

# Lab 03

In this week's lab we will work through the experimental design workflow for a one-way ANOVA.

## Learning outcomes

In this lab, you will work towards achieving learning outcomes...

## Lab objectives

At the end of the lab, students should be able to:

- Understand the workflow involved in conducting a one-way Analysis of Variance (ANOVA), including data exploration, model fitting, and interpretation of results.
- Analyse an experiment with a 1-way ANOVA and interpret the results.
- Explain situations when a 2-sample t-test gives the same results as a 1-way ANOVA.

## Prerequisites

The following packages are used for this lab: `readxl`, `ggplot2`, `dplyr` and `tidyverse`. To install these packages, run the following code **in the console**.

```
CODE  
install.packages(c("readxl", "ggplot2", "dplyr", "tidyverse"))
```

## Exercise 1 – diatoms

Medley & Clements (1998) sampled 34 locations along streams for diversity of diatoms. Each site was classified according to the Zn concentration in the water. There were 4 classes; background, low, medium and high. Were there *differences* between each of the groupings in term of diatom diversity? Let's find out.

## Practice 1 – Data import

Import the `Diatoms` worksheet from the `diatoms.xlsx` file into R.

If you have difficulty with this step please refer to this week's tutorial on importing MS Excel files.

```
CODE
library(readxl)

# You will need to look at the Excel file and work out
# the correct worksheet name and range:
diatoms ← read_excel(
  path = "diatoms.xlsx",
  sheet = ...,
  range = ...
)
```

## Workflow

We will briefly go through a typical analytical workflow which involves data exploration, model fitting, checking of model assumptions<sup>1</sup> and interpretation of the results.

## Data exploration

### Data structure

Checking the structure of the data is the first step in any data analysis. This is done using the `str()` function in R.

```
CODE
str(diatoms)
```

Of particular interest are the variables `Stream` and `Zinc`, which have been classified as characters (`chr`).

```
OUTPUT
tibble [34 × 4] (S3: tbl_df/tbl/data.frame)
$ Stream    : chr [1:34] "Eagle" "Blue" "Blue" "Blue" ...
$ Zinc      : chr [1:34] "BACK" "BACK" "BACK" "BACK" ...
$ Diversity: num [1:34] 2.27 1.7 2.05 1.98 2.2 1.53 0.76 1.89 1.4 2.18 ...
$ Group     : num [1:34] 1 1 1 1 1 1 1 2 2 ...
```

These variables are most likely factors and should be converted to such. This can be done using `as.factor()`.

```
CODE
```

---

<sup>1</sup>Assumptions will be covered formally next week when we look at model residuals.

```
diatoms$Zinc ← as.factor(diatoms$Zinc)  
diatoms$Stream ← as.factor(diatoms$Stream)
```

**The tidyverse approach** – We can use the `mutate()` function to convert the Zinc and Stream variables to factors.

```
CODE  
library(dplyr)  
diatoms <- diatoms %>%  
  mutate(  
    Zinc = as.factor(Zinc),  
    Stream = as.factor(Stream))  
}
```

We can then check if the conversion was successful by using the `str()` function again.

CODE

```
OUTPUT

tibble [34 x 4] (S3: tbl_df/tbl/data.frame)
$ Stream    : Factor w/ 6 levels "Arkan","Blue",..: 4 2 2 2 5 6 6 1 1 ...
$ Zinc       : Factor w/ 4 levels "BACK","HIGH",..: 1 1 1 1 1 1 1 1 3 3 ...
$ Diversity: num [1:34] 2.27 1.7 2.05 1.98 2.2 1.53 0.76 1.89 1.4 2.18 ...
$ Group     : num [1:34] 1 1 1 1 1 1 1 1 2 2 ...
```

## Summary statistics

Use `summary()` for a quick overview of common statistical measures for each variable in the data frame. However it is not informative when we are interested in differences between groups or factors as it only gives the summary statistics for the entire data set.

## CODE

Stream	Zinc	Diversity	Group
Arkan:7	BACK:8	Min. :0.630	Min. :1.000
Blue:7	HIGH:9	1st Qu.:1.377	1st Qu.:2.000
Chalk:5	LOW :8	Median :1.855	Median :3.000
Eagle:4	MED :9	Mean :1.694	Mean :2.559
Snake:5		3rd Qu.:2.058	3rd Qu.:3.750
Splat:6		Max. :2.830	Max. :4.000

It is more useful to calculate summary statistics for each level of Zn contamination, as we are interested in differences between the mean of each group.

We can use the `tapply()` function to calculate the mean and standard deviation for each level of Zinc. This has to be done separately for each summary statistic (mean and standard deviation).

The general structure of the `tapply()` function is 3 arguments which are described below based on the code above:

- the response variable on which we wish to apply the function, `diatoms$Diversity`;
- the categorical variable which indicates the groups we wish to separately apply the function to, `diatoms$Zinc`;
- the function we are using, `mean()`.

```
CODE
tapply(
  X = diatoms$Diversity,
  INDEX = diatoms$Zinc,
  FUN = mean
)

OUTPUT
BACK      HIGH      LOW      MED
1.797500 1.277778 2.032500 1.717778
```

```
CODE
tapply(
  X = diatoms$Diversity,
  INDEX = diatoms$Zinc,
  FUN = sd
)

OUTPUT
BACK      HIGH      LOW      MED
0.4852613 0.4268717 0.4449960 0.5030104
```

### 💡 Tip

**The tidyverse approach** – Use the `group_by()` and `summarise()` functions to calculate the mean and standard deviation for each level of Zinc. The code is longer, but readable as you can see the sequence of operations.

```
CODE
diatoms %>%
  group_by(Zinc) %>%
  summarise(
    mean = mean(Diversity),
    sd = sd(Diversity)
)
```

## Graphical summaries

Numerical summaries are nice, but visual summaries are often more informative and are the **summary of choice** for publication. They can also help us to check the assumptions of the statistical test (although sometimes this is not possible until we have fitted the model).

