Homework 4

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1 Equation

Since π is under the constraint $\sum_{j=1}^{k} \pi_j = 1$, we introduce the Lagrangian multiplier,

$$\tilde{\mathcal{L}}(\boldsymbol{\theta}, \lambda) = \tilde{\mathcal{L}}(\boldsymbol{\theta}) + \lambda \left(1 - \sum_{j=1}^{k} \pi_j \right)$$

$$= \sum_{i=1}^{n} \sum_{j=1}^{k} r_{i,j}^{(t)} (\log \pi_j + \log \mathcal{N}(\boldsymbol{x}_i; \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)) + \lambda \left(1 - \sum_{j=1}^{k} \pi_j \right)$$

We take the gradient w.r.t. π and get

$$\frac{\partial \tilde{\mathcal{L}}(\boldsymbol{\theta}, \lambda)}{\partial \pi_j} = \sum_{i=1}^n \frac{r_{i,j}^{(t)}}{\pi_j} - \lambda = 0 \implies \pi_j^{(t)} = \frac{1}{\lambda} \sum_{i=1}^n r_{i,j}^{(t)}$$

Using the constraint $\sum_{j=1}^{k} \pi_j = 1$ again, we finally have

$$\sum_{j=1}^{k} \pi_j^{(t)} = \frac{1}{\lambda} \sum_{j=1}^{k} \sum_{i=1}^{n} r_{i,j}^{(t)}$$

$$= \frac{1}{\lambda} \sum_{i=1}^{n} \sum_{j=1}^{k} \frac{\pi_j \mathcal{N}(\boldsymbol{x}_i; \boldsymbol{\mu}_j^{(t-1)}, \boldsymbol{\Sigma}_j^{(t-1)})}{\sum_{l=1}^{k} \pi_l \mathcal{N}(\boldsymbol{x}_i; \boldsymbol{\mu}_l^{(t-1)}, \boldsymbol{\Sigma}_l^{(t-1)})}$$

$$= \frac{1}{\lambda} \sum_{i=1}^{n} 1 = \frac{n}{\lambda} = 1 \implies \lambda = n$$

$$\therefore \pi_j^{(t)} = \frac{1}{n} \sum_{i=1}^{n} r_{i,j}^{(t)}$$

By taking the gradient w.r.t. μ , we get

$$\begin{split} \frac{\partial \tilde{\mathcal{L}}(\boldsymbol{\theta}, \lambda)}{\partial \boldsymbol{\mu}_{j}} &= \sum_{i=1}^{n} r_{i,j}^{(t)} \frac{\partial}{\partial \boldsymbol{\mu}_{j}} \left(-\frac{1}{2} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j})^{\top} \boldsymbol{\Sigma}_{j}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j}) \right) \\ &= \sum_{i=1}^{n} r_{i,j}^{(t)} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j})^{\top} \boldsymbol{\Sigma}_{j}^{-1} = \boldsymbol{0} \qquad (\because \boldsymbol{\Sigma}_{j} = \boldsymbol{\Sigma}_{j}^{\top}) \\ \therefore \boldsymbol{\mu}_{j}^{(t)} &= \frac{1}{\sum_{i=1}^{n} r_{i,j}^{(t)}} \sum_{i=1}^{n} r_{i,j}^{(t)} \boldsymbol{x}_{i} \end{split}$$

This time, taking the gradient w.r.t. Σ^{-1} , we get

$$\frac{\partial \tilde{\mathcal{L}}(\boldsymbol{\theta}, \lambda)}{\partial \boldsymbol{\Sigma}_{j}^{-1}} = \sum_{i=1}^{n} r_{i,j}^{(t)} \frac{\partial}{\partial \boldsymbol{\Sigma}_{j}^{-1}} \left(\frac{1}{2} \log(\det(\boldsymbol{\Sigma}_{j}^{-1})) - \frac{1}{2} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j})^{\top} \boldsymbol{\Sigma}_{j}^{-1} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j}) \right) \\
= \sum_{i=1}^{n} \frac{r_{i,j}^{(t)}}{2} \left(\boldsymbol{\Sigma}_{j} - (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j}) (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j})^{\top} \right) = \mathbf{0} \qquad (:: \boldsymbol{\Sigma}_{j} = \boldsymbol{\Sigma}_{j}^{\top}) \\
\therefore \boldsymbol{\Sigma}_{j}^{(t)} = \frac{1}{\sum_{i=1}^{n} r_{i,j}^{(t)}} \sum_{i=1}^{n} r_{i,j}^{(t)} \left(\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j}^{(t)} \right) \left(\boldsymbol{x}_{i} - \boldsymbol{\mu}_{j}^{(t)} \right)^{\top}$$

2 Implementation

```
import numpy as np
from scipy.stats import multivariate_normal as mn
import matplotlib.pyplot as plt
```

