1.1 Difference between deterministic and stochastic world

	deterministic world	stochastic world
Single variable: Temp of a sick man	$\mathbb{R}$ $T = 39°C$	random variable $\mathbb{E}$, Var, ...
Variables changing over time: T in first 3 days	$\mathbb{R}_+ \to \mathbb{R}$ $T(1) = 39$ $T(2) = 38.5$ $T(3) = 38$ $\vdots$	stochastic process

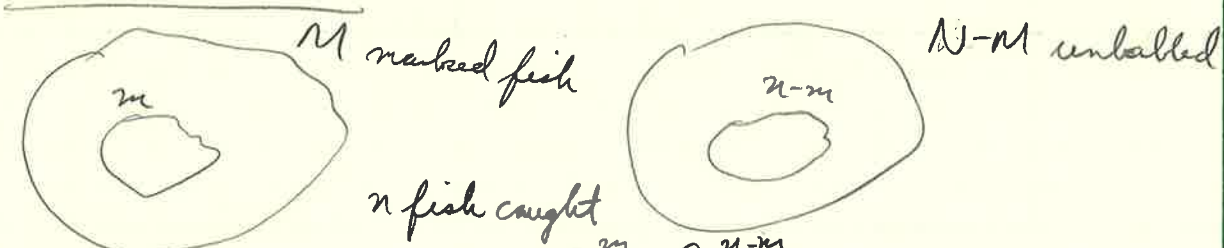
1.2 Difference between various fields of stochastics

Stochastics
- probability theory
- mathematical statistics
- stochastic processes

Consider a pond that contains fish

Prob theory : # of fish at some given time (N)
$\mathbb{E}$, Var, or limit laws

Mathematical Stats :



$$P\{\#\text{ marked} = m\} = \frac{C_M^m \, C_{N-M}^{n-m}}{C_N^M}$$

Repeat $m_1, m_2, \dots, m_q$

$$\left(\begin{array}{c}\text{log}\\ \text{likelihood}\end{array}\right) \sum_{k=1}^{q} P\{\#\text{ marked} = m_k\} \longrightarrow \max_N \quad (MLE)$$

1.3 Probability space $(\Omega, \mathcal{F}, P)$

General theory	Bernoulli Scheme $\begin{bmatrix} 1, \text{ success} \\ 0, \text{ failure} \end{bmatrix}$ $(a_1, \dots, a_n)$, $a_i \in \{0,1\}$	$[0,1]$ Select point from
Ω - sample space	$\#\Omega = 2^n$, set of all vectors with components $\in \{0,1\}$	$\Omega = [0,1]$
$\mathcal{F}$ - σ-algebra 1) $\Omega \in \mathcal{F}$ 2) $A \in \mathcal{F} \Rightarrow \Omega \setminus A \in \mathcal{F}$ 3) $A_1, \dots, A_n, \dots \in \mathcal{F}$ $\Downarrow$ $\bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$	$\mathcal{F}$ = power set $\#\mathcal{F} = 2^{\#\Omega} = 2^{2^n}$	$P\{X \in [\alpha, \beta]\}$ $\Rightarrow [\alpha, \beta), (\alpha, \beta]$, $(\alpha, \beta), [\alpha, \beta), \{\beta\} \in \mathcal{F}$ Borel σ-algebra
P - probability measure 1) $P(\Omega) = 1$ 2) $A_1, A_2, \dots \in \mathcal{F}$ (disjoint) $\Rightarrow P\{\bigcup_i A_i\} = \sum_i P(A_i)$ $P : \mathcal{F} \to [0,1]$	$P\{1\} = p$ $P\{0\} = 1 - p$	$P\{[\alpha, \beta]\} = \beta - \alpha$

1.4 Definition of a stochastic function. Types of stochastic functions.

$(\Omega, \mathcal{F}, P)$

Random variable - $\underbrace{\text{measurable}}_{\Downarrow}$ function $\xi : \Omega \to \mathbb{R}$.

$$\forall B \in \mathcal{B}(\mathbb{R}) : \xi^{-1}(B) \subset \mathcal{F}$$

T - time

$X : T \times \Omega \to \mathbb{R}$ - random function, if $\forall t \in T$: $X(t, \cdot)$ is a random variable on $(\Omega, \mathcal{F}, P)$, denoted X_t

