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Attilio Meucci

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Formulas and figures in this presentation refer to the book **Risk and Asset Allocation**, Springer.

The notation, say, (5.24) refers to Formula 24 in Chapter 5 of the book

The notation, say, (T4.12) refers to Formula 12 in the Technical Appendices for Chapter 4, which can be downloaded from www.symmys.com

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$$\mathbf{G}[f_{\mathbf{X}}] \equiv \text{"unknown truth"}. \quad (4.6)$$

$$\text{information } i_T \mapsto \text{number } \hat{\mathbf{G}} \quad (4.9)$$

$$\overset{\uparrow}{i_T} \equiv \{\mathbf{x}_1, \dots, \mathbf{x}_T\} \quad (4.8)$$

$$(4.13) \quad \hat{\mathbf{G}}[i_T] \approx \mathbf{G}[f_{\mathbf{X}}] \quad ?$$

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$$G[f_X] \equiv \int_{-\infty}^{+\infty} x f_X(x) dx. \quad (4.7)$$

$$\hat{G}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T x_t. \quad (4.10)$$

$$\hat{G}[i_T] \equiv x_1 x_T \quad (4.11)$$

$$\hat{G}[i_T] \equiv 3. \quad (4.12)$$

$$(4.13) \quad \hat{\mathbf{G}}[i_T] \approx \mathbf{G}[f_{\mathbf{X}}] \quad ?$$

$$\mathbf{G} [f_{\mathbf{X}}] \equiv \text{"unknown truth"}$$

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$$\text{information } i_T \mapsto \text{number } \widehat{\mathbf{G}}$$

(4.9)

$$\uparrow$$
$$i_T \equiv \{\mathbf{x}_1, \dots, \mathbf{x}_T\}$$

(4.8)

$$i_T \equiv \{\mathbf{x}_1, \dots, \mathbf{x}_T\} \mapsto I_T \equiv \{\mathbf{X}_1, \dots, \mathbf{X}_T\} \quad (4.14)$$

$$\widehat{\mathbf{G}} [i_T] \mapsto \widehat{\mathbf{G}} [I_T] \quad (4.15)$$



$$\text{Loss} \left(\widehat{\mathbf{G}}, \mathbf{G} \right) \equiv \left\| \widehat{\mathbf{G}} [I_T] - \mathbf{G} [f_{\mathbf{X}}] \right\|^2$$

(4.19)

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$$\hat{G}[i_T] \mapsto \hat{G}[I_T]$$

(4.15)

$$G[f_X] = \mu$$

(4.18)

$$X_t \sim N(\mu, \sigma^2)$$

(4.16)

$$\hat{G}[I_T] \equiv \frac{1}{T} \sum_{t=1}^T X_t \sim N\left(\mu, \frac{\sigma^2}{T}\right)$$

(4.17)



$$\text{Loss}(\hat{G}, G) \equiv \left\| \hat{G}[I_T] - G[f_X] \right\|^2$$

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(4.22)

$$\text{Loss}(\hat{G}, G) \equiv \left\| \hat{G}[I_T] - G[f_X] \right\|^2$$

(4.19)

$$\text{Err} \left(\hat{\mathbf{G}}, \mathbf{G} \right) \equiv \sqrt{\mathbf{E} \left\{ \left\| \hat{\mathbf{G}} \left(I_T \right) - \mathbf{G} \left[f_{\mathbf{X}} \right] \right\|^2 \right\}}. \quad (4.23)$$


$$\boxed{\text{Loss} \left(\hat{\mathbf{G}}, \mathbf{G} \right) \equiv \left\| \hat{\mathbf{G}} \left[I_T \right] - \mathbf{G} \left[f_{\mathbf{X}} \right] \right\|^2} \quad (4.19)$$

$$\text{Err}\left(\widehat{G}, G\right)=\frac{\sigma}{\sqrt{T}} \tag{4.24}$$

$$\text{Err}\left(\widehat{\mathbf{G}}, \mathbf{G}\right) \equiv \sqrt{\mathbb{E}\left\{\left\|\widehat{\mathbf{G}}\left(I_T\right)-\mathbf{G}\left[f_{\mathbf{X}}\right]\right\|^2\right\}} \tag{4.23}$$

$$G\left[f_X\right]=\mu. \tag{4.18}$$

$$X_t \sim \mathcal{N}\left(\mu, \sigma^2\right) \tag{4.16}$$

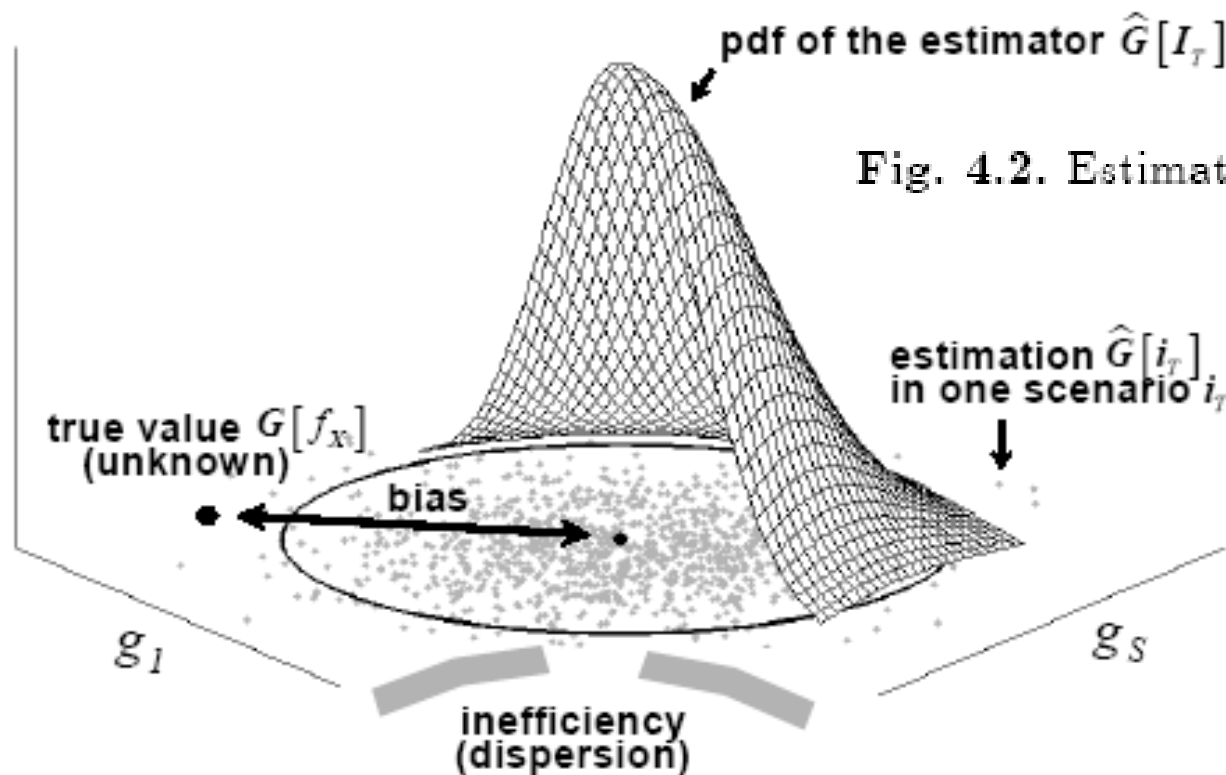
$$\widehat{G}\left[I_T\right] \equiv \frac{1}{T} \sum_{t=1}^T X_t \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{T}\right) \tag{4.17}$$

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$$\text{Inef}^2 [\hat{G}] \equiv E \left\{ \left\| \hat{G} [I_T] - E \left\{ \hat{G} [I_T] \right\} \right\|^2 \right\} \quad (4.26)$$

$$\text{Bias}^2 [\hat{G}, G] \equiv \left\| E \left\{ \hat{G} [I_T] \right\} - G [f_X] \right\|^2 \quad (4.25)$$



$$\text{Inef}^2 \left[\widehat{\mathbf{G}} \right] \equiv \mathbb{E} \left\{ \left\| \widehat{\mathbf{G}} [I_T] - \mathbb{E} \left\{ \widehat{\mathbf{G}} [I_T] \right\} \right\|^2 \right\} \quad (4.26)$$

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$$\text{Inef} \left[\widehat{G} \right] = \frac{\sigma}{\sqrt{T}} \tag{4.29}$$

$$\text{Bias} \left[\widehat{G}, G \right] = 0 \tag{4.28}$$


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$$\text{Err}^2 \left[\widehat{\mathbf{G}}, \mathbf{G} \right] = \text{Bias}^2 \left[\widehat{\mathbf{G}}, \mathbf{G} \right] + \text{Inef}^2 \left[\widehat{\mathbf{G}} \right]$$

(4.27)

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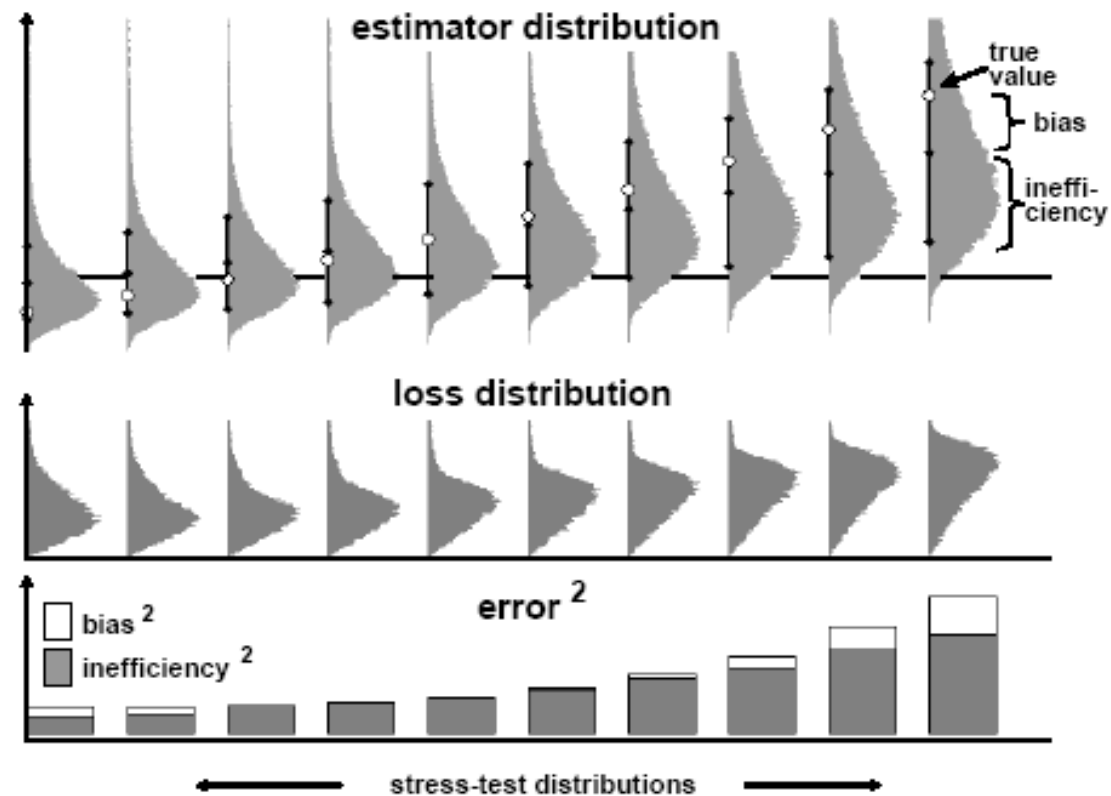
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$$f_{i_T}(\mathbf{x}) \equiv \frac{1}{T} \sum_{t=1}^T \delta^{(\mathbf{x}_t)}(\mathbf{x}) \quad (4.35)$$

$$\lim_{T \rightarrow \infty} F_{i_T}(\mathbf{x}) = F_{\mathbf{X}}(\mathbf{x}) \quad (4.34)$$

