## Conditioning, Inference and Probability Logic Bayesian Statistics and Machine Learning

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### Outline

- Conditioning and marginalizing
  - Independence
  - Products of elementary outcomes and possible worlds
  - Joint probabilities
  - Marginal probabilities, sum rule
  - Assigning joint probabilities, product rule
  - Extension of the product rule: the chain rule
  - Inverting conditioning, inference via Bayes' rule
  - Summary: conditioning, joint and marginal probabilities
- Probabilistic Machine Learning terminology
- Probability logic
  - Deductive vs. probabilistic reasoning
  - Justifying probability: logical perspective
  - Summary: probability logic

## Reminder: ingredients of probability

- Probability is built upon the concept of the *possible world* or *elementary outcome*  $\in W$ .
- Not all possible events need to be assigned a probability, only those that are members of a given  $\sigma$ -algebra  $\mathcal F$  .

**Definition**: A *probability space* is a tuple  $(W, \mathcal{F}, P)$ , where  $\mathcal{F}$  is a  $\sigma$ -algebra over W and  $P : \mathcal{F} \to [0, 1]$ , with the properties:

P1 
$$P(W) = 1$$

P2 If 
$$U, V \in \mathcal{F}$$
 and  $U \cap V = \emptyset$ , then  $P(U \cup V) = P(U) + P(V)$ .

**Important consequence**: if  $\mathcal{F}=2^W$ , then it is sufficent to specify P(w) for all  $w \in W$ . P2 then allows you to compute P(U) for any  $U \subseteq W$ .

From here on, we will assume that  $\mathcal{F} = 2^W$ !

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From here on, we will assume that  $\mathcal{F} = 2^W ! \bigcirc$ 



## Reminder: conditioning

Updating of probability when information in the form 'the real world is in U' or 'U just happended' becomes available.

**Conditioning**: If  $U, V \in \mathcal{F}$  and P(U) > 0, then  $P(V|U) = \frac{P(V \cap U)}{P(U)}$ .

Bayes' Rule:  $P(U|V) = P(V|U)\frac{P(U)}{P(V)}$ 

### Independence

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Two events  $U, V \in \mathcal{F}$  are independent if the probability of V is not updated on learning U and vice versa:

$$P(V|U) = P(V) = \frac{P(V \cap U)}{P(U)} \Rightarrow P(V \cap U) = P(V)P(U)$$

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Assume: the set W of possible worlds can be factorized into two sets,  $W = D \times H$ , i.e. the elements of W are tuples j = (d, h) where  $d \in D$  and  $h \in H$ .

- D be the possible elementary outcomes of rolling a die,  $D = \{d_1, \ldots, d_6\}.$
- $H = \{h_1, h_2\}$  be the set of hypotheses  $h_1 =$  'the die is fair', and  $h_2 =$  'the die will show only the numbers 1,2,3'.
- Then,  $(d_3, h_1) = '$ the die showed 3 and the die is fair'
- Let F='The die is fair'=  $\{(d_1,h_1),(d_2,h_1),\ldots,(d_6,h_1)\}$
- Let  $S_3$ ='The die showed 3'= { $(d_3, h_1), (d_3, h_2)$ }
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### Joint probabilities

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The set W of possible worlds can be factorized into two sets, W=D\times H. The elements of W are tuples j=(d,h) where d\in D and h\in H. F= "The die is fair' (i.e. h_1 is true) =\{(d_1,h_1),(d_2,h_1),\ldots,(d_6,h_1)\}. S_i="The die showed i'= \{(d_i,h_1),(d_i,h_2)\}. 'The die showed 3 and the die is fair'= (d_3,h_1)=S_3\cap F
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**Definition**: the probabilities

$$P(j) = P((d,h))$$

are called *joint probabilities* of (d, h), because they represent the probability of d and h happening together.

## Marginal probabilities, sum rule

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```

The probability of F can be computed via P2, because  $\{(d_i, h_1)\} \cap \{(d_j, h_1)\} = \emptyset$  for  $i \neq j$ .

**Definition**: the probability

$$P(F) = P(\bigcup_{i} \{(d_{i}, h_{1})\}) = \sum_{i=1}^{6} P(\{(d_{i}, h_{1})\})$$

P(F) is called marginal probability of  $h_1$  because the  $d_i$  have been marginalized in order to compute P(F). This is the sum rule of probability.

