

Introduction to hypothesis testing

School of Life and Environmental Sciences (SOLES)

Floris van Ogtrop
The University of Sydney

Mar 2024



THE UNIVERSITY OF
SYDNEY

Outline

- Inferential statistics
- Confidence intervals
- Hypothesis testing
- 1-sample tests

Inferential statistics

- You should be familiar with **EDA** by now
- Now we are going to take the next step into the world of **inferential statistics**
- This is an important and sometimes difficult step in statistics
- We are going to draw conclusions about a population from a “smaller” sample of that population

Confidence intervals

- A confidence interval is a range of values, derived from a sample, that is likely to contain the value of an unknown population parameter.
- Confidence are important and many people prefer them to formal hypothesis tests because they give a range of values rather than a single value.
- There are two types we are going to look at:
 - Confidence interval where we **know** the population variance
 - Confidence interval where we **do not know** the population variance

Confidence intervals - known variance

- Assuming our sample comes from a normal population, we can calculate the confidence interval for the mean of a population when the population **variance is known**. An example may be in manufacturing where we know the variance of a process.
- The formula for the confidence interval is:

$$CI = \bar{y} \pm Z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$$

for the 95% confidence interval, $\alpha = 0.05$ and $Z_{\alpha/2} = 1.96$ i.e. 2 standard errors from the mean.

Confidence intervals - known variance

- \bar{y} is a random variable with a sampling distribution.
- Because there is an infinite number of values of \bar{y} , there is an infinite number of intervals of the form $CI = \bar{y} \pm 1.96 \cdot \frac{\sigma}{\sqrt{n}}$.
- The probability statement says that 95% of these intervals will actually include μ between the limits.
- For any one interval, $CI = \bar{y} \pm 1.96 \cdot \frac{\sigma}{\sqrt{n}}$, we say that we are 95% confident that μ lies between these limits.

Confidence intervals - known variance

Example

Say we have measured the weights of 6 chocolate bars from Florry's chocolate factory. The readings were:

48 43 45 51 49 40

The mean for these six values is $\bar{y} = 46$ g. Let's assume that the bar weights follow a normal distribution and that $\sigma = 1$ g These assumptions allow us to calculate a 95% z-based confidence interval:

So the 95% CI for the current example is

$$CI = 46 \pm 1.96 \cdot \frac{1}{\sqrt{6}} \text{ or } CI = 46 \pm 0.80 \text{ or } CI = (45.20, 46.80)$$

So we are 95% confident that the true mean weight of the chocolate bars is between 45.20 and 46.80g.

- We will calculate confidence intervals with **unknown variance** in later on.

Hypothesis testing

- A hypothesis consists either of a suggested explanation for a phenomenon or of a reasoned proposal suggesting possible relationships between multiple phenomena

Why?

- To make evidence based decisions, we need to evaluate the evidence
- Hypothesis testing is a scientific method for weighing up the evidence given the data against the given hypothesis (model)

Evidence?

- The data is not consistent with the hypothesis if the gap between the observed value (data) and the expected value is too big (> 2 or 3 Standard Errors)

Hypothesis testing

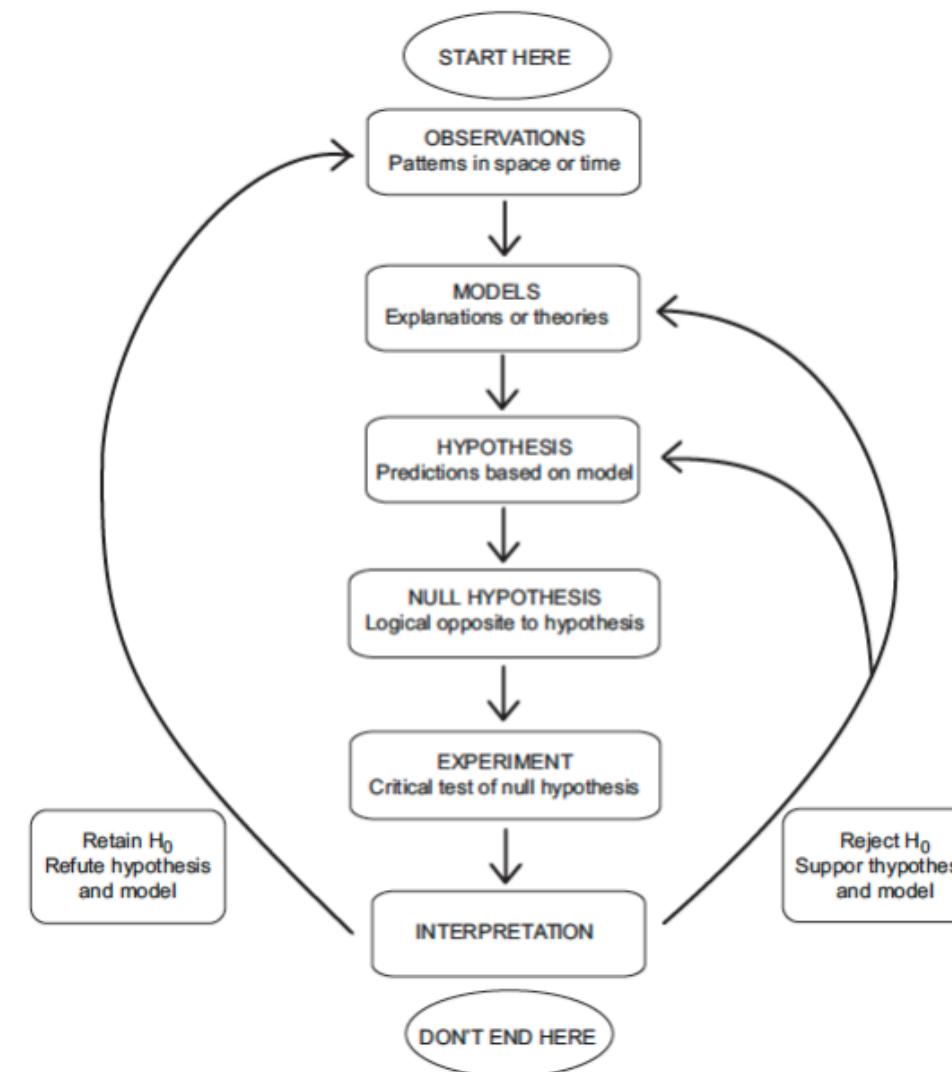
Framework

- Null Hypothesis (H_0) vs Alternative Hypothesis (H_1)
- Is the data consistent with H_0 ? Wait, but I'm interested in if there is a difference?
 - Falsifiability: only takes one contrary observation to falsify a statement (Karl Popper)
 - Can we falsify this statement? “There are invisible and undetectable fairies in my garden.” - it is non-falsifiable so we cannot test it with empirical data.
 - Can we falsify this statement? “All swans are black” - if we observe a single white swan, then the statement is falsified.



generated by DALL-E 3

Hypothesis testing



Underwood, A.J., 2009, April. Components of design in ecological field experiments. In *Annales Zoologici Fennici* (Vol. 46, No. 2, pp. 93-111). Finnish Zoological and Botanical Publishing Board.