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$$\int_{-\infty}^{q_p[f_X]} f_X (x) dx \equiv p, \quad (4.38)$$

$$\hat{q}_p [i_T] \equiv x_{[pT]:T} \quad (4.39)$$

NON-PARAMETRIC ESTIMATORS – ORDINARY LEAST SQUARES

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$$\mathbf{X} = \mathbf{B}\mathbf{F} + \mathbf{U}. \quad (4.50)$$

$$\mathbf{B}_r \equiv \mathbf{E} \{ \mathbf{X}\mathbf{F}' \} \mathbf{E} \{ \mathbf{F}\mathbf{F}' \}^{-1} \quad (3.121)$$

$$\hat{\mathbf{B}} [i_T] \equiv \left(\sum_t \mathbf{x}_t \mathbf{f}_t' \right) \left(\sum_t \mathbf{f}_t \mathbf{f}_t' \right)^{-1} \quad (4.52)$$

$$i_T \equiv \{ \mathbf{x}_1, \mathbf{f}_1, \dots, \mathbf{x}_T, \mathbf{f}_T \} \quad (4.51)$$

NON-PARAMETRIC ESTIMATORS – ORDINARY LEAST SQUARES

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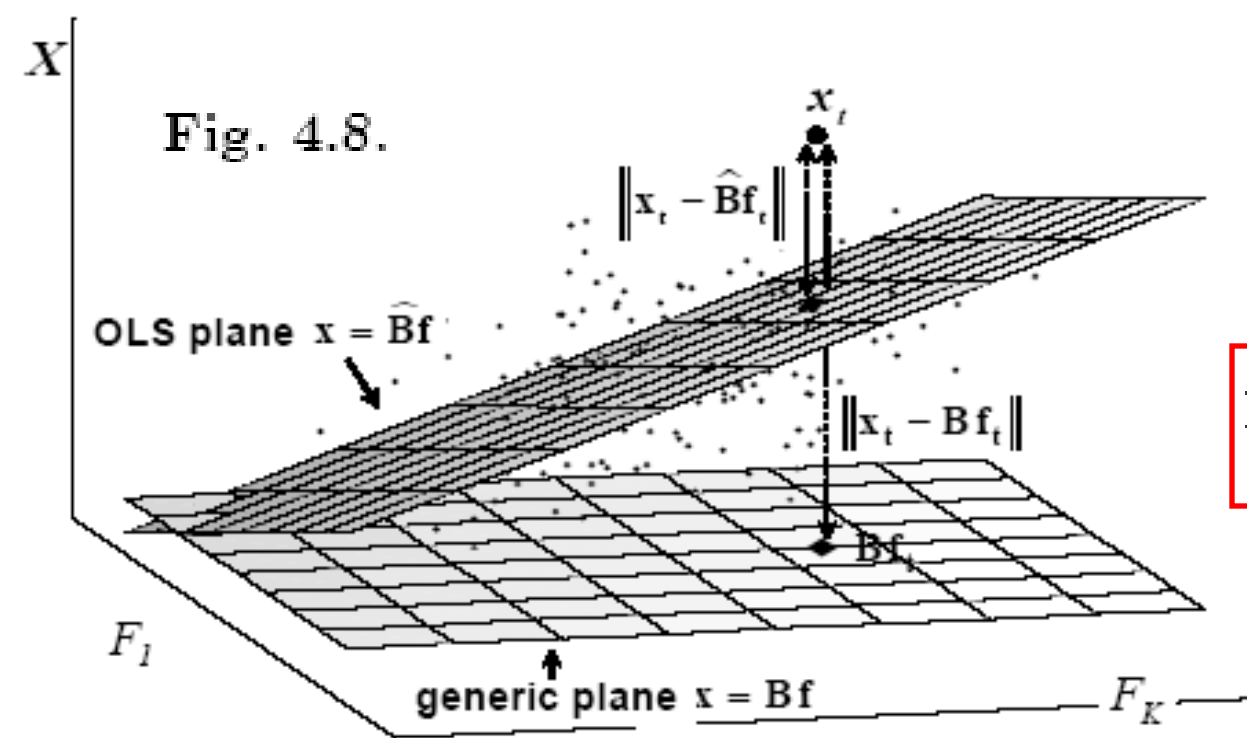
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$$\hat{\mathbf{B}} = \operatorname{argmin}_{\mathbf{B}} \sum_t \|\mathbf{x}_t - \mathbf{B}\mathbf{f}_t\|^2 \quad (4.53)$$

NON-PARAMETRIC ESTIMATORS – SAMPLE MEAN/COVARIANCE

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$$\mathbf{G}[f_{\mathbf{X}}] \equiv \text{"unknown truth"}. \quad (4.6)$$

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$$\text{Cov}\{\mathbf{X}\} \quad (2.67)$$

$$\text{information } i_T \mapsto \widehat{\mathbf{G}}[i_T] \equiv \mathbf{G}[f_{i_T}] \quad (4.36)$$

$$\widehat{\mathbf{E}}[i_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.41)$$

$$\widehat{\text{Cov}}[i_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\mathbf{E}}[i_T]) (\mathbf{x}_t - \widehat{\mathbf{E}}[i_T])' \quad (4.42)$$

NON-PARAMETRIC ESTIMATORS – SAMPLE MEAN/COVARIANCE

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$$\mathbf{G} [f_{\mathbf{X}}] \equiv \text{"unknown truth"} \quad (4.6)$$

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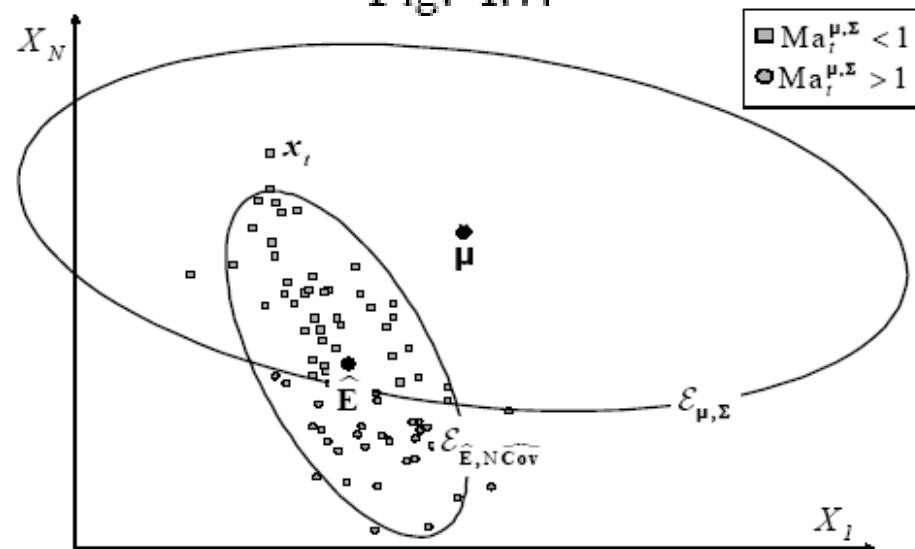
$$\widehat{\mathbf{E}} [i_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.41)$$

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$$\mathcal{E}_{\mu, \Sigma} \equiv \{ \mathbf{x} \in \mathbb{R}^N \text{ such that } (\mathbf{x} - \mu)' \Sigma^{-1} (\mathbf{x} - \mu) \leq 1 \} \quad (4.45)$$

$$\text{Ma}_t^{\mu, \Sigma} \equiv \text{Ma} (\mathbf{x}_t, \mu, \Sigma) \equiv \sqrt{(\mathbf{x}_t - \mu)' \Sigma^{-1} (\mathbf{x}_t - \mu)} \quad (4.46)$$

Fig. 4.7.



NON-PARAMETRIC ESTIMATORS – SAMPLE MEAN/COVARIANCE

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$$\text{information } i_T \mapsto \widehat{\mathbf{G}}[i_T] \equiv \mathbf{G}[f_{i_T}] \quad (4.36)$$

$$\widehat{\mathbf{E}}[i_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.41)$$

$$\widehat{\text{Cov}}[i_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\mathbf{E}}[i_T]) (\mathbf{x}_t - \widehat{\mathbf{E}}[i_T])' \quad (4.42)$$

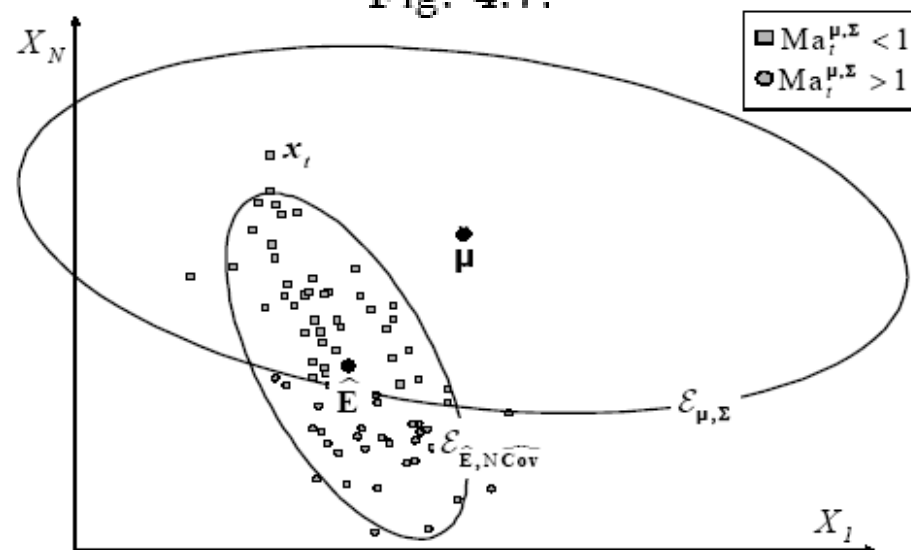
$$\mathcal{E}_{\mu, \Sigma} \equiv \{\mathbf{x} \in \mathbb{R}^N \text{ such that } (\mathbf{x} - \mu)' \Sigma^{-1} (\mathbf{x} - \mu) \leq 1\} \quad (4.45)$$

$$\text{Ma}_t^{\mu, \Sigma} \equiv \text{Ma}(\mathbf{x}_t, \mu, \Sigma) \equiv \sqrt{(\mathbf{x}_t - \mu)' \Sigma^{-1} (\mathbf{x}_t - \mu)} \quad (4.46)$$

$$\overline{r^2}(\mu, \Sigma) \equiv \frac{1}{T} \sum_{t=1}^T \left(\text{Ma}_t^{\mu, \Sigma} \right)^2 \quad (4.47)$$

$$\begin{aligned} (\widehat{\mathbf{E}}, N\widehat{\text{Cov}}) &= \underset{(\mu, \Sigma) \in \mathcal{C}}{\text{argmin}} [\text{Vol}\{\mathcal{E}_{\mu, \Sigma}\}] \\ \overline{r^2}(\mu, \Sigma) &\equiv 1 \end{aligned} \quad (4.48)$$

Fig. 4.7.