If $T=\mathbb{R}_+$, this is called a random process or stochastic process

$T=\mathbb{R}_+^n$, random field or stochastic field

$T=\mathbb{N}$ or $\mathbb{Z}$, discrete time stochastic process

$T=\mathbb{R}_+$ or $\mathbb{R}$, continuous time stochastic process

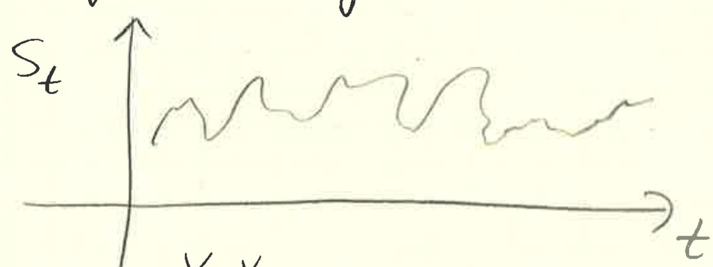
1.5 Trajectories and finite-dimensional distributions

$X: T\times\Omega \to \mathbb{R}$, $T=\mathbb{R}_+$

$\forall t \in T$: $X_t = X(t,\cdot)$ is a r.v. on $(\Omega, \mathcal{F}, P)$

Trajectory (= path)

X_t fix ω and get mapping $T \to \mathbb{R}$



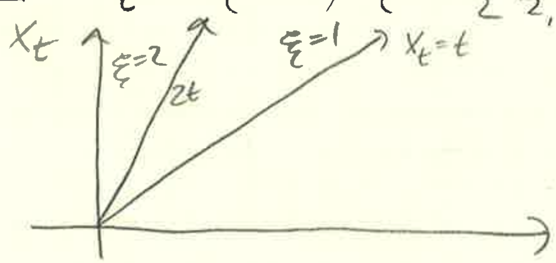
$X_1, X_2, \ldots,$

Finite-dimensional distribution $(X_{t_1}, X_{t_2}, \ldots, X_{t_n})$, $t_1, \ldots, t_n \in \mathbb{R}$

In mathematic stats, $X_{t_1}, X_{t_2} \ldots X_{t_n}$ are independent

In stochastic process, $X_{t_1}, X_{t_2}, \ldots, X_{t_n}$ are dependent

<u>Ex</u>: $X_t = \xi t$, $\xi = \begin{cases} 1, & \text{w.p. } 1/2 \\ 2, & \text{w.p. } 1/2 \end{cases}$



$P\{X_{t_1} \le x_1, X_{t_2} \le x_2\} =$
$\begin{cases} 0, & \min\left(\frac{x_1}{t_1}, \frac{x_2}{t_2}\right) < 1 \\ 1/2, & \quad '' \quad \in [1,2] \\ 1, & \quad '' \quad \ge 2 \end{cases}$

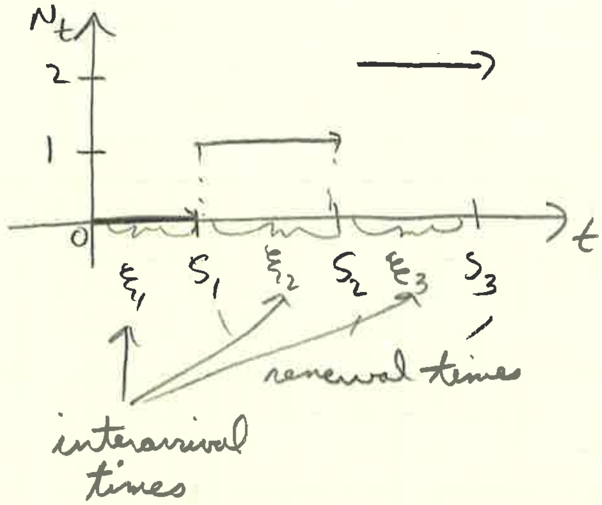
1.6 Renewal process. Counting process.

Renewal processes (discrete time)

$S_0 = 0,\ S_n = S_{n-1} + \xi_n$, where $\xi_1, \xi_2, \dots$ – iid > 0 a.s.

$P\{\xi_i > 0\} = 1 \iff F(0) = 0$

$N_t = \underset{k}{\operatorname{argmax}} \{S_k \le t\}$ (Counting process)



$\{S_n > t\} = \{N_t < n\}$

$F \Rightarrow \mathbb{E}N_t$ $\quad S_n = \xi_1 + \dots + \xi_n$

1.7. Convolution

Convolution $\quad X \perp\!\!\!\perp Y$

$X \sim F_X,\ Y \sim F_Y$

$F_{X+Y}(x) = \int_{\mathbb{R}} F_X(x-y)\, dF(y) =: \underbrace{F_X * F_Y}$ conv in terms of distribution functions

$X \sim p_X,\ Y \sim p_Y$ (If Y, X have densities)

$p_{X+Y}(x) = \int_{\mathbb{R}} p_X(x-y)\, p_Y(y)\, dy =: p_X * p_Y$ $\Big\}$ conv in terms of densities