Hypothesis testing

- The **null hypothesis** (H_0) assumes that the difference between the observation (data) and the expected value is due to chance
- The **alternative hypothesis** (H_1) assumes that the difference between the observed value (data) and expected value is NOT due to chance alone.

Test Statistic

Measures the difference between what is observed in the data and what is expected from the null hypothesis.

$$\text{Test statistic} = \frac{\text{Observed value} - \text{Expected value}}{\text{Standard error}}$$

P-Value

What is P-value?

- The P-value is a way of weighing up whether the sample (data) is consistent with H_0

More formally:

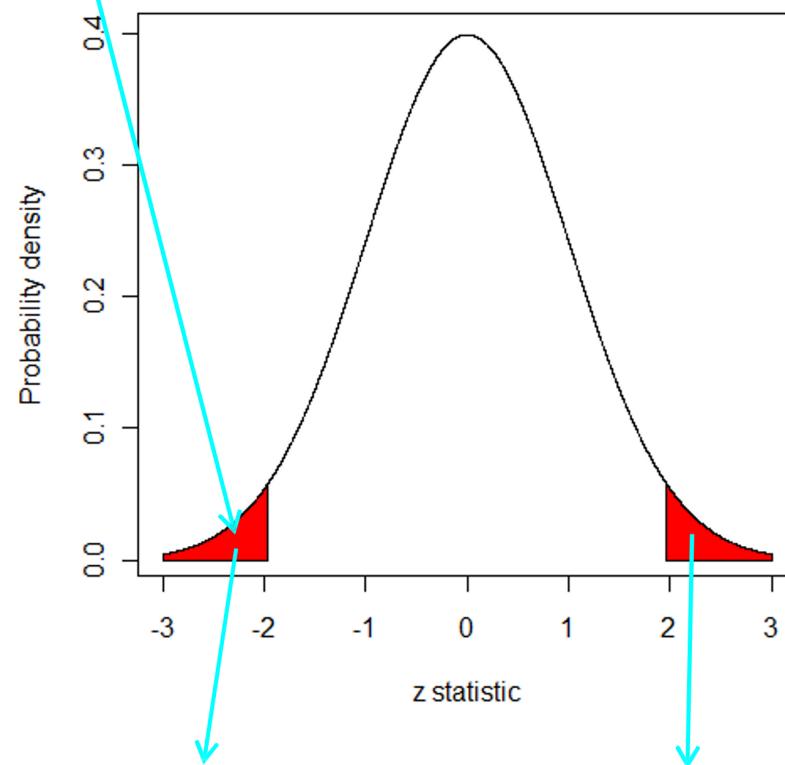
- The P-value is the chance of observing the test statistic (or something more extreme) if the H_0 is true.

P-Value

What does this have to do with sampling distributions?

Sample mean

2 tailed test



P = 0.025

P = 0.025

P-Value

“If the P-value is low the Null hypothesis must go.”

- So we calculate the P-value to assess the level of evidence against the Null hypothesis.
 - Generally if a P-value > 0.05 we say there is **no strong** evidence to reject the null hypothesis (we fail to reject it)
 - We never say we accept the null hypothesis!!
 - 0.05 is called the significance level
 - We may choose other significance levels such as 0.2, 0.01, 0.001 (for this course we will use 0.05)

P-Value

Common mistakes

- The p-value is not the chance that the NULL hypothesis is true
- A large p-value does not mean that H_0 is true

Size of p-value	What not to say	What to say
Small	Ho is not true	There is evidence against Ho
	Ho is false	We reject Ho
Large	We accept Ho	Data is consistent with Ho
		We retain Ho

- The use of 0.05 is not mandatory, it is a convention.

P-value

Controversy

Nature article 1

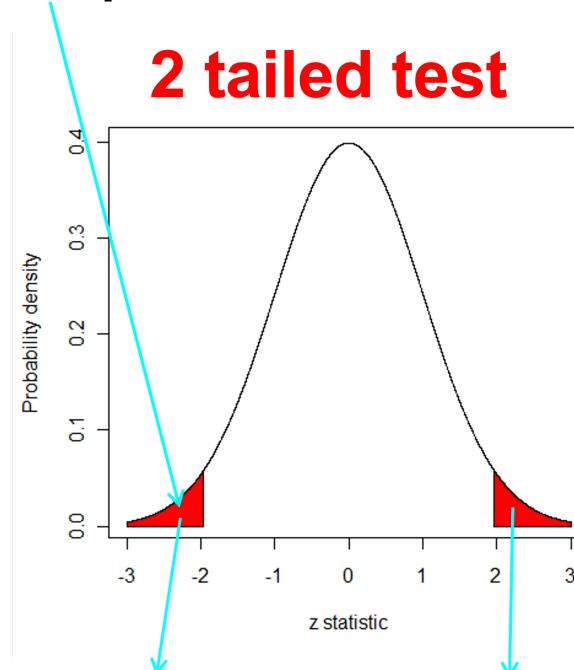
Nature article 2

- Should we move away from an over reliance on p-values towards more robust statistical methods and a greater emphasis on the size and reproducibility of effects?
- Look up terms like “**P Hacking**”...

Hypothesis testing

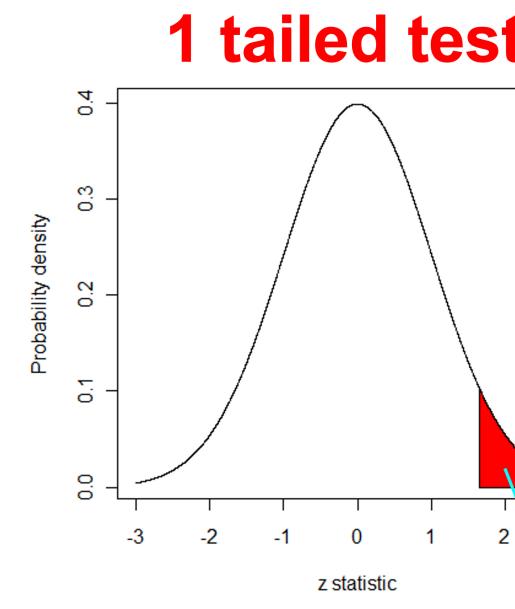
Tails??