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$$\text{information } i_T \mapsto \hat{\mathbf{G}}[i_T] \equiv \mathbf{G}[f_{i_T;\epsilon}] \quad (4.56)$$

$$f_{i_T} \mapsto f_{i_T;\epsilon} \equiv \frac{1}{T} \sum_{t=1}^T \frac{1}{(2\pi)^{\frac{N}{2}} \epsilon^N} e^{-\frac{1}{2\epsilon^2} (\mathbf{x} - \mathbf{x}_t)' (\mathbf{x} - \mathbf{x}_t)}. \quad (4.55)$$

MAXIMUM LIKELIHOOD ESTIMATORS

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MAXIMUM LIKELIHOOD ESTIMATORS

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{G}[f_{\mathbf{X}}] \equiv \text{"unknown truth"}. \quad (4.6)$$

$$\text{information } i_T \mapsto \text{number } \hat{\mathbf{G}} \quad (4.9)$$

$$i_T \equiv \{\mathbf{x}_1, \dots, \mathbf{x}_T\} \quad (4.8)$$

MAXIMUM LIKELIHOOD ESTIMATORS

Risk and Asset Allocation - Springer – symmys.com

$$G[f_X] \equiv \text{"unknown truth"} \quad (4.6)$$

$$\text{information } i_T \mapsto \text{number } \hat{G} \quad (4.9)$$

$$f_{\theta}(i_T) \equiv f_{\theta}(x_1) \cdots f_{\theta}(x_T). \quad (4.65)$$

$$\hat{\theta}[i_T] \equiv \operatorname{argmax}_{\theta \in \Theta} f_{\theta}(i_T) \quad (4.66)$$

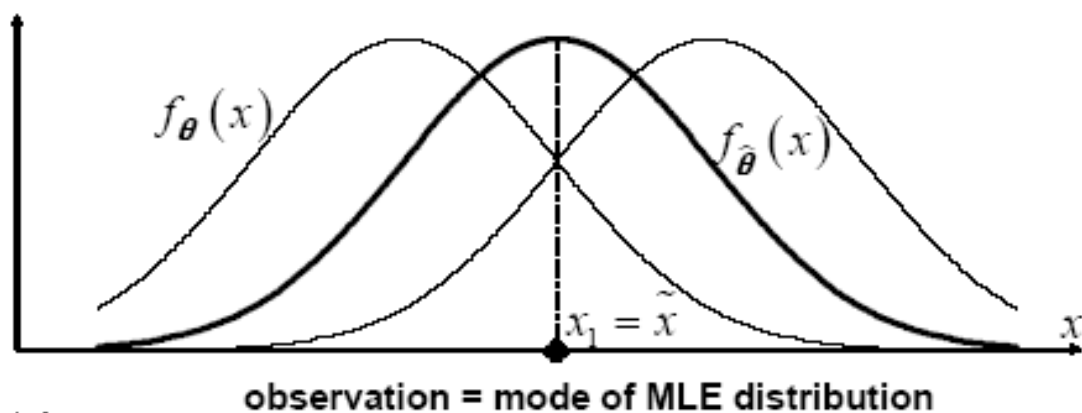


Fig. 4.10

MAXIMUM LIKELIHOOD ESTIMATORS

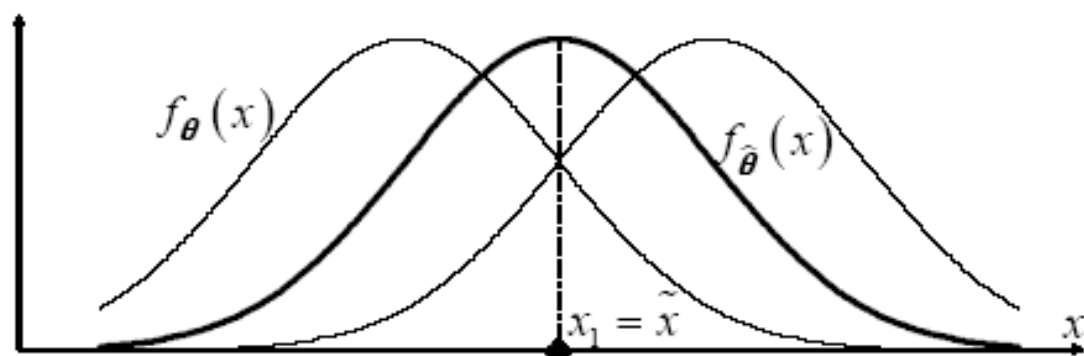
Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{G}[f_{\mathbf{X}}] \equiv \text{"unknown truth"} \quad (4.6)$$

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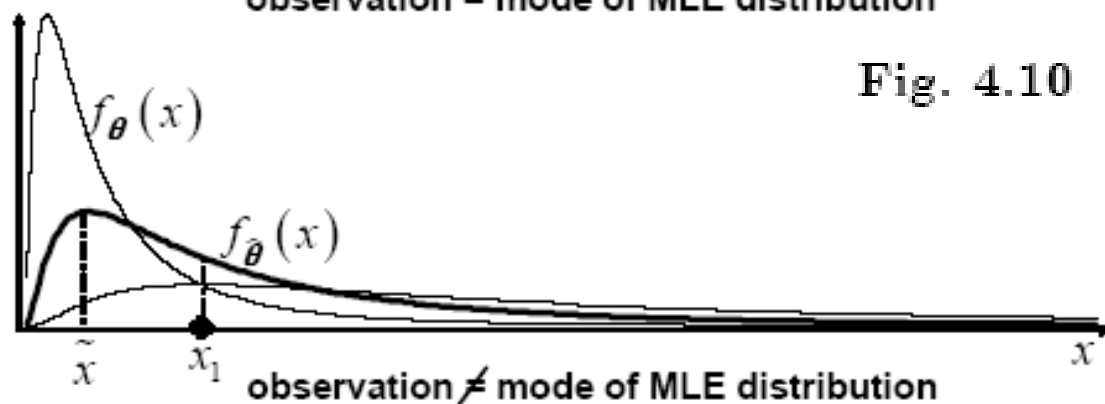
$$f_{\theta}(i_T) \equiv f_{\theta}(x_1) \cdots f_{\theta}(x_T) \quad (4.65)$$

$$\hat{\theta}[i_T] \equiv \operatorname{argmax}_{\theta \in \Theta} f_{\theta}(i_T) \quad (4.66)$$



observation = mode of MLE distribution

Fig. 4.10



observation \neq mode of MLE distribution

MAXIMUM LIKELIHOOD ESTIMATORS

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$$\mathbf{G} [f_{\mathbf{X}}] \equiv \text{"unknown truth"}. \quad (4.6)$$

$$\text{information } i_T \mapsto \text{number } \hat{\mathbf{G}} \quad (4.9)$$

$$f_{\boldsymbol{\theta}} (i_T) \equiv f_{\boldsymbol{\theta}} (\mathbf{x}_1) \cdots f_{\boldsymbol{\theta}} (\mathbf{x}_T). \quad (4.65)$$

$$\hat{\boldsymbol{\theta}} [i_T] \equiv \operatorname{argmax}_{\boldsymbol{\theta} \in \Theta} f_{\boldsymbol{\theta}} (i_T) \quad (4.66)$$

$$\bullet \quad \widehat{g(\boldsymbol{\theta})} = g(\hat{\boldsymbol{\theta}}) \quad (4.70)$$

$$\bullet \quad \hat{\boldsymbol{\theta}} [i_T] \sim \mathcal{N} \left(\boldsymbol{\theta}, \frac{\boldsymbol{\Gamma}}{T} \right) \quad (4.71)$$

$$\boldsymbol{\Gamma} \equiv \operatorname{Cov} \left\{ \frac{\partial \ln (f_{\boldsymbol{\theta}} (\mathbf{X}))}{\partial \boldsymbol{\theta}} \right\} \quad (4.72)$$

MAXIMUM LIKELIHOOD ESTIMATORS – ELLIPTICAL

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X} \sim \text{El}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, g) \quad (4.73)$$

$$f_{\theta}(\mathbf{x}) \equiv \frac{1}{\sqrt{|\boldsymbol{\Sigma}|}} g(\text{Ma}^2(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma})) \quad (4.74)$$

$$\text{Ma}(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \equiv \sqrt{(\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})} \quad (4.75)$$

MAXIMUM LIKELIHOOD ESTIMATORS – ELLIPTICAL

Risk and Asset Allocation - Springer – symmys.com

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$$\text{Ma}(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \equiv \sqrt{(\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})} \quad (4.75)$$

$$\hat{\boldsymbol{\mu}} = \sum_{t=1}^T \frac{w_t}{\sum_{s=1}^T w_s} \mathbf{x}_t \quad (4.81)$$

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=1}^T w_t (\mathbf{x}_t - \hat{\boldsymbol{\mu}}) (\mathbf{x}_t - \hat{\boldsymbol{\mu}})'. \quad (4.82)$$

$$w_t \equiv w(\text{Ma}^2(\mathbf{x}_t, \hat{\boldsymbol{\mu}}, \hat{\boldsymbol{\Sigma}})) \quad (4.80)$$

$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.79)$$

MAXIMUM LIKELIHOOD ESTIMATORS – ELLIPTICAL

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X} \sim \text{El}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, g) \quad (4.73)$$

$$f_{\theta}(\mathbf{x}) \equiv \frac{1}{\sqrt{|\boldsymbol{\Sigma}|}} g(\text{Ma}^2(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma})) \quad (4.74)$$

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$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.79)$$

$$g^N(z) \equiv \frac{e^{-\frac{z}{2}}}{(2\pi)^{\frac{N}{2}}}, \quad (4.96)$$

$$w(z) \equiv 1 \quad (4.97)$$

MAXIMUM LIKELIHOOD ESTIMATORS – ELLIPTICAL

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X} \sim \text{El}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, g) \quad (4.73)$$

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$$\hat{\boldsymbol{\mu}} = \sum_{t=1}^T \frac{w_t}{\sum_{s=1}^T w_s} \mathbf{x}_t \quad (4.81)$$

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$$w_t \equiv w(\text{Ma}^2(\mathbf{x}_t, \hat{\boldsymbol{\mu}}, \hat{\boldsymbol{\Sigma}})) \quad (4.80)$$

$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.79)$$

$$g^N(z) \equiv \frac{e^{-\frac{z}{2}}}{(2\pi)^{\frac{N}{2}}}, \quad (4.96)$$

$$w(z) \equiv 1 \quad (4.97)$$

$$g^{\text{Ca}}(z) = \frac{\Gamma(\frac{1+N}{2})}{\Gamma(\frac{1}{2}) (\pi)^{\frac{N}{2}}} (1+z)^{-\frac{1+N}{2}} \quad (4.83)$$

$$w_t = \frac{N+1}{1 + \text{Ma}^2(\mathbf{x}_t, \hat{\boldsymbol{\mu}}, \hat{\boldsymbol{\Sigma}})} \quad (4.84)$$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOCATION)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.95)$

estimator definition $\hat{\boldsymbol{\mu}}[i_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.98)$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOCATION)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim \mathbf{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\mu}}[I_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t$. (4.98)

estimator distribution
(replicability) $\hat{\boldsymbol{\mu}}[I_T] \sim \mathbf{N}\left(\boldsymbol{\mu}, \frac{\boldsymbol{\Sigma}}{T}\right)$ (4.102)

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOCATION)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\mu}}[i_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t$. (4.98)

estimator distribution
(replicability) $\hat{\boldsymbol{\mu}}[I_T] \sim N\left(\boldsymbol{\mu}, \frac{\boldsymbol{\Sigma}}{T}\right)$ (4.102)



Generalized t-tests...