$S_n = \xi_1 + \dots + \xi_n$

let $F^{n*} := \underbrace{F * \dots * F}_{n}$

1) $F^{n*}(x) \le F^{n}(x)$ if $F(0)=0$

$\xi_1, \dots, \xi_n \overset{iid}{\sim} F$

$\{\xi_1 + \dots + \xi_n \le x\} \subset \{\xi_1 \le x, \dots, \xi_n \le x\}$ Since $\xi_i > 0$ a.s.

$$\underbrace{P\{\xi_1 + \dots + \xi_n \le x\}}_{F^{n*}(x)} \le \prod_{k=1}^{n} \underbrace{P\{\xi_k \le x\}}_{F(x)}$$

2) $F^{n*}(x) \ge F^{(n+1)*}(x)$

$\{\xi_1 + \dots + \xi_n \le x\} \supset \{\xi_1 + \dots + \xi_{n+1} \le x\}$

Theorem: $S_n = S_{n-1} + \xi_n$ where $\xi_1, \xi_2, \dots \overset{iid}{\sim} F$, $F(0)=0$

(i) $\boxed{U(t) = \sum_{n=1}^{\infty} F^{n*}(t) < \infty}$

(2) $\boxed{\mathbb{E}N_t = U(t)}$

proof for (2)

$$\mathbb{E}N_t = \mathbb{E}\left[\#\{n : S_n \le t\}\right]$$

$$= \mathbb{E}\left[\sum_{n=1}^{\infty} \mathbb{1}\{S_n \le t\}\right] = \sum_{n=1}^{\infty} P\{S_n \le t\}$$

$$= \sum_{n=1}^{\infty} F^{n*}(t)$$

1.8 Laplace transform. Calculation of an expectation of a counting process (1)

Laplace transform

$f: \mathbb{R}_+ \to \mathbb{R}$: $\mathcal{L}_f(s) = \int_0^{\infty} e^{-sx} f(x)\,dx$

1) f - density of ξ, then $\mathcal{L}_f(s) = \mathbb{E}[e^{-s\xi}]$

2) f_1, f_2 : $\mathcal{L}_{\underbrace{f_1 * f_2}_{\text{densities}}}(s) = \mathcal{L}_{f_1}(s) \cdot \mathcal{L}_{f_2}(s)$

3) F - distribution function, $F(0)=0$, $p = F'$

$$\mathcal{L}_F(s) = \frac{\mathcal{L}_p(s)}{s}$$

$$\text{l.h.s.} = \int_{\mathbb{R}_+} F(x)\,\frac{d(e^{-sx})}{s} = \cancel{-\frac{F(x)\,e^{-sx}}{s}\Big|_0^{\infty}} + \frac{1}{s}\int_{\mathbb{R}_+} p(x)\,e^{-sx}\,dx$$

$$= \text{r.h.s.}$$

Ex

1) $$\mathcal{L}_{x^k}(s) = \int_{\mathbb{R}_+} x^n\,\frac{d(e^{-sx})}{s} = \frac{n}{s}\int_{\mathbb{R}_+} x^{n-1}e^{-sx}\,dx$$

$$= \frac{n}{s}\cdot\frac{n-1}{s}\cdots\frac{2}{s}\int_{\mathbb{R}_+} e^{-sx}\,dx = \frac{n!}{s^n}$$

2) $$\mathcal{L}_{e^{ax}}(s) = \frac{1}{s-a}, \text{ if } a<s$$

1.9 Laplace transform. Calculation of an expectation of a counting process (2)

$F \to \mathbb{E}N_t$

$$\mathbb{E}N_t = \mathcal{U}(t) = \sum_{n=1}^{\infty} F^{n*}(t) = F(t) + \left(\sum_{n=1}^{\infty} F^{n*}(t)\right) * F(t)$$

$$\Leftrightarrow \quad \mathcal{U} = F + \mathcal{U} \underset{\text{dist fcn}}{*} F = F + \mathcal{U} * \underset{\text{densities}}{p} \quad \text{If } F' = p \text{ exists}$$

$$\int_{\mathbb{R}} \mathcal{U}(x-y)\,dF(y) = \int_{\mathbb{R}} \mathcal{U}(x-y)\,p(y)\,dy$$

$$\mathcal{L}_{\mathcal{U}}(s) = \underbrace{\mathcal{L}_F(s)}_{= \frac{\mathcal{L}_p(s)}{s}} + \mathcal{L}_{\mathcal{U}}(s)\,\mathcal{L}_p(s)$$

$$\boxed{\mathcal{L}_{\mathcal{U}}(s) = \frac{\mathcal{L}_p(s)}{s(1-\mathcal{L}_p(s))}}$$

1. $F \to \mathcal{L}_p$
2. $\mathcal{L}_p \to \mathcal{L}_{\mathcal{U}}$
3. $\mathcal{L}_{\mathcal{U}} \to \mathcal{U}$

