

STA 215 Ass 5 12310401 王子恒

$$1. f_X(x) = \lambda e^{-\lambda x} \quad f_Y(y) = \mu e^{-\mu y}$$

$$\begin{aligned} P(Z=1) &= \iint_{x \leq y} f(x, y) dx dy \\ &= \int_0^{\infty} \int_0^y f_X(x) f_Y(y) dx dy \\ &= \int_0^{\infty} \int_0^y \lambda \mu e^{-\lambda x - \mu y} dx dy \\ &= \int_0^{\infty} \lambda \mu \left(-\frac{1}{\lambda} e^{-\lambda x - \mu y} \right) \Big|_0^y dy \\ &= \int_0^{\infty} \mu (e^{-\mu y} - e^{-\lambda y - \mu y}) dy \\ &= 1 - \frac{\mu}{\lambda + \mu} \end{aligned}$$

$$= \frac{\lambda}{\lambda + \mu}$$

$$P(Z=0) = 1 - P(Z=1) = \frac{\mu}{\lambda + \mu}$$

$$P(Z=z) = \begin{cases} \frac{\lambda}{\lambda + \mu} & z=1 \\ \frac{\mu}{\lambda + \mu} & z=0 \end{cases}$$

$$2. P(X \leq z) = \sum_{k=0}^{z-1} (1-p)^k p = 1 - (1-p)^z.$$

$$P(X \leq z, Y \leq z) = (1 - (1-p)^z)^2.$$

$$P(Z \leq z) = P(X \leq z, Y \leq z)$$

$$\begin{aligned} P(Z=z) &= P(X \leq z, Y \leq z) - P(X \leq z-1, Y \leq z-1) \\ &= (1 - (1-p)^z)^2 - (1 - (1-p)^{z-1})^2 \\ &= ((1-p)^z + (1-p)^{z-1} - 2) ((1-p)^z - (1-p)^{z-1}) \\ &= (1-p)^{z-1} ((1-p)^z + (1-p)^{z-1} - 2) (-p) \end{aligned}$$

$$P(Z=z) = \begin{cases} (1-p)^{z-1} ((1-p)^z + (1-p)^{z-1} - 2) (-p) & z=0 \\ 0 & \text{otherwise} \end{cases}$$

$$3. T = \min(X_1, X_2, X_3)$$

$$\begin{aligned} F_T(t) &= P(T \leq t) = P(\min(X_1, X_2, X_3) \leq t) \\ &= 1 - P(\min(X_1, X_2, X_3) > t) \\ &= 1 - P(X_1 > t) P(X_2 > t) P(X_3 > t) \\ &= 1 - (1 - (1 - e^{-\lambda t}))^3 \\ &= 1 - e^{-3\lambda t} \end{aligned}$$

$$f_T(t) = F_T(t)' = 3\lambda e^{-3\lambda t} \quad t \geq 0$$

$$f_T(t) = \begin{cases} 3\lambda e^{-3\lambda t} & t \geq 0 \\ 0 & t < 0 \end{cases}$$

$$4. \quad 1) E(X - Y) = E(X) - E(Y) = 0$$

$$\begin{aligned} \text{Var}(X - Y) &= \text{Var}(X) + \text{Var}(Y) - 2\text{Cov}(X, Y) \\ &= 1 + 1 - 2\rho \\ &= 2(1 - \rho) \end{aligned}$$

$$\therefore X - Y \sim N(0, 2(1 - \rho))$$

$$f_Z(z) = f_{X-Y}(x-y) = \frac{1}{\sqrt{2\pi(1-\rho)}} \cdot e^{-\frac{z^2}{4(1-\rho)}}$$

$$2) W = X - Y \quad V = XY$$

$$\begin{aligned} \text{Cov}(W, V) &= E[(X - Y)(XY)] - E(X - Y)E(XY) \\ &= E(X^2Y - XY^2) \\ &= E(X^2Y) - E(XY^2) \\ &= E(X^2)E(Y) - E(X)E(Y^2) \\ &= 0 \end{aligned}$$

$$\text{Cor}(W, V) = \frac{\text{Cov}(W, V)}{\sqrt{\text{Var}(W)\text{Var}(V)}} = 0.$$

$$\begin{aligned}
 5. \quad P_Y(y) &= \int_{-\infty}^{\infty} f(x, y) dx \\
 &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{x^2}{\sigma_x^2} - 2\rho\frac{xy}{\sigma_x\sigma_y} + \frac{y^2}{\sigma_y^2}\right]\right) dx \\
 &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{x}{\sigma_x} - \frac{y}{\sigma_y}\right]^2\right) dx \\
 &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{y^2(1-\rho^2)}{2\sigma_y^2}\right) \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2(1-\rho^2)}\left(\frac{x}{\sigma_x} - \rho\frac{y}{\sigma_y}\right)^2\right) dx \\
 &= \frac{1}{\sigma_y\sqrt{2\pi}} \exp\left(-\frac{y^2}{2\sigma_y^2}\right)
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow f_{X|Y}(x|y) &= \frac{f(x, y)}{f_Y(y)} = \frac{\frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{x^2}{\sigma_x^2} - 2\rho\frac{xy}{\sigma_x\sigma_y} + \frac{y^2}{\sigma_y^2}\right]\right)}{\frac{1}{\sigma_y\sqrt{2\pi}} \exp\left(-\frac{y^2}{2\sigma_y^2}\right)} \\
 &= \frac{1}{\sigma_x\sqrt{2\pi(1-\rho^2)}} \exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{(x - \rho\frac{\sigma_x}{\sigma_y}y)^2}{\sigma_x^2}\right]\right)
 \end{aligned}$$

thus the PDF of X given $Y=y$ is normal,

$$\begin{aligned}
 \mu(x|y) &= \rho\frac{\sigma_x}{\sigma_y}y \\
 \sigma^2(x|y) &= \sigma_x^2(1-\rho^2)
 \end{aligned}$$