Generalized p-values...

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOCATION)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim \mathbf{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\mu}}[I_T] = \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t$. (4.98)

estimator distribution
(replicability) $\hat{\boldsymbol{\mu}}[I_T] \sim \mathbf{N}\left(\boldsymbol{\mu}, \frac{\boldsymbol{\Sigma}}{T}\right)$ (4.102)



estimator evaluation
(global) $\text{Loss}(\hat{\boldsymbol{\mu}}, \boldsymbol{\mu}) \equiv [\hat{\boldsymbol{\mu}}[I_T] - \boldsymbol{\mu}]' [\hat{\boldsymbol{\mu}}[I_T] - \boldsymbol{\mu}]$ (4.108)

estimator evaluation
(summary) $\left\{ \begin{array}{l} \text{Err}^2(\hat{\boldsymbol{\mu}}, \boldsymbol{\mu}) = \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.109) \\ \text{Inef}^2(\hat{\boldsymbol{\mu}}) = \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.110) \\ \text{Bias}^2(\hat{\boldsymbol{\mu}}, \boldsymbol{\mu}) = 0. \quad (4.111) \end{array} \right.$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOCATION)

Risk and Asset Allocation - Springer – symmys.com

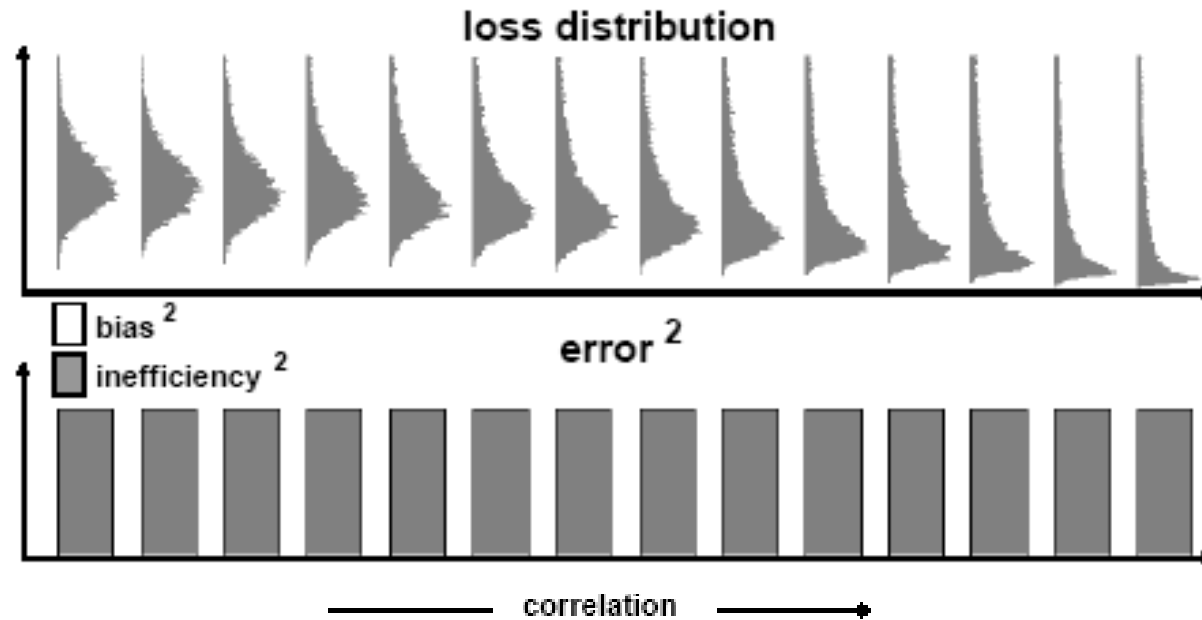


Fig. 4.11.

estimator evaluation
(global)

$$\text{Loss}(\hat{\mu}, \mu) \equiv [\hat{\mu}[I_T] - \mu]' [\hat{\mu}[I_T] - \mu] \quad (4.108)$$

estimator evaluation
(summary)

$$\left\{ \begin{array}{l} \text{Err}^2(\hat{\mu}, \mu) = \frac{1}{T} \text{tr}(\Sigma) \quad (4.109) \\ \text{Inef}^2(\hat{\mu}) = \frac{1}{T} \text{tr}(\Sigma) \quad (4.110) \\ \text{Bias}^2(\hat{\mu}, \mu) = 0. \quad (4.111) \end{array} \right.$$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (SCATTER)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\Sigma}}[i_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \hat{\boldsymbol{\mu}}) (\mathbf{x}_t - \hat{\boldsymbol{\mu}})'$ (4.99)

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (SCATTER)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\Sigma}}[i_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \hat{\boldsymbol{\mu}}) (\mathbf{x}_t - \hat{\boldsymbol{\mu}})'$ (4.99)

estimator distribution
(replicability) $T\hat{\boldsymbol{\Sigma}}[I_T] \sim W(T-1, \boldsymbol{\Sigma})$ (4.103)

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (SCATTER)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\Sigma}}[I_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \hat{\boldsymbol{\mu}})(\mathbf{x}_t - \hat{\boldsymbol{\mu}})'$ (4.99)

estimator distribution
(replicability) $T\hat{\boldsymbol{\Sigma}}[I_T] \sim W(T-1, \boldsymbol{\Sigma})$ (4.103)



Generalized t-tests...

Generalized p-values...

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (SCATTER)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ (4.95)

estimator definition $\hat{\boldsymbol{\Sigma}}[i_T] = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \hat{\boldsymbol{\mu}}) (\mathbf{x}_t - \hat{\boldsymbol{\mu}})'$ (4.99)

estimator distribution
(replicability) $T\hat{\boldsymbol{\Sigma}}[I_T] \sim W(T-1, \boldsymbol{\Sigma})$ (4.103)

estimator evaluation
(global) $\text{Loss}(\hat{\boldsymbol{\Sigma}}, \boldsymbol{\Sigma}) \equiv \text{tr} \left[\left(\hat{\boldsymbol{\Sigma}}[I_T] - \boldsymbol{\Sigma} \right)^2 \right]$ (4.118)

estimator evaluation
(summary) $\left\{ \begin{array}{l} \text{Err}^2(\hat{\boldsymbol{\Sigma}}, \boldsymbol{\Sigma}) = \frac{1}{T} \left[\text{tr}(\boldsymbol{\Sigma}^2) + \left(1 - \frac{1}{T}\right) [\text{tr}(\boldsymbol{\Sigma})]^2 \right] \quad (4.119) \\ \text{Inef}^2(\hat{\boldsymbol{\Sigma}}) = \frac{1}{T} \left(1 - \frac{1}{T}\right) \left[\text{tr}(\boldsymbol{\Sigma}^2) + [\text{tr}(\boldsymbol{\Sigma})]^2 \right] \quad (4.120) \\ \text{Bias}^2(\hat{\boldsymbol{\Sigma}}, \boldsymbol{\Sigma}) = \frac{1}{T^2} \text{tr}(\boldsymbol{\Sigma}^2) \quad (4.121) \end{array} \right.$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (SCATTER)

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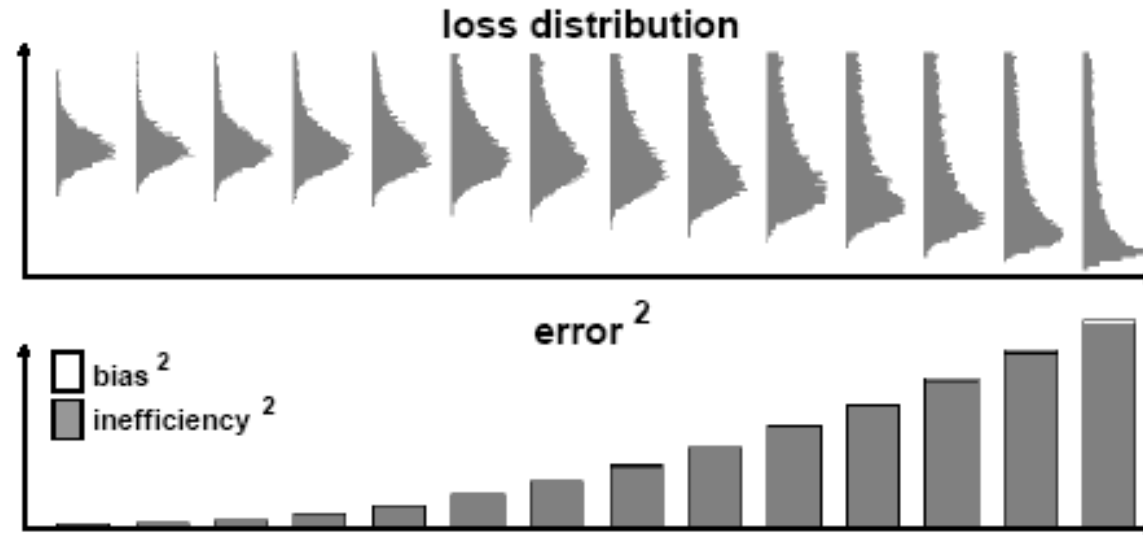


Fig. 4.12

estimator evaluation
(global)

$$\text{Loss} \left(\hat{\Sigma}, \Sigma \right) \equiv \text{tr} \left[\left(\hat{\Sigma} [I_T] - \Sigma \right)^2 \right] \quad (4.118)$$

estimator evaluation
(summary)

$$\text{Err}^2 \left(\hat{\Sigma}, \Sigma \right) = \frac{1}{T} \left[\text{tr} \left(\Sigma^2 \right) + \left(1 - \frac{1}{T} \right) [\text{tr} \left(\Sigma \right)]^2 \right] \quad (4.119)$$

$$\text{Inef}^2 \left(\hat{\Sigma} \right) = \frac{1}{T} \left(1 - \frac{1}{T} \right) \left[\text{tr} \left(\Sigma^2 \right) + [\text{tr} \left(\Sigma \right)]^2 \right] \quad (4.120)$$

$$\text{Bias}^2 \left(\hat{\Sigma}, \Sigma \right) = \frac{1}{T^2} \text{tr} \left(\Sigma^2 \right) \quad (4.121)$$

MAXIMUM LIKELIHOOD ESTIMATORS – NORMAL (LOADINGS)

Risk and Asset Allocation - Springer – symmys.com

normal assumption $\mathbf{X}|\mathbf{f} = \mathbf{B}\mathbf{f} + \mathbf{U}|\mathbf{f} \quad (4.88) \quad \mathbf{U}_t|\mathbf{f}_t \sim N(\mathbf{0}, \Sigma) \quad (4.123)$

estimator definition $\hat{\mathbf{B}}[i_T] = \hat{\Sigma}_{XF}[i_T] \hat{\Sigma}_F^{-1}[i_T] \quad (4.126)$

$$\hat{\Sigma}_{XF}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \mathbf{f}_t', \quad \hat{\Sigma}_F[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{f}_t \mathbf{f}_t'. \quad (4.127)$$

$$\hat{\Sigma}[i_T] = \frac{1}{T} \sum_{t=1}^T \left(\mathbf{x}_t - \hat{\mathbf{B}}[i_T] \mathbf{f}_t \right) \left(\mathbf{x}_t - \hat{\mathbf{B}}[i_T] \mathbf{f}_t \right)' \quad (4.128)$$

estimator distribution
(replicability) $\hat{\mathbf{B}}[I_T|\mathbf{f}_1, \dots, \mathbf{f}_T] \sim N\left(\mathbf{B}, \frac{\Sigma}{T}, \hat{\Sigma}_F^{-1}\right) \quad (4.129)$

$$T\hat{\Sigma}[I_T|\mathbf{f}_1, \dots, \mathbf{f}_T] \sim W(T - K, \Sigma) \quad (4.130)$$



Generalized t-tests, generalized p-values, estimator evaluation

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

Attilio Meucci

www.symmys.com

Formulas and figures in this presentation refer to the book **Risk and Asset Allocation**, Springer.