In the case of ANOVA model we are interested in the distribution of the response variable for each level of the factor. The table below gives heuristic rules about which graphical summary to use based on the number of observations.

observations	graphics	command
1-5	plot raw data	<code>stripchart()</code> or <code>geom_jitter()</code>
6-20	boxplot	<code>boxplot()</code> or <code>geom_boxplot()</code>
20 or more	histogram	<code>hist()</code> or <code>geom_histogram()</code>

We will show examples of both the histogram and boxplot for the diatom data. We recommend that you use `ggplot2` for your graphical summaries as it is more flexible and has a more consistent syntax than base R graphics. However, we will show both approaches here.

### Histogram

As there are less than 10 observations per group the histogram may not be very informative. However, it may still be useful for checking the normality of the data. You may be able to see the limitations of the histogram for small sample sizes.

#### Tip

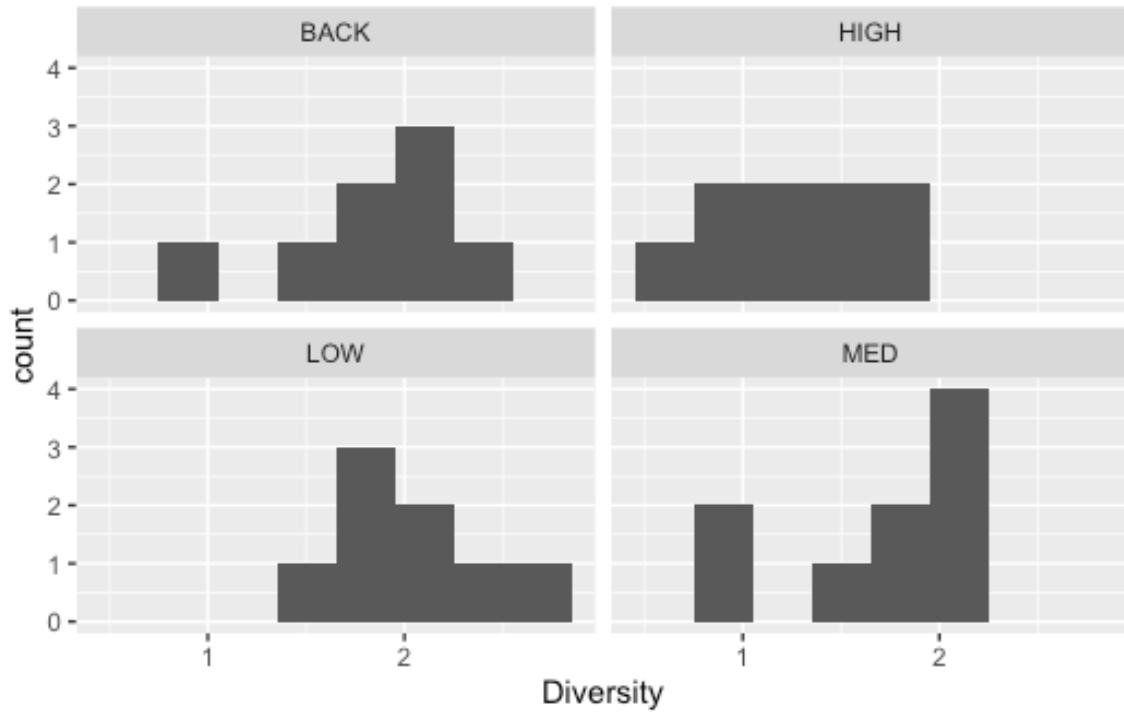
To calculate the number of observations per group:

```
CODE
tapply(
  X = diatoms$Diversity,
  INDEX = diatoms$Zinc,
  FUN = length
)
```

## ggplot2

```
CODE
library(ggplot2)
ggplot(diatoms, aes(x = Diversity)) +
```

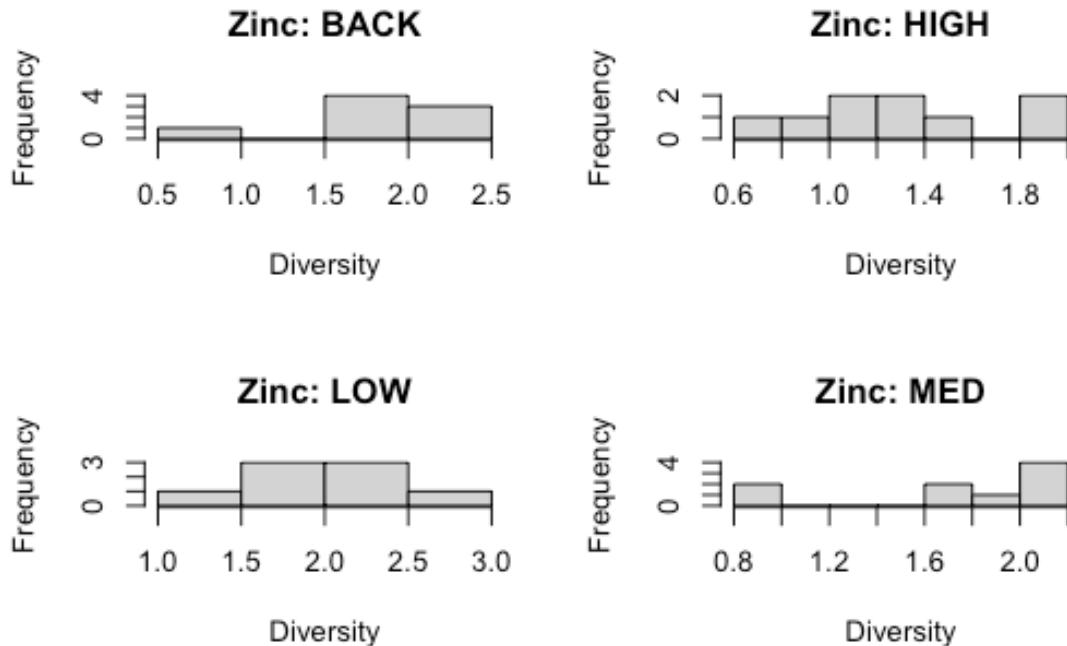
```
geom_histogram(binwidth = .3) +  
facet_wrap(~Zinc)
```