Sample mean



$$P = 0.025$$

2 tailed test



$$P = 0.025$$

1 tailed test

$$P = 0.05$$

- In this course we will mostly use two-tailed tests because we are interested in whether the observed value is different to the expected value.
- One tailed tests are used when we are specifically interested in whether the observed value is greater or less than the expected value. Examples are in quality control and production processes.

Hypothesis testing

- Let say we would like to test whether a sample or population behaves according to a known or expected value. For example, ENVX1002 students have a mean heart rate of 70 beats per minute (c). We could test the relationship between our population or sample mean (μ) or (y) and our expected mean (c).
 - $H_0 : \mu = c$ $H_0 : \mu \neq c$ *Two tailed*
 - $H_0 : \mu \geq c$ $H_0 : \mu < c$ *One tailed*
 - $H_0 : \mu \leq c$ $H_0 : \mu > c$ *One tailed*
- We usually use the first (two tailed) form unless we are specifically interested in testing whether the difference is in one direction or the means of two variables are equivalent (i.e. equivalence testing in production)
- In this course we are assuming that the population/sample is normally distributed.

Hypothesis testing

Type I error

- False positive: we reject H_0 when it is true

Type II error

- False negative: we accept H_0 when it is false

		DECISION
		H_0 rejected H_0 accepted / retained
H_0 is false	H_0 rejected Correct decision $1-\beta$	False negative Type II Error ($P = \beta$ usually unknown)
H_0 is true	False positive Type I Error ($P = \alpha$ usually 0.05)	Correct decision $1-\alpha$

Note: Decreasing the significance level (α) to say 0.01 will increase the probability of type II error

Hypothesis testing

Sample size & Power

- While we won't do these calculations, it is important to think about sample size and power. For example, when comparing two means we can estimate the sample size required using the following

$$n = \left(\frac{(Z_{\alpha/2} + Z_{\beta})^2 \cdot (SD_1^2 + SD_2^2)}{\delta^2} \right)$$

- n represents the required sample size for each group.
- $Z_{\alpha/2}$ denotes the critical value of the normal distribution at the desired significance level (α). For example, for a 95% confidence level ($\alpha = 0.05$, $Z_{\alpha/2} = 1.96$).
- Z_{β} is the Z-score corresponding to the desired power of the test ($1 - \beta$). For 80% power ($\beta = 0.2$), you would use the Z-score that corresponds to 80% in the normal distribution table.
- SD_1 and SD_2 are the standard deviations of the two groups being compared.
- δ is the minimum difference in means between the two groups that the study is designed to detect, also known as the effect size.

Hypothesis testing

Sample size & Power

In a nut shell...

High Power: If a test has high power (close to 1), it means that the probability of committing a Type II error is low. This indicates that the test is very likely to detect an effect or difference when one exists.

Low Power: Conversely, if a test has low power, the probability of committing a Type II error is high. This means the test is less likely to detect an effect or difference, even if one exists.

Sample size: Generally power increases as sample size increases, but so does time and costs!

Hypothesis testing

Steps in hypothesis testing

- Choose level of significance (usually 0.05)
- Write null and alternate hypotheses
- Check assumptions
- Calculate test statistic
- Obtain P-value or critical value
- Make statistical conclusion
- Write a scientific (biological) conclusion

Hypothesis testing

Steps in hypothesis testing

Sometimes we use HATPC to remember these steps - HAT Prevent Cancer - come up with your own

- **H:** Hypothesis
- **A:** Assumptions
- **T:** Test statistic
- **P:** P-Value
- **C:** Conclusion

Example: One sample t-test



Hypothesis

Say our hippie friend here has started feeding his mung bean crop CBD oil and claims it has significantly increased the yield compared to the national average of 1.1 t/ha.

State the Null and Alternate hypothesis!

- $H_0 = 1.1 \text{ t/ha}$ (Null hypothesis)
- $H_1 \neq 1.1 \text{ t/ha}$ (Alternate hypothesis)



THE UNIVERSITY OF SYDNEY

Example: One sample t-test

Assumptions

- This time we are testing a data set where the **variance is unknown**.
- A key assumption for the one-sample t-test is that the sample is “reasonably” normally distributed.
- Why do we need to test the assumption of normality?
 - Think about if our distribution is right skewed, then our *mean* may not reflect the central tendency of our sample population.

Example: One sample t-test

Assumptions

Small Samples: With small samples (typically $n < 30$), even slight deviations from normality can significantly impact the t-test's validity because the t-test relies heavily on the assumption of normality in this scenario. However, if the data are symmetrically distributed and don't have outliers, the t-test can be robust to mild violations of normality. Note that the t-test adjusts for the sample size, so the larger the sample, the less important the normality assumption becomes.

Large Samples: For larger samples ($n > 30$), the Central Limit Theorem (CLT) suggests that the distribution of the sample means tends to be normal, regardless of the population distribution. This means that for sufficiently large samples, the t-test can be robust to violations of normality. However, extreme skewness, heavy tails, or outliers can still affect the test's performance.

