

7.5.2 LOGISTIC REGRESSION FOR $K > 2$ CLASSES

➡ Following the case where $K = 2$, in the logistic case, it may seem that we could assume that, for $k = 1, \dots, K$,

$$m_k(\mathbf{x}) = E(Y_k | \mathbf{X} = \mathbf{x}) = P(G = k | \mathbf{X} = \mathbf{x}) = \frac{\exp(\beta_{k0} + \beta_k^T \mathbf{x})}{1 + \exp(\beta_{k0} + \beta_k^T \mathbf{x})}.$$

➡ However as seen earlier, we must have $m_1(\mathbf{x}) + \dots + m_K(\mathbf{x}) = 1$.

➡ This can be satisfied if we use the following slightly different model: for $k = 1, \dots, K - 1$:

$$m_k(\mathbf{x}) = \frac{\exp(\beta_{k0} + \beta_k^T \mathbf{x})}{1 + \sum_{\ell=1}^{K-1} \exp(\beta_{\ell 0} + \beta_{\ell}^T \mathbf{x})}.$$

➡ Then we take

$$\begin{aligned} m_K(\mathbf{x}) &= 1 - \sum_{k=1}^{K-1} m_k(\mathbf{x}) = 1 - \frac{\sum_{k=1}^{K-1} \exp(\beta_{k0} + \beta_k^T \mathbf{x})}{1 + \sum_{\ell=1}^{K-1} \exp(\beta_{\ell 0} + \beta_{\ell}^T \mathbf{x})} \\ &= \frac{1}{1 + \sum_{\ell=1}^{K-1} \exp(\beta_{\ell 0} + \beta_{\ell}^T \mathbf{x})}. \end{aligned}$$

☞ For $k = 1, \dots, K - 1$, estimate the β_{k0} 's and β_k 's by their maximum likelihood estimators $\hat{\beta}_{k0}$, $\hat{\beta}_k$ (again can be applied to PLS components instead) and deduce estimators \hat{m}_k of m_k . Then estimate $m_K(\mathbf{x})$ by

$$\hat{m}_K(\mathbf{x}) = 1 - \sum_{k=1}^{K-1} \hat{m}_k(\mathbf{x}).$$

☞ Classify the new individual in group \hat{G} which gives the max value among $\hat{m}_1(\mathbf{x}), \dots, \hat{m}_K(\mathbf{x})$:

$$\hat{G} = \arg \max_{k=1, \dots, K} \hat{m}_k(\mathbf{x})$$

☞ Note that this is again a linear method. To see why, note that to find the max it suffices to compute, for $k = 1, \dots, K - 1$

$$\log \frac{\hat{P}(G = k | \mathbf{X} = \mathbf{x})}{\hat{P}(G = K | \mathbf{X} = \mathbf{x})} = \log \frac{\hat{m}_k(\mathbf{x})}{\hat{m}_K(\mathbf{x})} = \hat{\beta}_{k0} + \hat{\beta}_1^T \mathbf{x},$$

and compare the groups 2 by 2 based on those log ratios.

👉 Ex: suppose $K > 5$ and $\hat{m}_5(\mathbf{x})$ is the max. Then

$$\hat{m}_5(\mathbf{x}) > \hat{m}_K(\mathbf{x}) \iff \log\{\hat{m}_5(\mathbf{x})/\hat{m}_K(\mathbf{x})\} > 0$$

and for all $k \neq 5$ and $\neq K$,

$$\hat{m}_k(\mathbf{x}) < \hat{m}_5(\mathbf{x}) \iff \frac{\hat{m}_k(\mathbf{x})}{\hat{m}_K(\mathbf{x})} < \frac{\hat{m}_5(\mathbf{x})}{\hat{m}_K(\mathbf{x})} \iff \log \frac{\hat{m}_k(\mathbf{x})}{\hat{m}_K(\mathbf{x})} < \log \frac{\hat{m}_5(\mathbf{x})}{\hat{m}_K(\mathbf{x})}$$

so we are able to make our decision based only on these log ratios which are linear in \mathbf{x} .

LD, QD based on normality assumption.

7.5.3 LD AND QD METHODS FOR $K > 2$ GROUPS (K>2时LD与QD的实施)

➡ Using the same ideas, LD and QD methods can be generalised to the case where data come from $K > 2$ groups.