The notation, say, (5.24) refers to Formula 24 in Chapter 5 of the book

The notation, say, (T4.12) refers to Formula 12 in the Technical Appendices for Chapter 4, which can be downloaded from www.symmys.com

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.132)$$

$$\hat{\boldsymbol{\mu}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.135)$$

$$\hat{\boldsymbol{\mu}}[I_T] \sim N\left(\boldsymbol{\mu}, \frac{\boldsymbol{\Sigma}}{T}\right) \quad (4.102)$$

$$\text{Err}^2(\hat{\boldsymbol{\mu}}, \boldsymbol{\mu}) = \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.136)$$

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.132)$$

$$\hat{\boldsymbol{\mu}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.135)$$

$$\text{Err}^2(\hat{\boldsymbol{\mu}}^S, \boldsymbol{\mu}) < \frac{1}{T} \text{tr}(\boldsymbol{\Sigma})$$

? (4.137)

$$\text{Err}^2(\hat{\boldsymbol{\mu}}, \boldsymbol{\mu}) = \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.136)$$

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.132)$$

$$\hat{\boldsymbol{\mu}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.135)$$

$$\text{Err}^2(\hat{\boldsymbol{\mu}}^S, \boldsymbol{\mu}) < \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.137)$$

$\hat{\boldsymbol{\mu}}^S \equiv (1 - \alpha) \hat{\boldsymbol{\mu}} + \alpha \mathbf{b}.$

(4.138)

$$\alpha \equiv \frac{1}{T} \frac{N\bar{\lambda} - 2\lambda_1}{(\hat{\boldsymbol{\mu}} - \mathbf{b})'(\hat{\boldsymbol{\mu}} - \mathbf{b})}. \quad (4.139)$$

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.132)$$

$$\hat{\boldsymbol{\mu}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.135)$$

$$\text{Err}^2(\hat{\boldsymbol{\mu}}^S, \boldsymbol{\mu}) < \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.137)$$

$$\hat{\boldsymbol{\mu}}^S \equiv (1 - \alpha) \hat{\boldsymbol{\mu}} + \alpha \mathbf{b}. \quad (4.138)$$

$$\alpha \equiv \frac{1}{T} \frac{N\bar{\lambda} - 2\lambda_1}{(\hat{\boldsymbol{\mu}} - \mathbf{b})'(\hat{\boldsymbol{\mu}} - \mathbf{b})}. \quad (4.139)$$

$$\boldsymbol{\Sigma} \mapsto \hat{\boldsymbol{\Sigma}}.$$

$$\mathbf{b} \mapsto \frac{\mathbf{1}' \hat{\boldsymbol{\mu}}}{N} \mathbf{1} \quad (4.141)$$

$$\mathbf{b} \mapsto \frac{\mathbf{1}' \hat{\boldsymbol{\Sigma}}^{-1} \hat{\boldsymbol{\mu}}}{\mathbf{1}' \hat{\boldsymbol{\Sigma}}^{-1} \mathbf{1}} \mathbf{1}. \quad (4.142)$$

SHRINKAGE ESTIMATORS – LOCATION PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.132)$$

$$\hat{\boldsymbol{\mu}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t. \quad (4.135)$$

$$\text{Err}^2(\hat{\boldsymbol{\mu}}^S, \boldsymbol{\mu}) < \frac{1}{T} \text{tr}(\boldsymbol{\Sigma}) \quad (4.137)$$

$$\hat{\boldsymbol{\mu}}^S \equiv (1 - \alpha) \hat{\boldsymbol{\mu}} + \alpha \mathbf{b}. \quad (4.138)$$

$$\alpha \equiv \frac{1}{T} \frac{N\bar{\lambda} - 2\lambda_1}{(\hat{\boldsymbol{\mu}} - \mathbf{b})'(\hat{\boldsymbol{\mu}} - \mathbf{b})}. \quad (4.139)$$

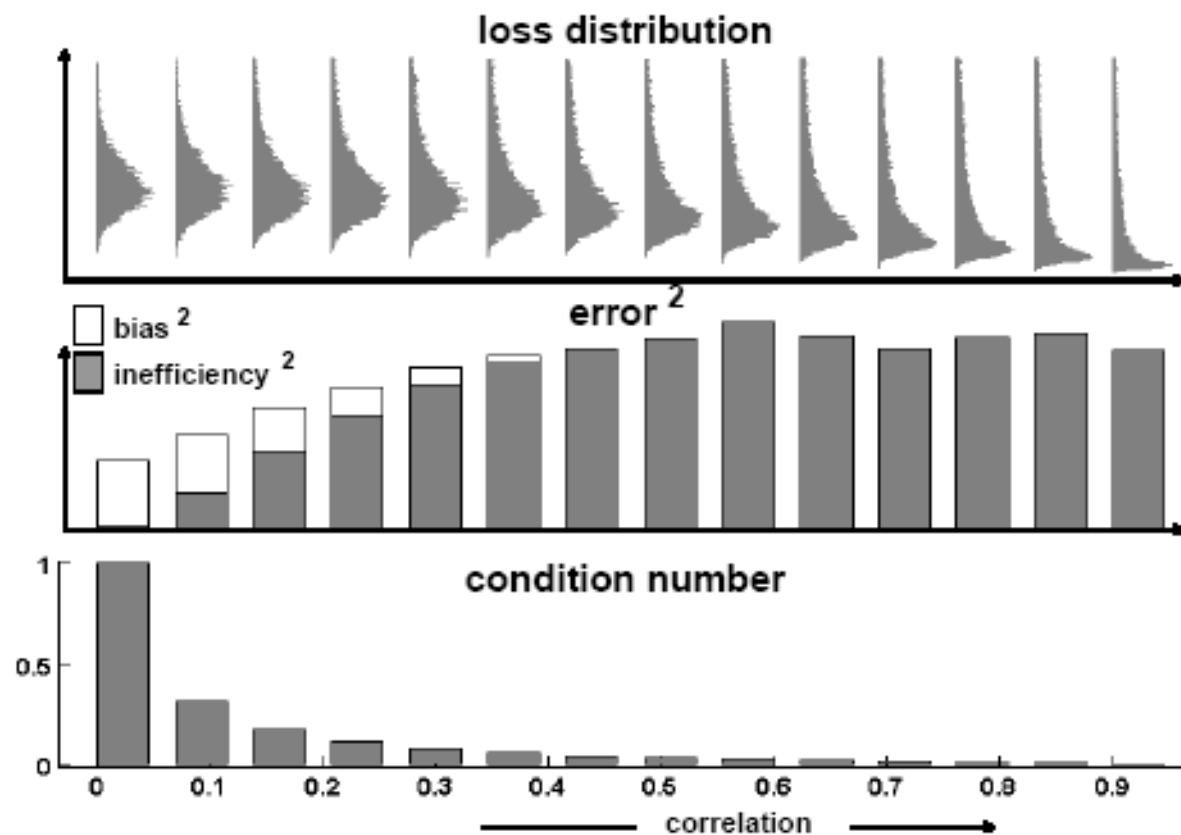


Fig. 4.13. Shrinkage estimator of mean: evaluation

$$\boldsymbol{\Sigma} \mapsto \hat{\boldsymbol{\Sigma}}.$$

$$\mathbf{b} \mapsto \frac{\mathbf{1}' \hat{\boldsymbol{\mu}}}{N} \mathbf{1} \quad (4.141)$$

$$\mathbf{b} \mapsto \frac{\mathbf{1}' \hat{\boldsymbol{\Sigma}}^{-1} \hat{\boldsymbol{\mu}}}{\mathbf{1}' \hat{\boldsymbol{\Sigma}}^{-1} \mathbf{1}} \mathbf{1}. \quad (4.142)$$

SHRINKAGE ESTIMATORS – SCATTER PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.143)$$

$$\hat{\boldsymbol{\Sigma}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]] [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]]' \quad (4.146)$$

$$T\hat{\boldsymbol{\Sigma}}[I_T] \sim W(T-1, \boldsymbol{\Sigma}) \quad (4.103)$$

SHRINKAGE ESTIMATORS – SCATTER PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.143) \qquad \hat{\boldsymbol{\Sigma}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]] [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]]' \quad (4.146)$$

$$\widehat{\text{CN}}\{\mathbf{X}\} \equiv \frac{\hat{\lambda}_N}{\hat{\lambda}_1} < \frac{\lambda_N}{\lambda_1} \equiv \text{CN}\{\mathbf{X}\} \quad (4.156) \qquad T\hat{\boldsymbol{\Sigma}}[I_T] \sim W(T-1, \boldsymbol{\Sigma}) \quad (4.103)$$

