

Homework 4

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Question 1: Hoff 6.1 (d)

d) Describe the effects of the prior distribution for γ on the results.

```
## Q1: Hoff 6.1

# data and prior
man_A <- c(1, 0, 0, 1, 2, 2, 1, 5, 2, 0, 0, 0, 0, 0, 0, 1, 1, 1, 0, 0, 0, 1, 1, 2, 1, 3,
           2, 0, 0, 3, 0, 0, 0, 2, 1, 0, 2, 1, 0, 0, 1, 3, 0, 1, 1, 0, 2, 0, 0, 2, 2, 1,
           3, 0, 0, 0, 1, 1)
man_B <- c(2, 2, 1, 1, 2, 2, 1, 2, 1, 0, 2, 1, 1, 2, 0, 2, 2, 0, 2, 1, 0, 0, 3, 6, 1, 6,
           4, 0, 3, 2, 0, 1, 0, 0, 0, 3, 0, 0, 0, 0, 0, 1, 0, 4, 2, 1, 0, 0, 1, 0, 3, 2,
           5, 0, 1, 1, 2, 1, 2, 1, 2, 0, 0, 0, 2, 1, 0, 2, 0, 2, 4, 1, 1, 1, 2, 0, 1, 1,
           1, 1, 0, 2, 3, 2, 0, 2, 1, 3, 1, 3, 2, 2, 3, 2, 0, 0, 0, 1, 0, 0, 0, 1, 2, 0,
           3, 3, 0, 1, 2, 2, 2, 0, 6, 0, 0, 0, 2, 0, 1, 1, 1, 3, 3, 2, 1, 1, 0, 1, 0, 0,
           2, 0, 2, 0, 1, 0, 2, 0, 0, 2, 2, 4, 1, 2, 3, 2, 0, 0, 0, 1, 0, 0, 1, 5, 2, 1,
           3, 2, 0, 2, 1, 1, 3, 0, 5, 0, 0, 2, 4, 3, 4, 0, 0, 0, 0, 0, 2, 2, 0, 0, 2,
           0, 0, 1, 1, 0, 2, 1, 3, 3, 2, 2, 0, 0, 2, 3, 2, 4, 3, 3, 4, 0, 3, 0, 1, 0, 1,
           2, 3, 4, 1, 2, 6, 2, 1, 2, 2)
sum_A <- sum(man_A)
sum_B <- sum(man_B)
n_A <- length(man_A)
n_B <- length(man_B)

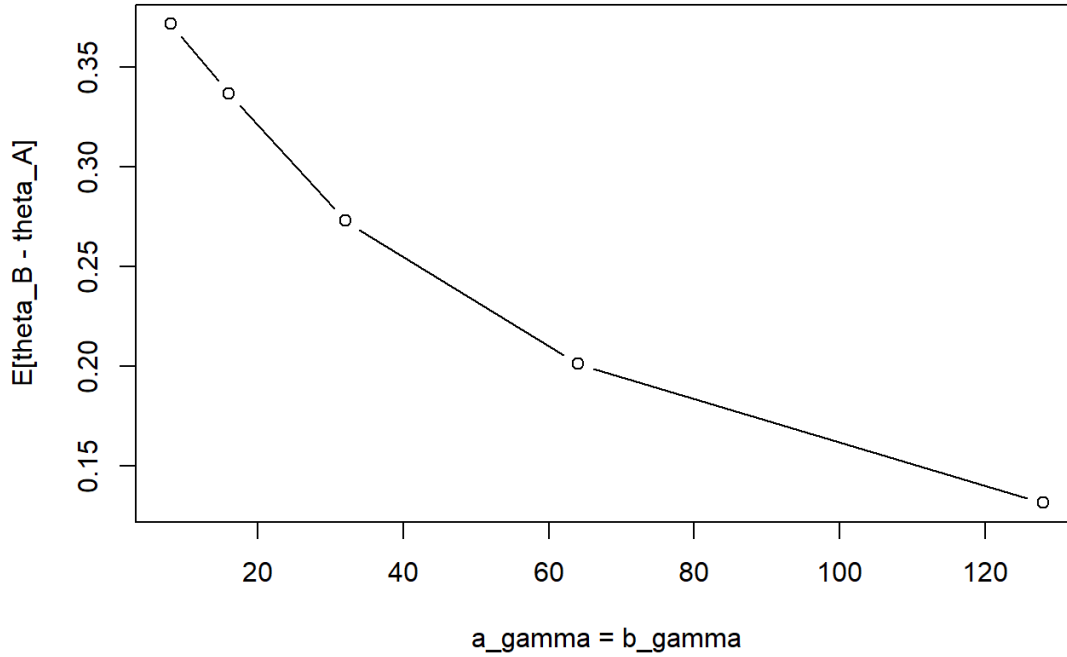
a_theta <- 2; b_theta <- 1
ab_gamma <- c(8,16,32,64,128)
m <- length(ab_gamma)
S <- 5000

# starting point
theta_0 = a_theta / b_theta
gamma_0 = 1
gibbs_theta = matrix(nrow = m, ncol = S + 1)
gibbs_gamma = matrix(nrow = m, ncol = S + 1)
gibbs_theta[,1] = theta_0
gibbs_gamma[,1] = gamma_0

# gibbs sampler
for (s in 2:(S+1)) {
  gibbs_theta[,s] = rgamma(m, shape = sum_A + sum_B + a_theta,
                           rate = n_A + n_B * gibbs_gamma[,s-1] + b_theta)
  gibbs_gamma[,s] = rgamma(m, shape = sum_B + ab_gamma,
                           rate = n_B * gibbs_theta[,s] + ab_gamma)
}

# analysis
theta_A = gibbs_theta
theta_B = gibbs_theta * gibbs_gamma
E_BmA = rowMeans(theta_B - theta_A)
plot(ab_gamma, E_BmA, "b",
     xlab = "a_gamma = b_gamma",
     ylab = "E[theta_B - theta_A]",
     main = "Effects of Different Priors on the Results")
```

