

Problem Set 1

Submission:

Thursday, 02/10/2022, until 1 PM, to be uploaded on the NYU Brightspace course homepage.

1. Expectation and cumulative distribution function

[4 Points]

A real random variable X is *Rayleigh-distributed* with parameter $\sigma > 0$ if its law is characterized by the probability density function

$$f_X(x) = \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right) \mathbb{1}_{(0,\infty)}(x).$$

We write $X \sim Ra(\sigma)$.

- (a) Calculate the cumulative distribution function of a $Ra(\sigma)$ -distributed random variable X .
- (b) Let $U \sim \mathcal{U}([0, 1])$. Using the result from (a), show that for $\sigma > 0$ one has $Y = \sigma \sqrt{-2 \log(U)} \sim Ra(\sigma)$.
Hint: Calculate $F_Y(x) = \mathbf{P}[Y \leq x]$ for $x \in \mathbb{R}$.
- (c) Calculate $\mathbf{E}[X]$ and $\text{Var}[X]$ for $X \sim Ra(\sigma)$.

2. Sums and quotients of random variables

[4 Points]

In this problem we recall the important notion of **convolution** for two probability distributions.

- If X and Y are independent discrete real random variables with values in $\Omega_X, \Omega_Y \subseteq \mathbb{R}$ respectively, then $Z = X + Y$ has probability mass function

$$p_Z(k) = \sum_{\ell \in \Omega_Y} p_X(k - \ell) p_Y(\ell),$$

for $k \in \Omega_X + \Omega_Y$, where p_X and p_Y are the probability mass functions of X and Y , and we set $p_X(r) = 0$ if $r \notin \Omega_X$.

- If X and Y are independent continuous real random variables then $Z = X + Y$ has probability density function

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x - y) f_Y(y) dy.$$

where f_X and f_Y are the probability density functions of X and Y respectively.

- (a) Let $\lambda, \mu > 0$. Show that if $X \sim \text{Pois}(\lambda)$ and $Y \sim \text{Pois}(\mu)$ are independent, then $X + Y \sim \text{Pois}(\lambda + \mu)$.
- (b) Let X_1, \dots, X_n be i.i.d. real random variables with $X_1 \sim \Gamma(\alpha, \beta)$ and $\alpha, \beta > 0$. Show that $\sum_{i=1}^n X_i \sim \Gamma(n\alpha, \beta)$.
Hint: You can use that

$$\int_0^1 u^{a-1} (1-u)^{b-1} du = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}, \quad a, b > 0.$$

- (c) For the quotient of two independent, continuous, positive real random variables X and Y one can also show that $Z = \frac{X}{Y}$ has density

$$f_Z(z) = \int_0^\infty y f_X(zy) f_Y(y) dy, \quad z > 0.$$

Using this, determine the law of the quotient of two independent, $\mathcal{U}([0, 1])$ -distributed random variables.

3. Variances and covariances

[4 Points]

- (a) Let X and Y be jointly continuous with joint density

$$f_{X,Y}(x, y) = (x + y) \mathbb{1}_{\{(x,y) \in [0,1]^2\}}.$$

Calculate the covariance matrix Σ and the correlation $\rho(X, Y)$. Then use Σ to calculate $\text{Var}[X - 2Y]$.

- (b) Let X_1, \dots, X_n be i.i.d. real random variables with $\mu = \mathbf{E}[X_1]$ and $\sigma^2 = \text{Var}[X_1]$. We define

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \text{ (sample mean)}, \quad S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \text{ (sample variance)}.$$

Show that $\mathbf{E}[\bar{X}_n] = \mu$, $\text{Var}[\bar{X}_n] = \frac{\sigma^2}{n}$ and $\mathbf{E}[S_n^2] = \sigma^2$.

4. Moment generating functions

[4 Points]

In this problem we recall the **moment generating functions** for random variables. For a real random variable X , the moment generating function is defined by

$$\psi_X(t) = \mathbf{E}[e^{tX}], \quad t \in \mathbb{R},$$

whenever this expression exists. One has:

- Whenever $\psi_X(t) = \psi_Y(t)$ for $t \in (-\varepsilon, \varepsilon)$, $\varepsilon > 0$, for two random variables X and Y , then $X \stackrel{d}{=} Y$.
- $\psi'_X(0) = \mathbf{E}[X]$.

- (a) Calculate the moment generating function of $X \sim \mathcal{N}(0, 1)$.
- (b) Suppose that X and Y are independent. Explain why $\psi_{X+Y}(t) = \psi_X(t)\psi_Y(t)$ (assuming all these expressions exist for a given t). Use this to show that if $X \sim \mathcal{N}(0, \sigma_1^2)$ and $Y \sim \mathcal{N}(0, \sigma_2^2)$, then $X + Y \sim \mathcal{N}(0, \sigma_1^2 + \sigma_2^2)$, where $\sigma_1, \sigma_2 > 0$.
- (c) The moment generating function of $X \sim \text{Geo}(p)$ with $p \in (0, 1)$ is given by

$$\psi_X(t) = \frac{pe^t}{1 - (1-p)e^t}, \quad t < -\log(1-p).$$

Use this to verify that $\mathbf{E}[X] = \frac{1}{p}$.