## Lecture 5: Nonlinear Regression Functions

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- Nonlinear Regression Functions:
- Nonlinear in Xs
- Polynomials in X
- 4 Logarithms
- 5 Interactions Between Independent Variables

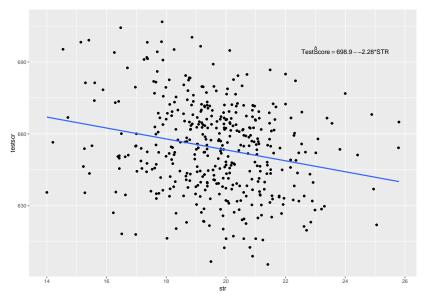
#### Nonlinear Regression Functions:

#### Introduction

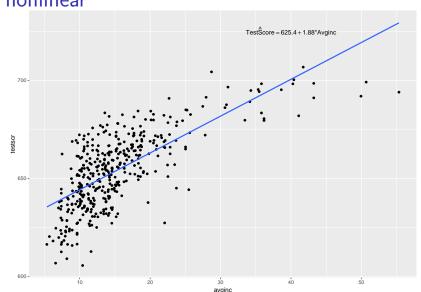
- Everything so far has been linear in the X's. But the linear approximation is not always a good one.
- We extend nonlinear into two cases
- nonlinear in Xs
  - Polynomials, Logarithms and Interactions
  - The multiple regression framework can be extended to handle regression functions that are nonlinear in one or more X.
  - the difference from a standarad multiple OLS regression ishow to explain estimating coefficients.
- nonlinear in function
  - Discrete Dependent Variables or Limited Dependent Variables
  - Linear function is not a good prediciton function.

#### Nonlinear in Xs

### The TestScore – STR relation looks linear (maybe)



## But the TestScore – Income relation looks nonlinear



## Nonlinear Regression Regression Functions – General Ideas

Our regression model is

$$Y_i = \beta_0 + \beta_1 X_{1,i} + \ldots + \beta_k X_{k,i} + u_i$$

- The effect of Y on a change in  $X_j$  by 1 (unit) is constant and equals  $\beta_j$ :
- If a relation between Y and X is nonlinear:
  - The effect on Y of a change in X depends on the value of X that is, the marginal effect of X is not constant.
  - A linear regression is misspecified the functional form is wrong
  - The estimator of the effect on Y of X is biased(a special case of OVB)
- The solution to this is to estimate a regression function that is nonlinear in X.

#### **OLS Assumptions Still Hold**

General formula for a nonlinear population regression model:

$$Y_i = f(X_{1,i}, X_{2,i}, ..., X_{k,i}) + u_i \\$$

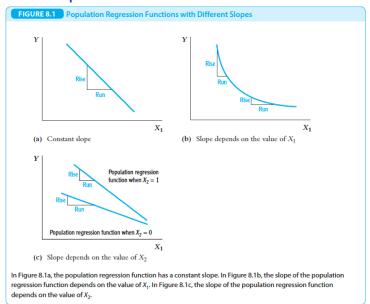
#### Assumptions:

- $E[u_i|X_{1,i},X_{2,i},...,X_{k,i}]=0$  implies that f is the conditional expectation of Y given the X's.
- $\ \ \, \textbf{2} \ \, (X_{1,i},X_{2,i},...,X_{k,i}) \ \, \text{are i.i.d.}$
- Large outliers are rare.
- No perfect multicollinearity.

#### Two Cases:

- lacksquare The effect of change in  $X_1$  on Y depends on  $X_1$  itself.
  - for example: the effect of a change in class size is bigger when initial class size is small
- $\ensuremath{\mathbf{2}}$  The effect of change in  $X_1$  on Y depends on another variable, like  $X_2$ 
  - For example: the effect of class size on test score depends on the percentage of disadvantaged pupils in the class

#### Different Slops



# The Effect on Y of a Change in X in a Nonlinear Specifications

### The Expected Change on Y of a Change in $X_1$ in the Nonlinear Regression Model (8.3)

KEY CONCEPT

8.1

The expected change in Y,  $\Delta Y$ , associated with the change in  $X_1$ ,  $\Delta X_1$ , holding  $X_2, \ldots, X_k$  constant, is the difference between the value of the population regression function before and after changing  $X_1$ , holding  $X_2, \ldots, X_k$  constant. That is, the expected change in Y is the difference:

$$\Delta Y = f(X_1 + \Delta X_1, X_2, \dots, X_k) - f(X_1, X_2, \dots, X_k). \tag{8.4}$$