Effects of Different Priors on the Results



- d. The plot shows that the expectation of the difference between θ_A and θ_B decreases to 0 as a_γ, b_γ increases to infinity. This could be seen from the posterior mean of the full conditional distribution of γ , which is $\frac{n_B}{n_B\theta + b_\gamma} \frac{\sum y_{Bi}}{n_B} + \frac{b_\gamma}{n_B\theta + b_\gamma} \frac{a_\gamma}{b_\gamma}$. This converges to 1 when prior sample size $a_\gamma = b_\gamma \rightarrow \infty$. Since $\frac{\theta_B}{\theta_A} = \gamma$, $E(\theta_A|y_A, y_B) = E(\theta_B|y_A, y_B)$ and thus $E(\theta_B - \theta_A|y_A, y_B) = 0$.

Question 2:

****a)** Show the exact steps involved in an algorithm for sampling from the posterior distribution for λ and γ .*8

We know from homework 3 that $p(\lambda|\gamma, x, y) = \text{Gamma}(\sum y_i + 1, n_0 + n_1\gamma + 1)$ given n_0, n_1 as number of observations when $x_i = 0$ and $x_i = 1$ respectively.

We just need to derive the full conditional of γ .

$$\begin{aligned}
 p(\gamma|\lambda, x, y) &\propto p(\gamma, \lambda, x, y) \\
 &\propto p(x, y|\lambda, \gamma)p(\lambda, \gamma) \\
 &\propto p(x, y|\lambda, \gamma)p(\gamma) \\
 &= p(y|x, \lambda, \gamma)p(\gamma) \\
 &\propto \prod_{i=1}^n (\lambda \gamma^{x_i})^{y_i} e^{-\lambda \gamma^{x_i}} \gamma^{1-1} e^{-\gamma} \\
 &= \prod_{i=1}^n \lambda^{y_i} \gamma^{x_i y_i} e^{-\lambda \gamma^{x_i}} e^{-\gamma} \\
 &\propto \gamma^{\sum x_i y_i} e^{-\lambda(n_0 + n_1 \gamma)} e^{-\gamma} \\
 &\propto \gamma^{\sum x_i y_i} e^{-(\lambda n_1 \gamma + \gamma)} \\
 &= \gamma^{\sum x_i y_i} e^{-\gamma(\lambda n_1 + 1)} \\
 &= \text{Gamma}(\sum x_i y_i + 1, \lambda n_1 + 1)
 \end{aligned}$$

