# Applied Statistics in R

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Master's Program

Game Theory and Operations Research

Saint Petersburg State University

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Lecture 5 1 / 32

# Agenda

Binary logistic regression

Lecture 5 2 / 32

Lecture 5 3 / 32

 $y \in \{0,1\}$ : dependent variable,

 $x_1, \ldots, x_k$ : predictor (explanatory) variables.

$$x = (1, x_1, \dots, x_k)^{\mathsf{T}}$$

Sample  $(1, x_{i1}, \ldots, x_{ik}, y_i)$ ,  $i = 1, \ldots, n$ . The model

$$\mathsf{P}\{y_i = 1\} = F(\beta^T x_i).$$

Let  $y_i^* = \beta^T x_i + \varepsilon_i$  where  $\varepsilon_1, \dots, \varepsilon_n$ : i.i.d. with  $E\varepsilon_i = 0$  and  $var \varepsilon_i = \sigma^2$ .

re 5 4 / 32

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$$y_i = \begin{cases} 1, & y_i^* \geqslant 0, \\ 0, & y_i^* < 0. \end{cases}$$

re 5 4 / 32

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 $x_1, \ldots, x_k$ : predictor (explanatory) variables.

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Sample  $(1, x_{i1}, ..., x_{ik}, y_i), i = 1, ..., n$ . The model

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Let  $y_i^* = \beta^T x_i + \varepsilon_i$  where  $\varepsilon_1, \dots, \varepsilon_n$ : i.i.d. with  $E\varepsilon_i = 0$  and  $\operatorname{var} \varepsilon_i = \sigma^2$ .

$$y_i = \begin{cases} 1, & y_i^* \geqslant 0, \\ 0, & y_i^* < 0. \end{cases}$$

$$P\{y_i = 1\} = P\{y_i^* \ge 0\} = P\{\beta^T x_i + \varepsilon_i \ge 0\} = P\{\varepsilon_i \ge -\beta^T x_i\} = 1 - F\left(-\frac{\beta^T x_i}{\sigma}\right) = F\left(\frac{\beta^T x_i}{\sigma}\right),$$

assuming F(x) is a continuous CDF of  $\varepsilon_i/\sigma$  and F(-x)=1-F(x).

$$\bullet \ \, \text{Probit model} \, \, \Phi(u) = \frac{1}{\sqrt{2\pi}} \int\limits_{-\infty}^{u} e^{-\frac{z^2}{2}} dz.$$

Lecture 5 5 / 32

5/32

- Probit model  $\Phi(u) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{u} e^{-\frac{z^2}{2}} dz$ .
- Logit model  $\Lambda(u) = \frac{e^u}{1 + e^u}$ .

How do we estimate  $\beta$ ? Consider the likelihood

$$L(y_1, ..., y_n) = \prod_{i:y_i=0} (1 - F(\beta^T x_i)) \prod_{i:y_i=1} F(\beta^T x_i)$$
$$= \prod_{i=1}^n F^{y_i}(\beta^T x_i) (1 - F(\beta^T x_i))^{1-y_i},$$

Lecture 5 5 / 32

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$$\ln L(y_1, ..., y_n) = \sum_{i=1}^n (y_i \ln F(\beta^T x_i) + (1 - y_i) \ln(1 - F(\beta^T x_i))),$$

5/32

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$$L(y_{1},...,y_{n}) = \prod_{i:y_{i}=0} (1 - F(\beta^{T}x_{i})) \prod_{i:y_{i}=1} F(\beta^{T}x_{i})$$

$$= \prod_{i=1}^{n} F^{y_{i}}(\beta^{T}x_{i})(1 - F(\beta^{T}x_{i}))^{1-y_{i}},$$

$$\ln L(y_{1},...,y_{n}) = \sum_{i=1}^{n} (y_{i} \ln F(\beta^{T}x_{i}) + (1 - y_{i}) \ln(1 - F(\beta^{T}x_{i}))),$$

$$\frac{\partial \ln L}{\partial \beta} = \sum_{i=1}^{n} \left( \frac{y_{i}f(\beta^{T}x_{i})}{F(\beta^{T}x_{i})} - \frac{(1 - y_{i})f(\beta^{T}x_{i})}{1 - F(\beta^{T}x_{i})} \right) x_{i} = 0,$$

where f is a pdf.