Example: One sample t-test

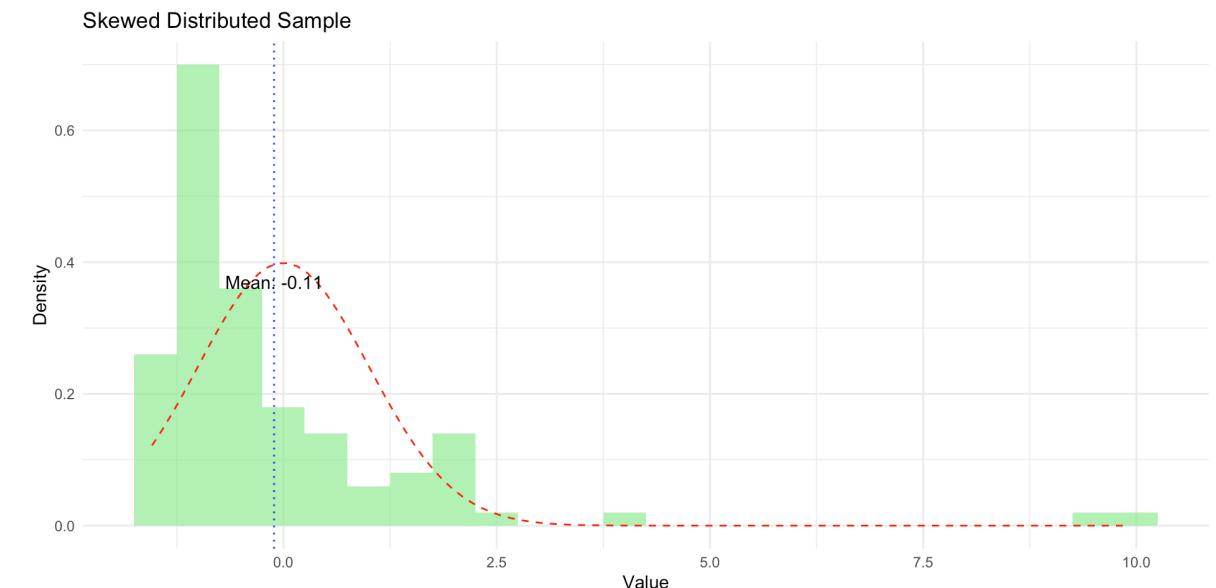
Assumptions

```

1 library(ggplot2)
2
3 # Assuming 'skewed_sample' is your skewed data. For this example
4 set.seed(0) # For reproducibility
5 skewed_sample <- rlnorm(100, meanlog = 0, sdlog = 1) - exp(-0.11)
6
7 # Create a data frame for ggplot
8 data <- data.frame(value = skewed_sample)
9
10 # Plot
11 p <- ggplot(data, aes(x=value)) +
12   geom_histogram(aes(y=..density..), binwidth = 0.5, fill="lightgreen", color="black") +
13   stat_function(fun=dnorm, args=list(mean=0, sd=1), color="red", linetype="dashed") +
14   labs(title="Skewed Distributed Sample", x="Value", y="Density") +
15   theme_minimal() +
16   geom_vline(aes(xintercept=mean(value)), color="blue", linetype="solid") +
17   annotate("text", x=mean(skewed_sample), y=0.35, label=p[[1]])

```

```
1 print(p)
```



Example: One sample t-test

Assumptions

- Let's look at our mung bean trial data: We have data from 21 plots from a *completely randomised design* (look this up) done on our hippie friends farm. CBD treatment was used on all plots.

```
1 yield <- c(0.6, 0.8, 0.8, 0.9, 1.0, 1.1, 1.2, 1.2, 1.3, 1.3, 1.4, 1.4, 1.5, 1.6, 1.6, 1.6, 1.7, 1.9, 1.9, 1.9, 2.0, 2.1, 2.2)
```

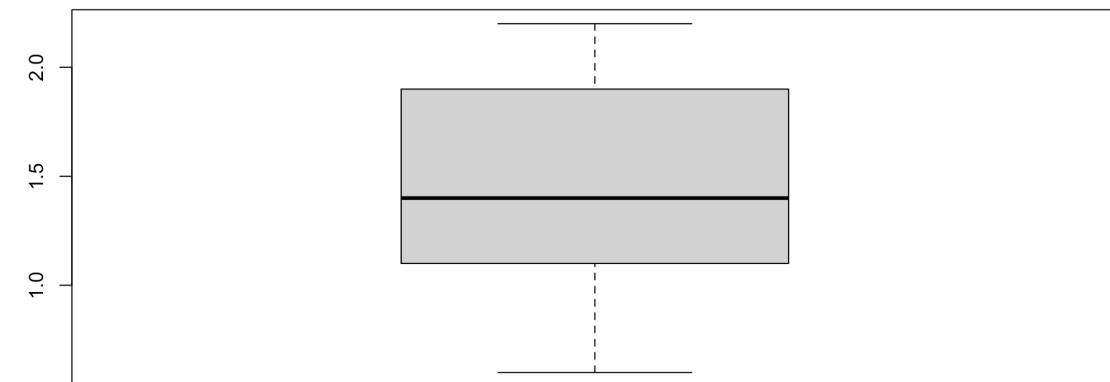
- We have a small sample size so often we use a *stem plot* or *box plot* to look at the distribution. Both look quite symmetrical!

```
1 stem(yield)
```

The decimal point is at the |

0		6889
1		0122334
1		5667999
2		012

```
1 boxplot(yield)
```



Example: One sample t-test

Assumptions

- We can also look at the 6 number summary and the histogram. The histogram looks quite symmetrical and the 6 number summary is quite symmetrical.

```
1 summary(yield)
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.600	1.100	1.400	1.429	1.900	2.200

- We can see that the mean and the median are quite similar. This is a good sign that the data is symmetrical and perhaps normally distributed.
- We can also start to think how our hippies mean yield compares to the national average of 1.1 t/ha.

Example: One sample t-test

Assumptions

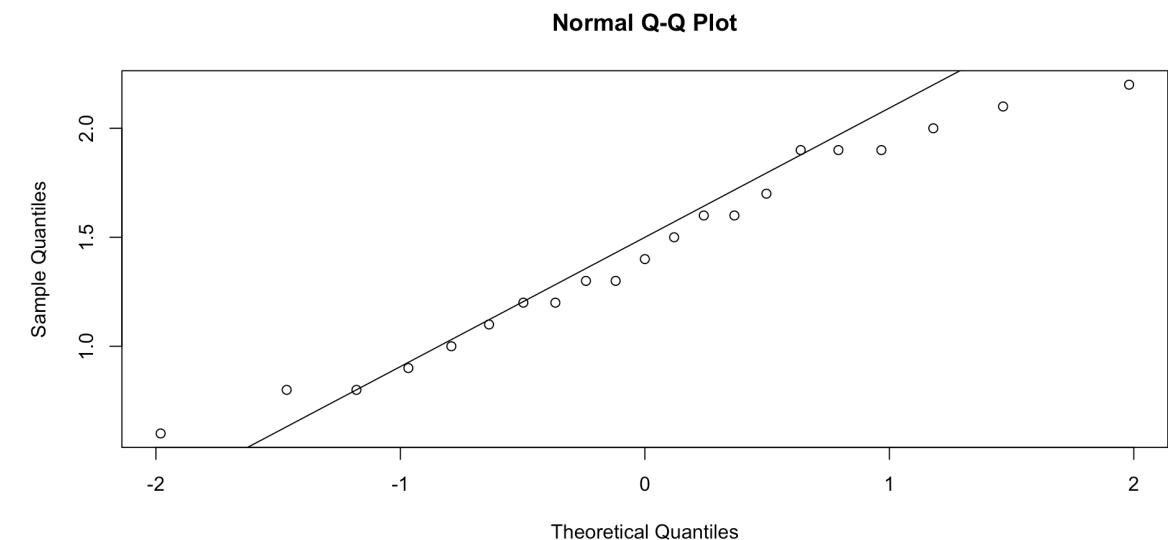
- We can also use a hypothesis test called the *Shapiro Test* and use a plot called a *qqplot* (“Quantile Quantile plot”).
- We can see that the p-value of the shapiro test > 0.05 - no strong evidence that the distribution is not normal.
- We can see that the points on the Q-Q plot are reasonably straight and follow the “line”.