Therefore, the exact steps involved in gibbs sampler are the following:

- Start with initial value λ_0 and γ_0
- For iterations $t=1, \dots, 1000$

- Sample $\lambda^{(t)}$ from the conditional posterior distribution

$$p(\lambda|\gamma, x, y) = \text{Gamma}(\sum y_i + 1, n_0 + n_1 \gamma^{t-1} + 1)$$

- Sample $\gamma^{(t)}$ from the conditional posterior distribution

$$p(\gamma|\lambda, x, y) = \text{Gamma}(\sum x_i y_i + 1, \lambda^{t-1} n_1 + 1)$$

- This generates a dependent sequence of parameter values.

b) Simulate data and generate samples from the posterior distribution for these data.

```
# simulate data and prior
lambda = gamma = 1
set.seed(1)
treated = rpois(n=50, lambda=1)
control = rpois(n=50, lambda=1)

sum_tr <- sum(treated)
sum_con <- sum(control)
n_tr <- length(treated)
n_con <- length(control)

S = 1000 # number of samples to draw

# starting point
lambda_0 = 1
gamma_0 = 1
gibbs_lambda = matrix(nrow = 1, ncol = S + 1)
gibbs_gamma = matrix(nrow = 1, ncol = S + 1)
gibbs_lambda[,1] = lambda_0
gibbs_gamma[,1] = gamma_0

# gibbs sampler
set.seed(1234)
for (s in 2:(S+1)){
  gibbs_lambda[,s] = rgamma(1, shape = sum_tr + sum_con + 1,
                           rate = n_con + n_tr * gibbs_gamma[,s-1] + 1)
  gibbs_gamma[,s] = rgamma(1, shape = sum_tr+1,
                           rate = n_tr * gibbs_lambda[,s] + 1)
}
```

c) Use your code to - estimate the posterior mean and a 95% credible interval for $\log(\lambda)$ - estimate the predictive distribution for subjects having $x_i = 0$ and $x_i = 1$. Are these predictive distributions different?

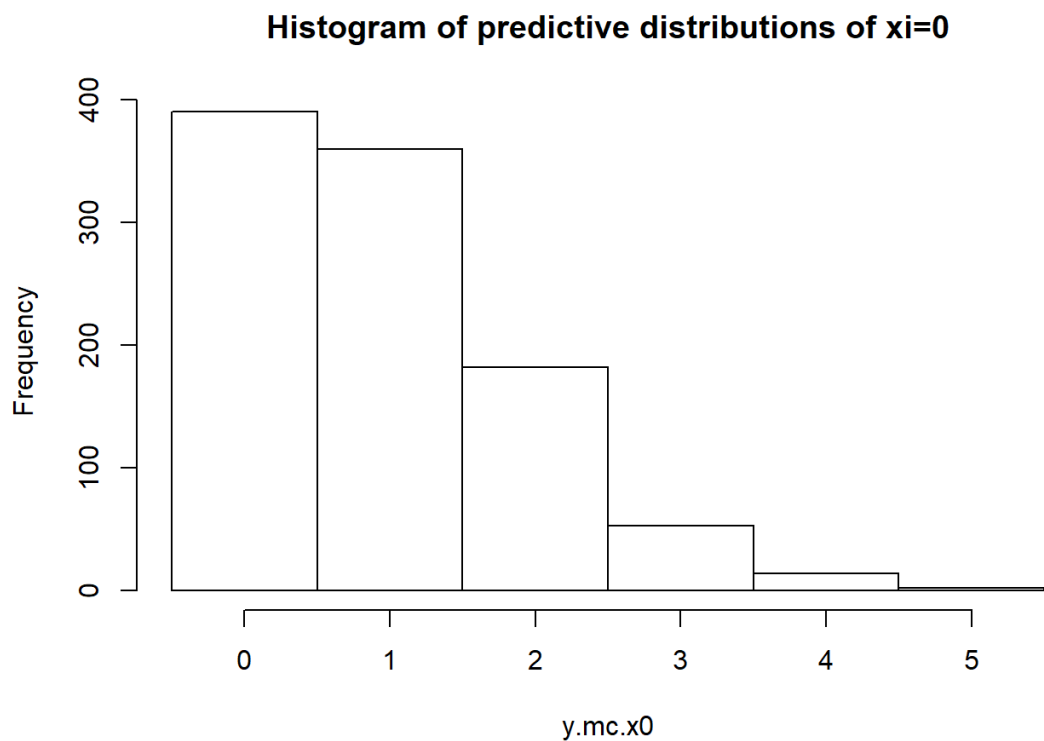
```
# posterior mean for log(lambda)
mean(log(gibbs_lambda[1,])) #-0.0721
```

```
## [1] -0.07213057
```

```
# 95% credible interval for log(lambda)
quantile(log(gibbs_lambda[1,]), c(0.025,0.975))
```