## base R (no loops)

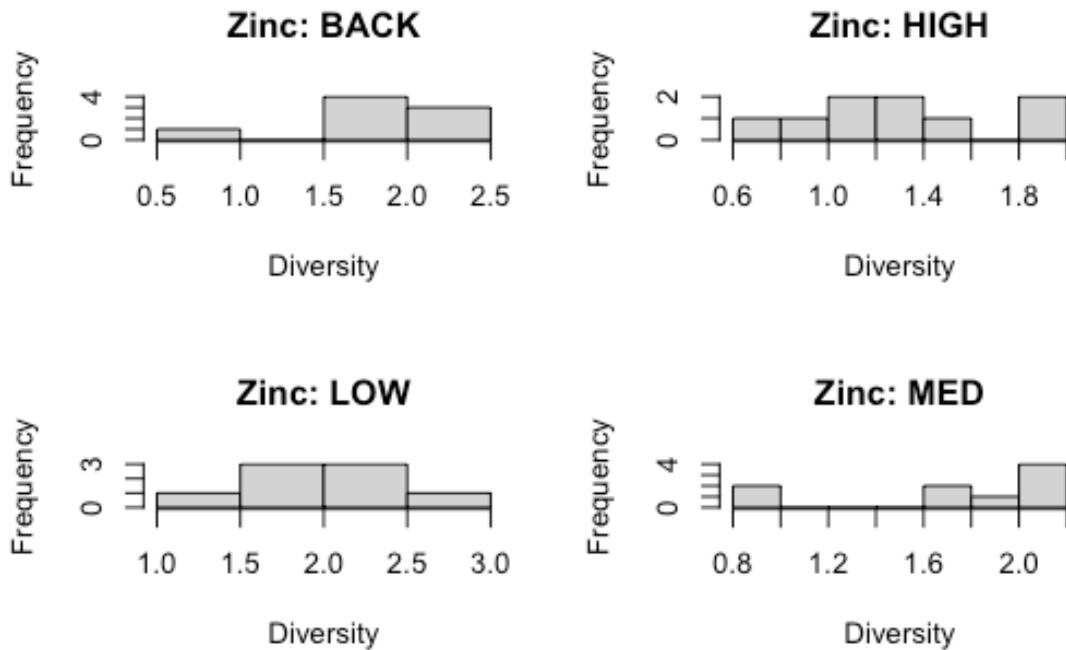
```
CODE  
par(mfrow = c(2, 2))  
hist(  
  x = diatoms$Diversity[diatoms$Zinc == "BACK"],  
  main = "Zinc: BACK",  
  xlab = "Diversity"  
)  
hist(  
  x = diatoms$Diversity[diatoms$Zinc == "HIGH"],  
  main = "Zinc: HIGH",  
  xlab = "Diversity"  
)  
hist(  
  x = diatoms$Diversity[diatoms$Zinc == "LOW"],  
  main = "Zinc: LOW",  
  xlab = "Diversity"  
)  
hist(  
  x = diatoms$Diversity[diatoms$Zinc == "MED"],  
  main = "Zinc: MED",  
  xlab = "Diversity")
```

```
    main = "Zinc: MED",
    xlab = "Diversity"
)
```



## base R (loops)

```
CODE
par(mfrow = c(2, 2))
for (i in levels(diatoms$Zinc)) {
  hist(
    x = diatoms$Diversity[diatoms$Zinc == i],
    main = paste("Zinc:", i),
    xlab = "Diversity"
  )
}
```



### Boxplot

A more appropriate plot for this data is the boxplot. If we want to format the plot for publication we should make the plot clear by adding axis labels and a figure caption.

## base R

```
CODE
boxplot(
  Diversity ~ Zinc,
  data = diatoms,
  ylab = "Diversity",
  xlab = "Zinc concentration"
)
```

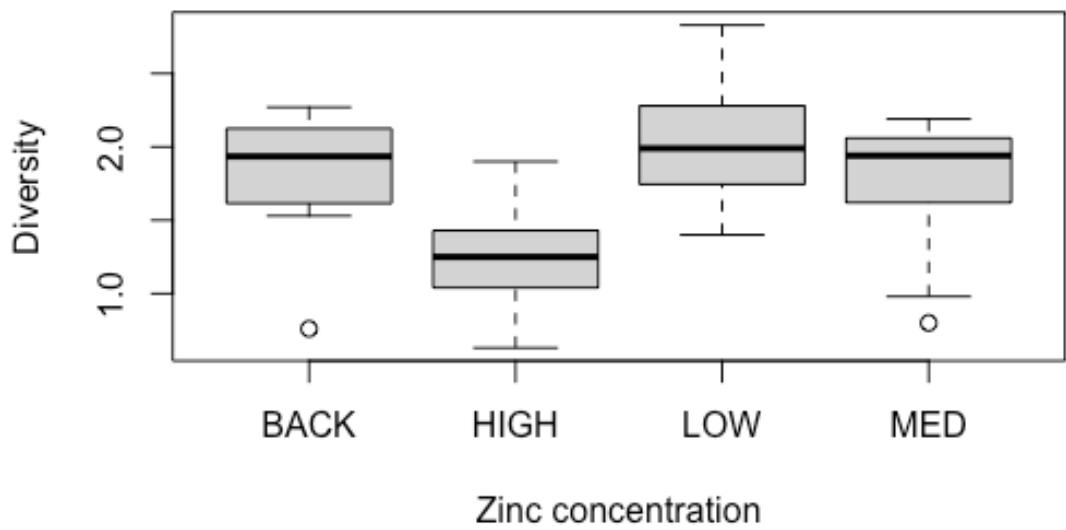


Figure 1: Boxplot of diatom diversity by Zinc concentration.