The estimator of this unknown population difference is the difference between the predicted values for these two cases. Let  $\hat{f}(X_1, X_2, \dots, X_k)$  be the predicted value of Y based on the estimator  $\hat{f}$  of the population regression function. Then the predicted change in Y is

$$\Delta \hat{Y} = \hat{f}(X_1 + \Delta X_1, X_2, \dots, X_k) - \hat{f}(X_1, X_2, \dots, X_k). \tag{8.5}$$

# A General Approach to Modeling Nonlinearities Using Multiple Regression

- Identify a possible nonlinear relationship.
- Specify a nonlinear function and estimate its parameters by OLS.
- Determine whether the nonlinear model improves upon a linear model.
- Plot the estimated nonlinear regression function.
- Estimate the effect on Y of a change in X.

#### Two complementary approaches:

- Polynomials in X
  - The population regression function is approximated by a quadratic, cubic, or higher-degree polynomial.
- Logarithmic transformations
  - Y and/or X is transformed by taking its logarithm
  - this gives a percentages interpretation that makes sense in many applications

## Polynomials in $\boldsymbol{X}$

#### Polynomials in X

Approximate the population regression function by a polynomial:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X^2 ... + \beta_r X_i^r + u_i$$

- This is just the linear multiple regression model except that the regressors are powers of X!
- Estimation, hypothesis testing, etc. proceeds as in the multiple regression model using OLS.
- The coefficients are difficult to interpret, but the regression function itself is interpretable.

# Testing the null hypothesis that the population regression function is linear

$$H_0: \beta_2=0, \beta_3=0,..., \beta_r=0 \ and \ H_1: at \ least \ one \ \beta_j \neq 0$$

• it can be tested using the **F-statistic** 

#### Which degree polynomial should I use?

- How many powers of X should be included in a polynomial regression?
- The answer balances a trade-off between flexibility and statistical precision.
- In many applications involving economic data, the nonlinear functions are smooth, that is, they do not have sharp jumps, or "spikes."
- If so, then it is appropriate to choose a small maximum degree for the polynomial, such as 2, 3, or 4.

#### Example: the TestScore-Income relation

Quadratic specification:

$$TestScore_i = \beta_0 + \beta_1 Income_i + \beta_2 (Income_i)^2 + u_i$$

• Cubic specification:

$$TestScore_i = \beta_0 + \beta_1 Income_i + \beta_2 (Income_i)^2 + \beta_3 (Income_i)^3 + \beta_3 (Income_i)^3 + \beta_4 (Income_i)^3 + \beta_4 (Income_i)^4 + \beta_5 (Income_i)$$

### Estimation of the quadratic specification in R

```
##
## Call:
     felm(formula = testscr ~ avginc + I(avginc^2), data
##
##
## Residuals:
      Min 1Q Median
                             3Q
##
                                    Max
## -44.416 -9.048 0.440 8.348 31.639
##
## Coefficients:
              Estimate Robust s.e t value Pr(>|t|)
##
## (Intercept) 607.30174 2.90175 209.288 <2e-16 ***
## avginc 3.85100 0.26809 14.364 <2e-16 ***
## I(avginc^2) -0.04231 0.00478 -8.851 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0
##
```

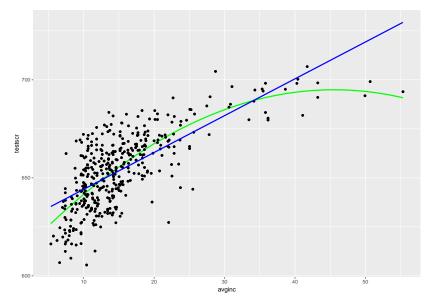
#### Interpreting the estimated regression function

• The OLS regression in quadratic Xs yields

$$\widehat{TestScore} = 607.3 + 3.85 Income - 0.0423 (Income)^{2}$$

$$(2.9) \qquad (0.27) \qquad (0.0048)$$

#### Linear and Quadratic Regression in figure



#### Quadratic vs Linear

- Is the quadratic model better than the linear model?
- We can test the null hypothesis that the regression function is linear against the alternative hypothesis that it is quadratic:

$$H_0:\beta_2=0~and~H_1:\beta_2\neq 0$$

the t-statistic

$$t = \frac{(\hat{\beta}_2 - 0)}{SE(\hat{\beta}_2)} = \frac{-0.0423}{0.0048} = -8.81$$