➡ Let $\pi_1 = P(G = 1), \dots, \pi_K = P(G = K)$.

➡ LD: Assume that for $k = 1, \dots, K$, $\mathbf{X}|G = k \sim N_p(\mu_k, \Sigma)$ and classify new obs \mathbf{x} in group k that maximises \hat{m}_k , or equivalently,

$$\delta_k(\mathbf{x}) = \log \hat{\pi}_k + \mathbf{x}^T \hat{\Sigma}^{-1} \hat{\mu}_k - \frac{1}{2} \hat{\mu}_k^T \hat{\Sigma}^{-1} \hat{\mu}_k.$$

➡ QD: Assume that for $k = 1, \dots, K$, $\mathbf{X}|G = k \sim N_p(\mu_k, \Sigma_k)$ and classify new obs \mathbf{x} in group k that maximises

$$\delta_k(\mathbf{x}) = \log(\hat{\pi}_k) - \frac{1}{2} \log\{\det(\hat{\Sigma}_k)\} - \frac{1}{2}(\mathbf{x} - \hat{\mu}_k)^T \hat{\Sigma}_k^{-1}(\mathbf{x} - \hat{\mu}_k).$$

👉 As in case $K = 2$ groups, for simplicity, relabel the data \mathbf{X}_i making the distinction of which group they come from: $\mathbf{X}_{i,k}$, $i = 1, \dots, n_k$ denote all the \mathbf{X}_i 's that are from group k ;

👉 $\hat{\pi}_k = n_k/n$ is the proportion of training data that are from group k or $\hat{\pi}_k = 1/K$ depending on our beliefs,

$$\hat{\mu}_k = \bar{\mathbf{X}}_k = \frac{1}{n_k} \sum_{i=1}^{n_k} \mathbf{X}_{i,k}$$
$$\hat{\Sigma}_k = \frac{1}{n_k - 1} \sum_{i=1}^{n_k} (\mathbf{X}_{i,k} - \hat{\mu}_k)(\mathbf{X}_{i,k} - \hat{\mu}_k)^T$$

and

$$\hat{\Sigma} = \frac{1}{n_1 + \dots + n_K - K} \sum_{k=1}^K \sum_{i=1}^{n_k} (\mathbf{X}_{i,k} - \hat{\mu}_k)(\mathbf{X}_{i,k} - \hat{\mu}_k)^T.$$

then \mathcal{P}

logistic, LQR, linear regression 都是参数模型, 要估计参数, 还有很多假设 (正态, 独立)

8 CLASSIFICATION AND REGRESSION TREES AND RELATED METHODS

8.1 CLASSIFICATION AND REGRESSION TREES (CART)

Hastie et al. (2017), section 9.2 (second edition, 12th printing).

↓
不现实 我们只能
是一种 approximation

8.1.1 INTRODUCTION

Methods from previous chapter rely on strong parametric assumptions (linear model, logistic model or normality assumptions).

When these assumptions are too far from the truth, the performance of classifiers can be poor; e.g. recall the example at page 253.

We need more flexible models that are less driven by strong parametric assumptions.

Instead of using linear or logistic regression assumptions, one possibility is to use regression trees.

based on non-parametric assumption

8.1.2 REGRESSION TREES

Suppose we observe an i.i.d. sample $(\mathbf{X}_1, Y_1), \dots, (\mathbf{X}_n, Y_n)$ coming from the regression model

$$E(Y_i | \mathbf{X}_i = \mathbf{x}) = m(\mathbf{x}),$$

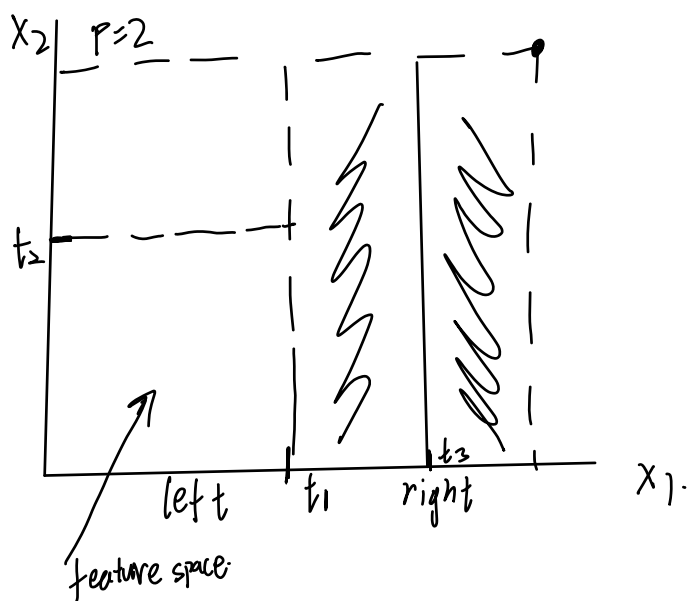
where $Y_i \in \mathbb{R}$ and $\mathbf{X}_i = (X_{i1}, \dots, X_{ip})^T \in \mathbb{R}^p$ is continuous.