## ggplot2

```
CODE
ggplot(diatoms, aes(x = Zinc, y = Diversity)) +
  geom_boxplot() +
  ylab("Diversity") +
  theme_minimal(base_size = 12)
```

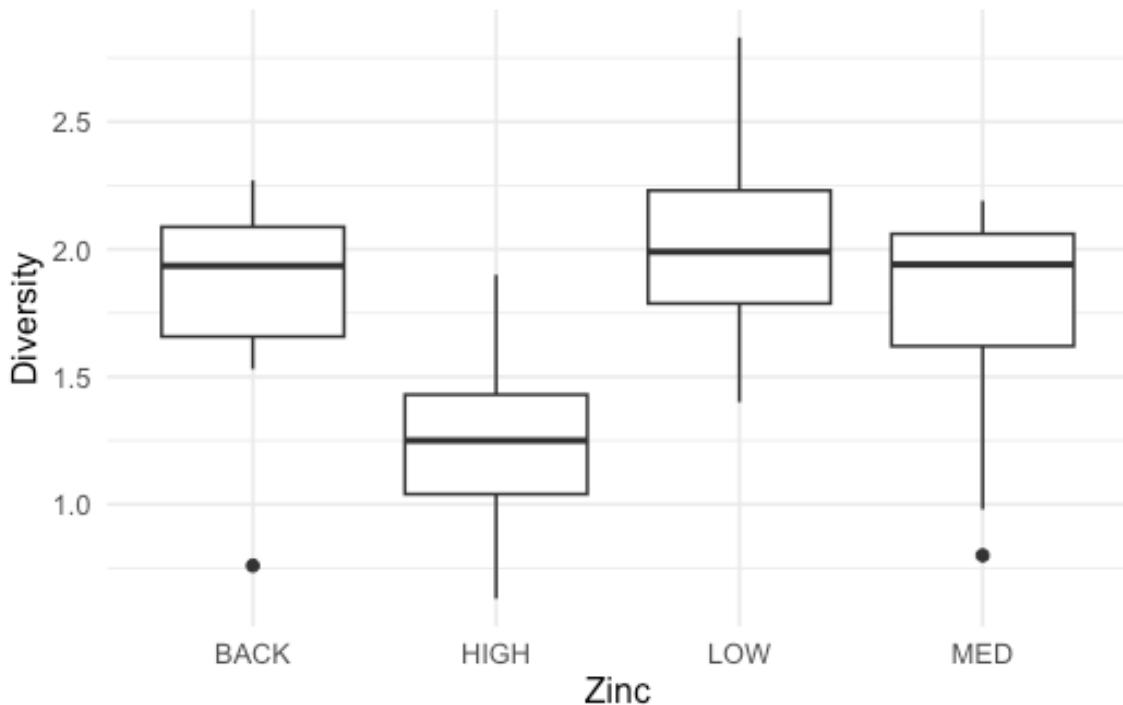


Figure 2: Boxplot of diatom diversity by Zinc concentration.

## Question 1

What can you say about the different levels of Zinc from looking at the mean and standard deviation values when running the following?

```

CODE
tapply(
  X = diatoms$Diversity,
  INDEX = diatoms$Zinc,
  FUN = mean
)

tapply(
  X = diatoms$Diversity,
  INDEX = diatoms$Zinc,
  FUN = sd
)

```

## Solution 1

The mean diatom diversity for the HIGH level of Zn concentration is much smaller than the other groupings indicating that there is something happening, more formal hypothesis testing is needed to determine whether this is significant.

In terms of the variance of each group these are quite uniform.

Note: the assumption of constant variance (more on this next week) is likely to be met. For example, the ratio of the largest standard deviation to the smallest standard deviation (0.503/0.427) is less than 2.

## Question 2

Can you interpret and describe the boxplot above?

## Solution 2

- Based on the median values the smallest diversity is associated with the HIGH level of Zn concentration.
- The  $Q_1$  and  $Q_3$  (the “box part”) for the HIGH level of Zn do not overlap with the others levels of Zn, but the boxes overlap with each other for the other levels. This is some initial evidence that it will be likely that all the different levels of Zn will have the same diversity with the exception of the HIGH level of Zn.

## Model fitting

A 1-way ANOVA involves one treatment (or grouping) factor. The model we are fitting is:

$$y_{i,j} = \mu_i + \epsilon_{i,j}$$

where:

1.  $y_{i,j}$  is the response for observation  $j$  in treatment (or group)  $i$ ,
2.  $\mu_j$  is the mean of treatment (or group)  $i$ ,
3.  $\epsilon_{i,j}$  is the residual term which is an independent random variable that has a mean of 0, constant variance and is normally distributed. This can be expressed shorthand as  $\sim N(0, \sigma^2)$ . The residual MS estimates  $\sigma^2$ .

The statistical hypotheses we are testing are:

$$H_0 : \mu_1 = \mu_2 = \dots = \mu_t$$

$$H_1 : \text{not all } \mu_j \text{ are equal}$$

where  $\mu_j$  is the mean diatom diversity for each level of Zn concentration.

The code below fits the ANOVA model using the `aov()` function and saves the it to an object called `anova.diatoms`. We can then extract the ANOVA table using the `summary` function.