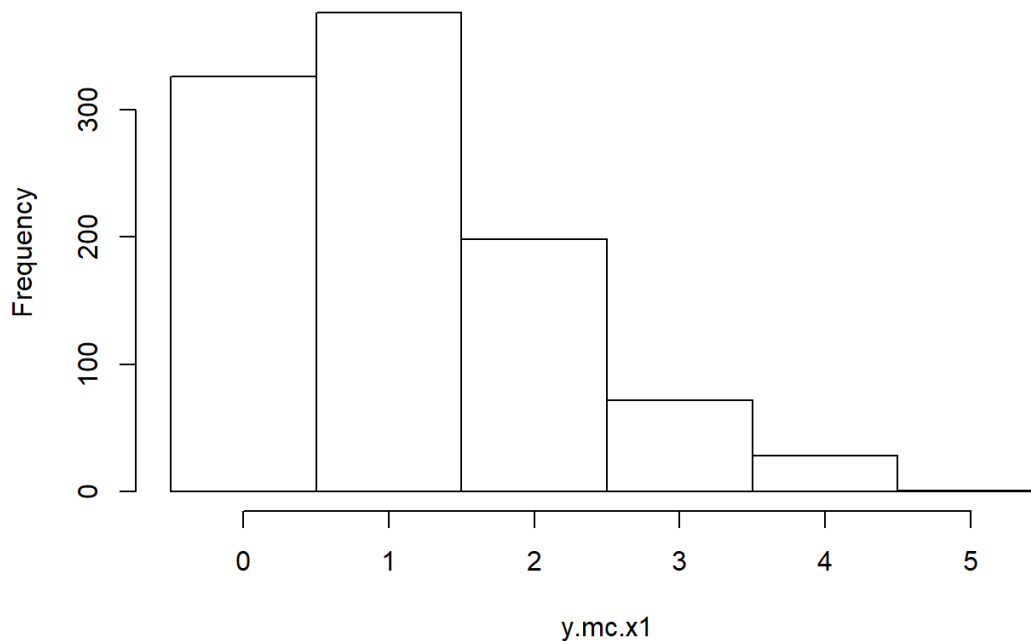
```
##      2.5%      97.5%
## -0.3423654  0.1977386
```

```
# predictive distribution
#1) for  $x_i = 0$ 
y.mc.x0 = rpois(1001, gibbs_lambda[1,])
hist(y.mc.x0, breaks = seq(-0.5, max(y.mc.x0)+0.5,1), main="Histogram of predictive distributions of  $x_i=0$ ")
```



```
#2) for  $x_i = 1$ 
y.mc.x1 = rpois(1001, gibbs_lambda[1,]*gibbs_gamma[1,])
hist(y.mc.x1, breaks = seq(-0.5, max(y.mc.x1)+0.5,1), main="Histogram of predictive distributions of  $x_i=1$ ")
```

Histogram of predictive distributions of $x_i=1$



Yes, these predictive distributions are slightly different in terms of the frequency of 0 and 1. When $x_i = 0$, the frequency of 0 is the highest, while $x_i = 1$, the frequency of 1 is the highest.