```
1 shapiro.test(yield)
```

```
1 qqnorm(yield)
2 qqline(yield)
```

Shapiro-Wilk normality test

```
data: yield
W = 0.96569, p-value = 0.6371
```



Example: One sample t-test

Test statistic

t-test

- The formula is the same as the z-test but we use the t-distribution instead of the normal distribution and we use the sample standard deviation instead of the population standard deviation.

$$t = \frac{\bar{y} - \mu}{s/\sqrt{n}} \text{ population variance unknown (estimate } s^2\text{)}$$

- \bar{y} is the sample mean
- μ is the population mean
- s is the sample standard deviation
- n is the sample size

Example: One sample t-test

CI & T-Statistic & P-value

Luckily for us we have a function in RStudio that does all of the above in one function!

```
1 t.test(yield, mu=1.1,  
2         alternative="two.sided",  
3         conf.level=0.95)
```

One Sample t-test

```
data: yield  
t = 3.2238, df = 20, p-value = 0.004258  
alternative hypothesis: true mean is not equal to 1.1  
95 percent confidence interval:  
 1.215969 1.641174  
sample estimates:  
mean of x  
 1.428571
```

We can see the following output:

- the t-statistic = 3.2238
- the degrees of freedom = 20 ($n-1$)
- the p-value = 0.004
- the 95% confidence interval = (1.22, 1.64)
- the sample mean = 1.43

Example: One sample t-test

Conclusion



Created using DALL.E 3

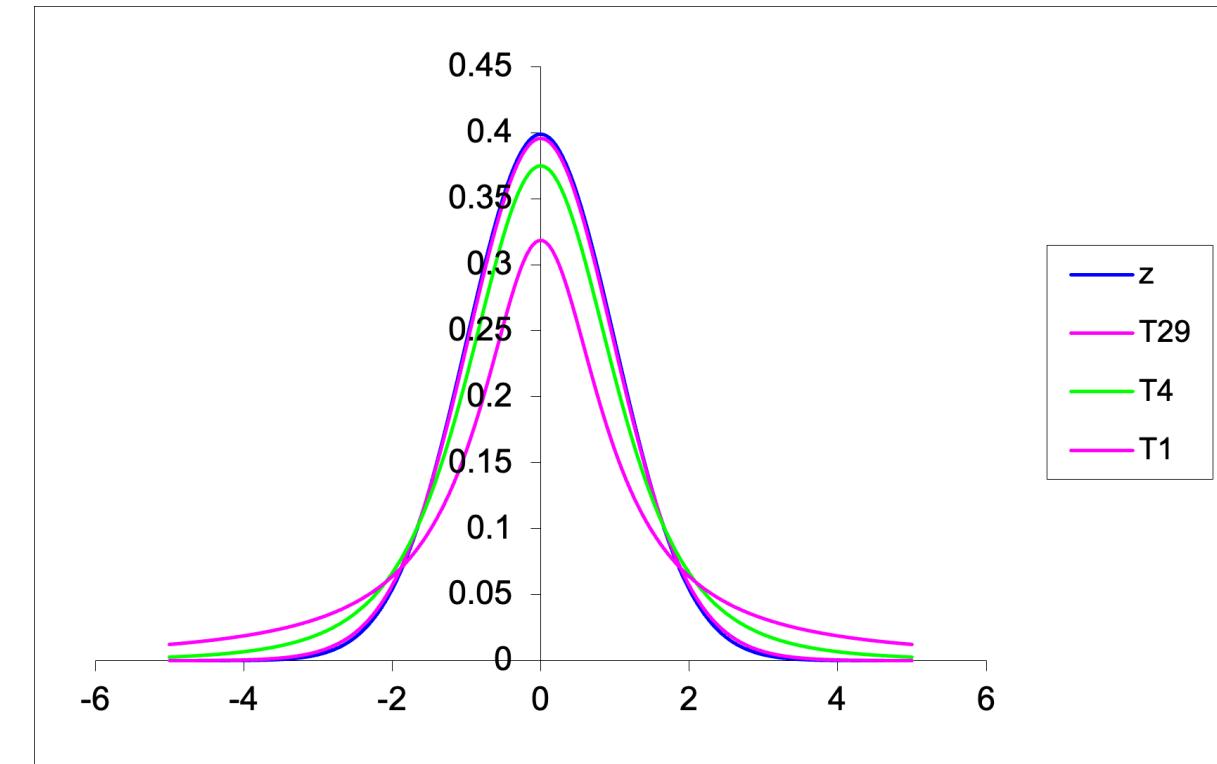
Let's revisit our hippie friend's hypothesis

- $H_0 = 1.1 \text{ t/ha}$ (Null hypothesis)
- $H_1 \neq 1.1 \text{ t/ha}$ (Alternate hypothesis)
- The p-value is 0.004 which is less than 0.05. We reject the null hypothesis.
- There is strong evidence that the mean yield of the mung beans is different to 1.1 t/ha. Given that the mean yield is 1.43 t/ha and we can be 95% confident that the true mean yield is between 1.22 and 1.64 t/ha. We are pretty confident that the CBD oil has increased the yield such that the mean yield is significantly greater than the national average.

The t-distribution and degrees of freedom?

The T-distribution has the following properties:

- It is bell-shaped, symmetrical about the mean, median and mode, which are all of equal value. [It is a little flatter than the normal distribution.]
- The area under the curve = 1, as is the case for all continuous probability distributions.
- The probability density function is defined by three parameters, the mean μ , the standard deviation σ and the sample size n .
- The exact shape of the t distribution depends on the quantity called degrees of freedom, df. The $df = n - 1$ for any t distribution.



Comparing the shapes of the Student's T and the Z (normal) curve

A little more on the P-value

Why is the p-value 0.05?

- The p-value is the probability of observing a test statistic as extreme as the one calculated, assuming the null hypothesis is true.
- It is based on the idea we 1 in 20 times we will reject the null hypothesis when it is true.

A little more on confidence intervals - Unknown variance

- We can also calculate a confidence interval when we do not know the population variance.
- The formula for the confidence interval is:

$$CI = \bar{y} \pm t_{\alpha/2} \cdot \frac{s}{\sqrt{n}}$$

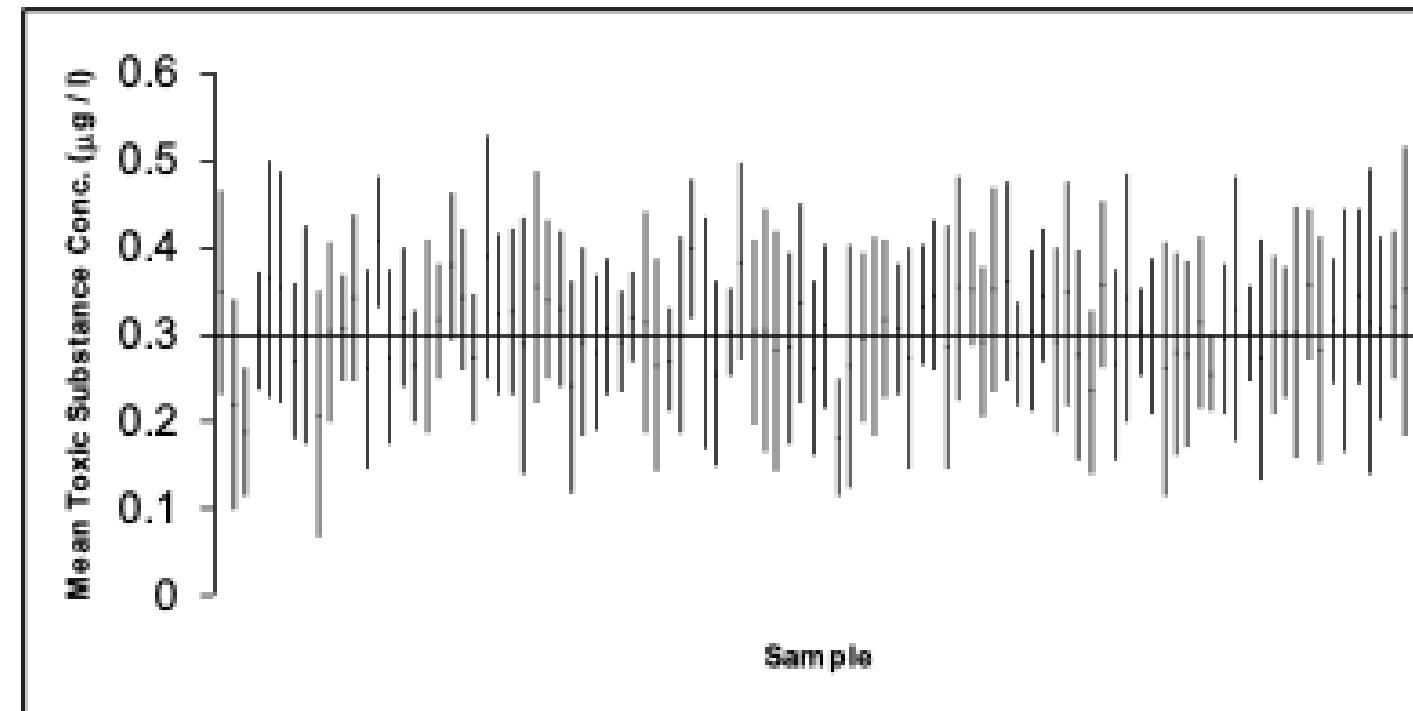
- For the 95% confidence interval, $\alpha = 0.05$ and $t_{\alpha/2} = 2.093$ (for 20 degrees of freedom)
- The degrees of freedom is $n - 1$