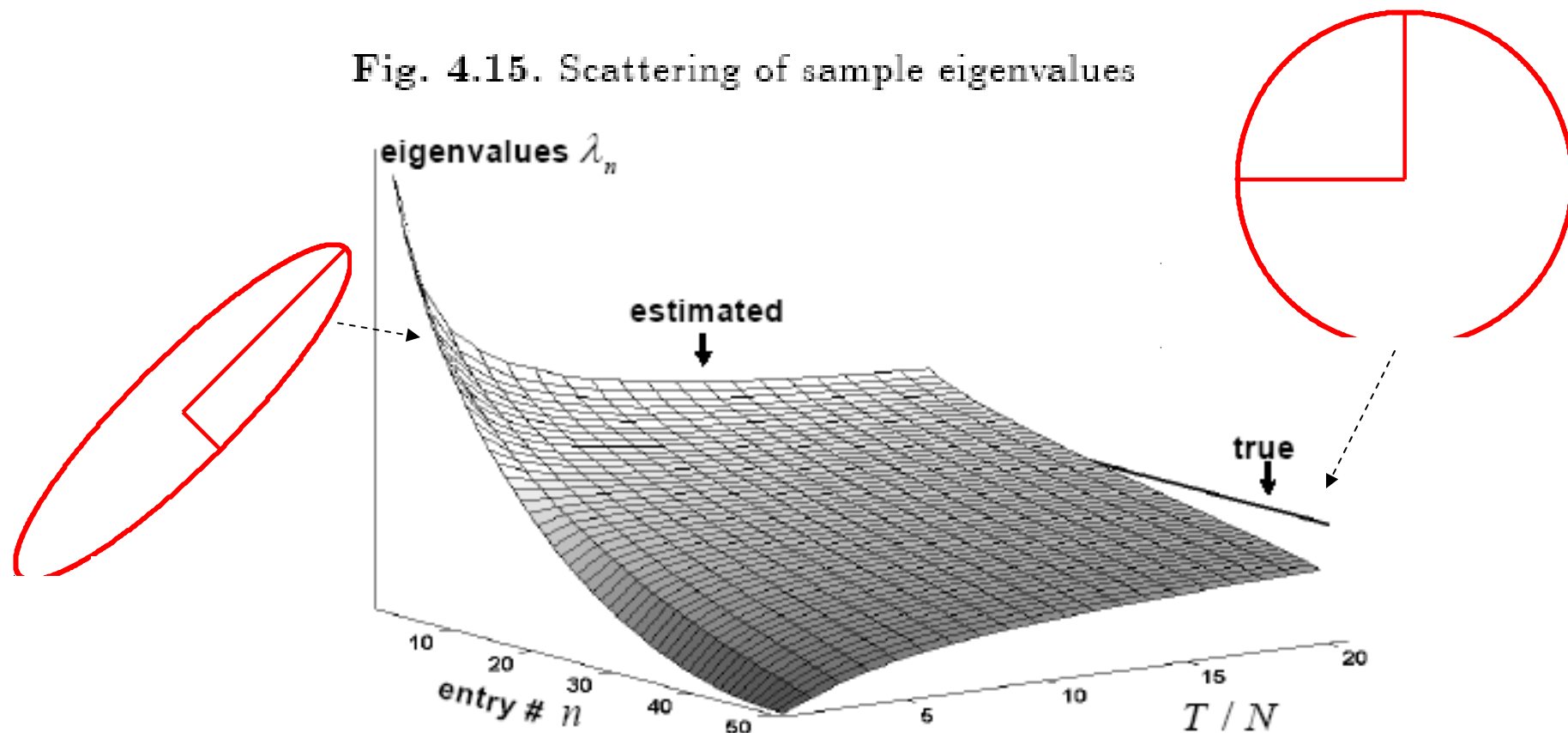
SHRINKAGE ESTIMATORS – SCATTER PARAMETER

Risk and Asset Allocation - Springer – symmys.com

$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.143) \quad \widehat{\boldsymbol{\Sigma}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T [\mathbf{x}_t - \widehat{\boldsymbol{\mu}}[i_T]] [\mathbf{x}_t - \widehat{\boldsymbol{\mu}}[i_T]]' \quad (4.146)$$

$$\widehat{\text{CN}}\{\mathbf{X}\} \equiv \frac{\widehat{\lambda}_N}{\widehat{\lambda}_1} < \frac{\lambda_N}{\lambda_1} \equiv \text{CN}\{\mathbf{X}\} \quad (4.156)$$

Fig. 4.15. Scattering of sample eigenvalues



SHRINKAGE ESTIMATORS – SCATTER PARAMETER

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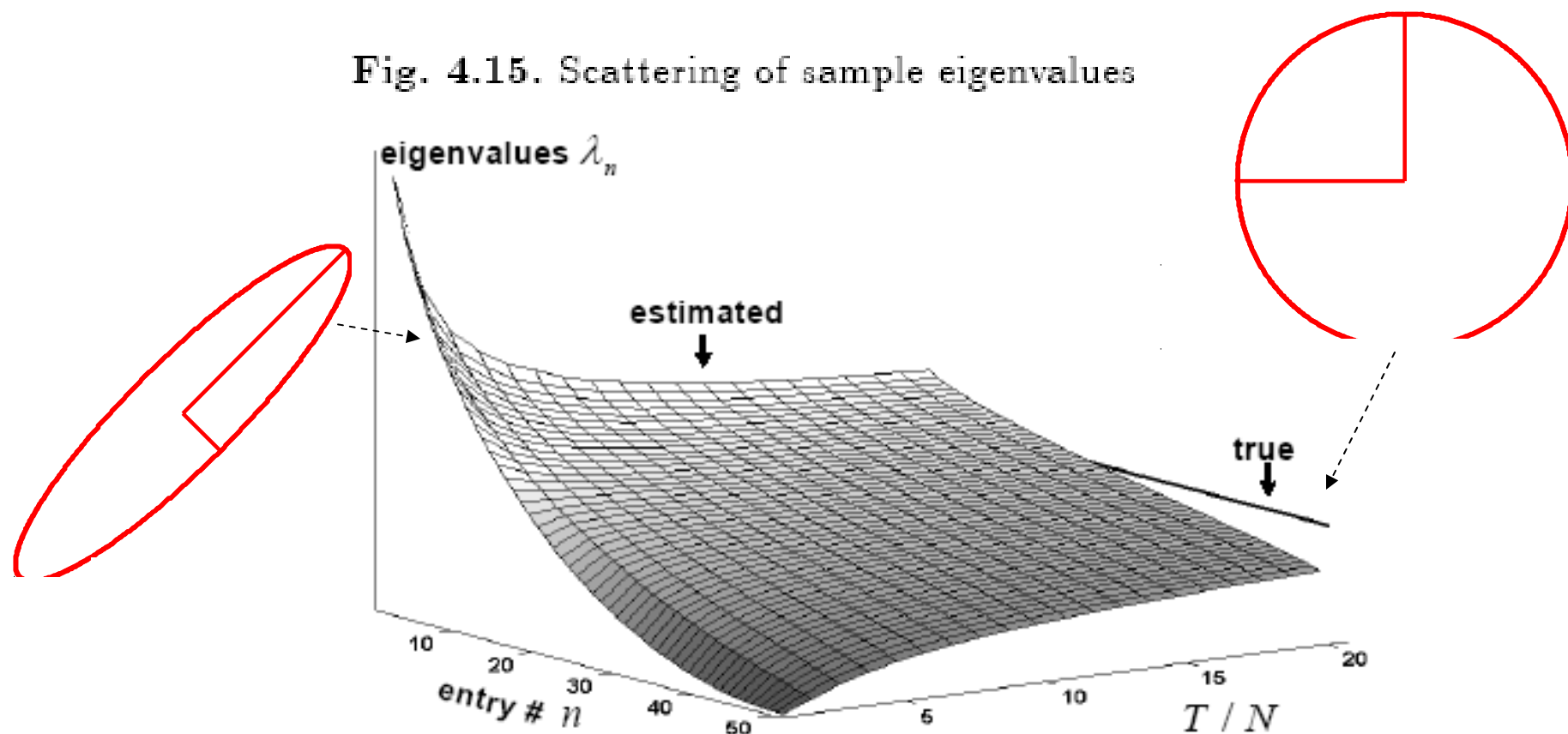
$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.143)$$

$$\widehat{\boldsymbol{\Sigma}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T [\mathbf{x}_t - \widehat{\boldsymbol{\mu}}[i_T]] [\mathbf{x}_t - \widehat{\boldsymbol{\mu}}[i_T]]' \quad (4.146)$$

$$\widehat{\boldsymbol{\Sigma}}^S \equiv (1 - \alpha) \widehat{\boldsymbol{\Sigma}} + \alpha \widehat{\mathbf{C}} \quad (4.160)$$

$$\widehat{\mathbf{C}} \equiv \frac{\sum_{n=1}^N \widehat{\lambda}_n}{N} \mathbf{I} \quad (4.159)$$

Fig. 4.15. Scattering of sample eigenvalues



SHRINKAGE ESTIMATORS – SCATTER PARAMETER

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$$\mathbf{X}_t \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (4.143)$$

$$\hat{\boldsymbol{\Sigma}}[i_T] \equiv \frac{1}{T} \sum_{t=1}^T [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]] [\mathbf{x}_t - \hat{\boldsymbol{\mu}}[i_T]]' \quad (4.146)$$

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$$\hat{\mathbf{C}} \equiv \frac{\sum_{n=1}^N \hat{\lambda}_n}{N} \mathbf{I} \quad (4.159)$$

$$\alpha \equiv \frac{\frac{1}{T} \sum_{t=1}^T \text{tr} \left\{ \left(\mathbf{x}_t \mathbf{x}_t' - \hat{\boldsymbol{\Sigma}} \right)^2 \right\}}{\text{tr} \left\{ \left(\hat{\boldsymbol{\Sigma}} - \hat{\mathbf{C}} \right)^2 \right\}} \quad (4.161)$$

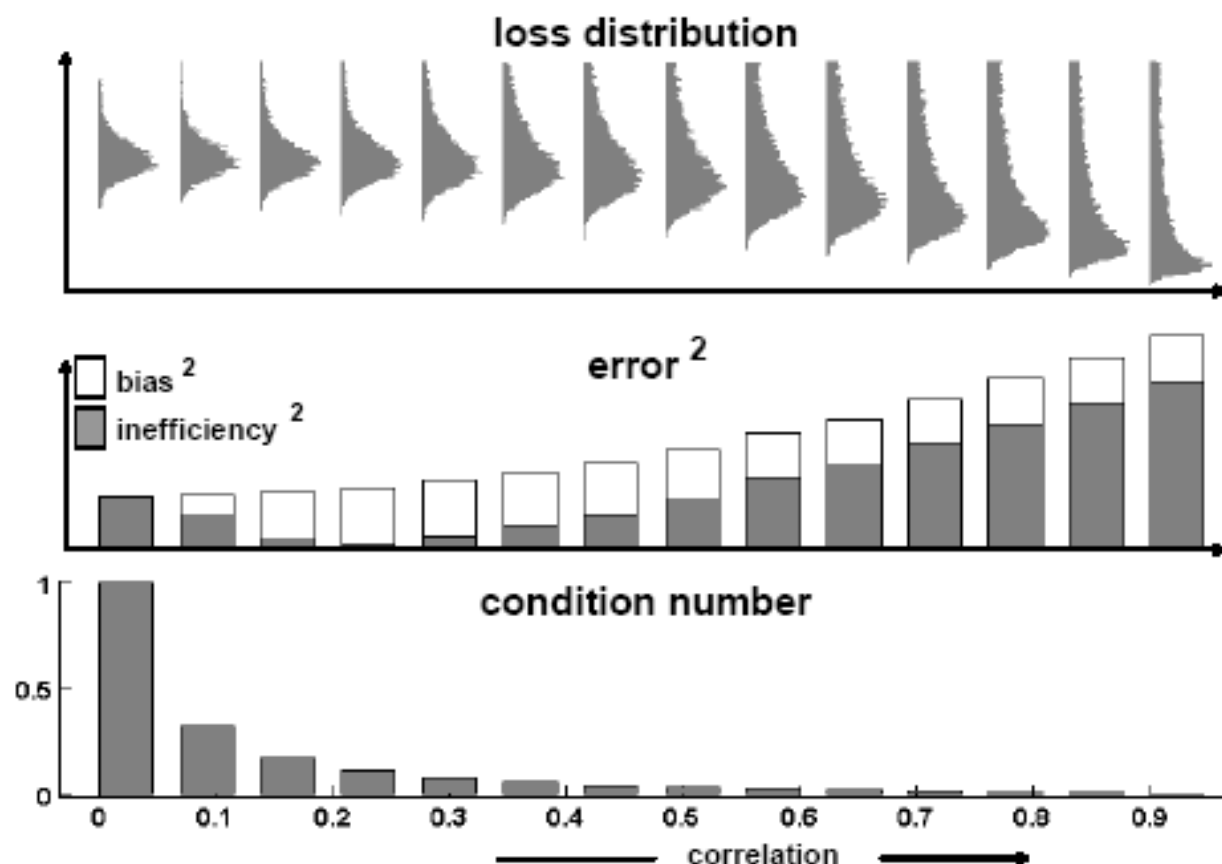


Fig. 4.16. Shrinkage estimator of covariance: evaluation

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Formulas and figures in this presentation refer to the book **Risk and Asset Allocation**, Springer.