- Since 8.81>2.58 we reject the null hypothesis (the linear model) at a 1% significance level.
- the F-test

$$F - statistic_{q=2,d=417} = 261.3, p - value \approx 0.00$$

#### Interpreting the estimated regression function

- Predict Change in TestScore for a change in income
- from \$10,000 per capita to \$11,000 per capita:

$$\Delta TestScore = 607.3 + 3.85 \times 11 - 0.0423 \times (11)^{2}$$
$$- [607.3 + 3.85 \times 10 - 0.0423 \times (10)^{2}]$$
$$= 2.96$$

• from \$40,000 per capita to \$41,000 per capita:

$$\Delta TestScore = 607.3 + 3.85 \times 41 - 0.0423 \times (41)^{2}$$
$$- [607.3 + 3.85 \times 40 - 0.0423 \times (40)^{2}]$$
$$= 0.42$$

### Logarithms

#### Logarithmic functions of Y and/or X

- Another way to specify a nonlinear regression model is to use the natural logarithm of Y and/or X.
- Ln(X) = the natural logarithm of X
- Logarithmic transforms permit modeling relations in "percentage" terms (like elasticities), rather than linearly.

#### Review of the Basic Logarithmic functions

$$\begin{split} &ln(1/x) = -ln(x) \\ &ln(ax) = ln(a) + ln(x) \\ &ln(x/a) = ln(x) - ln(a) \\ &ln(x^a) = aln(x) \end{split}$$

#### Logarithms and percentages

Because

$$\begin{split} \ln(x + \Delta x) - \ln(x) &= \ln\left(\frac{x + \Delta x}{x}\right) \\ &\cong \frac{\Delta x}{x} \left( when \ \frac{\Delta x}{x} \ is \ very \ small \right) \end{split}$$

For example:

$$ln(1+0.01) = ln(101) - ln(100) = 0.00995 \cong 0.01$$

#### The three log regression specifications:

linear-log

$$Y_i = \beta_0 + \beta_1 ln(X_i) + u_i$$

Opening the second s

$$ln(Y_i) = \beta_0 + \beta_1 X_i + u_i$$

log-log

$$ln(Y_i) = \beta_0 + \beta_1 ln(X_i) + u_i$$

- The interpretation of the slope coefficient differs in each case.
- The interpretation is found by applying the general "before and after" rule: "figure out the change in Y for a given change in X."

#### I. Linear-log population regression function

• the Regression Model:

$$Y_i = \beta_0 + \beta_1 ln(X_i) + u_i$$

• Change X:

$$\begin{split} \Delta Y &= [\beta_0 + \beta_1 ln(X + \Delta X)] - [\beta_0 + \beta_1 ln(X)] \\ &= \beta_1 [ln(X + \Delta X) - ln(X)] \\ &\cong \beta_1 \frac{\Delta X}{X} \end{split}$$

• Now  $100 \times \frac{\Delta X}{X} = percentage \ change \ in \ X$ , then

$$\beta_1 \cong \frac{\Delta Y}{\frac{\Delta X}{X}}$$

• Interpretation of  $\beta_1$ : a 1 percent increase in X (multiplying X by 1.01 or  $100 \times \frac{\Delta X}{X}$ ) is associated with a  $0.01\beta_1$  or  $\frac{\beta_1}{100}$  change in

#### Example: the TestScore – log(Income) relation

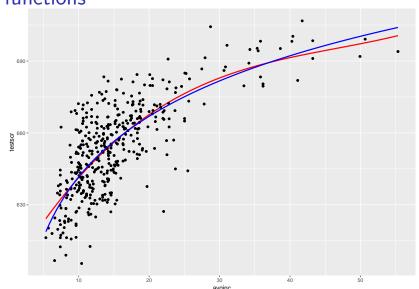
The OLS regression of In(Income) on Testscore yields

$$TestScore = 557.8 + 36.42 \times ln(Income)$$

$$(3.8) \quad (1.4)$$

• Interpretation of  $\beta_1$ : a **1%** increase in Income is associated with an increase in TestScore of **0.36** points on the test.

