 1$

Estimating MA coefficients (cont'd)

- In that case if $|\theta| < 1$

$$\frac{Y_t}{(1 - \theta L)} = Y_t + \theta L Y_t + \theta^2 L^2 Y_t + \theta^3 L^3 Y_t + \dots = \epsilon_t$$

- Isolating Y_t we get

$$Y_t = -\underbrace{\theta L Y_t}_{Y_{t-1}} - \underbrace{\theta^2 L^2 Y_t}_{Y_{t-2}} - \underbrace{\theta^3 L^3 Y_t}_{Y_{t-3}} - \dots + \epsilon_t$$

From the definition of operator L

- Therefore

$$Y_t = -\theta Y_{t-1} - \theta^2 Y_{t-2} - \theta^3 Y_{t-3} - \dots + \epsilon_t$$

Now you can use the same AR methods

Example of Parameter Estimation

□ Temperature Data

Series: temp
ARIMA(2,0,0) with non-zero mean

Coefficients:

ar1	ar2	mean
0.3190	0.2711	-0.4614
s.e. 0.2803	0.2907	2.0245

sigma^2 estimated as 14.31: log likelihood=-31.45
AIC=70.9 AICc=76.62 BIC=72.84

Call:
lm(formula = Y[, 1] ~ Y[, 2] + Y[, 3])

Coefficients:

(Intercept)	Y[, 2]	Y[, 3]
0.6157	0.2317	0.3206

[1] "Yule Walker results:"

	[,1]	[,2]
[1,]	0.3219669	0.249063

Y_{t-1}

Y_{t-2}

□ Electricity Prices Data

Series: random_price
ARIMA(2,0,0) with non-zero mean

Coefficients:

ar1	ar2	mean
-0.0711	0.0335	0.0000
s.e. 0.0706	0.0710	0.0059

sigma^2 estimated as 0.007577: log likelihood=208.03
AIC=-408.06 AICc=-407.85 BIC=-394.82

Call:
lm(formula = Y[, 1] ~ Y[, 2] + Y[, 3])

Coefficients:

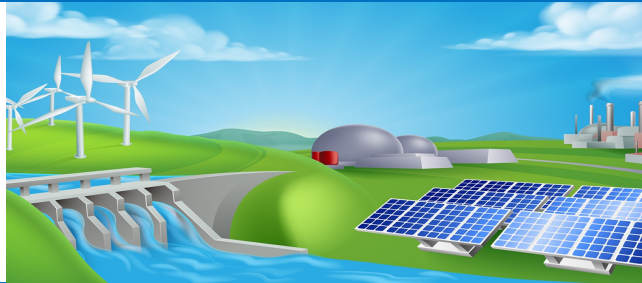
(Intercept)	Y[, 2]	Y[, 3]
-0.001259	0.045078	0.043987

[1] "Yule Walker results:"

	[,1]	[,2]
[1,]	-0.0705510	0.03261288

Y_{t-1}

Y_{t-2}



THANK YOU !

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