1.10 Laplace transform. Calculation of an expectation of a counting process (3)

Example: $S_n = S_{n-1} + \xi_n$, $\xi_1, \xi_1, \dots$ have density $p(x)$

$$p(x) = \frac{e^{-x}}{2} + e^{-2x}, \quad x > 0$$

$\mathbb{E}N_t = ?$

① $p \to \mathcal{L}p$: $\mathcal{L}p(s) = \frac{1}{2}\mathcal{L}_{e^{-x}}(s) + \mathcal{L}_{e^{-2x}}(s)$

$$= \frac{1}{2(s+1)} + \frac{1}{s+2} = \frac{3s+4}{2(s+1)(s+2)}$$

② $\mathcal{L}p \to \mathcal{L}_u$: $\mathcal{L}_u(s) = \frac{\mathcal{L}p(s)}{s(1-\mathcal{L}p(s))} = \frac{3s+4}{s^2(2s+3)}$

③ $\mathcal{L}_u \to u$: $\mathcal{L}_u(s) = \frac{A}{s^2} + \frac{B}{s} + \frac{C}{2s+3}$

$$= \frac{A(2s+3) + B(2s^2+3s) + Cs^2}{s^2(2s+3)}$$

$$3s+4 = (2B+C)s^2 + (2A+3B)s + 3A$$

$A = \frac{4}{3}$, $2A+3B = 3 \Leftrightarrow B = \frac{1}{9}$, $2B+C=0 \Leftrightarrow C = -\frac{2}{9}$

$$\boxed{u(t) = \frac{4}{3}t + \frac{1}{9}(1) - \frac{1}{9}e^{-3/2\, t}}$$

1.11 Limit theorems for renewal processes

$S_n = S_{n-1} + \xi_n$; $\xi_1, \xi_2, \dots$ iid >0 a.s.

Thm 1 $\mu = \mathbb{E}\xi_1 < \infty \Rightarrow \frac{N_t}{t} \xrightarrow{t\to\infty} \frac{1}{\mu}$ a.s.

(Analog to SLLN)

$$\boxed{\text{SLLN: } \frac{\xi_1 + \dots + \xi_n}{n} \xrightarrow[n\to\infty]{} \mu \text{ a.s.}}$$

Thm 2: (Analog of CLT) $\sigma^2 = \mathrm{Var}\,\xi_1 < \infty$

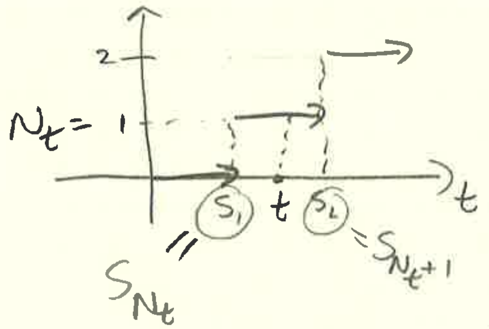
Then, $Z_t := \frac{N_t - t/\mu}{\sigma\sqrt{t}/\mu^{3/2}} \xrightarrow[t\to\infty]{d} N(0,1)$

$$P\{Z_t \le x\} \to \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} e^{-u^2/2}\,du$$

$$\boxed{\text{CLT: } \frac{\xi_1 + \dots + \xi_n - n\mu}{\sigma\sqrt{n}} \xrightarrow{d} N(0,1)}$$

proof (thm 1)

$$S_{N_t} \le t \le S_{N_t+1}$$



$$\frac{N_t}{S_{N_t+1}} \le \frac{N_t}{t} \le \frac{N_t}{S_{N_t}}$$

$$\lim_{t\to\infty} \frac{N_t}{S_{N_t}} = \lim_{n\to\infty} \frac{n}{S_n} = \frac{1}{\mu} \text{ by SLLN}$$

$$\lim_{t\to\infty} \frac{N_t}{S_{N_t+1}} = \underbrace{\lim_{t\to\infty} \frac{N_t}{N_t+1}}_{=1} \cdot \underbrace{\lim_{t\to\infty} \frac{N_t+1}{S_{N_t+1}}}_{=1/\mu} = \frac{1}{\mu}$$

proof (thm 2)

$$P\left\{\frac{S_n - n\mu}{\sigma\sqrt{n}} \le x\right\} \to \phi(x), \quad x \in R$$

$$P\{S_n \le \underbrace{n\mu + \sigma\sqrt{n}\, x}_{=t}\} \to \phi(x)$$

$\Leftrightarrow P\{N_t \ge n\}$ (set complements)