8d) Run convergence diagnostics - is your chain mixing well? What is the effective sample size? Does the mixing differ for λ and γ ?

```
library(truncnorm)
library(coda)

# summary statistics
summary(mcmc(gibbs_lambda[1,], start = 1))
```

```
##
## Iterations = 1:1001
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 1001
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##      Mean      SD      Naive SE Time-series SE
## 0.939502 0.131277 0.004149 0.007157
##
## 2. Quantiles for each variable:
##
## 2.5% 25% 50% 75% 97.5%
## 0.7101 0.8469 0.9324 1.0213 1.2186
```

```
summary(mcmc(gibbs_gamma[1,], start = 1))
```

```
##
## Iterations = 1:1001
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 1001
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean           SD       Naive SE Time-series SE
##      1.177383      0.230249      0.007277      0.013213
##
## 2. Quantiles for each variable:
##
##    2.5%   25%   50%   75%  97.5%
## 0.7888 1.0130 1.1591 1.3158 1.6862
```

```
# effective sample size
effectiveSize(mcmc(gibbs_lambda[1,], start = 1))
```

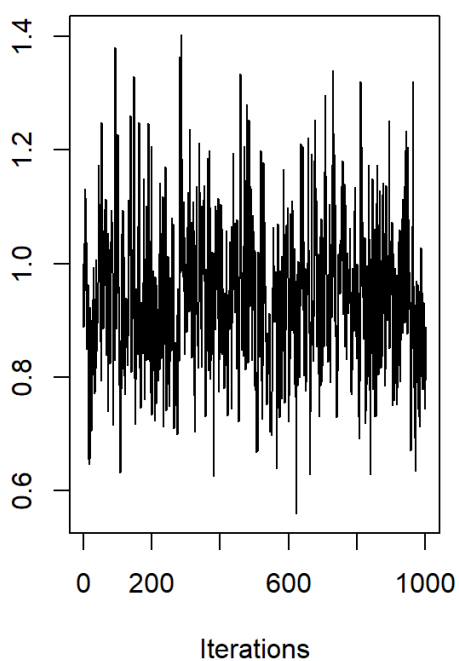
```
##      var1
## 336.4329
```

```
effectiveSize(mcmc(gibbs_gamma[1,], start = 1))
```

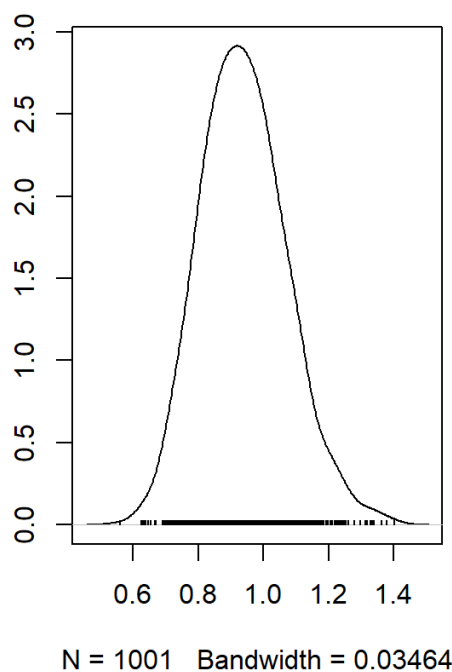
```
##      var1
## 303.6749
```

```
# trace plots
plot(mcmc(gibbs_lambda[1,], start = 1))
```