When we don't want to highlight the dependence on the sample, we use the notation (\mathbf{X}, Y) , where

$$E(Y | \mathbf{X} = \mathbf{x}) = m(\mathbf{x}),$$

$Y \in \mathbb{R}$ and $\mathbf{X} = (X_1, \dots, X_p)^T \in \mathbb{R}^p$ is continuous.

Note the difference between the notation \mathbf{X}_i for the i th training vector, and X_i for the i th component of the vector \mathbf{X} . We will use both notations a lot in this chapter.



1st split: divide $(x_1 < t_1 \quad x_1 > t_3)$ into 2 regions

2nd split $(x_2 < t_2 \quad x_2 > t_2)$

$t_1 < x_1 < t_3$ $x_1 > t_3$... 不断地划分区域来找 x 的范围
构造超平面.

gradually \rightarrow divide big rectangles into ^{some} smaller ~~rectangles~~ rectangles each time. and each time work on one rectangle. (either vertically or horizontally).

$(x_1 \text{ less/greater to } t_2)$
 x_2

☞ The main idea of regression trees is to partition the “feature space”, i.e. the domain of \mathbf{X} , i.e. the subset of \mathbb{R}^p of all possible values of \mathbf{X} .

☞ On each partition we approximate $m(\mathbf{x})$ by a constant.

☞ To understand regression trees, consider the case where $p = 2$. Here we need to estimate the regression curve

$$m(\mathbf{x}) = E(Y|\mathbf{X} = \mathbf{x})$$

where $Y \in \mathbb{R}$ and $\mathbf{X} = (X_1, X_2)^T \in \mathbb{R}^2$ is continuous.

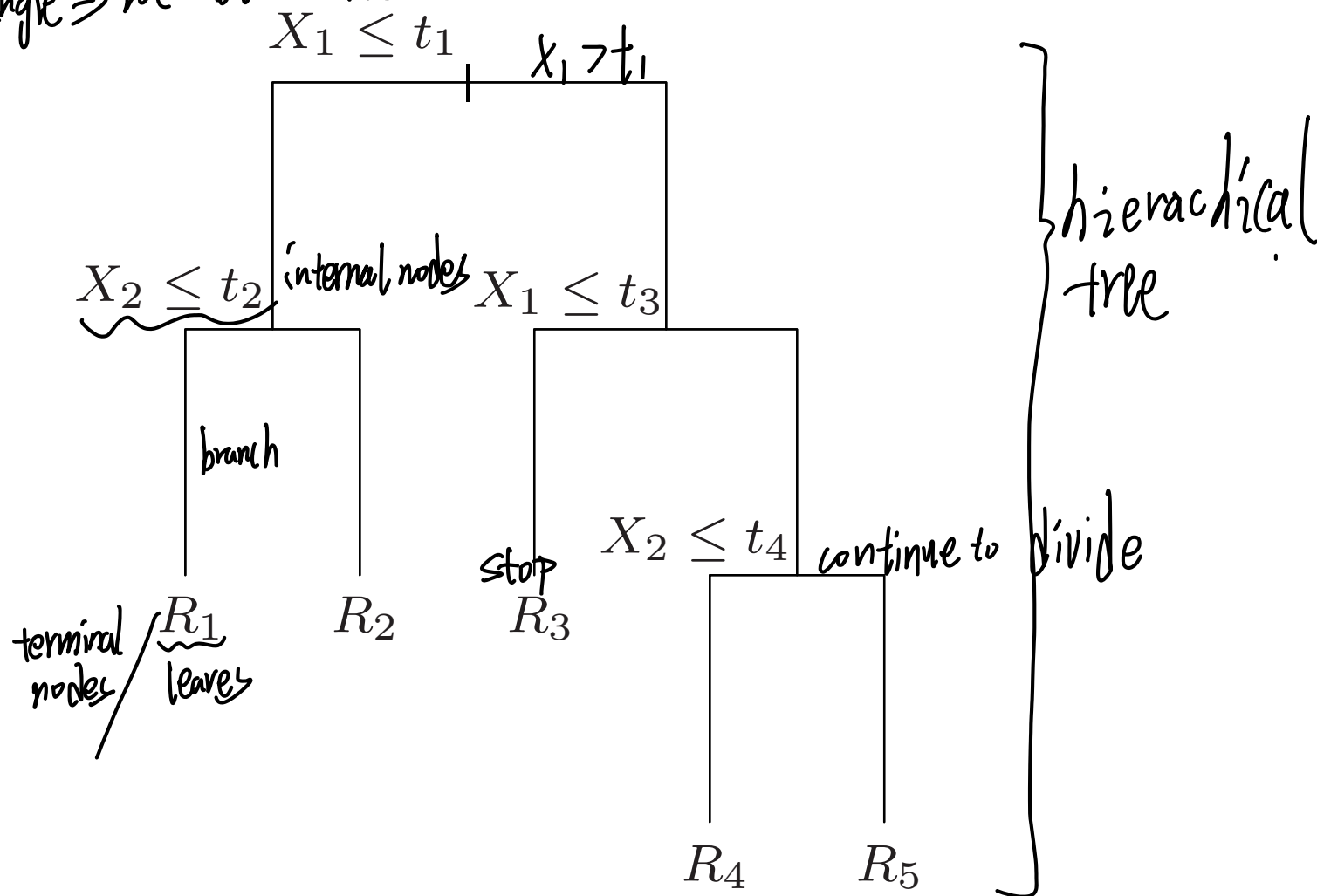
☞ Then we construct a sequence of partitions of the type:

- $\{X_1 \leq t_1\}, \{X_1 > t_1\}$
- $\{X_2 \leq t_2\}, \{X_2 > t_2\}$
- $\{X_1 \leq t_3\}, \{X_1 > t_3\}$

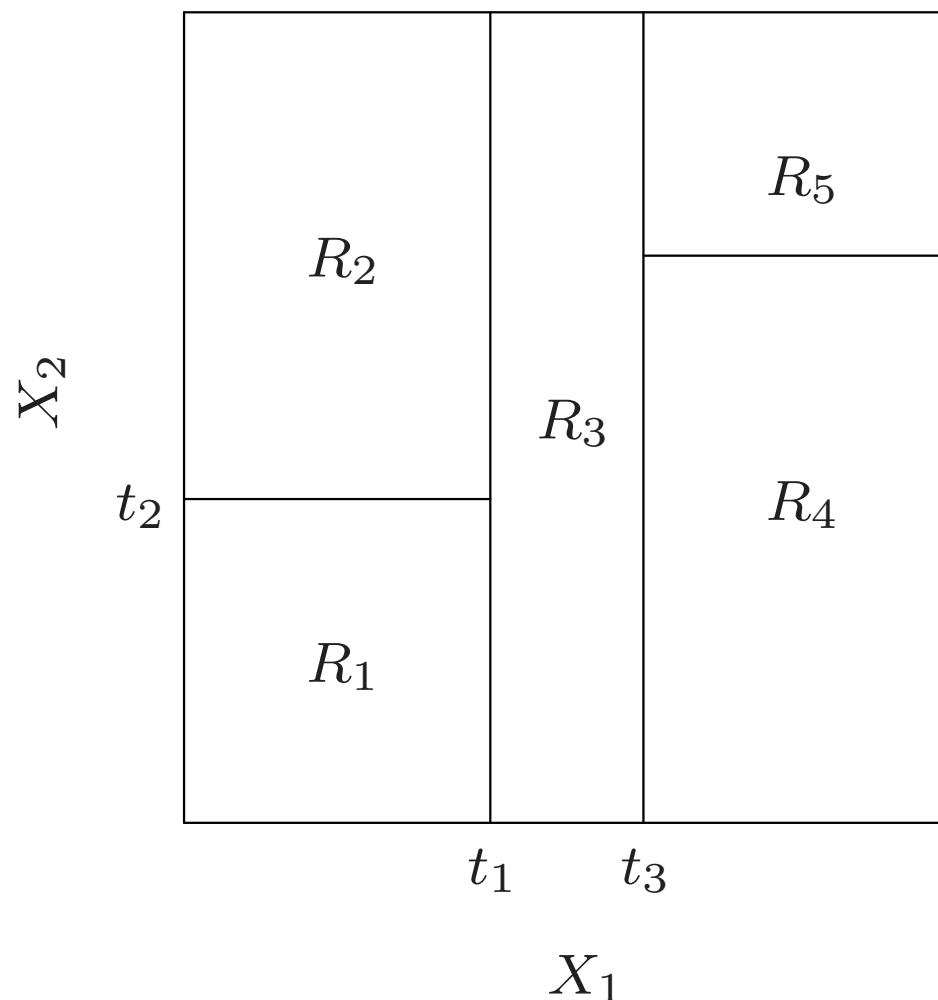
- etc. See handwritten construction on video.