# Test scores: linear-log and cubic regression functions



#### Case II. Log-linear population regression function

• the regression model is

$$ln(Y_i) = \beta_0 + \beta_1 X_i + u_i$$

• Change X:

$$\begin{split} \ln(\Delta Y + Y) - \ln(Y) &= [\beta_0 + \beta_1(X + \Delta X)] - [\beta_0 + \beta_1 X] \\ \ln(1 + \frac{\Delta Y}{Y}) &= \beta_1 \Delta X \end{split}$$

then

$$\frac{\Delta Y}{Y} \cong \beta_1 \Delta X$$

• Now  $100\frac{\Delta Y}{Y} = percentage \ change \ in \ Y$ , so a change in X by one unit is associated with a  $\beta_1\% = \frac{\beta_1}{100}$  change in Y.

#### Earning function: log-linear functions

The OLS regression of age on earnings yields

$$ln(\widehat{Earnings}) = 2.811 + 0.0096 Age$$

$$(0.018) \quad (0.0004)$$

-According to this regression, when one more year old, earnings are predicted to increase by  $100\times0.0096=0.96\%$ 

#### Case III. Log-linear population regression function

• the regression model is

$$ln(Y_i) = \beta_0 + \beta_1 ln(X_i) + u_i$$

• Change X:

$$\begin{split} \ln(\Delta Y + Y) - \ln(Y) &= [\beta_0 + \beta_1 ln(X + \Delta X)] - [\beta_0 + \beta_1 ln(X)] \\ ln(1 + \frac{\Delta Y}{Y}) &= ln(1 + \frac{\Delta X}{X}) \\ \frac{\Delta Y}{Y} &\cong \beta_1 \frac{\Delta X}{Y} \end{split}$$

- Now  $100\frac{\Delta Y}{Y} = percentage \ change \ in \ Y$  and  $100\frac{\Delta X}{X} = percentage \ change \ in \ X$
- so a 1% change in X by one unit is associated with a  $\beta_1$ % change in Y,thus  $\beta_1$  has the interpretation of an **elasticity**.

#### Test scores and income: log-log specifications

```
##
## t test of coefficients:
##
##
              Estimate Std. Error t value Pr(>|t|)
  (Intercept) 6.3363494 0.0059105 1072.056 < 2.2e-16 **
## loginc 0.0554190 0.0021395 25.903 < 2.2e-16 **
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0
```

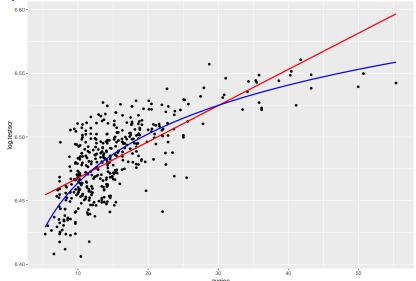
$$ln(\widehat{TestScore}) = 6.336 + 0.055 \times ln(Income)$$

$$(0.006) \quad (0.002)$$

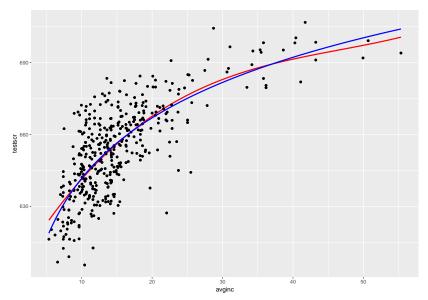
 An 1% increase in Income is associated with an increase of .0554% in TestScore.

# Test scores: The log-linear and log-log

specifications:



### linear-log and cubic regression functions



## Choice of specification should be guided

- By economic logic or theories(which interpretation makes the most sense in your application?),
- t-test and F-test are enough, other formal tests seldom are used in reality.
- ullet Plotting predicted values and use  $\overline{R^2}$  or SER

## Summary

- We already have a very powerful tool for detecting misspecified functional form: the F test for joint exclusion restrictions.
- We can add quadratic terms of any significant variables to a model and to perform a joint test of significance. If the additional quadratics are significant, they can be added to the model.
- It can be difficult to pinpoint the precise reason for functional form misspecification.
- Fortunately, using logarithms of certain variables and adding quadratic or cubic functions are sufficient for detecting many important nonlinear relationships in economics.