The notation, say, (5.24) refers to Formula 24 in Chapter 5 of the book

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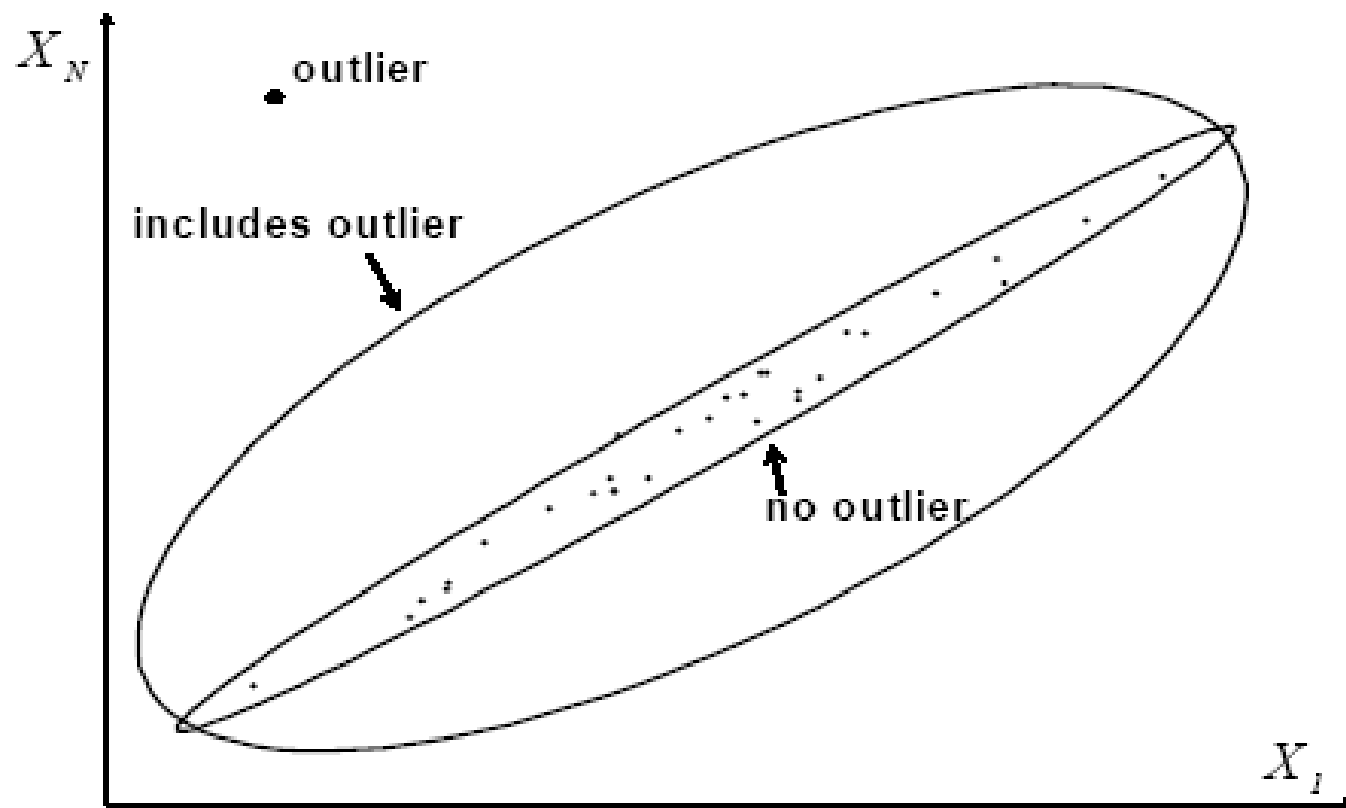


Fig. 4.18. Sample estimators: lack of robustness

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$$SC(x, \hat{G}) \equiv T\hat{G}(x_1, \dots, x_T, x) - T\hat{G}(x_1, \dots, x_T) \quad (4.166)$$

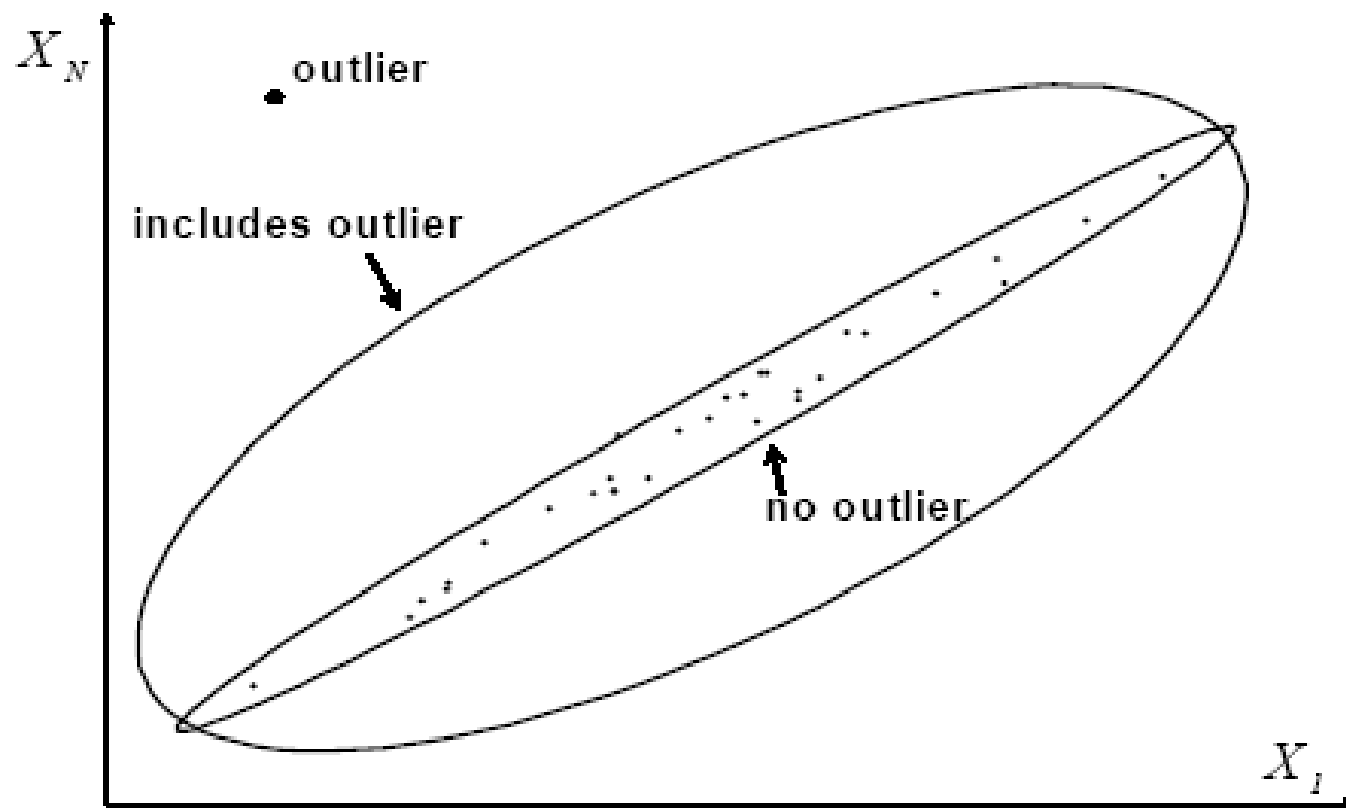


Fig. 4.18. Sample estimators: lack of robustness

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- $$\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv T\hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T, \mathbf{x}) - T\hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T) \quad (4.166)$$

$$f_{i_T} \mapsto (1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \quad (4.183)$$

$$f_{i_T} \equiv \frac{1}{T} \sum_{t=1}^T \delta^{(\mathbf{x}_t)} \quad \epsilon \equiv 1/(T+1)$$

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- $$\boxed{\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv T\hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T, \mathbf{x}) - T\hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T)} \quad (4.166)$$

$$f_{i_T} \mapsto (1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \quad (4.183)$$

$$f_{i_T} \equiv \frac{1}{T} \sum_{t=1}^T \delta^{(\mathbf{x}_t)} \quad \epsilon \equiv 1/(T+1)$$

- $$\boxed{\hat{\mathbf{G}} \equiv \tilde{\mathbf{G}}[f_{i_T}]} \quad (4.167)$$

- non-parametric**

$$\boxed{\hat{\mathbf{G}} \equiv \mathbf{G}[f_{i_T}]} \quad (4.169)$$

- maximum likelihood**

$$\psi(\mathbf{x}, \boldsymbol{\theta}) \equiv \frac{\partial}{\partial \boldsymbol{\theta}} \ln(f_{\boldsymbol{\theta}}(\mathbf{x})) \quad (4.176)$$

$$\tilde{\boldsymbol{\theta}}[h] : \int_{\mathbb{R}^N} \psi(\mathbf{x}, \tilde{\boldsymbol{\theta}}) h(\mathbf{x}) d\mathbf{x} \equiv \mathbf{0}. \quad (4.175)$$

$$\boxed{\hat{\boldsymbol{\theta}} \equiv \tilde{\boldsymbol{\theta}}[f_{i_T}]} \quad (4.177)$$

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- $$\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv T \hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T, \mathbf{x}) - T \hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T) \quad (4.166)$$

$$f_{i_T} \mapsto (1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \quad (4.183)$$

$$f_{i_T} \equiv \frac{1}{T} \sum_{t=1}^T \delta^{(\mathbf{x}_t)} \quad \epsilon \equiv 1/(T+1)$$

- $$\hat{\mathbf{G}} \equiv \tilde{\mathbf{G}}[f_{i_T}] \quad (4.167)$$

$$\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv \frac{1 - \epsilon}{\epsilon} \left\{ \tilde{\mathbf{G}} \left[(1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \right] - \tilde{\mathbf{G}}[f_{i_T}] \right\} \quad (4.184)$$

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- $$\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv T \hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T, \mathbf{x}) - T \hat{\mathbf{G}}(\mathbf{x}_1, \dots, \mathbf{x}_T) \quad (4.166)$$

$$f_{i_T} \mapsto (1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \quad (4.183)$$

$$f_{i_T} \equiv \frac{1}{T} \sum_{t=1}^T \delta^{(\mathbf{x}_t)} \quad \epsilon \equiv 1/(T+1)$$

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$$\text{SC}(\mathbf{x}, \hat{\mathbf{G}}) \equiv \frac{1 - \epsilon}{\epsilon} \left\{ \tilde{\mathbf{G}} \left[(1 - \epsilon) f_{i_T} + \epsilon \delta^{(\mathbf{x})} \right] - \tilde{\mathbf{G}}[f_{i_T}] \right\} \quad (4.184)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \hat{\mathbf{G}}) \equiv \lim_{\epsilon \rightarrow 0} \frac{1}{\epsilon} \left(\tilde{\mathbf{G}} \left[(1 - \epsilon) f_{\mathbf{X}} + \epsilon \delta^{(\mathbf{x})} \right] - \tilde{\mathbf{G}}[f_{\mathbf{X}}] \right) \quad (4.185)$$