I used some functionalities of NumPy and SciPy to implement the EM algorithm, and Matplotlib for visualization.

```
1 X=np.loadtxt('X.txt')
```

First, X is set to be 2D data written in X.txt.

```
n,d=X.shape
k=5 # number of components
random seed=0
```

 \boldsymbol{X} is composed of n observations, where each observation has d values. In the case of X.txt, n=2000 and d=2. The number of components is fixed k=5. I additionally used random_seed hyperparameter. This is to analyze the sensitivity of the EM algorithm with respect to the initialization and for reproducibility of the results.

```
np.random.seed(random_seed)
pi=np.random.rand(k)
pi/=sum(pi) #normalization

mu=np.random.randn(k,d)
sigma=[]
for _ in range(k):
    A=np.random.randn(d,d)
sigma.append(AGA.T)
```

Before running the EM algorithm, parameter θ should be initialized. I used np.random.rand() to get π . Sicne π is under the constraint $\sum_{j=1}^k \pi_j = 1$, it requires normalization.

While μ can be easily initialized by using np.random.randn, Σ needs a few steps. Since each covariance matrix Σ_j should be symmetric and positive semi-definite, the easiest way is to initialize Σ_j to AA^{\top} where $A \in \mathbb{R}^{d \times d}$ is just a random matrix.

```
def E_step(pi, mu, sigma):
                                                         # Responsibility values
                                                        r=np.array([[pi[j]*mn.pdf(X[i],mu[j],sigma[j]))
                                                                                                                                                           for j in range(k) | for i in range(n) |)
                                                        for i in range(n):
                                                                                         s=sum(r[i])
                                                                                         for j in range(k):
                                                                                                                         r[i][j]/=s
                                                         # Complete-data log-likelihood
    9
                                                         L=sum (sum (r[i][j]*(np.log(pi[j])+mn.logpdf(X[i],mu[j],sigma[j])) \\ \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j],sigma[j])) \\ \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \\ \setminus (p.log(pi[j])+mn.logpdf(X[i],mu[j])) \\ \setminus (p.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])) \\ \setminus (p.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.log(pi[j])+mn.l
10
                                                                                                                                                                                                                                                                                             for j in range(k)) for i in range(n))
11
                                                        return r, L
12
```

In the E-step, the algorithm computes the responsibility values $\{r_{i,j}^t\}$ and the complete-data log-likelihood $\tilde{\mathcal{L}}$. The function gets π , μ and Σ as the inputs and outputs r and $\tilde{\mathcal{L}}$. I used ndarray of NumPy for the parameters, and this enables such compact implementation.

In the M-step, the algorithm updates the parameters. The function just computes the new parameter values according to the equations that we derived in the Q.1. It only takes r as the input and outputs π , μ and Σ .

```
1  t=1
2  Ls=[]
3  L_prev=0
4  eps=1e-3
5  while True:
6    r, L=E_step(pi, mu, sigma)
7    Ls.append(L)
8    if t>1  and  abs(L_prev-L)<eps:</pre>
```

```
break

L_prev=L

pi,mu,sigma=M_step(r)

Logging

if t%5==1:

print(L)

t+=1
```

Finally, this is the main part of the EM algorithm. Using E_step and M_step, the algorithm runs until the log-likelihood converges. I set eps=1e-3 as the threshold of convergence.

Figure 1: The first(left) and last(right) log-likelihood values

Fig. 1 and 2 are the results. The results show that the EM algorithm is sensitive with respect to the initialization. It is quite interesting that the algorithm reached to the almost same value with random_seed of 0, 1 and 777. However, when random_seed is 42, the algorithm converged to very different value. Also, it took 10 times more iterations to converge, which looks quite wrong.

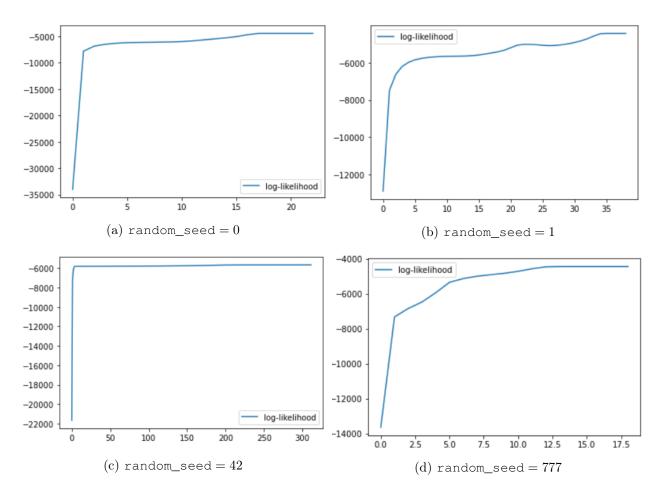


Figure 2: The trace plot of the log-likelihood