# Interactions Between Independent Variables

### Interactions Between Two Binary Variables

- Assume we would like to study the earnings of worker in the labor market
- ullet The population linear regression of  $Y_i$  is

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + u_i$$

- the Dependent Variable:  $\log \ earnings(Y_i, where Y_i = ln(Earnings))$
- Independent Variables: two binary variables
  - $D_1 i = 1$  if the person graduate from college
  - $D_2i = 1$  if the worker's gender is female
- So  $\beta_1$  is the effect on log earnings of having a college degree, holding gender constant, and  $\beta_2$  is the effect of being female, holding schooling constant.

### Interactions Between Two Binary Variables

- The effect of having a college degree in this specification, holding constant gender, is the same for men and women. No reason that this must be so.
- $\bullet$  the effect on  $Y_i$  of  $D_{1i}$  , holding  $D_{2i}$  constant, could depend on the value of  $D_{2i}$
- there could be an interaction between having a college degree and gender so that the value in the job market of a degree is different for men and women.
- ullet The new regression model of  $Y_i$  is

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} \times D_{2i}) + u_i$$

• The new regressor, the product  $D_{1i} \times D_{2i}$ , is called an interaction term or an interacted regressor.

### Interactions Between Two Binary Variables:

ullet The regression model of  $Y_i$  is

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} \times D_{2i}) + u_i$$

 $\bullet$  Then the conditional expectation of Yi for  $D_{1i}=0$  , given a value of  $D_{2i}$ 

$$E(Y_i|D_{1i} = 0, D_{2i} = d_2) = \beta_0 + \beta_1 \times 0 + \beta_2 d_2 + \beta_3 (0 \times d_2) = \beta_0 + \beta_2 d_2 + \beta_3 d_3 + \beta_$$

• Then the conditional expectation of Yi for  $D_{1i}=1$ , given a value of  $D_{2i}$ 

$$E(Y_i|D_{1i}=1,D_{2i}=d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 (1 \times d_2) = \beta_0 + \beta_1 \times 1 + \beta_1 \times 1 + \beta_2 d_2 + \beta_2 \times 1 + \beta_1 \times 1 + \beta_2 d_2 + \beta_3 \times 1 + \beta_2 d_2 + \beta_3 \times 1 + \beta_3 \times$$

### Interactions Between Two Binary Variables:

 The effect of this change is the difference of expected values, which is

$$E(Y_i|D_{1i}=1,D_{2i}=d_2) - E(Y_i|D_{1i}=0,D_{2i}=d_2) = \beta_1 + \beta_3 d_2$$

- In the binary variable interaction specification, the effect of acquiring a college degree (a unit change in  $D_{1i}$ ) depends on the person's gender.
- $\bullet$  If the person is male,thus  $D_{2i}=d_2=0$  ,then the effect is  $\beta_1$
- If the person is female,thus  $D_{2i}=d_2=1$  ,then the effect is  $\beta_1+\beta_3$
- So the coefficient  $\beta_3$  is the difference in the effect of acquiring a college degree for women versus men.

# Application to the STR and the percentage of English learners

- ullet Let  $HiSTR_i$  be a binary variable for STR
  - $HiSTR_i = 1$  if the STR > 20
  - $HiSTR_i = 0$  otherwise
- ullet Let  $HiEL_i$  be a binary variable for English learner
  - $HiEL_i = 1$  if the  $el_pct > 10percent$
  - $HiEL_i = 0$  otherwise

# Application to the STR and the percentage of English learners

- one binary variable, whether the worker has a college degree
- the individual's years of work experience  $(X_i)$ ,
- the first population model is

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + u_i$$

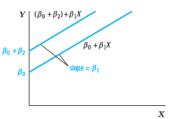
• the second population model is

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 (D_i \times X_i) + u_i$$

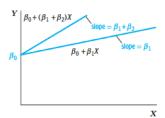
• the third population model is

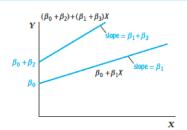
$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 (D_i \times X_i) + u_i$$

#### FIGURE 8.8 Regression Functions Using Binary and Continuous Variables



(a) Different intercepts, same slope





(b) Different intercepts, different slopes

(c) Same intercept, different slopes

#### Interactions Between Two Continuous Variables

- Now suppose that both independent variables  $(X_{1i} \text{ and } X_{2i})$  are continuous.
- ullet  $X_{1i}$  is his or her years of work experience
- ullet  $X_{2i}$  is the number of years he or she went to school.
- there might be an interaction between these two variables so that the effect on wages of an additional year of experience depends on the number of years of education.
- the population regression model