```
1 qt(0.975, df=20)
```

```
[1] 2.085963
```

A little more on confidence intervals - Unknown variance

- Another interesting result to contemplate can be demonstrated through the following simulation of 100 studies, each containing $n = 6$ observations of a fictitious toxic substance concentration ($\mu\text{g/l}$) assumed to be $\sim N(0.3, 0.12)$. For each sample the 95% confidence interval calculated:



Depiction of confidence intervals from each of 100 simulated samples

A little more on confidence intervals - Unknown variance

- Approximately 95% of these simulated samples have a confidence interval that includes the true value of 0.3 ($\mu\text{g/l}$). In the graph above, a confidence interval includes the true mean value of 0.3 if the vertical line (representing the width of the CI) crosses the horizontal line.
- We have to accept that 5% of the time we will not capture the true mean.
- We can widen the confidence interval to 99% so we are more confident but this can increase the chance of a type II error.

Example: One sample test

What happens when the assumptions are not met?

1. We can try and transform the data to make it more normal (and equal variance for later)
2. We can use a non-parametric test (later)

Data transformations

My data is not normal :(what can I do? - Sometimes we can transform our data to make it more normal. - Let's look at an example: - Total nitrogen (TN) levels @ Wallacia in western Sydney on the Nepean River - According to the ANZECC guidelines, the maximum acceptable level of TN in an lowland river is $500 \mu\text{g}/L$ and for an upland river = $250 \mu\text{g}/l$. (see Table 3.3.2 in Australian and New Zealand Guidelines for Fresh and Marine Water Quality)

```
1 TN <- read.csv("data/TN_Wallacia.csv")
2 str(TN)
```

```
'data.frame': 29 obs. of 1 variable:
 $ TN: int 1020 1120 1170 920 920 1010 850 910 800 710 ...
```

```
1 mean(TN$TN, na.rm=TRUE) # na.rm=TRUE if we have missing values
```

```
[1] 855.8621
```

```
1 median(TN$TN, na.rm=TRUE)
```

```
[1] 800
```

State Hypothesis

- $H_0 : \mu = 500 \mu\text{g/l}$ (Null hypothesis)
- $H_1 : \mu \neq 500 \mu\text{g/l}$ (Alternate hypothesis)



Check assumptions