Q6. 1) $\mu_x = 4500$, $\sigma_x = 1500$, $\mu_y = 5500$, $\sigma_y = 2000$
 $\rho = 0.65$

$$(X|Y=y) = y \sim N(\mu_{X|Y}, \sigma_{X|Y}^2)$$

$$\text{where } \mu_{X|Y} = \mu_x + \rho\frac{\sigma_x}{\sigma_y}(y - \mu_y) = 4500 + 0.65 \times \frac{1500}{2000}(6800 - 5500) = 5133.75$$

$$\sigma_{X|Y}^2 = \sigma_x^2(1-\rho^2) = \sqrt{1500^2(1-0.65^2)} = 1139.85$$

$$Z = \frac{6800 - \mu_{X|Y}}{\sigma_{X|Y}} = \frac{6800 - 5133.75}{1140} = 1.46$$

$$\Rightarrow P(Z < 1.46) = 0.9279$$

$$\Rightarrow P(X > Y | X+Y=12000) = P(X-Y > 0 | X+Y=12000)$$

$$W = X - Y \quad V = X + Y$$

$$\begin{pmatrix} W \\ V \end{pmatrix} = \begin{pmatrix} X - Y \\ X + Y \end{pmatrix} = \begin{pmatrix} 1 & -1 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} X \\ Y \end{pmatrix} \sim N\left(\begin{pmatrix} -1000 \\ 1000 \end{pmatrix}, \begin{pmatrix} 1533^2 & -1750000 \\ -1750000 & 3186^2 \end{pmatrix}\right)$$

$$\rho = \frac{-1750000}{1533 \times 3186} = -0.358$$

$$W|V=12000 \sim N\left(-1000 - 0.358 \times \frac{1533}{3186} \times 2000, (1-0.358^2) \times 1533^2\right) = N(-1344.5, 1431.4^2)$$

$$P(W > 0 | V=12000) = 1 - \Phi\left(\frac{1344.5}{1431.4}\right) = 1 - 0.8264 = 0.1736$$

7.1) Assume $X \sim \text{Poisson}(\lambda)$

$$P(X=k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad \text{according to CLT}$$

$$\Rightarrow \mu = \lambda, \quad \sigma^2 = \lambda \Rightarrow X \sim N(\lambda, \lambda)$$

(2)

```
import numpy as np
import scipy.stats as stats
import matplotlib.pyplot as plt

lambdas = [2, 5, 10, 20]

plt.figure(figsize=(12, 8))

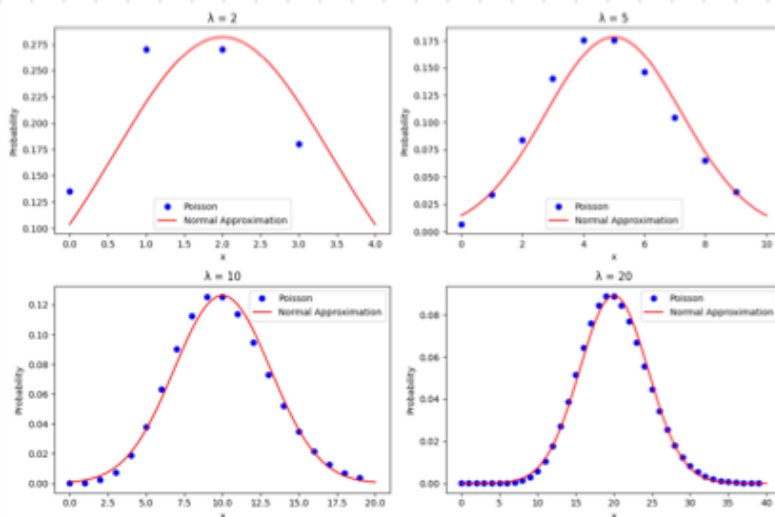
for i, lambda_ in enumerate(lambdas):

    poisson = stats.poisson(lambda_)
    x_poisson = np.arange(0, 2 * lambda_)
    y_poisson = poisson.pmf(x_poisson)

    normal = stats.norm(lambda_, np.sqrt(lambda_))
    x_normal = np.linspace(0, 2 * lambda_, 1000)
    y_normal = normal.pdf(x_normal)

    plt.subplot(2, 2, i + 1)
    plt.plot(x_poisson, y_poisson, 'bo', label='Poisson')
    plt.plot(x_normal, y_normal, 'r-', label='Normal Approximation')
    plt.title(f' $\lambda = \{lambda_\}$ ')
    plt.xlabel('x')
    plt.ylabel('Probability')
    plt.legend()

plt.tight_layout()
plt.savefig('/Users/garendelh/Codes/Python/poisson_normal_approximation.png')
```



when λ is small, the normal approximation is not great.
when λ is larger the normal approximation is better.