page 306 of Hastie et al. (2017): Constructing a tree by a sequence of binary partitions of the type $\{X_j \leq t\}$, $\{X_j > t\}$. After a number of partitions, stop the splitting process and obtain regions R_1, \dots, R_L .

by dividing rectangle \Rightarrow we obtain a tree.



👉 page 306 of Hastie et al. (2017): At the end of the sequence of binary partitions, we have partitioned the feature space in rectangles, say R_1, R_2, \dots, R_L . Example when $p = 2$ and $L = 5$:



divided
 x_j greater/less than

☛ The regions R_1, \dots, R_L obtained at the end of the process are called terminal nodes or leaves of the tree.
 终点 叶子

☛ The splits such as $\{X_1 \leq t_1\}$, inside the tree, are called internal nodes.

☛ The segments of the tree that connect the nodes are called branches of the tree.

☛ Once we have partitioned the space into regions R_1, \dots, R_L , on each region we approximate the regression curve m by a constant:
 sum of all of my regions (这个)

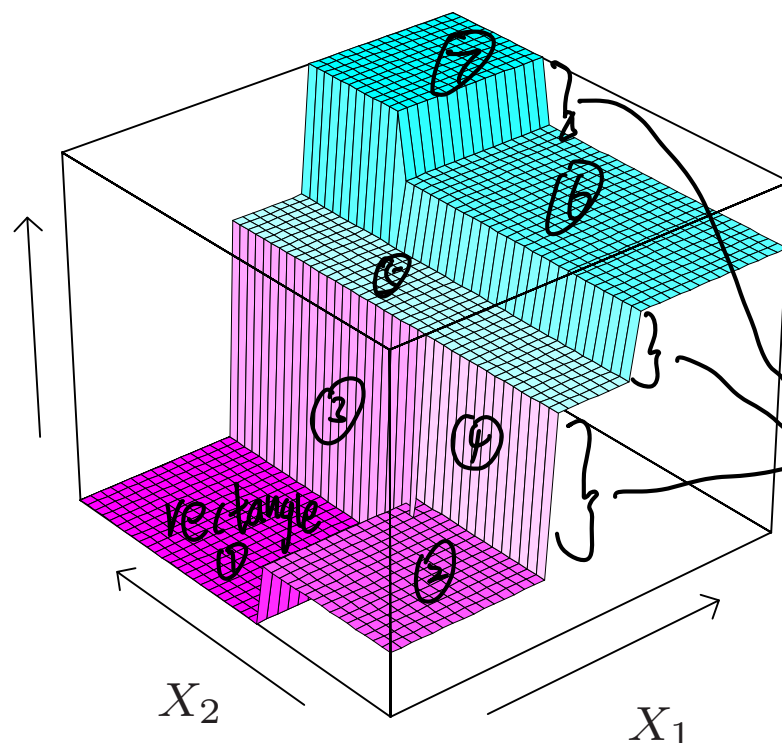
For all \mathbf{x} in feature space, $m(\mathbf{x}) \approx \sum_{\ell=1}^L c_{\ell} \underbrace{I\{\mathbf{x} \in R_{\ell}\}}_{\text{Indicator Function.}}$
 where

$$\underbrace{c_{\ell} I\{\mathbf{x} \in R_{\ell}\}} = \begin{cases} c_{\ell} & \text{if } \mathbf{x} \in R_{\ell} \\ 0 & \text{otherwise.} \end{cases}$$

constant used in that particular region indicator of \mathbf{x} in the particular region.

page 306 of Hastie et al., 2017: piecewise constant approximations on regions R_1, R_2, \dots, R_5 of \mathbb{R}^2 constructed earlier:

by each of region we approximate the regression curve by a constant
 不同的 rectangle 不同的 constant



回到平面的一意思

height is different on each rectangle

☞ Why is this flexible? As long as we partition a feature space in small enough pieces, we can always approximate well a regression curve by constants on each piece.

just use constant in each region instead of using sophisticated functions.