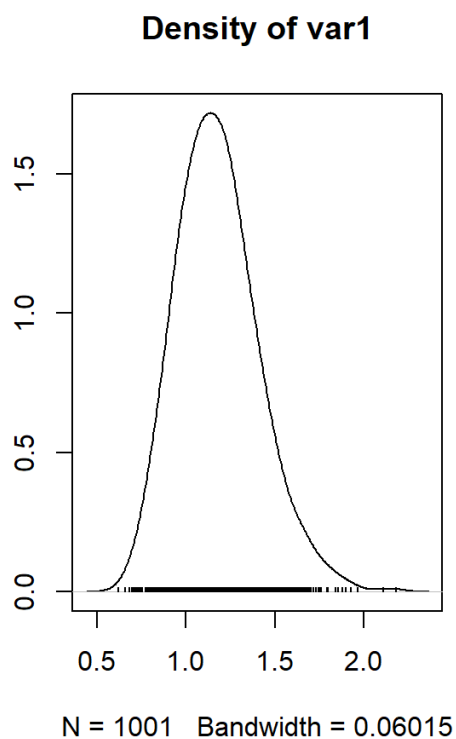
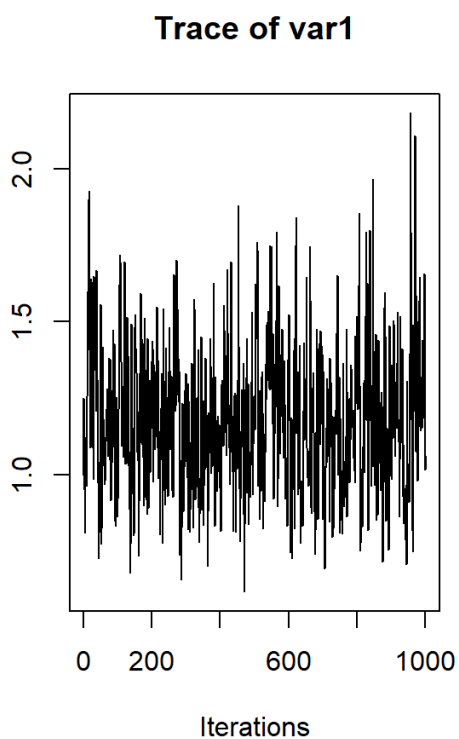
Trace of var1



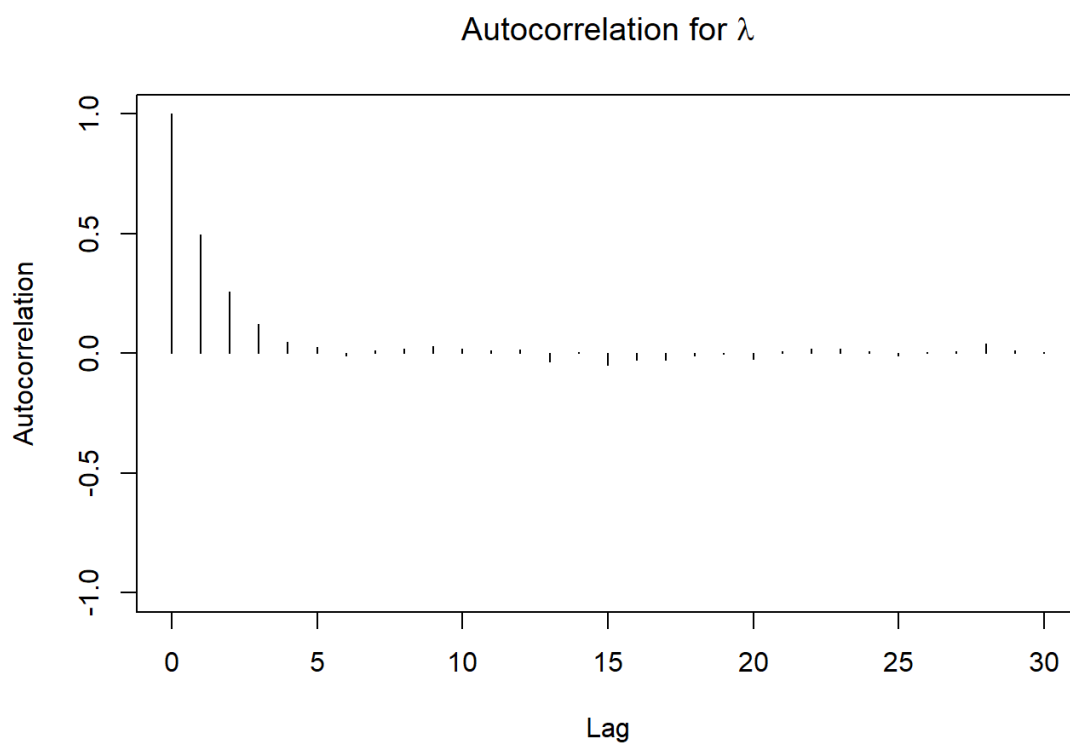
Density of var1



```
plot(mcmc(gibbs_gamma[1,], start = 1))
```