CODE

```
anova.diatoms ← aov(formula = Diversity ~ Zinc, data = diatoms)
summary(anova.diatoms)
```

OUTPUT

	Df	Sum Sq	Mean Sq	F value	Pr(>F)						
Zinc	3	2.567	0.8555	3.939	0.0176 *						
Residuals	30	6.516	0.2172								
---											
Signif. codes:	0	'***'	0.001	'**'	0.01	'*'	0.05	'..'	0.1	' '	1

## Assumptions and interpretation of results

The assumptions of the ANOVA model are:

- The residuals are independent,
- The residuals are normally distributed,
- The residuals have constant variance.

Based on the boxplots and histograms during data exploration, the assumptions of normality and equal variances are met. We will discuss assumptions again in greater detail next week when we start to look at residuals.

**We can only interpret the results of the ANOVA model if the assumptions are met.** We can report the results the following way:

**Reporting:** The results indicate that there are significant differences between the levels of Zn concentration in terms of diatom diversity ( $F = 3.9$ ,  $df = 3, 30$ ,  $P = 0.02$ ).

## Post-hoc testing

The ANOVA test only tells us that there are differences between the groups, but it does not tell us which groups are different. We can use the `emmeans` package to perform post-hoc testing.

```
CODE
library(emmeans)
```

OUTPUT

Welcome to emmeans.  
Caution: You lose important information if you filter this package's results.  
See '? tidy'

```
CODE
posthoc ← emmeans(anova.diatoms, "Zinc")
posthoc
```

OUTPUT

```

Zinc emmean    SE df lower.CL upper.CL
BACK    1.80 0.165 30    1.461    2.13
HIGH    1.28 0.155 30    0.961    1.60
LOW     2.03 0.165 30    1.696    2.37
MED     1.72 0.155 30    1.401    2.04

```

Confidence level used: 0.95

The output of the `emmeans` function gives us the estimated marginal means for each level of Zinc and the 95% confidence intervals. The confidence intervals for the HIGH level of Zn concentration do not overlap with the other levels of Zn concentration, indicating that the HIGH level of Zn concentration is significantly different from the other levels.

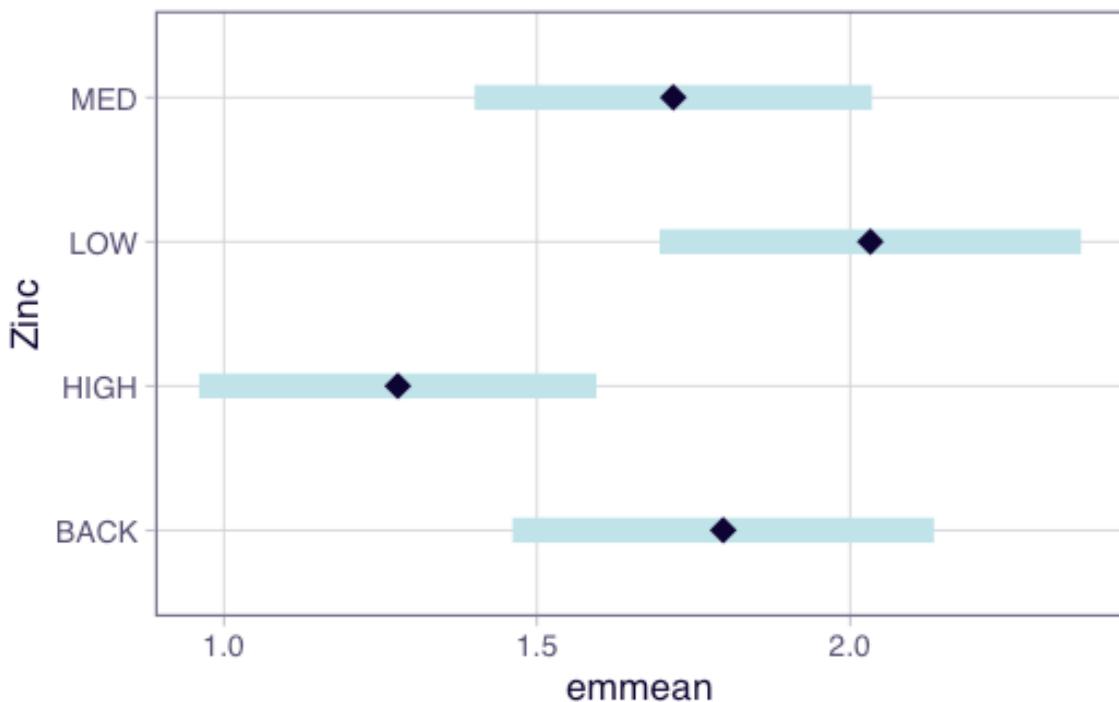
**Reporting:** The post-hoc test indicates that the HIGH level of Zn concentration is significantly different from the other levels of Zn concentration. There are no other significant differences.

We can visualise the results of the post-hoc testing using the `plot` function, if you prefer a visual representation.

```

CODE
plot(posthoc)

```



## Exercise 2 – chicks

An experiment was designed to compare 15-day mean comb weights (g) of two lots of male chicks, one receiving sex hormone A (testosterone), the other C (dehydroandrosterone). While we usually analyse these data by using a (pooled) two-sample t-test, a single factor analysis of variance approach could be used (with two levels of the treatment factor). We will compare the results from both analyses.

The data is found in the *Comb* worksheet of the *chick\_marigold.xlsx* file. Download it below:

Read the data and save it as comb. It should look like below:

```
CODE
comb

OUTPUT
# A tibble: 22 x 2
  CombWt Hormone
     <dbl>   <chr>
1      57    A
2     120    A
3      101   A
4      137   A
5      119   A
6      117   A
7      104   A
8       73    A
9       53    A
10      68    A
# i 12 more rows
```

## Question 3

Perform some checks to verify the two-sample t-test (or one-way ANOVA) is appropriate, i.e. investigate the shape of the distributions, and the standard deviations.