☞ The finer the partition of the feature space, the better is the approximation of m by constants c_1, c_2, c_3, \dots on the regions R_1, R_2, R_3, \dots

☞ To understand this, here is an example in the case where $p = 1$.

分得越细, regression 就越精准.

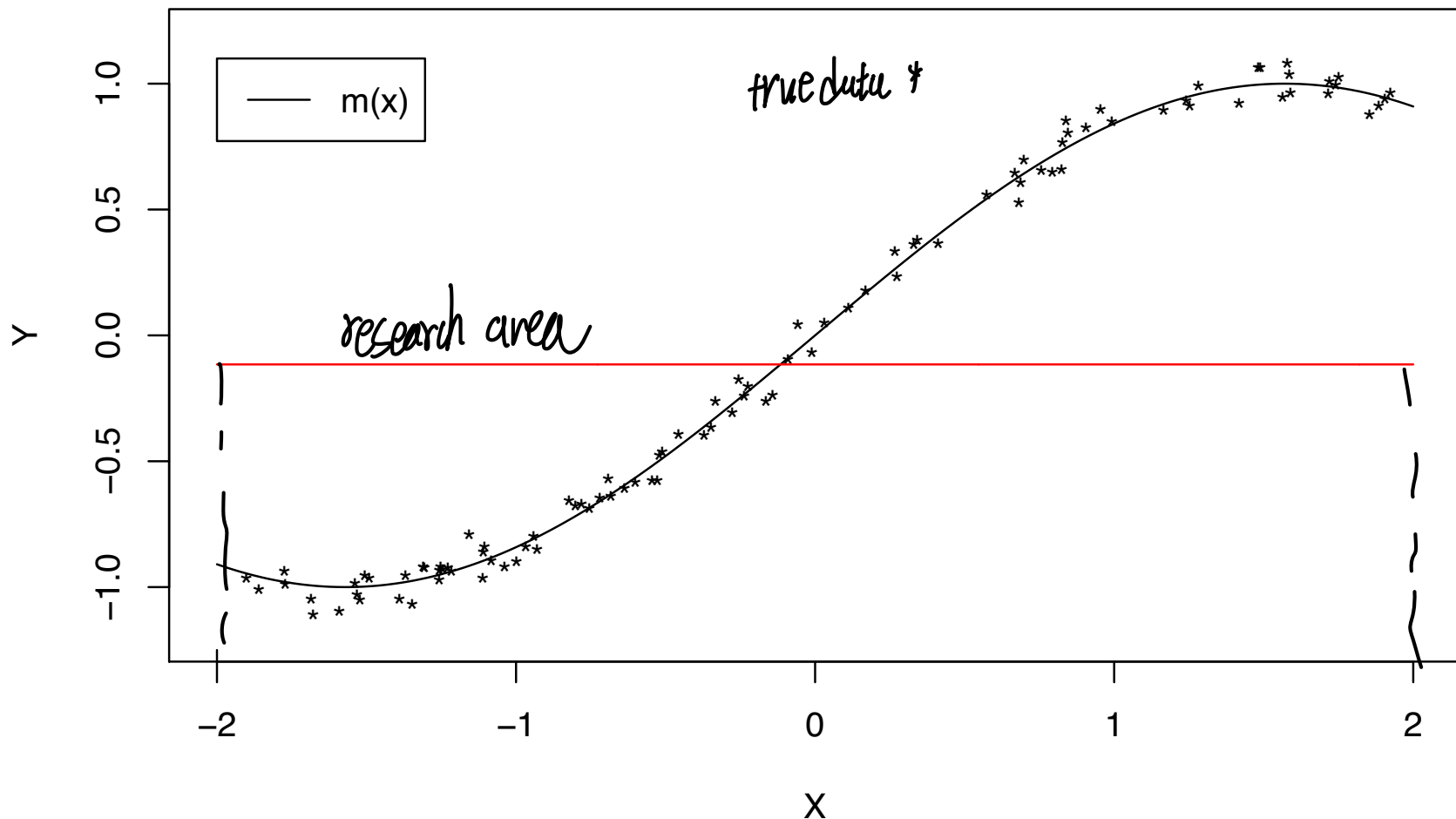
只有一个变量

Example with $p = 1$. Here, partitioning the feature space means splitting the range of values of the variable X into intervals.