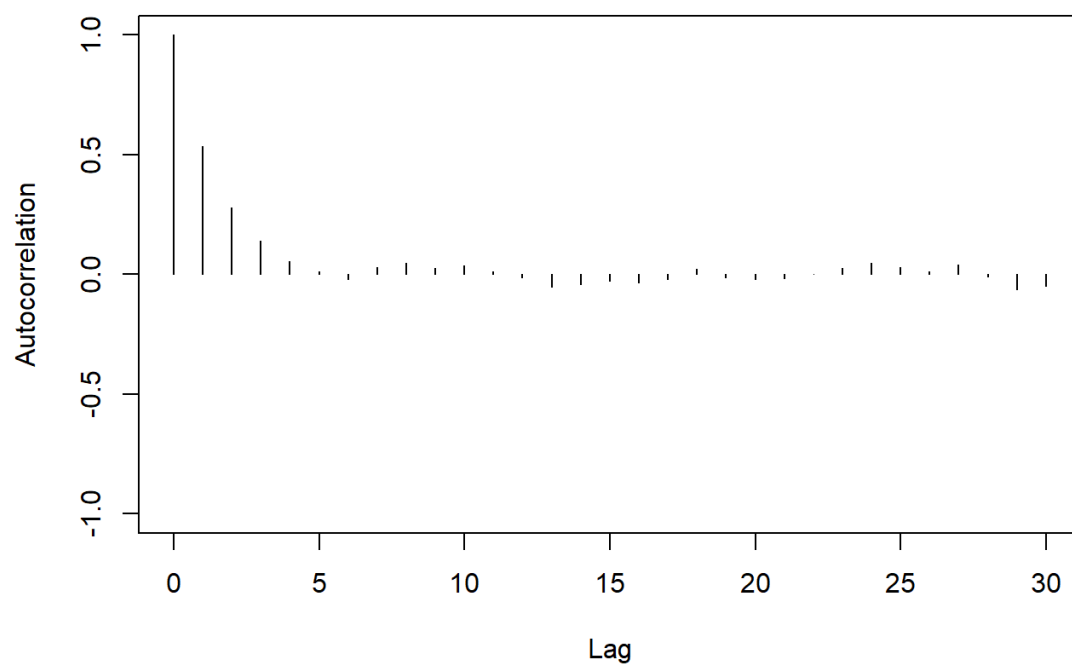


```
# autocorrelation plot
autocorr.plot(mcmc(gibbs_lambda[1,]), main=expression(paste("Autocorrelation for ",lambda)))
```



```
autocorr.plot(mcmc(gibbs_gamma[1,]), main=expression(paste("Autocorrelation for ",gamma)))
```

Autocorrelation for γ



From the trace plots of both λ and γ , we can see that there isn't cyclic local trends in the mean. Also from the autocorrelation plots, we can see that after 6-7 lags, the autocorrelation dies out, which means a pretty good mixing from observation.