# For the logit model:

$$\frac{\partial \ln L}{\partial \beta} = \sum_{i=1}^{n} (y_i - \Lambda(\beta^T x_i)) x_i = 0.$$

ure 5 6 / 32

For the logit model:

$$\frac{\partial \ln L}{\partial \beta} = \sum_{i=1}^{n} (y_i - \Lambda(\beta^T x_i)) x_i = 0.$$

For the probit model:

$$\frac{\partial \ln L}{\partial \beta} = \sum_{i=1}^{n} \frac{(2y_i - 1)\varphi(\beta^T x_i)}{\Phi((2y_i - 1)\beta^T x_i)} x_i = 0, \quad \varphi(x) = \Phi'(x).$$

re 5 6 / 32

# Using R...

# glm: logit model

```
\begin{split} & \mathsf{glm}(\mathsf{formula},\,\mathsf{family}=\mathsf{binomial}("\mathsf{logit}"),\,\mathsf{data},\,\mathsf{weights},\,\mathsf{subset},\,\mathsf{na.action},\\ & \mathsf{start}=\mathsf{NULL},\,\mathsf{etastart},\,\mathsf{mustart},\,\mathsf{offset},\,\mathsf{control}=\mathsf{list}(...),\,\mathsf{model}=\mathsf{TRUE},\\ & \mathsf{method}="\mathsf{glm.fit}",\,\mathsf{x}=\mathsf{FALSE},\,\mathsf{y}=\mathsf{TRUE},\,\mathsf{contrasts}=\mathsf{NULL},\,...) \end{split}
```

## glm: probit model

```
\begin{split} & \mathsf{glm}(\mathsf{formula},\,\mathsf{family}=\mathsf{binomial}("\,\mathsf{probit}"),\,\mathsf{data},\,\mathsf{weights},\,\mathsf{subset},\\ & \mathsf{na.action},\,\mathsf{start}=\mathsf{NULL},\,\mathsf{etastart},\,\mathsf{mustart},\,\mathsf{offset},\,\mathsf{control}=\mathsf{list}(...),\\ & \mathsf{model}=\mathsf{TRUE},\,\mathsf{method}="\,\mathsf{glm.fit}",\,\mathsf{x}=\mathsf{FALSE},\,\mathsf{y}=\mathsf{TRUE},\,\mathsf{contrasts}=\mathsf{NULL},\,\ldots) \end{split}
```

re 5 7 / 32

# Using R...

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```

#### summary

summary(object)

Lecture 5 7 / 32

# confint

confint(object)

Lecture 5 8 / 32

## confint

confint(object)

#### wald.test

```
library(aod)
```

wald.test(Sigma, b, Terms = NULL, L = NULL, H0 = NULL, df = NULL,

verbose = FALSE)

vcov(object)

coef(object)

Lecture 5 8 / 32

 $Consider\ example\ [http://www.ats.ucla.edu/stat/r/dae/logit.htm].$ 

There are 4 variables.

There are 3 independent variables: mark at the exam (gre), mean mark (at school) (gpa), rank of the school (rank).

The dependent variable is the results of the exam: accept or reject at the University (admit).

Variables gre and gpa are continuous, variable rank is discrete (values from 1 (best school) to 4 (worst school).