Data points:*. Approximate m by a constant (red line) without partitioning feature space does not give a good approximation of m :

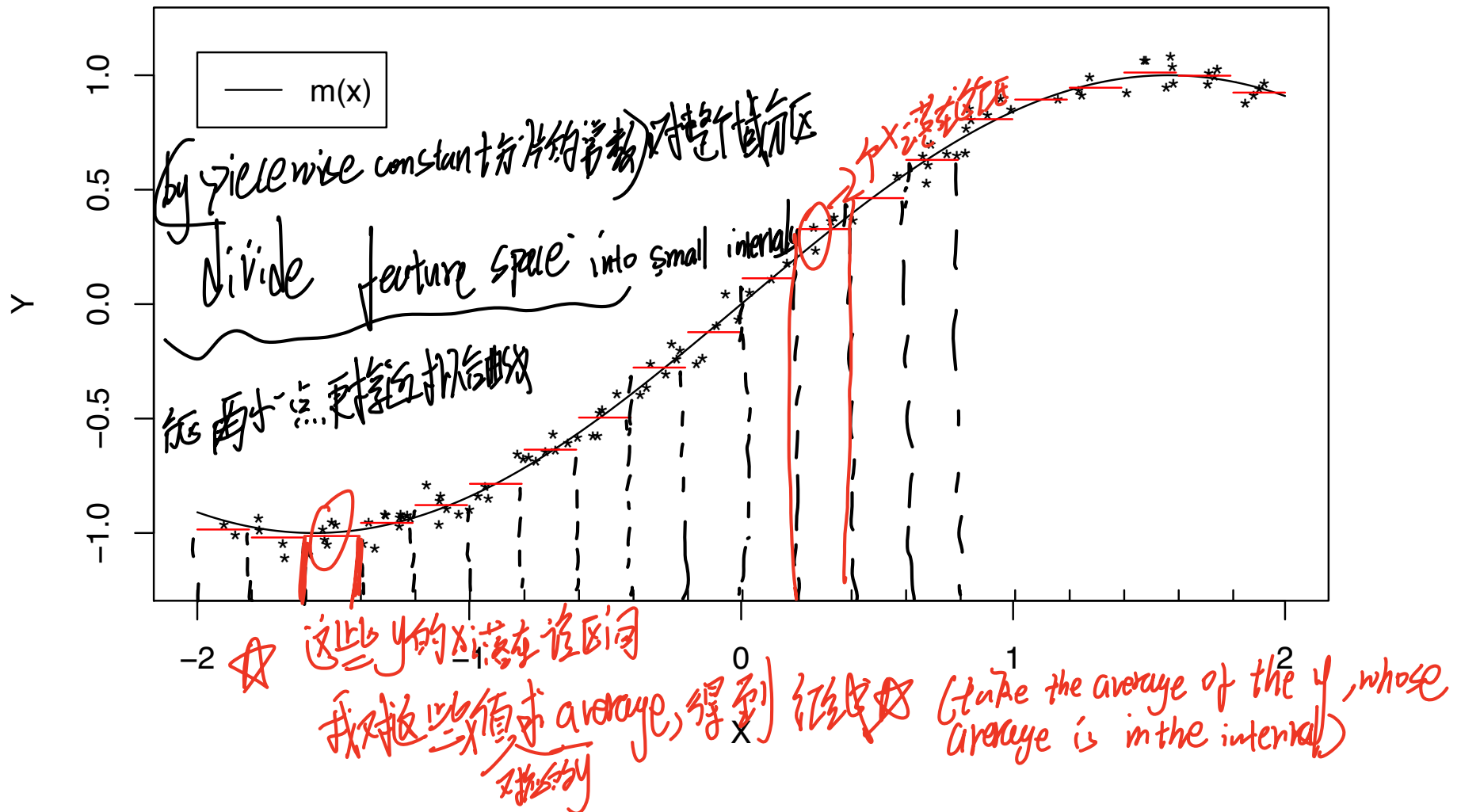
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no partition



Partition the real line into small intervals and approximate m by a well chosen constant on each interval gives a much better approximation of m

reasonable partition



在数据中 in each piece is (best) possible approximation to my curve
 FITTING THE TREE (CHOOSING THE CONSTANTS)

In practice don't know m and can't choose c_j that approximates m "the best" on each R_j .