```

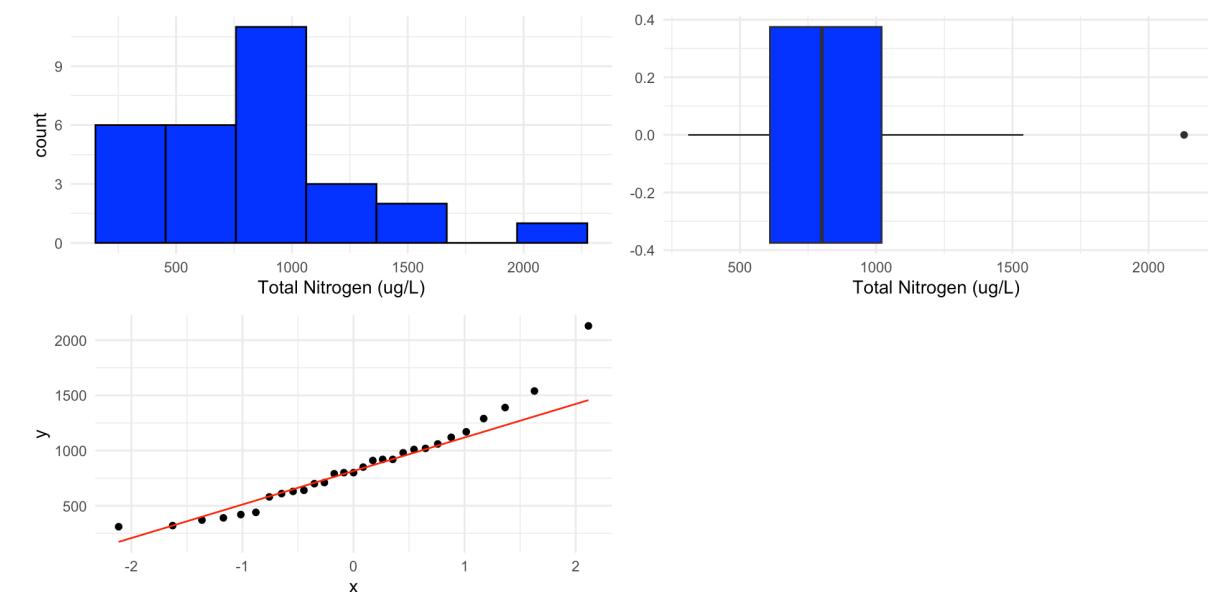
1 library(ggplot2)
2 library(gridExtra)
3
4 # Histogram
5 p1 <- ggplot(TN, aes(x = TN)) +
6   geom_histogram(fill = "blue", color = "black",
7                 bins=7) +
8   xlab("Total Nitrogen (ug/L)") +
9   theme_minimal()
10
11 # Boxplot
12 p2 <- ggplot(TN, aes(x = TN)) +
13   geom_boxplot(fill = "blue") +
14   xlab("Total Nitrogen (ug/L)") +
15   theme_minimal()
16
17 # Q-Q plot
18 p3 <- ggplot(TN, aes(sample = TN)) +
19   stat_qq() +
20   stat_qq_line(col = "red") +
21   theme_minimal()

```

```

1 # Arrange the plots in a 2x2 layout
2 grid.arrange(p1, p2, p3, ncol = 2)

```



Check assumptions

Shapiro-Wilk normality test

- For smaller sample sizes, use the Shapiro-Wilk test (say < 50)

```
1 shapiro.test(TN$TN)
```

Shapiro-Wilk normality test

```
data: TN$TN
W = 0.92582, p-value = 0.04293
```

- We see P<0.05 so we **reject** the null hypothesis that the data is normally distributed.
- For really large samples (>>30) the central limit theorem suggests that the distribution of the sample means tends to be normal, regardless of the population distribution. This means that for sufficiently large samples, the t-test can be robust to violations of normality. However, extreme skewness, heavy tails, or outliers can still affect the test's performance.

Transformations

For right (positive) skewed data, we can use a

- $1/x$ inverse transformation for highly skewed data.
- \log_{10} or \log_e transformation for very skewed data.
- \sqrt square root transformation for moderately skewed data.

For left (negative) skewed data (which is rare to find), we can “try” a

- x^2 or x^3 transformation.

Transformations & Recheck assumptions

```
1 # Let's try a log10 transformation
2 # We start by creating a new column in our data frame called log10_TN
3 # We then take the log10 of the TN column and store it in the new column
4 TN$log10_TN <- log10(TN$TN)
```

Geometric mean

Note that now we are looking at the geometric mean as opposed the the arithmetic mean which was 855.86 in our case