Get data from website and print first 6 strings of dataset:

```
mydata <- read.csv("http://www.ats.ucla.edu/stat/data/
binary.csv")
head(mydata)</pre>
```

Lecture 5 9 / 32

# Result

	admit	gre	gpa	rank
1	0	380	3.61	3
2	1	660	3.67	3
3	1	800	4.00	1
4	1	640	3.19	4
5	0	520	2.93	4
6	1	760	3.00	2

For descriptive statistics we use function summary:

summary(mydata)

Lecture 5 10 / 32

#### Result

```
admit
                                                rank
                    gre
                                  gpa
               Min. :220.0
                              Min. :2.260
Min. :0.0000
                                            Min. :1.000
1st Qu.:0.0000
               1st Qu.:520.0
                             1st Qu.:3.130
                                            1st Qu.:2.000
Median :0.0000
               Median :580.0
                             Median :3.395
                                            Median :2.000
Mean :0.3175
              Mean:587.7 Mean:3.390
                                             Mean :2.485
3rd Qu.:1.0000
               3rd Qu.:660.0
                             3rd Qu.:3.670
                                            3rd Qu.:3.000
Max. :1.0000
               Max. :800.0
                              Max. :4.000
                                            Max. :4.000
```

sapply(mydata, sd)

#### Result

admit gre gpa rank 0.4660867 115.5165364 0.3805668 0.9444602

Lecture 5 11 / 32

For binary regression we use function glm. Before this, we convert rank using function factor to make it categoric variable.

The name of the model is mylogit.

Main arguments are "formula": admit  $\sim$  gre + gpa + rank, name of dataset: data = mydata, and type of regression model: family = binomial("logit"):

```
mydata$rank <- factor(mydata$rank)
mylogit <- glm(admit ~ gre + gpa + rank, data = mydata,
family = binomial("logit"))</pre>
```

Lecture 5 12 / 32

# summary(mylogit)

```
Results:
Call:
glm(formula = admit \sim gre + gpa + rank, family = binomial("logit"),
data = mydata
Deviance Residuals:
  Min
                  Median
                             3Q
                                     Max
           1Q
 -1.6268 -0.8662 -0.6388
                            1.1490
                                    2.0790
Coefficients:
             Estimate Std. Error
                                   z value
                                            Pr(>|z|)
 (Intercept)
             -3.989979
                        1.139951
                                   -3.500
                                            0.000465
             0.002264
                        0.001094
                                    2.070
                                            0.038465
    gre
             0.804038
                        0.331819 2.423
                                            0.015388
    gpa
                                                        *
 rank[T.2]
            -0.675443
                        0.316490
                                   -2.134
                                            0.032829
                                                       ***
 rank[T.3] -1.340204
                        0.345306
                                   -3.881
                                            0.000104
                                                       ***
 rank[T.4]
            -1.551464
                        0.417832
                                   -3.713
                                            0.000205
```

Lecture 5 13 / 32

— Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 499.98 on 399 degrees of freedom

Residual deviance: 458.52 on 394 degrees of freedom

AIC: 470.52

Number of Fisher Scoring iterations: 4.

Lecture 5 14 / 32

Estimates of coefficients of logistic regression are presented in Table, where one can find the values of Wald statistics and corresponding p-values. By default,  $\alpha=0.05$ .

All coefficients are significant for the example.

We may use function confint for construction of confidence intervals with confidence level 0.95.

## confint(mylogit)

#### Results

	2.5 %	97.5 %
(Intercept)	-6.271620	-1.79255
gre	0.000138	0.00444
gpa	0.160296	1.46414
rank2	-1.300889	-0.05675
rank3	-2.027671	-0.67037
rank4	-2.400027	-0.75354

Lecture 5 15 / 32

Maximal likelihood method is used for confidence intervals.

Likelihood ratio statistics is equal Null deviance — Residual deviance and it is 499.98-458.52=41.46.

If the null hypotesis (about non-significance of logistic regression) is true, statistics of the test is distributed by  $\chi^2$ -distribution with 399-394=5 degrees of freedom.

Critical range is  $(11.0705, \infty)$ .

Value AIC can be used for model fitness. The less the value, the better the model.