## Assigning joint probabilities

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To apply the sum rule, we need the probabilities  $P((d_i, h_1))$ . Assigning them directly might be difficult:

- Principle of indifference:  $P((d, h)) = \frac{1}{|D||H|} = \frac{1}{12}$ ??
- But then,  $P((d_5, h_2)) > 0$ , but it should be impossible.
- ⇒ this makes no sense.

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We can write  $(d_3, h_1) = S_3 \cap F$ 

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This is known as the product rule.

**Question**: can we assign joint probabilities via the product rule i.e. first determine P(U|V) and then P(V)?

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## **Question**: can we assign probabilities via conditioning and the product rule?

- Given that  $h_1$  is true, we assign  $P(d_i|h_1 = \text{true}) = P(S_i|F) = \frac{1}{6}$ .
- Given that h<sub>2</sub> is true, we assign
  - $P(d_i|h_2 = \text{true}) = P(S_i|F) = \frac{1}{3}$  for i = 1, 2, 3
  - $P(d_i|h_2 = \text{true}) = P(S_i|\overline{F}) = 0 \text{ for } i = 4, 5, 6$

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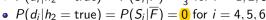
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- Not knowing anything else, we assign  $P(F) = P(\overline{F}) = \frac{1}{2}$ .
- Applying the principle of indifference seems justified here.

Now we can compute the joint probabilities

$$P((d_2, h_1)) = P(S_2|F)P(F) = \frac{1}{6} \cdot \frac{1}{2}$$
  
 $P((d_4, h_2)) = P(S_4|\overline{F})P(\overline{F}) = 0 \cdot \frac{1}{2}$ 

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Product rule so far: for  $U, V \in W$ , P(U) > 0:

 $P(V \cap U) = P(V|U)P(U)$ 

**Question**: can this be generalized to more than 2 events?

**Answer**: assume:  $U = R \cap S$ . Then

$$P(V \cap R \cap S) = P(V \cap U) = P(V|U)P(U)$$
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Generally, for a sequence of events or propositions  $S_i,\ i=1,\ldots,N$  we obtain the **chain rule of probability** 

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with  $P(S_N | \cap_{j=N+1}^N S_j) = P(S_n)$ , which holds for any ordering of the  $S_i$ .

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A consequence is (assume k < N):

$$P(\bigcap_{i=1}^{N} S_{i}) = \prod_{i=1}^{k} P(S_{i} | \bigcap_{j=i+1}^{N} S_{j}) \prod_{i=k+1}^{N} P(S_{i} | \bigcap_{j=i+1}^{N} S_{j})$$
  
=  $P(\bigcap_{i=1}^{k} S_{i} | \bigcap_{i=k+1}^{N} S_{i}) P(\bigcap_{i=k+1}^{N} S_{i})$ 

E.g.:  $P(R \cap S \cap T) = P(R|S \cap T)P(S \cap T) = P(R \cap S|T)P(T)$ 

For a sequence of events or propositions  $S_i$ ,  $i=1,\ldots,N$  we obtain the **chain rule of probability** 

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## Inverting conditioning, inference via Bayes' rule

### Reminder:

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The set W of possible worlds can be factorized into two sets, W=D\times H. The elements of W are tuples j=(d,h) where d\in D=\{d_1,\ldots,d_6\} and h\in H=\{h_1,h_2\}. F= The die is fair' (i.e. h_1 is true) = \{(d_1,h_1),(d_2,h_1),\ldots,(d_6,h_1)\}. \overline{F}= The die shows only 1,2,3' (i.e. 'h_2 is true) = \{(d_1,h_2),\ldots,(d_6,h_2)\}=(D\times H)\setminus F. S_j= The die showed i'= \{(d_j,h_1),(d_j,h_2)\}.
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### Assume we observe that the die showed a 3, i.e. $S_3$ happened.

**Question**: what does this tell us about whether the die is fair (F) or not  $(\overline{F})$ ?

**Answer**: in probabilistic terms: what is  $P(F|S_3)$ ? This answer could be provided by Bayes' rule

$$P(F|S_3) = P(S_3|F) \frac{P(F)}{P(S_3)} = \frac{P(S_3 \cap F)}{P(S_3)}$$

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Assume we observe that the die showed a  $\frac{3}{1}$ , i.e.  $\frac{5}{3}$  happened.

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### **Question**: what is the marginal probability $P(S_3)$ ?