$n\mu \approx t$
$n \approx t/\mu$ (for n large enough)

$$n = \frac{t}{\mu} - \frac{\sigma\sqrt{n}}{\mu} x \approx \frac{t}{n} - \frac{\sigma\sqrt{t}}{\mu^{3/2}} x$$

$$\Rightarrow P\{Z_t \ge -x\} \to \phi(x) \quad \Leftrightarrow P\{Z_t \le x\} = 1 - P\{Z_t \ge -x\} \to 1 - \phi(-x) = \phi(x)$$

Poisson Processes

2.1 Definition of a Poisson process as a special example of a renewal process. Exact forms of the distributions of the renewal process and the counting process (1)

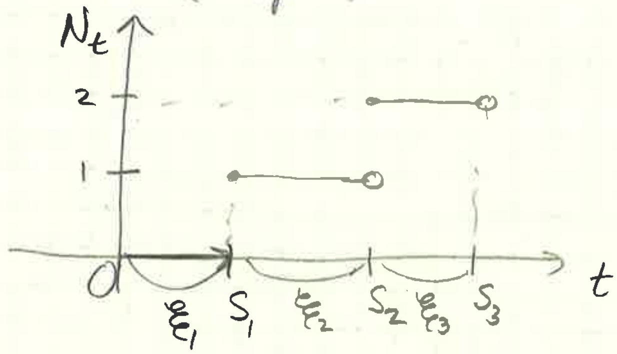
Renewal process

$S_0 = 0, \quad S_n = S_{n-1} + \xi_n, \quad \xi_1, \xi_2, \dots$ - i.i.d > 0 a.s., $\xi_i \sim F$

$N_t = \underset{k}{\operatorname{argmax}} \{ S_k \le t \}$ (Counting process)

$$U(t) = \mathbb{E} N_t = \sum_{n=1}^{\infty} F^{n*}(t)$$

$$\mathcal{L}_U(s) = \frac{\mathcal{L}_p(s)}{s(1-\mathcal{L}_p(s))} \quad : \quad p \to \mathcal{L}_p \to \mathcal{L}_U \to U \qquad (p = F')$$



$$\mathcal{L}_U(s) = \int_{\mathbb{R}_+} e^{-sx} U(x)\, dx$$

2.2 .... (2)

Poisson process

<u>Def 1</u>: A Poisson process is a renewal process s.t.

$$\xi_i \sim p(x) = \lambda e^{-\lambda x} \mathbb{I}\{x > 0\}, \quad \lambda\text{- intensity or rate}$$

<u>Thm</u> (i): A distribution function of S_n

$$F_{S_n}(x) = \begin{cases} 1 - e^{-\lambda x} \sum_{k=0}^{n-1} \frac{(\lambda x)^k}{k!}, & x > 0 \\ 0, & x < 0 \end{cases}$$

$$p_{S_n}(x) = \lambda \frac{(\lambda x)^{n-1}}{(n-1)!} e^{-\lambda x} \mathbb{I}\{x > 0\}$$

(ii) $\mathbb{P}\{N_t = n\} = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$, $N_t \sim \text{Poisson}(\lambda t)$

2.3 ··· (3)

Proof (i) $n=1$: $S_1 = \xi_1$

$$p_{S_1}(x) = \lambda e^{-\lambda x}, \; x>0$$

$n \to n+1$

$$p_{S_{n+1}}(x) = \int_0^x p_{S_n}(x-y)\, p_{\xi_{n+1}}(y)\,dy$$

$$= \int_0^x \frac{\lambda^n (x-y)^{n-1}}{(n-1)!} e^{-\lambda(x-\cancel{y})} \lambda \cancel{e^{-\lambda y}}\,dy$$

$$= \frac{\lambda^{n+1}}{(n-1)!} e^{-\lambda x} \int_0^x (x-y)^{n-1} dy = \frac{\lambda^{n+1}}{(n-1)!} e^{-\lambda x} \frac{x^n}{n}$$

$$= \lambda \frac{(\lambda x)^n}{n!} e^{-\lambda x} \qquad \square$$

2.4 ··· (4)

proof (ii)

$$P\{N_t = n\} = P\{S_n \le t\} - P\{S_{n+1} \le t\} \quad =$$

$$\{N_t = n\} = \underbrace{\{S_n \le t\}}_{=A} \cap \underbrace{\{S_{n+1} > t\}}_{=B}$$

$$\left.\begin{array}{l} A \cap B = A \setminus B^c \\ \text{Here: } B^c \subset A \end{array}\right\} \Rightarrow P\{A \cap B\} = P\{A\} - P\{B^c\}$$

$$= \left(1 - e^{-\lambda t} \sum_{k=0}^{n-1} \frac{(\lambda t)^k}{k!}\right) - \left(1 - e^{-\lambda t} \sum_{k=0}^{n} \frac{(\lambda t)^k}{k!}\right)$$