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + u_i$$

#### Interactions Between Two Continuous Variables

 $\bullet$  Thus the effect on Y of a change in  $X_1$  , holding  $X_2$  constant, is

$$\frac{\Delta Y}{\Delta X_1} = \beta_1 + \beta_3 X_2$$

• A similar calculation shows that the effect on Y of a change  $\Delta X_1$  in  $X_2$ , holding  $X_1$  constant, is

$$\frac{\Delta Y}{\Delta X_2} = \beta_2 + \beta_3 X_1$$

 $\bullet$  That is, if  $X_1$  changes by  $\Delta X_1$  and  $X_2$  changes by  $\Delta X_2$  , then the expected change in Y

$$\Delta Y = (\beta_1 + \beta_3 X_2) \Delta X_1 + (\beta_2 + \beta_3 X_1) \Delta X_2 + \beta_3 \Delta X_1 \Delta X_2$$

# Application to the STR and the percentage of English learners

• The estimated interaction regression

$$ln(\widehat{TestScore}) = 686.3 - 1.12STR - 0.67PctEL + 0.0012(STR - 0.012) + 0.0012(STR - 0.$$

• when the percentage of English learners is at the median(PctEL=8.85), the slope of the line relating test scores and the STR is

$$\frac{\Delta Y}{\Delta X_1} = \beta_1 + \beta_3 X_2 = -1.12 + 0.0012 \times 8.85 = -1.11$$

• when the percentage of English learners is at the 75th percentile(PctEL=23.0), the slope of the line relating test scores and the STR is

$$\frac{\Delta Y}{\Delta Y} = \beta_1 + \beta_2 X_2 = -1.12 + 0.0012 \times 23.0 = -1.09$$

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# Application : STR and the percentage of English

| rners   |                     |                     |                 |                     |                     |                     |                    |  |  |  |
|---|---------------------|---------------------|-----------------|---------------------|---------------------|---------------------|--------------------|--|--|--|
| TABLE 8.3 Nonlinear Re  | gression Mod        | lels of Test        | Scores          |                     |                     |                     |                    |  |  |  |
| Dependent variable: average test score in district; 420 observations. |                     |                     |                 |                     |                     |                     |                    |  |  |  |
| Regressor   | (1)                 | (2)                 | (3)             | (4)                 | (5)                 | (6)                 | (7)                |  |  |  |
| Student-teacher ratio (STR)   | -1.00**<br>(0.27)   | -0.73**<br>(0.26)   | -0.97<br>(0.59) | -0.53<br>(0.34)     | 64.33**<br>(24.86)  | 83.70**<br>(28.50)  | 65.29**<br>(25.26) |  |  |  |
| STR <sup>2</sup>  |                     |                     |                 |                     | -3.42**<br>(1.25)   | -4.38**<br>(1.44)   | -3.47**<br>(1.27)  |  |  |  |
| STR <sup>3</sup>  |                     |                     |                 |                     | 0.059**<br>(0.021)  | 0.075**<br>(0.024)  | 0.0608<br>(0.021)  |  |  |  |
| % English learners  | -0.122**<br>(0.033) | -0.176**<br>(0.034) |                 |                     |                     |                     | -0.166<br>(0.034)  |  |  |  |
| % English learners<br>≥ 10%? (Binary, <i>HiEL</i> )                   |                     |                     | 5.64<br>(19.51) | 5.50<br>(9.80)      | -5.47**<br>(1.03)   | 816.1*<br>(327.7)   |                    |  |  |  |
| HiEL × STR  |                     |                     | -1.28<br>(0.97) | -0.58<br>(0.50)     |                     | -123.3*<br>(50.2)   |                    |  |  |  |
| $HiEL \times STR^2$   |                     |                     |                 |                     |                     | 6.12*<br>(2.54)     |                    |  |  |  |
| $HiEL \times STR^3$   |                     |                     |                 |                     |                     | -0.101*<br>(0.043)  |                    |  |  |  |
| % Eligible for subsidized lunch                                       | -0.547**<br>(0.024) | -0.398**<br>(0.033) |                 | -0.411**<br>(0.029) | -0.420**<br>(0.029) | -0.418**<br>(0.029) | -0.4028<br>(0.033) |  |  |  |