**Answer**: apply the sum rule

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$$= \sum_{i=1}^{2} P(d_3, h_i)$$

$$= P(S_3|F)P(F) + P(S_3|\overline{F})P(\overline{F})$$

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Thus, we obtain

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Assume we observed  $S_4$ , i.e. the die shows a 4. Now we find

$$P(S_4) = P(\{(d_4, h_1), (d_4, h_2)\})$$

$$= P(S_4|F)P(F) + P(S_4|\overline{F})P(\overline{F})$$

$$= \frac{1}{6} \cdot \frac{1}{2} + 0 \cdot \frac{1}{2}$$

$$= \frac{1}{12} \qquad (1)$$

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Assume we observed  $S_4$ , i.e. the die shows a 4. The probabilites  $P(F|S_4)$  and  $P(\overline{F}|S_4)$  now are

$$(F|S_4) = P(S_4|F)\frac{P(F)}{P(S_4)} = \frac{1}{6} \cdot \frac{\frac{1}{2}}{\frac{1}{12}} = \mathbf{1}$$

$$(\overline{F}|S_4) = P(S_4|\overline{F})\frac{P(\overline{F})}{P(S_4)} = 0 \cdot \frac{\frac{1}{2}}{\frac{1}{12}} = 0$$

(2)

⇒ we are now *certain* that the die is fair.

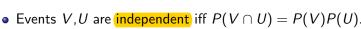
## Summary: conditioning, joint and marginal probabilities

- Events V, U are independent iff  $P(V \cap U) = P(V)P(U)$ .
- If U, V independent: P(V|U) = P(V) and P(U|V) = P(U).
- Assume the set W of possible worlds can be factorized into two sets, W = D × H.
  - Elementary outcomes are tuples  $w = (d, h), d \in D, h \in H$ .
  - Joint probability P((d, h)) that d and h happen together.
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  - Product rule:  $P(U \cap V) = P(U|V)P(V)$
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### Probabilistic Machine Learning terminology



- D: observed variables, the data.
- H: hypotheses or hidden variables or latent variables.
  often assigned via Principle of Indifference.
- P(F): **prior** probability of F='the coin is fair', i.e. before observing data.
- $P(S_3|F)$ : **likelihood** of F. Conditional prob. of data  $(S_3 = '$  die shows three') given hidden/latent  $F \subseteq D \times H$ .
- $P(S_3)$  (if  $S_3$  was observed): marginal probability of data. Also: evidence.
- $P(F|S_3)$ : **posterior** probability of F, i.e. after observing data.
- Hierarchical model:  $P(S_3, F) = P(S_3|F)P(F)$ .
  - Decompose difficult-to-assign  $P(S_i, F)$  into more accessible  $P(S_i|F)$  and P(F).



#### Reminder:

Probability space  $(W, \mathcal{F}, P)$ , where  $\mathcal{F} = 2^W$ .

### Propositional logic

- Propositions A,B
- $A \in \{\text{true}, \text{false}\}$
- Logical operators
  - AND: A ∧ B
  - OR:  $A \vee B$
  - NOT: ~ A
- Deductive reasoning.
   Validity, soundness.

### Probability

- Propositions  $A, B \in \mathcal{F}$
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### Deductive vs. probabilistic reasoning: example

Let the propositions A, B. Assume you are doing a Neuroscience experiment.

- A ="stimulus X is presented"
- B ="neuron Y fires"
- If stimulus X is presented, then neuron Y fires.
- Stimulus X is presented.
- Therefore, neuron Y fires.

$$\begin{array}{c}
A \Rightarrow B \\
A \\
\hline
B
\end{array}$$

This is an example of modus ponens

The  $\Rightarrow$  is a (material) implication.

The conclusion B is true *given* that the premises  $A \Rightarrow B$  and A are both true.

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 $A \Rightarrow B$  and A are both true.

**Probability**:  $P(B|A \cap (A \Rightarrow B)) = 1$ .

**Thus:** need to express the implication  $A \Rightarrow B$  with set operators.

Truth table of $A \Rightarrow B$					
A	В	$A \Rightarrow B$	$\sim A \vee B$		
true	true	true	true		
true	false	false	false		
false	true	true	true		
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$$A \Rightarrow B = \sim A \vee B$$
  
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### Uncertain premises

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# **Question**: what if we are not sure of the premises, i.e. they are not *given*?

Propositional logic: we're stuck. No answer.