$$= e^{-\lambda t} \frac{(\lambda t)^n}{n!} \qquad \square$$

2.5 Memoryless property

A r.v. X possesses the memoryless property iff $P\{X > u+v\} = P\{X>u\}\, P\{X>v\}$. If $P\{X>v\}>0$, then

$$\boxed{P\{X > u+v \mid X > v\} = P\{X>u\}}$$

Thm 2: Let X be a r.v. with density $p(x)$, then

X- memoryless $\iff p(x) = \lambda e^{-\lambda x}$

Ex busses arrive every 20 ± 2 minutes

$v = 19$ min, $u = 10$ min

(l.h.s.) $P\{X > 29 \mid X > 19\} = 0$ given the data

(r.h.s) $P\{X > 10\} = 1$

Thus, Poisson process is not appropriate

2.6 Other definitions of Poisson processes (1)

Def 2 N_t – an integer value process s.t.

0) $N_0 = 0$ a.s.

1) N_t has independent increments: $\forall t_0 < t_1 < \dots < t_n$,

$N_{t_1} - N_{t_0}, \dots, N_{t_n} - N_{t_{n-1}}$ are independent

2) N_t has stationary increments

$$N_t - N_s \overset{d}{=} N_{t-s}$$

3) $N_t - N_s \sim \text{Poisson}(\lambda(t-s))$, $t > s$

3) $\Rightarrow$ 2)

2.7 Other definitions of Poisson processes (2)

$$P\{N_{t+h} - N_t = 0\} = 1 - \lambda h + o(h), \quad h \to 0$$

$$P\{N_{t+h} - N_t = 1\} = \lambda h + o(h), \quad h \to 0$$

$$P\{N_{t+h} - N_t \geq 2\} = o(h), \quad h \to 0$$

$$\lim_{h \to 0} \frac{1 - P\{N_{t+h} - N_t = 0\}}{h} = \lim_{h \to 0} \frac{1 - e^{-\lambda h}}{h} = \lambda$$

Def 3 N_t is a Poisson process, if

0) $N_0 = 0$

1) N_t has independent increments

2) N_t has stationary increments

3') $$\lim_{h \to 0} \frac{P\{N_{t+h} - N_t \geq 2\}}{P\{N_{t+h} - N_t = 1\}} = 0$$

2.8 Non-homogeneous Poisson processes (1)

$N_t \sim Pois(\lambda t) \Rightarrow \mathbb{E}N_t = \lambda t$

Def: Let $\Lambda(t)$ be a differentiable, increasing function s.t. $\Lambda(0)=0$. Then, $X_t = N_t$ is a non-homogeneous Poisson process if,

0) $N_0 = 0$

1) N_t has independent increments

2) $\boxed{N_t - N_s \sim Pois(\Lambda(t) - \Lambda(s))}$

2.9 Non-homogeneous Poisson processes (2) (NHPP)

$\lambda(t) = \Lambda'(t)$ - intensity function

Properties NHPP:

1) $\mathbb{E}N_t = \Lambda(t)$

$\Lambda(t) = \alpha t^{\beta}, \alpha > 0, \beta > 0$ (for example)

2) if $\lambda(t) = const \Rightarrow \Lambda(t) = const \cdot t$

3) $\left.\begin{array}{l}\Lambda(t)\text{-differentiable} \Rightarrow \Lambda(t)\text{-continuous} \\ \Lambda(t)\text{-increasing}\end{array}\right\} \Rightarrow$

$\Rightarrow \exists \Lambda^{-1}(t)$. If $\text{Image}\,\Lambda(t) = \mathbb{R}_+$, $N_{\Lambda^{-1}(t)}$ - homogeneous P.P.

2.10 Relation between renewal theory and NHPP (1)

$S_n = \underset{t}{\operatorname{argmin}} \{N_t = n\}$, $\xi_n = S_n - S_{n-1}$

$\xi_1, \xi_2, \ldots$ - iid ?

1) $p_{\xi_1}(x) = \lambda(x) e^{-\Lambda(x)}$

$P\{\xi_1 \le x\} = P\{S_1 \le x\} \underset{\{S_n > t\} = \{N_t < n\}}{=} P\{N_x \ge 1\} = 1 - P\{N_x = 0\} =$

$= 1 - e^{-\Lambda(x)}$

Take derivatives of both sides to finish proof $\square$.