# An Good Application: Kung and Ma(2014)

- Main Question: examine whether the cultural norms associated with Confucianism served to attenuate the effect of economic shocks in triggering peasant rebellions in the Qing dynasty.
- Dependent variable: **social conflict**(peasant rebellions)
- Independt variable: economic shocks(crop failure)
- Interaction Variable: **cultural norms**(Confucianism)

# Kung and Ma(2014): OLS Model

• The effect of economic shocks on peasant rebellions

$$Rebellion_{it} = \beta_1 shock_{it-1} + country_i + year_t + u_{it}$$

•  $\hat{\beta_1}$  is the estimate coefficient of  $\beta_1$ , which represents the effect of economic shock in the previous year on peasant rebellions.

#### Kung and Ma(2014): OLS Results

**Table 2** Economic shocks and peasant rebellions.

|                           | Dependent variable is number of peasant rebellions |                        |                       |                       |  |  |
|---------------------------|--|------------------------|-----------------------|-----------------------|--|--|
|                           | (1)  | (2)                    | (3)                   | (4)                   |  |  |
| Crop failure <sub>t</sub> |  |                        | -0.00054<br>(0.00462) |                       |  |  |
| Crop failure $t-1$        | 0.04541***<br>(0.00570)                            | 0.00973**<br>(0.00395) | 0.00979*** (0.00358)  |                       |  |  |
| Crop failure $_{t-2}$     | (0.00010)  | (0.0000)               | 0.00022 (0.00452)     |                       |  |  |
| $Drought_{t-1}$           |  |                        | (0.00132)             | 0.00697*<br>(0.00355) |  |  |
| $Waterlog_{t-1}$          |  |                        |                       | -0.00190 $(0.00178)$  |  |  |
| Year fixed-effects        | No   | Yes                    | Yes                   | (0.00178)<br>Yes      |  |  |
| County fixed-effects      | No   | Yes                    | Yes                   | Yes                   |  |  |
| R-squared                 | 0.01   | 0.18                   | 0.18                  | 0.18                  |  |  |
| Number of observations    | 27,713   | 27,713                 | 27,606                | 27,713                |  |  |

Notes: OLS results. Robust standard errors clustered at the county level are reported in parentheses.

- \* Significant at 10%.
- \*\* Significant at 5%.
- \*\*\* Significant at 1%.

# Kung and Ma(2014): OLS with interaction terms

 The effect of Confucian norms on the effect of economic shock on peasant rebellions

$$\begin{split} Rebellion_{it} = \\ \beta_2 shock_{it-1} + \beta_3 shock_{it-1} \times Confucianism_i \\ + country_i + year_t + u_{it} \end{split}$$

# Kung and Ma(2014): the result of interaction terms

Table 3 The mitigating effect of Confucianism: baseline results.

|  | Dependent variable is number of peasant rebellions |            |            |            |  |  |
|--|--|------------|------------|------------|--|--|
|  | (1)  | (2)        | (3)        | (4)        |  |  |
| Crop failure <sub>t - 1</sub>                                    | 0.03639*   | 0.02313**  | 0.11610**  | 0.07348**  |  |  |
|  | (0.01857)  | (0.01137)  | (0.04568)  | (0.02851)  |  |  |
| Crop failure <sub>t - 1</sub> × ln(temples/area)                 | -0.01092   | -0.00872** | (          | (          |  |  |
|  | (0.00696)  | (0.00409)  |            |            |  |  |
| Crop failure <sub><math>t-1</math></sub> × ln(chaste women/area) | , ,  |            | -0.01621** | -0.01090** |  |  |
|  |  |            | (0.00669)  | (0.00416)  |  |  |
| Year fixed-effects   | Yes  | Yes        | Yes        | Yes        |  |  |
| County fixed-effects   | Yes  | Yes        | Yes        | Yes        |  |  |
| County-specific time trend                                       | No   | Yes        | No         | Yes        |  |  |
| R-squared  | 0.18   | 0.21       | 0.18       | 0.21       |  |  |
| Number of observations   | 27,713   | 27,713     | 27,713     | 27,713     |  |  |

Notes: OLS results. Robust standard errors clustered at the county level are reported in parentheses. \* Significant at 10%.

<sup>\*\*</sup> Significant at 5%.