**Probability:** what can we say about the *marginal* probability of the conclusion, P(B)?

### Assume

$$P(A \Rightarrow B) = p_I$$

• 
$$P(A) = p_A$$

**Question**: What is 
$$P(B) = P(A \cap B) + P(\bar{A} \cap B)$$

 $A \Rightarrow B$ 

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**Question**: What is  $P(B) = P(A \cap B) + P(\bar{A} \cap B)$ 

### Uncertain premises

#### Reminder:

Propositions A = "stimulus X is presented", B = "neuron Y fires". Modus ponens:  $\frac{A \Rightarrow B}{A} = \bar{A} \cup B.$   $P(B|A \cap (A \Rightarrow B) = P(B|A \cap B) = 1.$ 

**Question**: what if we are not sure of the premises, i.e. they are not *given*?

**Propositional logic:** we're stuck. *No answer.* 

**Probability:** what can we say about the *marginal* probability of the conclusion, P(B)?

### Assume

- $P(A \Rightarrow B) = p_I$
- $\bigcirc$

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 $A \Rightarrow B = \overline{A} \cup B$ .  $P(B) = P(A \cap B) + P(\overline{A} \cap B)$  $P(A \Rightarrow B) = p_I \text{ and } P(A) = p_A$ .

We compute  $P(A \cap B) = P(B|A)P(A)$  via the product rule.

Note that

$$P(A \Rightarrow B) = P(\bar{A} \cup B)$$

$$= P(\bar{A} \cup B)$$

$$= P(\bar{A} \cap \bar{B})$$

$$= 1 - P(A \cap \bar{B})$$

Thus 
$$P(\bar{B} \cap A) = P(\bar{B}|A)P(A) = 1 - p_I$$
 and  $P(B|A) = 1 - \frac{1 - p_I}{p_A}$ .  
Whence  $P(B \cap A) = P(B|A)P(A) = p_A + p_I - 1$ .

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P(A \Rightarrow B) & = & P(\bar{A} \cup B) \\
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We cannot compute  $P(\bar{A} \cap B)$  from our premises, but we know  $P(\bar{A} \cap B) \geq 0$ . Thus

$$P(B) \geq p_I + p_A - 1$$

i.e. we know a lower bound on P(B).

In the limit of certainty, i.e.  $p_I = p_A = 1$ , we obtain P(B) = 1. I.e. the deductive result is recovered.

Reminder:  $A \Rightarrow B = \overline{A} \cup B$ .  $P(B) = P(A \cap B) + \underbrace{P(\overline{A} \cap B)}_{P(A \Rightarrow B) = p_I}$  and  $P(A) = p_A$ .  $P(B \cap A) = P(B|A)P(A) = p_A + p_I - 1$ .



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### Justifying probability: the logical perspective

**Goal:** extend *propositional* logic to a *probability* logic. Basic ingredients:

- Propositions with truth values: T,F.
- Logical connectives:  $\sim$ ,  $\vee$ ,  $\wedge$ .  $\vee$  is not necessary.
- Inference schemes: modus ponens, tollens etc.

We'd like a more fine-grained representation of uncertainty, which should reduce to propositional reasoning in the limit of certainty.

**Question:** [Cox 1948, Jaynes 2003] what kind of axioms/desiderata are required for this?

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- Not the only possible choice.
- Further convention: greater plausibility corresponds to a greater number.
- Continuity property: an infintesimally greater plausibility corresponds to an infinitesimally greater number.
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  - Conditional plausibility of A given B: AB.

### D2: Qualitative correspondence with common sense.

- This desideratum is about reasoning with conjunctions.
- Assume plausibilities (A|C') > (A|C) and  $(B|A \wedge C') = (B|A \wedge C)$ .
- Then  $(A \wedge B|C') > (A \wedge B|C)$ , never the other way around!
- Likewise, if (A|C') = (A|C) and  $(B|A \wedge C') > (B|A \wedge C)$ .
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### **D3**: Consistency.

- If a conclusion can be reasoned out in more than one way, then every possible way must lead to the same result.
- Equivalent states of knowledge are assigned the same plausibilities.
- No available information (relevant to the conclusion) is ignored.

### The quantitative rules of probabilistic logic

The desiderata D1-3 can be shown to result in the following quantitative rules for plausible reasoning, where the plausibility is denote by P, and is (of course) a probability:



Conditioning and marginalizing

 $P(A \land B|C) = P(B|A \land C)P(A|C) = P(A|B \land C)P(B|C)$ 



2  $P(A|B) + P(\sim A|B) = 1.$ 

These rules for conjunction and negation are enough to construct all other logical connectives. Inference can be done by conditioning the conclusion on the premises.



## Summary: probability logic

- Propositions <u>A,B</u> in propositional logic correspond to propositions in probability theory.
- Logical operators correspond to set operations.
- Deductive reasoning can be expressed probabilistically.
- Probability provides answers beyond propositional logic.
- In the limit of certainty, deductive results are recovered.