One may use Wald test for estimation of logit model with function wald.test from package aod:

```
install.packages("aod")
library(aod)
wald.test(b = coef(mylogit), Sigma = vcov(mylogit), Terms = 1:6)
```

By Terms the list of independent variables to be tested is given.

Lecture 5 16 / 32

#### wald.test:

Wald test:

Chi-squared test:

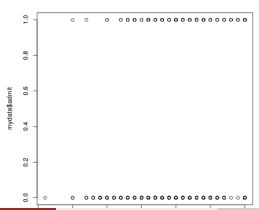
$$X2 = 73.4$$
,  $df = 6$ ,  $P(> X2) = 8.2e - 14$ 

Wald statistics is 73.4. The p-value is 8.2e-14, less than 0.05. We use coefficients of logistic regression as b, and vcov(mylogit) as covariance matrix. As Terms we use independent variables 1-6.

Lecture 5 17 / 32

For graph of admit as a function of gre, and admit as a function of gpa we use plot:

```
plot(mydata$admit,mydata$gre)
plot(mydata$admit,mydata$gpa)
```



Lecture 5 18 / 32

# Probit model

```
\label{eq:myprobit} \mbox{myprobit} < - \mbox{glm(admit} \sim \mbox{gre} + \mbox{gpa} + \mbox{rank, data} = \mbox{mydata, family} = \mbox{binomial("probit"))} \\ \mbox{summary(myprobit)}
```

#### Results:

```
\label{eq:glm} $$ glm(formula = admit \sim gre + gpa + rank, family = binomial("probit"), $$ data = mydata)$$ Deviance Residuals: $$ Min 1Q Median 3Q Max $$ -1.6163 -0.8710 -0.6389 1.1560 2.1035 $$
```

Lecture 5 19 / 32

#### Coefficients:

	Estimate	Std. Error	z value	Pr(> z )	
(Intercept)	-2.386836	0.673946	-3.542	0.000398	***
gre	0.001376	0.000650	2.116	0.034329	*
gpa	0.477730	0.197197	2.423	0.015410	*
rank[T.2]	-0.415399	0.194977	-2.131	0.033130	*
rank[T.3]	-0.812138	0.208358	-3.898	9.71e-05	***
rank[T.4]	-0.935899	0.245272	-3.816	0.000136	***

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 (Dispersion parameter for binomial family taken to be 1)

Null deviance: 499.98 on 399 degrees of freedom

Residual deviance: 458.41 on 394 degrees of freedom

AIC: 470.41

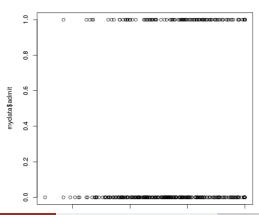
Number of Fisher Scoring iterations: 4

Lecture 5 20 / 32

The model is significant.

For probit model the maximal likelihood statistics is 41.57.

If the null hypotesis (about non-significance of probit regression) is true, statistics of the test is distributed by  $\chi^2$ -distribution with 5 degrees of freedom. Critical range is  $(11.0705, \infty)$ .



Lecture 5 21 / 32

# VIF (test for multicollinearity)

Like in case of linear regression, we should check for multicollinearity in the model.

```
vif(mylogit)
vif(myprobit)
```

Lecture 5 22 / 32

Value AIC for probit model is less than for logit model, therefore, probit model is better according to this criteria.

Confidence intervals for probit model:

# confint(myprobit)

# Result:

	2.5 %	97.5 %
(Intercept)	-3.7201050682	-1.076327713
gre	0.0001104101	0.002655157
gpa	0.0960654793	0.862610221
rank[T.2]	-0.7992113929	-0.032995019
rank[T.3]	-1.2230955861	-0.405008112
rank[T.4]	-1.4234218227	-0.459538829

23 / 32

Verify the probit model using function wald.test:

library(aod) wald.test(b = coef(myprobit), Sigma = vcov(myprobit), Terms = 1:6)

## Result:

Wald test:

Chi-squared test:

$$X2 = 83.6$$
, df = 6,  $P(> X2) = 6.7e - 16$ 

The value of Wald statistics is 83.6. The p-value is 6.7e-16 which is less 0.05.