## Solution 3

```
CODE  
library(readxl)  
combs ← read_excel("data/chick_marigold.xlsx", sheet = "Comb")  
str(combs)
```

CODE

```
combs$Hormone ← as.factor(combs$Hormone)
```

CODE

```
aggregate(CombWt ~ Hormone, summary, data = combs)
```

OUTPUT

Hormone	CombWt.Min.	CombWt.1st Qu.	CombWt.Median	CombWt.Mean	CombWt.3rd Qu.
1 A	53.0	70.5	104.0	97.0	118.5
2 C	22.0	31.5	50.0	56.0	85.0

CombWt.Max.

1 A	137.0
2 C	96.0

CODE

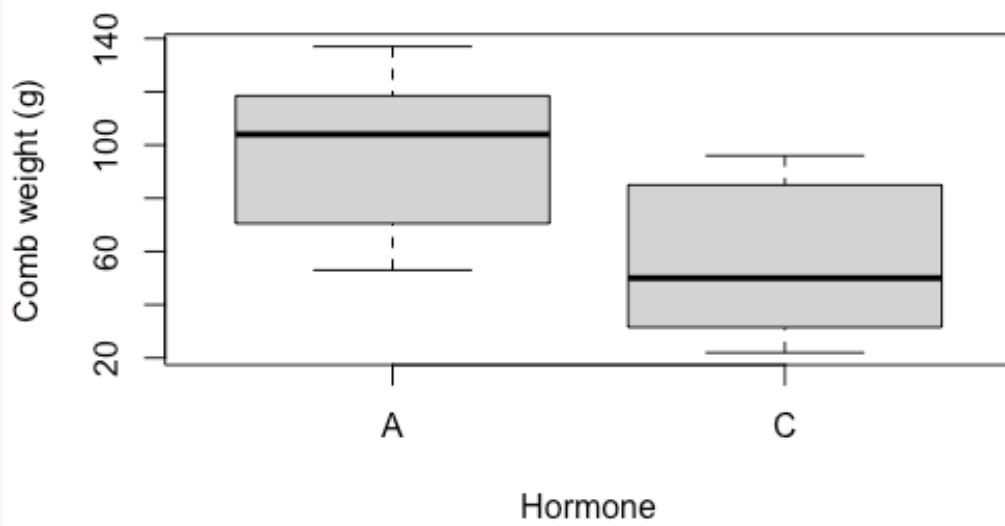
```
aggregate(CombWt ~ Hormone, sd, data = combs)
```

OUTPUT

Hormone	CombWt
1 A	29.10670
2 C	27.83523

CODE

```
boxplot(CombWt ~ Hormone, ylab = "Comb weight (g)", data = combs)
```



Descriptive statistics and boxplot would indicate that the two-sample t-test (or one-way ANOVA) would be appropriate as they appear to meet model assumptions of normality and equal variance.

## Question 4

In R, perform a 2-sample t-test using the `t.test()` function, and use the `var.equal = TRUE` argument to ensure a standard 2-sample t-test is performed. Interpret the output.

## Solution 4

### CODE

```
t.test(CombWt ~ Hormone, var.equal = TRUE, data = combs)
```

### OUTPUT

```
Two Sample t-test

data: CombWt by Hormone
t = 3.3764, df = 20, p-value = 0.003
alternative hypothesis: true difference in means between group A and group C is not equal to 0
95 percent confidence interval:
15.66997 66.33003
sample estimates:
mean in group A mean in group C
      97          56
```

- The t-test ( $t = 3.38$ ,  $df = 20$ ,  $P = 0.003$ ) indicates that the means are significantly different, with the chicks given Hormone A (mean = 97 g) having a significantly higher weight than those given Hormone C (mean 56 g).
- The estimated difference is 41 g, but the 95% CI for this estimate is between 15.7 and 66.3 g (A above C).
- Since this 95% CI does not include zero, this also indicates the means differ significantly.

## Question 5

Next, perform a one-way ANOVA using the `aov` function, followed by `summary`, and interpret the results.

## Solution 5

### CODE

```
combs.aov ← aov(CombWt ~ Hormone, data = combs)
summary(combs.aov)
```

### OUTPUT

```

Df Sum Sq Mean Sq F value Pr(>F)
Hormone      1    9245    9245    11.4   0.003 **
Residuals   20   16220     811
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

The results ( $F = 11.4$ ,  $df = 1, 20$ ,  $P = 0.003$ ) indicates that the means are significantly different, with the chicks given Hormone A (mean = 97 g) having a significantly higher weight than those given Hormone C (mean 56 g).

## Question 6

Compare the t-test and ANOVA outputs. What do you notice about:

- the degrees of freedom;
- the P-value?

## Solution 6

- The  $df$  of  $t$  are 20, the denominator  $df$  for the ANOVA  $F$  statistic is also 20.
- They are the same ( $P$ -value = 0.03) for both tests.

## Question 7

When there are only two treatment groups, the observed  $F$  and  $t$  values are related by  $F = t^2$ . Demonstrate this for the observed values in this exercise.

## Solution 7

When there are only two treatment groups, the observed  $F$ - and  $t$ -values are related by  $F = t^2$ .

Demonstration:  $(t_{obs})^2 = (3.38)^2 = F_{obs}$  of 11.4.

## Exercise 3 – lambs

**Work on this exercise in your own time.**

The levels of immunoglobulin (Ig) in blood serum (g/100 ml) in 3 breeds of newborn lambs have been investigated. A total of 44 lambs were sampled, with approximately equal numbers per breed.

The researcher wants to know whether or not there are significant differences in immunoglobulin levels between the breeds.

The data is found in `lambs.csv`. Download the data if you have not done so:

**Read the data into R and save it as `lambs`.** Use the `read_csv()` function which should be available when you load the `tidyverse` package.

It should look like below:

CODE
<pre>lambs</pre>

OUTPUT
<pre># A tibble: 44 × 2   Ig Breed   &lt;dbl&gt; &lt;dbl&gt; 1 1.1   1 2 2.2   1 3 1.7   1 4 1.4   1 5 1.6   1 6 2.3   1 7 1.4   1 8 1.9   1 9 0.8   1 10 1.6  1 # i 34 more rows</pre>

Based on the demonstrator walkthrough, analyse the `lambs` dataset and test the hypothesis that:

$$H_0 : \mu_1 = \mu_2 = \mu_3$$

$$H_1 : \text{not all } \mu_j \text{ are equal}$$

where  $\mu_j$  is the mean Ig level for breed  $j$ .

Remember to:

- State your null and alternate hypotheses.
- Plot an appropriate summary graph for the data.
- Demonstrate the model fit using the `aov()` function.
- Test assumptions, by checking sd values, or by looking at your exploratory plots (checking residuals is not necessary at this point).
- Report your test statistic, degrees of freedom and p-value.
- Report the statistical conclusion by addressing the null hypothesis.
- Explain the results within a biological context to the data.

## Answer

First we read the data into R and then change the Breed variable to a factor after assessing the structure of the data.

CODE