2.11 ··· (2)

$S_k = \underset{t}{\operatorname{argmin}} \{N_t = k\}$

$\xi_k = S_k - S_{k-1}$

1) $\boxed{p_{\xi_1}(t) = \lambda(t) e^{-\Lambda(t)}}$

2) $\boxed{p_{\xi_2|\xi_1}(t|s) = \lambda(t+s) e^{-\Lambda(t+s)+\Lambda(s)}}$

$$F_{(\xi_1,\xi_2)}(s,t) = P\{\xi_1 \le s, \xi_2 \le t\} = \int_0^s P\{\underbrace{\cancel{\xi_1 \le s}}_{\text{Since } y<s}, \xi_2 \le t \mid \xi_1 = y\} p_{\xi_1}(y)\,dy$$

$$= \int_0^s P\{\underbrace{N_{t+y} - N_y \ge 1 \mid \xi_1 = y}_{\text{independent}}\} p_{\xi_1}(y)\,dy$$

$$= \int_0^s \left(1 - e^{-\Lambda(t+y)+\Lambda(y)}\right) \lambda(y) e^{-\Lambda(y)}\,dy$$

$$p_{(\xi_1,\xi_2)}(s,t) = \frac{\partial}{\partial t}\left(\frac{\partial}{\partial s} F_{(\xi_1,\xi_2)}(s,t)\right)$$

$$= \frac{\partial}{\partial t}\left(\left(1 - e^{-\Lambda(t+s)+\Lambda(s)}\right) \lambda(s) e^{-\Lambda(s)}\right)$$

$$= \lambda(t+s) e^{-\Lambda(t+s)+\Lambda(s)} \lambda(s) e^{-\Lambda(s)}$$

Then $p_{\xi_2|\xi_1}(t|s) = \dfrac{p_{(\xi_1,\xi_2)}(s,t)}{p_{\xi_1}(s)}$ finishes the proof $\square$

2.12 ··· (3)

$\xi_1, \xi_2, \ldots$ – i.i.d. ? (NHPP can be obtained from renewal process iff NHPP is homogeneous PP)

$$p_{\xi_1}(t) = p_{\xi_2|\xi_2}(t|s), \quad \forall t, s > 0$$

$$\lambda(t) e^{-\Lambda(t)} = \lambda(t+s) e^{-\Lambda(t+s)+\Lambda(s)}$$

$$\left(\int_0^T \cdots dt\right): \ \cancel{e^{-\Lambda(0)}} - e^{-\Lambda(T)} = \cancel{e^{0}} - e^{-\Lambda(T+S)+\Lambda(S)}$$

$$\boxed{\begin{array}{l} \Lambda(T) = \Lambda(T+S) - \Lambda(S), \quad \forall S, T > 0 \\ \Lambda(t) - \text{increasing} \end{array}} \Rightarrow \Lambda(t) = \overset{\text{const}}{\overset{\parallel}{\lambda}} t$$

2.13 Elements of queuing theory. M/G/k systems (1)

$$\boxed{\begin{aligned} P\{N_{t+h}-N_t=0\} &= 1-\lambda h + o(h) \quad (\lambda \to \lambda(t)) \\ P\{N_{t+h}-N_t=1\} &= \lambda h + o(h) \quad (\lambda \to \lambda(t)) \\ P\{N_{t+h}-N_t\geq 2\} &= o(h) \end{aligned}}$$

$M/G/k$

I) Arrival Process: M - memoryless (Poisson)
D - deterministic
G - general

II) Service time (M, D, G)

III) A number of services $(1, 2, \dots, \infty)$

$\underline{M/G/\infty}$ $\quad \tau > 0$ (time moment)

$\xrightarrow[\text{Customer arrivals}]{N(t)}$

$\to N_1(t)$ - still being served at τ : $\lambda_1(t)=\lambda(1-G(\tau-t))$

$\to N_2(t)$ - already completed by τ : $\lambda_2(t)=\lambda G(\tau-t)$

$$P\{N_1(t+\delta)-N_1(t)=1\} = P\{N(t+\delta)-N(t)=1\}\cdot\left(P\{Y>\tau-t\}+o(\delta)\right)$$
$$=(\delta\lambda+o(\delta))(1-G(\tau-t)+o(\delta))$$
$$=\boxed{\lambda\delta(1-G(\tau-t)+o(\delta))}$$

2.14 ... (2)

bernoulli scheme

$$P\{N_1(t)=n_1,\ N_2(t)=n_2\} = P\{N_1(t)=n_1,\ N_2(t)=n_2 \mid N(t)=n_1+n_2\}$$
$$\cdot P\{N(t)=n_1+n_2\}$$
$$=C^{n_1}_{n_1+n_2}(1-G(\tau-t))^{n_1}G(\tau-t)^{n_2}\cdot e^{-\lambda t}\frac{(\lambda t)^{n_1+n_2}}{(n_1+n_2)!}$$
$$=\frac{\lambda t(1-G(\tau-t))^{n_1}}{n!}e^{-\lambda(1-G(\tau-t))}\cdot\frac{\lambda t(G(\tau-t))^{n_2}}{n_2!}e^{-\lambda G(\tau-t)}$$
$$=P\{N_1(t)=n_1\}\cdot P\{N_2(t)=n_2\}$$