```
1 10^mean(TN$log10_TN)
```

[1] 770.0666

Recheck assumptions

```

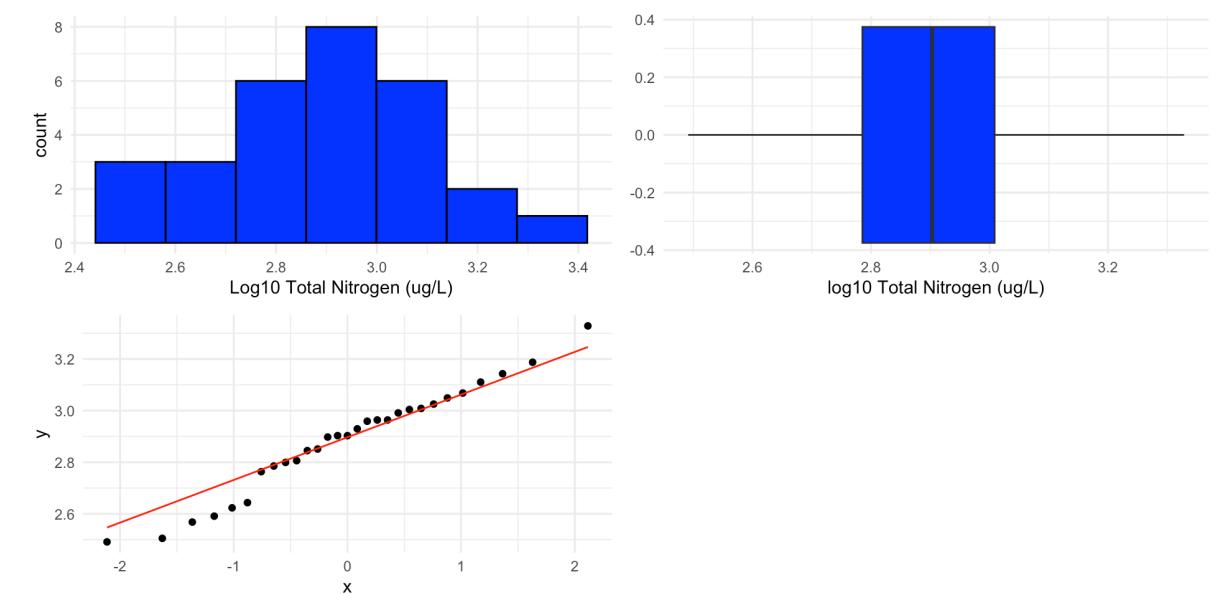
1 library(ggplot2)
2 library(gridExtra)
3
4 # Histogram
5 p1 <- ggplot(TN, aes(x = log10_TN)) +
6   geom_histogram(fill = "blue", color = "black",
7                 bins=7) +
8   xlab("Log10 Total Nitrogen (ug/L)") +
9   theme_minimal()
10
11 # Boxplot
12 p2 <- ggplot(TN, aes(x = log10_TN)) +
13   geom_boxplot(fill = "blue") +
14   xlab("log10 Total Nitrogen (ug/L)") +
15   theme_minimal()
16
17 # Q-Q plot
18 p3 <- ggplot(TN, aes(sample = log10_TN)) +
19   stat_qq() +
20   stat_qq_line(col = "red") +
21   theme_minimal()

```

```

1 # Arrange the plots in a 2x2 layout
2 grid.arrange(p1, p2, p3, ncol = 2)

```



Recheck assumptions

Shapiro-Wilk normality test

```
1 shapiro.test(TN$log10_TN)
```

```
Shapiro-Wilk normality test
```

```
data: TN$log10_TN  
W = 0.97279, p-value = 0.6375
```

- From our previous slide, our data seems to be quite symmetrical
- Now $P>0.05$ so we **fail to reject** the null hypothesis that the data is normally distributed.
- We can now proceed with our t-test using the log transformed data.
- **NOTE:** *We will also need to transform the hypothesised mean of 500 to the log10 scale.*

T-test & P-Value

```
1 t.test(TN$log10_TN, mu=log10(500),  
2       alternative="two.sided",  
3       conf.level=0.95)
```

One Sample t-test

```
data: TN$log10_TN  
t = 4.8768, df = 28, p-value = 3.884e-05  
alternative hypothesis: true mean is not equal to 2.69897  
95 percent confidence interval:  
 2.807748 2.965309  
sample estimates:  
mean of x  
 2.886528
```

- The p-value is <0.001 (3.d.p.) which is less than 0.05. Therefore, we reject the null hypothesis.
- There is strong evidence that the mean value of TN in Nepean River is different to 500 ug/L.
- Can we say something about the direction??

Confidence intervals

- We need to back transform the confidence interval to the original scale.
- To back-transform the confidence interval we can use the following:

$10^{CI_{low}}$ and $10^{CI_{high}}$

```
1 10^(2.81)
```

[1] 645.6542

```
1 10^(2.97)
```

[1] 933.2543

- This means that we are 95% confident that the geometric mean (back-transformed mean) of the sample is between approximately 646 and 933 ug/l. This is higher than the hypothesised mean.

Conclusion

1. The data was log10 transformed to meet the assumptions of the t-test.
2. We have strong evidence that the mean value of Total Nitrogen in Nepean River is different to 500 ug/L.
3. We are 95% confident that the geometric mean (back-transformed mean) of the sample is between approximately 646 and 933 ug/l.
4. Looking at the confidence interval, we can say that the mean value of TN in Nepean River is significantly greater than 500 ug/L.

One tailed test

- We can also use a one tailed test if we are specifically interested in whether the observed value is greater or less than the expected value.
- In this case we only want to know if the TN concentration is greater than the ANZECC guidelines of $500 \mu\text{g/l}$.
- $H_0 : \mu = 500 \mu\text{g/l}$ (Null hypothesis)
- $H_1 : \mu > 500 \mu\text{g/l}$ (Alternate hypothesis)

One tailed test

```
1 t.test(TN$log10_TN, mu=log10(500),  
2       alternative="greater",  
3       conf.level=0.95)
```

```
One Sample t-test

data: TN$log10_TN
t = 4.8768, df = 28, p-value = 1.942e-05
alternative hypothesis: true mean is greater than 2.69897
95 percent confidence interval:
 2.821104      Inf
sample estimates:
mean of x
 2.886528
```

- $P < 0.001$ (3.d.p.) so we reject the null hypothesis.
- Because we are doing a one tailed test we can now conclude that the mean value of TN in Nepean River is greater than 500 ug/L.
- Take a look at the confidence interval and the p-value. What do you notice when comparing to the two tailed test?

One tailed test

Ruxton and Neuhäuser 2010

Methods in Ecology and Evolution



British Ecological Society

Methods in Ecology and Evolution 2010, 1, 114–117

doi: 10.1111/j.2041-210X.2010.00014.x

FORUM

When should we use one-tailed hypothesis testing?

Graeme D. Ruxton^{1*} and Markus Neuhäuser²

¹*Division of Ecology and Evolutionary Biology, Faculty of Biomedical & Life Sciences, University of Glasgow, Glasgow G12 8QQ, UK; and* ²*Department of Mathematics and Technique, RheinAhrCampus, Koblenz University of Applied Sciences, Südallee 2, 53424 Remagen, Germany*

References

- Quinn G. P. & Keough M. J. (2002) Experimental design and data analysis for biologists. Cambridge University Press, Cambridge, UK.
- Logan, M. (2010). Biostatistical design and analysis using R a practical guide. Hoboken, N.J., Wiley-Blackwell.
- Fox, G. A., S. Negrete-Yankelevich, and V. J. Sosa. (2015). Ecological statistics: contemporary theory and application. Oxford University Press, USA
- Ruxton, G.D. and Neuhauser, M., 2010. When should we use one-tailed hypothesis testing? Methods in Ecology and Evolution, 1 (2), 114-117.

Thanks!

This presentation is based on the [SOLES Quarto reveal.js template](#) and is licensed under a [Creative Commons Attribution 4.0 International License](#).



THE UNIVERSITY OF SYDNEY