```
library(readxl)
lambs ← readr::read_csv("data/lambs.csv")
str(lambs)
```

OUTPUT

```
spc_tbl_ [44 × 2] (S3: spec_tbl_df/tbl_df/tbl/data.frame)
$ Ig    : num [1:44] 1.1 2.2 1.7 1.4 1.6 2.3 1.4 1.9 0.8 1.6 ...
$ Breed: num [1:44] 1 1 1 1 1 1 1 1 1 1 ...
- attr(*, "spec")=
  .. cols(
  ..   Ig = col_double(),
  ..   Breed = col_double()
  .. )
- attr(*, "problems")=<externalptr>
```

CODE

```
lambs$Breed ← as.factor(lambs$Breed)
```

Now we do some exploratory data analysis. The main purpose here is to test the assumptions for performing the ANOVA. The `summary()` function selects the appropriate format to summarise for each variable in the data frame `lambs`: quartiles etc for Ig, a continuous variable, and a frequency distribution for levels of Breed, a factor. Note the unequal number of observations per breed.

CODE

```
summary(lambs)
```

OUTPUT

Ig	Breed
Min. :0.800	1:16
1st Qu.:1.600	2:13
Median :2.000	3:15
Mean :1.998	
3rd Qu.:2.400	
Max. :3.000	

We wish to see the summary statistics for each Breed.

CODE

```
tapply(
  X = lambs$Ig,
  INDEX = lambs$Breed,
  FUN = mean
)
```

OUTPUT

```

1      2      3
1.587500 2.230769 2.233333

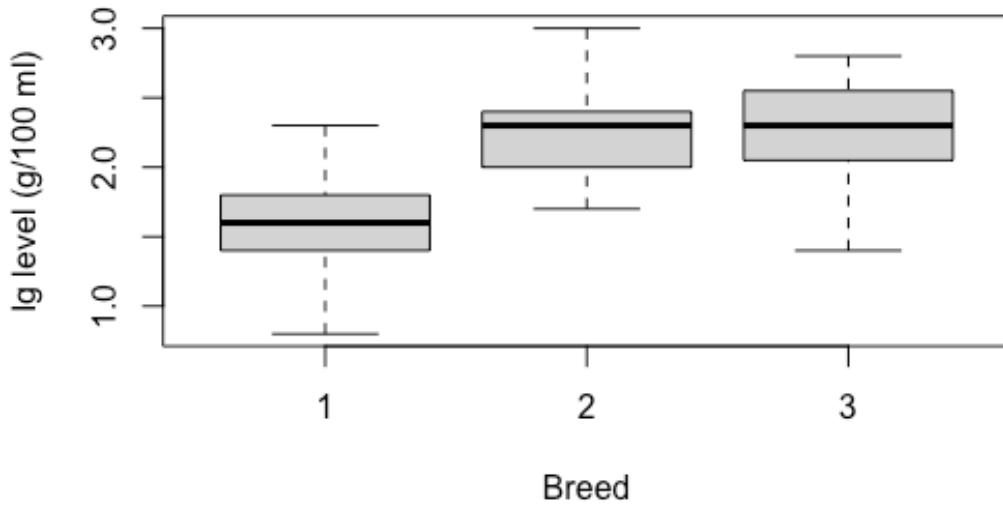
```

Note the median and mean for each Breed are similar which is an indicator of a symmetrical distribution. It is best to look at boxplots and below are these for each Breed. The equal length of the whiskers indicate that we can assume the data is normally distributed.

```

CODE
boxplot(Ig ~ Breed, ylab = "Ig level (g/100 ml)", data = lambs)

```



Now to assess whether the variances are equal.

```

CODE
vars ← tapply(
  X = lambs$Ig,
  INDEX = lambs$Breed,
  FUN = sd
)
vars

```

```

OUTPUT
1      2      3
0.3810074 0.3520999 0.4047339

```

```

CODE

```

```
# ratio of largest to smallest  
max(vars) / min(vars)
```

```
OUTPUT  
[1] 1.149486
```

The ratio of the largest (0.405) to smallest (0.352) is less than 2 so we can assume they are equal.

Now we can perform the ANOVA. The statistical null and alternative hypotheses are:

$$H_0 : \mu_1 = \mu_2 = \mu_3$$

$$H_1 : \text{not all } \mu_j \text{ are equal}$$

where  $\mu_j$  is the mean Ig level for breed  $j$ .