Lecture 5 24 / 32

#### Predictions

predicted < - predict(mylogit, testData, type="response")</pre>

When we use the predict function on this model, it will predict the log(odds) of the Y variable. This is not what we ultimately want because, the predicted values may not lie within the 0 and 1 range as expected. So, to convert it into prediction probability scores that is bound between 0 and 1, we use the plogis().

#### Predictions

predicted < - plogis(predict(mylogit, testData))</pre>

Lecture 5 25 / 32

# Decide on optimal prediction probability cutoff for the model

- The default cutoff prediction probability score is 0.5 or the ratio of 1's and 0's in the training data. But sometimes, tuning the probability cutoff can improve the accuracy in both the development and validation samples.
- The InformationValue::optimalCutoff function provides ways to find the optimal cutoff to improve the prediction of 1's, 0's, both 1's and 0's and to reduce the misclassification error.
- Compute the optimal score that minimizes the misclassification error for the above model.

## CutOff

 $\label{library-libra$ 

Lecture 5 26 / 32

# Confusion Matrix

Predictions vs. Observed	1	0
1	TP	FP
0	FN	TN

- TP: True Positive,
- FP: False Positive,
- FN: False Negative,
- TN: True Negative.

# confusion\_matrix(mylogit,DATA=NA)

where DATA is a data frame on which the confusion matrix will be made. If omitted, the confusion matrix is on the data used in mylogit.

Lecture 5 27 / 32

# Confusion Matrix

If specified, the data frame must have the same column names as the data used to build the model in mylogit. This function makes classifications on the data used to build a logistic regression model by predicting dependent variable when the predicted probability exceeds 50%.

confusionMatrix(testData\$admit, predicted, threshold = optCutOff)

Lecture 5 28 / 32

# Specificity and Sensitivity

Sensitivity (or True Positive Rate) is the percentage of 1's (actuals) correctly predicted by the model:

$$Sensitivity = \frac{TP}{TP + FN}$$

Specificity is the percentage of 0's (actuals) correctly predicted. Specificity can also be calculated as 1 - FalsePositiveRate.

$$Specificity = \frac{TN}{TN + FP}$$

## Sensitivity and specificity

sensitivity(testData\$admit, predicted, threshold = optCutOff) specificity(testData\$admit, predicted, threshold = optCutOff)

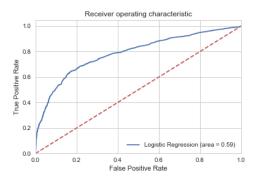
Lecture 5 29 / 32

The receiver operating characteristic (ROC) curve is another common tool used with binary classifiers. Receiver Operating Characteristic(ROC) summarizes the model's performance by evaluating the trade offs between true positive rate (sensitivity) and false positive rate (1- specificity). For plotting ROC, it is advisable to assume p > 0.5 since we are more concerned about success rate. ROC summarizes the predictive power for all possible values of p > 0.5. The area under curve (AUC), referred to as index of accuracy(A) or concordance index, is a perfect performance metric for ROC curve. Higher the area under curve, better the prediction power of the model. Below is a sample ROC curve. The ROC of a perfect predictive model has TP equals 1 and FP equals 0. This curve will touch the top left corner of the graph.

#### **ROC Curve**

plotROC(testData\$admit, predicted)

# **ROC** Curve



Lecture 5 31 / 32

## Dataset

#### binary regression.xls

Row: student.

First column: the number of hours the student slept before exam.

Second column: the number of hours the student studied before exam.

Third column: the exam is passed (1) or failed (0).

Make logit and probit models, write the equation of the model, test the

model with different statistics.

Lecture 5 32 / 32