Instead, for each j , using only the data $(\mathbf{X}_1, Y_1), \dots, (\mathbf{X}_n, Y_n)$, approximate "the best" c_j by an estimator \hat{c}_j . How?

Suppose we have partitioned the feature space of \mathbf{X} into R_1, \dots, R_L .
 For $\ell = 1, \dots, L$, for $\mathbf{x} \in R_\ell$, we approximate $m(\mathbf{x})$ by

把特征空间分成 L 个
 的函数或 estimator of regression curve (常数) \Rightarrow 在 R_j 内的 \mathbf{x}_i 对应的 y_i 的值

$\hat{m}(\mathbf{x}) = \hat{c}_\ell = \text{average of } Y_i \text{'s whose } \mathbf{X}_i \in R_\ell$

$$= \sum_{i=1}^n Y_i \cdot I(\mathbf{X}_i \in R_\ell) / \sum_{i=1}^n I(\mathbf{X}_i \in R_\ell)$$

Sum of indicators
 Indicator of whether \mathbf{x}_i are in the region or not

选择 \hat{c}_j Same as choosing \hat{c}_j 's that minimise residual sum of squares (RSS)

$$\sum_{i=1}^n \left\{ Y_i - \sum_{\ell=1}^L c_\ell I(\mathbf{X}_i \in R_\ell) \right\}^2 = \sum_{\ell=1}^L \sum_{\mathbf{X}_i \in R_\ell} (Y_i - c_\ell)^2$$

每个区域

wrt c_1, \dots, c_L

BUILDING THE TREE (CHOOSING CONSECUTIVE SPLITS)

☞ Recall that for $\mathbf{X} = (X_1, \dots, X_p)^T$, a split is of the type $\{X_j \leq t\}, \{X_j > t\}$ for some $1 \leq j \leq p$ and some value t .

☞ Ideally, choose splits so that final tree minimises RSS but **not computationally possible**.

☞ Why? Would involve constructing all possible sequences of splits $\{X_j \leq t\}, \{X_j > t\}$ for all $j = 1, \dots, p$ and all values of t .

☞ Note: in fact, for split on X_j , only need to consider $t \in \{X_{1j}, \dots, X_{nj}\}$.

Why? To find \hat{c}_ℓ on R_ℓ , only need to know $I(\mathbf{X}_i \in R_\ell)$. Changing t for split $\{X_j \leq t\}, \{X_j > t\}$ used in R_ℓ affects $I(\mathbf{X}_i \in R_\ell)$ through $I(X_{ij} \leq t)$, whose value only changes when t changes from $t < X_{ij}$ to $t \geq X_{ij}$ (or vice versa).

☞ Even with this simplification, too time consuming to consider all partitions.

about tree: at the end, the entire feature space was divided into a number of regions.

实际情况, 我们只有树没有图 \Rightarrow 有很多变量时 \rightarrow 一棵树

做 regression tree 时: we approximate a regression curve by piecewise constant.
(一个区域一个常数)

在每个区域 \Rightarrow we approximate regression curve by a constant (one constant per region).
In every small region

对于每个区域的 constant: (the fit that we use in each region)
= average of y_i whose x_i belongs to that region ($x_i \in R_j$)
[we need to check which data falls in a particular region. \Rightarrow we take the average of these y_i of these particular individuals whose x_i was in the region]
选择数据点并求平均 (ERS)
$$\sum_{i: x_i \in R_j} (y_i - \hat{c}_j)^2$$

How do we choose the splitting point. ^{two regions}
 $X = \{x_1, \dots, x_p\} \Rightarrow \{x_j \leq t\} \{x_j > t\}$ (how do we choose j & t)

split is based on residue sum of square

choose the tree that can minimize the $RSS \rightarrow$ 但树太多

所以无法计算大树 (计算资源)

we use the Indicator of variable to fit the tree (x 在哪个区域 就是 1 不在是 0)