```
CODE  
lambs.aov ← aov(Ig ~ Breed, data = lambs)  
summary(lambs.aov)
```

```
OUTPUT  
Df Sum Sq Mean Sq F value    Pr(>F)  
Breed      2  4.231  2.1156   14.56 1.67e-05 ***  
Residuals  41  5.959  0.1453  
---  
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

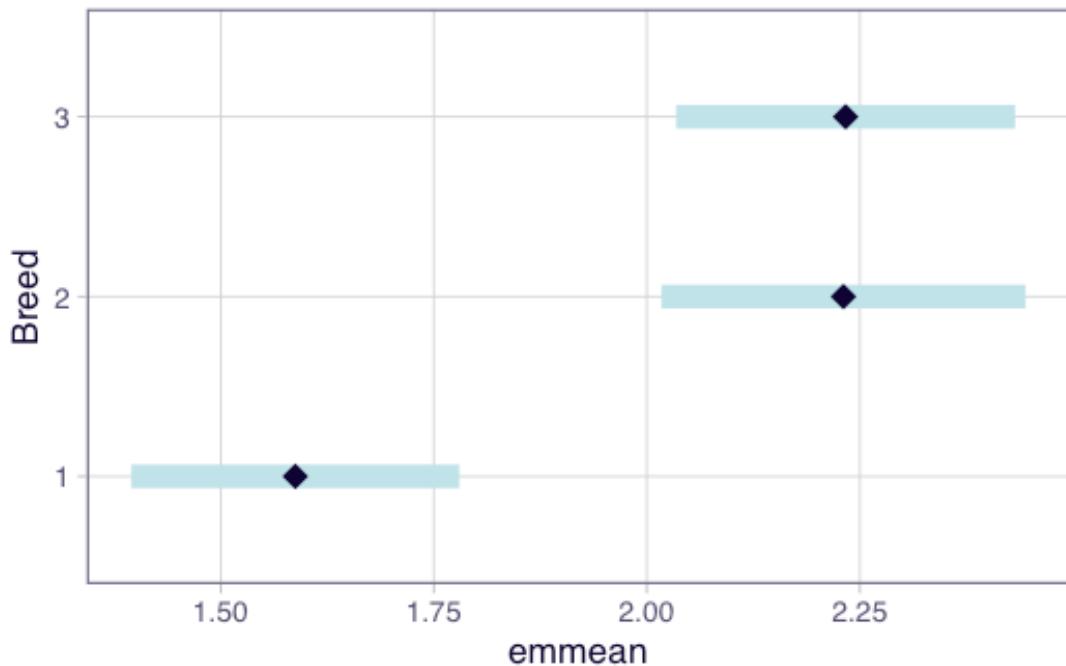
The  $P$ -value is below 0.05 so we reject the null hypothesis. Now explore which pair(s) of means are different.

```
CODE  
library(emmeans)  
emmeans(lambs.aov, "Breed")
```

```
OUTPUT  
Breed emmean     SE df lower.CL upper.CL  
1       1.59 0.0953 41     1.40     1.78  
2       2.23 0.1060 41     2.02     2.44  
3       2.23 0.0984 41     2.03     2.43
```

```
Confidence level used: 0.95
```

```
CODE  
plot(emmeans(lambs.aov, "Breed"))
```



- Based on the ANOVA table our conclusions are there are significant differences between the Ig means between the three breeds of lambs ( $F = 14.6$ ;  $df = 2, 41$ ;  $P < 0.0001$ ).
- The 95% CI of the the mean for Breed 1 does not overlap with the 95% CI of the other means.
- Therefore, Breed 1 has substantially less immunoglobulin ( $1.59 \pm 0.10$  g/100ml) than the other two: Breed 2:  $2.23 \pm 0.11$  g/100ml, and Breed 3:  $2.23 \pm 0.10$  g/100ml (expressed as mean  $\pm$  standard error).

## Thanks!

Did you know you can also knit to PDF? Check the documentation for [R Markdown](#) or [Quarto](#) for more information.

## Attribution

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[Click here for session information](#)

```

CODE
sessionInfo()

OUTPUT
R version 4.5.2 (2025-10-31)
Platform: aarch64-apple-darwin20
Running under: macOS Tahoe 26.2

Matrix products: default
BLAS:  /System/Library/Frameworks/Accelerate.framework/Versions/A/Frameworks/vecLib.framework/Versions/
A/libBLAS.dylib
LAPACK: /Library/Frameworks/R.framework/Versions/4.5-arm64/Resources/lib/libRlapack.dylib; LAPACK
version 3.12.1

locale:
[1] en_AU.UTF-8/en_AU.UTF-8/en_AU.UTF-8/C/en_AU.UTF-8/en_AU.UTF-8

time zone: Australia/Sydney
tzcode source: internal

attached base packages:
[1] stats      graphics    grDevices utils      datasets   methods    base

other attached packages:
[1] emmeans_2.0.0 ggplot2_4.0.2 readxl_1.4.5

loaded via a namespace (and not attached):
[1] bit_4.6.0        rematch_2.0.0     gtable_0.3.6      jsonlite_2.0.0
[5] crayon_1.5.3     dplyr_1.2.0       compiler_4.5.2    tidyselect_1.2.1
[9] parallel_4.5.2    scales_1.4.0      yaml_2.3.12      fastmap_1.2.0
[13] lattice_0.22-7   coda_0.19-4.1   readr_2.1.6      R6_2.6.1
[17] labeling_0.4.3    generics_0.1.4   knitr_1.51       tibble_3.3.1
[21] tzdb_0.5.0       pillar_1.11.1    RColorBrewer_1.1-3 rlang_1.1.7
[25] utf8_1.2.6       xfun_0.56       S7_0.2.1        bit64_4.6.0-1
[29] otlr_0.2.0       estimability_1.5.1 cli_3.6.5      withr_3.0.2
[33] magrittr_2.0.4    digest_0.6.39    grid_4.5.2       vroom_1.6.7
[37] xtable_1.8-4     mvtnorm_1.3-3   hms_1.1.4       lifecycle_1.0.5
[41] vctrs_0.7.1      evaluate_1.0.5  glue_1.8.0      farver_2.1.2
[45] cellranger_1.1.0 rmarkdown_2.30   tools_4.5.2      pkgconfig_2.0.3
[49] htmltools_0.5.9

```