Therefore $N_1 \perp\!\!\!\perp N_2$

2.15 Compound Poisson Processes (1)

$$X_t = \sum_{k=1}^{N_t} \xi_k \ , \quad \xi_1, \xi_2, \dots \text{ - iid. }, N_t \text{ - P.P. with intensity } \lambda$$

and $\xi_1, \xi_2, \dots$ and N_t are independent

$\xi_1, \xi_2, \dots$ claim sizes

N_t - amount of claims until time t (Insurance interpretation)

X_t - aggregated claim amount

1) Probability generating function (PGF)

ξ - integer, ≥ 0 values

$$\boxed{\phi_\xi(u) = \mathbb{E}[u^\xi]}, \ |u| < 1$$

$$\xi_1 \perp\!\!\!\perp \xi_2 \Rightarrow \phi_{\xi_1+\xi_2}(u) = \phi_{\xi_1}(u)\,\phi_{\xi_2}(u)$$

2) Moment-generating function (MGF)

$$\boxed{\mathcal{L}_\xi(u) = \mathbb{E}[e^{-u\xi}]}, \ \xi \geq 0, \ u > 0$$

2.16 ... (2)

3) Characteristic function

$$\phi_\xi(u) = \mathbb{E}[e^{iu\xi}], u \in \mathbb{R}, \forall \xi \ , \quad \phi_\xi : \mathbb{R} \to \mathbb{C} \ , \quad \xi_1 \perp\!\!\!\perp \xi_2 \Rightarrow \phi_{\xi_1+\xi_2}(u) = \phi_{\xi_1}(u)\phi_{\xi_2}(u)$$

Thm $$\boxed{\phi_{X_t - X_s}(u) = e^{\lambda(t-s)(\phi_{\xi_1}(u)-1)}}$$

Proof: $$\text{lhs} = \mathbb{E}e^{iu(X_t - X_s)} = \sum_{k=0}^{\infty} \mathbb{E}[\underbrace{e^{iu(X_t-X_s)}}_{\substack{\| d \\ \xi_1 + \dots + \xi_k}} \mid N_t - \cancel{N_s} = k]\, P\{N_t - N_s = k\}$$

from $\perp\!\!\!\perp$

$$= \sum_{k=0}^{\infty} (\phi_{\xi_1}(u))^k e^{-\lambda(t-s)} \frac{[\lambda(t-s)]^k}{k!}$$

$\square$

2.17 ⋯ (3)

$X_t = \sum_{k=1}^{N_t} \xi_k$, ξ can be any random variable

ξ: $\phi_\xi(u) = E[e^{iu\xi}]$

$\phi: \mathbb{R} \to \mathbb{C}$

$\xi_1 \perp\!\!\!\perp \xi_2 \Rightarrow \phi_{\xi_1+\xi_2}(u) = \phi_{\xi_1}(u)\,\phi_{\xi_2}(u)$

Thm: $\boxed{\phi_{X_t - X_s}(u) = e^{\lambda(t-s)(\phi_{\xi_1}(u)-1)}, \quad t > s \geq 0}$

proof

$$lhs = \mathbb{E}[e^{iu(X_t - X_s)}]$$

$$= \sum_{k=0}^{\infty} \mathbb{E}[e^{iu\underbrace{(X_t-X_s)}_{\overset{d}{=}\ \xi_1+\dots+\xi_k}} \mid N_t - \cancel{N_s = k}]\cdot \underbrace{P\{N_t - N_s = k\}}_{\sim Pois(\lambda(t-s))}$$

since $\xi_1, \xi_2, \dots \perp\!\!\!\perp N_t$

$$= \sum_{k=0}^{\infty} [\phi_{\xi_1}(u)]^k \cdot e^{-\lambda(t-s)} \frac{[\lambda(t-s)]^k}{k!}$$

□

2.18 ⋯ (4)

Corollary: $\boxed{\begin{aligned} \mathbb{E}X_t &= \lambda t \mathbb{E}\xi_1 \\ Var X_t &= \lambda t \mathbb{E}\xi_1^2 \end{aligned}}$

proof $E[\xi^r] < \infty \Rightarrow \phi(u)$ is r-times differentiable at 0 and $\phi^{(r)}(0) = i^r \mathbb{E}\xi^r$

$$\mathbb{E}X_t = \frac{\phi_{X_t}'(0)}{i} = \frac{\lambda t\,\phi_\xi'(0)\cdot \overbrace{\phi_{X_t}(0)}^{=1}}{\underbrace{i}_{=\mathbb{E}\xi}} = \lambda t \mathbb{E}\xi_1$$

□