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1. OLS estimator for ARU) is blacked, but consistent.

Consider
$$X_{t+1} = \phi, \chi_t + \xi_{t+1}, \xi_{t+1} \sim N(0, \sigma^2)$$

square loss $L = \frac{1}{\xi_t} (\hat{\chi}_t - \chi_t)^2$

36, 20 3 P, = Ster XX-1 XX+

Loss
$$L = \sum_{t=1}^{T} (\widehat{\chi}_{t} - \chi_{t})^{2}$$

$$= \sum_{t=1}^{T} (\widehat{\rho}_{t} \chi_{t+1} - \chi_{t})^{2}$$

$$= \sum_{t=1}^{T} (\widehat{\rho}_{t} \chi_{t+1} - \chi_{t})^{2}$$

If under the assumption that (P,) <1, which

means ARU) is stationary, by ergodic thm, OLS estimator is consistent.

 $\sum_{t=1}^{T} \chi_{t-1}^{2}$

2.
$$I_{t}=0.01+0.1I_{t-2}+\xi_{t}$$
, $\xi_{t}\sim N(0,0.02)$
(a) $E(I_{t})=0.01+0.1E(I_{t-2})$

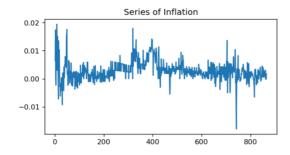
$$\Rightarrow v.99 Var(17) = v.v2$$
 $\Rightarrow Var(17) = \frac{2}{99}$

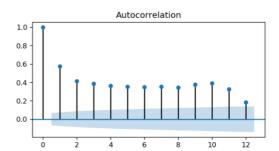
$$Var(N) = \frac{2}{99}$$
(b) $Y = (N(N) / N)$

$$\begin{array}{ll}
\text{Sol} & \text{So$$

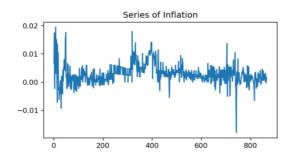
rz = Cov (rt, rt-2)

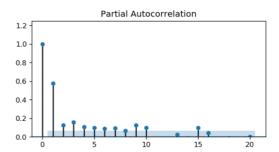
3. (a)





(b) Plot the PACF for the inflation series:





From the graph, we see that lag-1 is highly significant, so choose an AR(1) model.

AutoReg Model Results

| Dep. Variable: Model: Method: Date: Time: Sample: | _ | AutoRegonditional n, 26 Apr 2 20:43 | y(1) Log MLE S.D. 2021 AIC | Observations: Likelihood of innovations | | 865 3852.184 0.003 -11.748 -11.731 -11.742 |
|---|------------------|---|----------------------------------|---|----------------|---|
| ========= | coef | std err | Z | P> z | [0.025 | 0.975] |
| intercept pai.L1 | 0.0012 0.5755 | 0.000 0.028 | 9.723 20.704 Roots | 0.000 0.000 | 0.001 0.521 | 0.001 0.630 |
| ========== | Real | In | aginary | Modulus | | Frequency |
| AR.1 | 1.7376 | | ⊦0.0000j | 1.7376 | | 0.0000 |

Using AIC:

AutoReg Model Results

| Dep. Variable | | | | Observations: | | ======== 86 |
|----------------------|---------|---------------|--------------|--------------------|------------|----------------|
| Model: | | AutoReg(1 | .5) Log | Likelihood | | 3919.85 |
| Method: | (| Conditional M | | of innovation | ıs | 0.00 |
| Date: | Mo | n, 26 Apr 20 | 21 AIC | | | -12.02 |
| Time: | | 20:43: | | | | -11.92 |
| Sample: | | | 15 HQI | C | | -11.98 |
| | | | 865 | | | |
| | coef | std err | Z | P> z | [0.025 | 0 . 975 |
| intercept | 0.0004 | 0.000 | 3.191 | 0.001 | 0.000 | 0.00 |
| pai.L1 | 0.4414 | 0.034 | 13.123 | 0.000 | 0.376 | 0.50 |
| pai.L2 | 0.0763 | 0.036 | 2.098 | 0.036 | 0.005 | 0.14 |
| pai.L3 | 0.0621 | 0.036 | 1.710 | 0.087 | -0.009 | 0.13 |
| pai.L4 | 0.0153 | 0.036 | 0.427 | 0.669 | -0.055 | 0.08 |
| pai.L5 | 0.0547 | 0.035 | 1.542 | 0.123 | -0.015 | 0.12 |
| pai.L6 | 0.0118 | 0.035 | 0.335 | 0.738 | -0.057 | 0.08 |
| pai.L7 | 0.0575 | 0.035 | 1.641 | 0.101 | -0.011 | 0.12 |
| pai.L8 | 0.0279 | 0.034 | 0.812 | 0.417 | -0.039 | 0.09 |
| pai.L9 | 0.0601 | 0.034 | 1.752 | 0.080 | -0.007 | 0.12 |
| pai.L10 | 0.0766 | 0.034 | 2.229 | 0.026 | 0.009 | 0.14 |
| pai.L11 | 0.0702 | 0.034 | 2.039 | 0.041 | 0.003 | 0.13 |
| pai.L12 | -0.1683 | 0.034 | -4.897 | 0.000 | -0.236 | -0.10 |
| pai.L13 | -0.0068 | 0.035 | -0.197 | 0.844 | -0.075 | 0.06 |
| pai.L14 | -0.0350 | 0.034 | -1.025 | 0.305 | -0.102 | 0.03 |
| pai.L15 | 0.0948 | 0.031 | 3.010 | 0.003 | 0.033 | 0.15 |
| Jsing BIC: ====== | ======= | AutoReg | Model Re | esults ======== | ======= | ====== |
| Dep. Variable: | | | | Observations: | | 86 |
| Model: | _ | AutoReg(1 | | Likelihood | | 3907.25 |
| Method: | | onditional M | | of innovation | S | 0.00 |
| Date: | Мо | n, 26 Apr 20 | | | | -11.96 |
| Time: | | 20:55: | | | | -11.88 |
| Sample: | | | 12 HQIC | | | -11.93 |
| ========= | | | 65 ====== | :======== | | |
| | coef | std err | z | P> z | [0.025 | 0.975 |
| intercept | 0.0005 | 0.000 | 3.771 | 0.000 | 0.000 | 0.00 |
| pai.L1 | 0.4287 | 0.034 | 12.742 | 0.000 | 0.363 | 0.49 |
| pai.L2 | 0.0509 | 0.036 | 1.396 | 0.163 | -0.021 | 0.12 |
| pai.L3 | 0.0472 | 0.036 | 1.301 | 0.193 | -0.024 | 0.11 |
| pai.L4 | 0.0596 | 0.036 | 1.658 | 0.097 | -0.011 | 0.13 |
| pai.L5 | 0.0507 | 0.035 | 1.438 | 0.150 | -0.018 | 0.12 |
| pai.L6 | -0.0126 | 0.035 | -0.358 | 0.720 | -0.082 | 0.05 |
| pai.L7 | 0.1013 | 0.035 | 2.878 | 0.004 | 0.032 | 0.17 |
| pai.L8 | 0.0230 | 0.035 | 0.651 | 0.515 | -0.046 | 0.09 |
| pai . L9 | 0.0749 | 0.035 | 2.132 | 0.033 | 0.006 | 0.14 |
| pai.L10 | 0.1126 | 0.035 | 3.227 | 0.001 | 0.044 | 0.18 |
| pai.L11 | 0.0484 | 0.035 | 1.403 | 0.161 | -0.019 | 0.11 |
| | | | | | | |

Python ar_select_order function selects AR(15) under AIC, and selects AR(12) under BIC.

-5.519

0.000

-0.238

-0.113

0.032

-0.1759

pai.L12

4.

- (a) Economically, the first difference is essentially the difference of current season earnings compared with the past season, while the seasonal difference is compared with the same season last year.
- (b) The model is built in codes. The estimated model is:

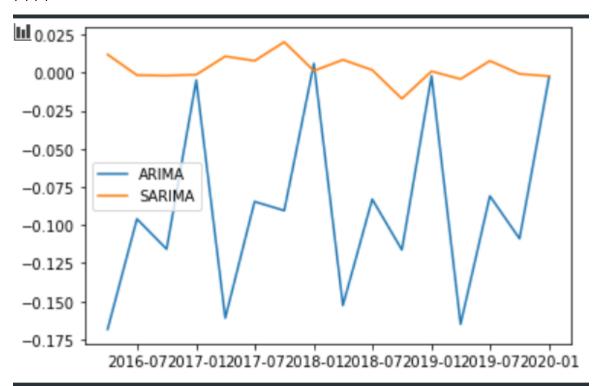
$$x_{t+1} - x_t = 0.0506 + \epsilon_{t+1} - 0.4920\epsilon_t$$

(c) The model is built in codes. The estimated model is:

$$(x_t - x_{t-1}) - (x_{t-4} - x_{t-5}) = (\epsilon_t - 0.1593\epsilon_{t-1}) - 0.4830(\epsilon_{t-4} - 0.1593\epsilon_{t-5})$$

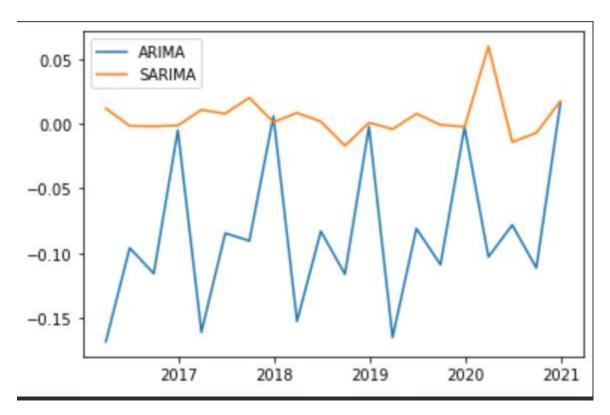
In the model $\theta_1=0.1593$, $\theta_2=0.4830$, which means the past earning increase shocks (both the past season and the same season for the last year) is negatively correlated with predicted next season earnings.

(d) (e)



ARIMA MSE: 0.01134832807889095 SARIMA MSE: 7.346810242072138e-05

The seasonal arima model (airline model) performs better.



Both models perform worse in 2020. This is because the COVID pandemic is an exogeneous event outside of the description of the model. To improve the forecast accuracy, we might include macroeconomic predictors such as expected inflation, GDP growth, treasury rate, etc.