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$$\widehat{\mathbf{E}} \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.196)$$

$$\widehat{\mathbf{Cov}} \equiv \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\mathbf{E}}) (\mathbf{x}_t - \widehat{\mathbf{E}})' \quad (4.197)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\mathbf{E}}) = \mathbf{x} - \mathbf{E}\{\mathbf{X}\} \quad (4.198)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\mathbf{Cov}}) = (\mathbf{x} - \mathbf{E}\{\mathbf{X}\}) (\mathbf{x} - \mathbf{E}\{\mathbf{X}\})' - \mathbf{Cov}\{\mathbf{X}\} \quad (4.199)$$

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$$\widehat{\mathbf{E}} \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.196)$$

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$$\widehat{\text{Cov}} \equiv \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\mathbf{E}}) (\mathbf{x}_t - \widehat{\mathbf{E}})' \quad (4.197)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\text{Cov}}) = (\mathbf{x} - \mathbf{E}\{\mathbf{X}\}) (\mathbf{x} - \mathbf{E}\{\mathbf{X}\})' - \text{Cov}\{\mathbf{X}\} \quad (4.199)$$

$$\widehat{\boldsymbol{\mu}} = \sum_{t=1}^T \frac{w\left(\text{Ma}^2\left(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}\right)\right)}{\sum_{s=1}^T w\left(\text{Ma}^2\left(\mathbf{x}_s, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}\right)\right)} \mathbf{x}_t \quad (4.203)$$

$$\widehat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\boldsymbol{\mu}}) (\mathbf{x}_t - \widehat{\boldsymbol{\mu}})' w\left(\text{Ma}^2\left(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}\right)\right) \quad (4.204)$$

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$$\widehat{\mathbf{E}} \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.196)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\mathbf{E}}) = \mathbf{x} - \mathbf{E}\{\mathbf{X}\} \quad (4.198)$$

$$\widehat{\text{Cov}} \equiv \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\mathbf{E}}) (\mathbf{x}_t - \widehat{\mathbf{E}})' \quad (4.197)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\text{Cov}}) = (\mathbf{x} - \mathbf{E}\{\mathbf{X}\}) (\mathbf{x} - \mathbf{E}\{\mathbf{X}\})' - \text{Cov}\{\mathbf{X}\} \quad (4.199)$$

$$\widehat{\boldsymbol{\mu}} = \sum_{t=1}^T \frac{w(\text{Ma}^2(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}))}{\sum_{s=1}^T w(\text{Ma}^2(\mathbf{x}_s, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}))} \mathbf{x}_t \quad (4.203)$$

$$\widehat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\boldsymbol{\mu}}) (\mathbf{x}_t - \widehat{\boldsymbol{\mu}})' w(\text{Ma}^2(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}})) \quad (4.204)$$

$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.205)$$

$$f_{\theta}(\mathbf{x}) \equiv \frac{1}{\sqrt{|\boldsymbol{\Sigma}|}} g(\text{Ma}^2(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma})) \quad (4.201) \quad \leftarrow \text{Ma}^2(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \equiv (\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \quad (4.202)$$

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$$\widehat{\mathbf{E}} \equiv \frac{1}{T} \sum_{t=1}^T \mathbf{x}_t \quad (4.196)$$

$$\text{IF}(\mathbf{x}, f_{\mathbf{X}}, \widehat{\mathbf{E}}) = \mathbf{x} - \mathbf{E}\{\mathbf{X}\} \quad (4.198)$$

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$$\widehat{\boldsymbol{\mu}} = \sum_{t=1}^T \frac{w(\text{Ma}^2(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}))}{\sum_{s=1}^T w(\text{Ma}^2(\mathbf{x}_s, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}))} \mathbf{x}_t \quad (4.203)$$

$$\|\text{IF}(\mathbf{x}, f_{\mathbf{X}}, (\widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}}))\| \propto \|\boldsymbol{\psi}\| \quad (4.208)$$

$$\widehat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \widehat{\boldsymbol{\mu}})(\mathbf{x}_t - \widehat{\boldsymbol{\mu}})' w(\text{Ma}^2(\mathbf{x}_t, \widehat{\boldsymbol{\mu}}, \widehat{\boldsymbol{\Sigma}})) \quad (4.204)$$

$$\boldsymbol{\psi}(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \equiv \begin{pmatrix} w(\text{Ma}^2(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}))(\mathbf{x} - \boldsymbol{\mu}) \\ w(\text{Ma}_{\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}}^2) \text{vec}[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})'] - \text{vec}[\boldsymbol{\Sigma}] \end{pmatrix} \quad (4.207)$$

$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.205)$$

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$$\hat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=1}^T (\mathbf{x}_t - \hat{\boldsymbol{\mu}}) (\mathbf{x}_t - \hat{\boldsymbol{\mu}})' w(\text{Ma}^2(\mathbf{x}_t, \hat{\boldsymbol{\mu}}, \hat{\boldsymbol{\Sigma}})) \quad (4.204)$$

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$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.205)$$

normal

$$w \equiv 1$$

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$$w(z) \equiv -2 \frac{g'(z)}{g(z)} \quad (4.205)$$

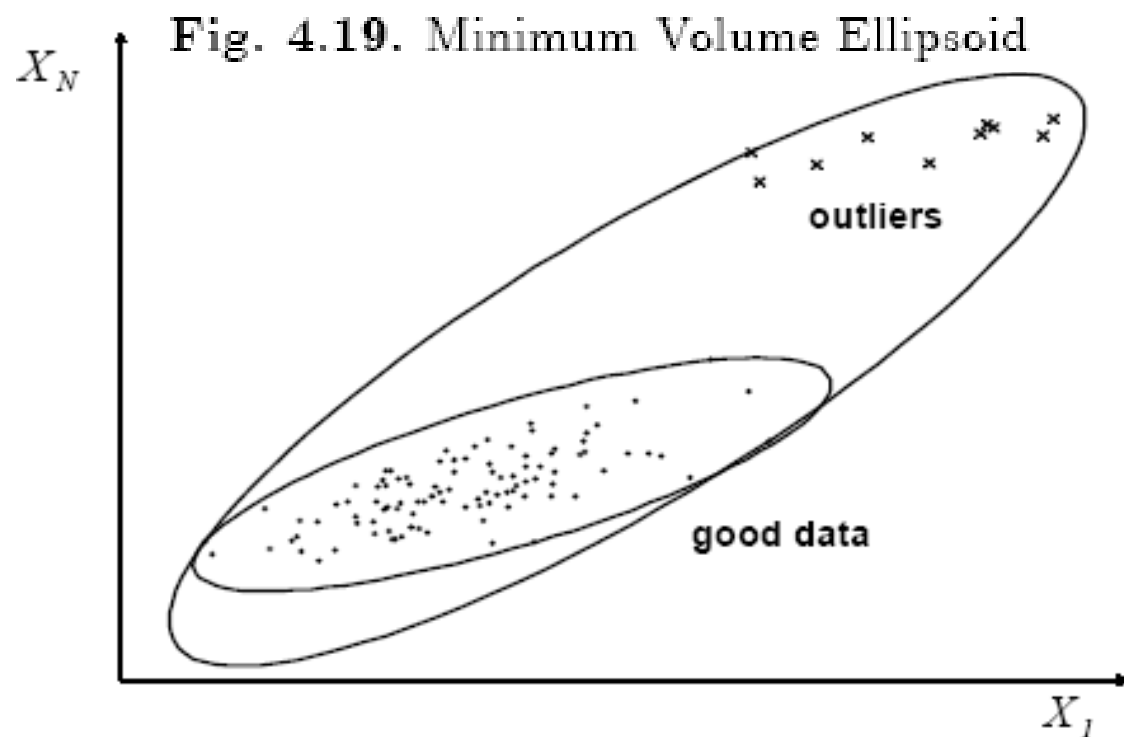
normal
 $w \equiv 1$

$$\text{Cauchy} \\ w(z) = \frac{N+1}{1+z} \quad (4.209)$$

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$$\mathcal{E}_{\mu, \Sigma}^q \equiv \{ \mathbf{x} \in \mathbb{R}^N \text{ such that } (\mathbf{x} - \mu)' \Sigma^{-1} (\mathbf{x} - \mu) \leq q^2 \} \quad (4.231)$$



ROBUST ESTIMATORS – HIGH BREAKDOWN

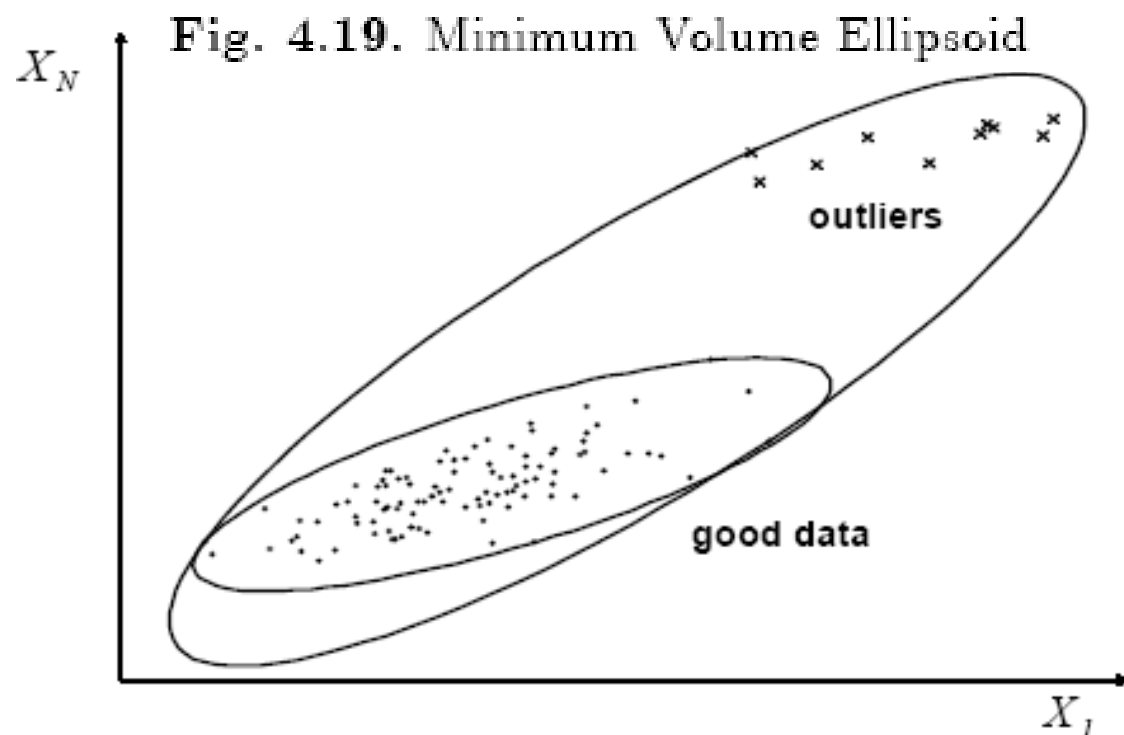
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$$\mathcal{E}_{\mu, \Sigma}^q \equiv \{x \in \mathbb{R}^N \text{ such that } (x - \mu)' \Sigma^{-1} (x - \mu) \leq q^2\} \quad (4.231)$$

$$\text{Ma}_t^{\mu, \Sigma} \equiv \sqrt{(x_t - \mu)' \Sigma^{-1} (x_t - \mu)}. \quad (4.234)$$

$$\text{Vol} \left\{ \mathcal{E}_{\mu, \Sigma}^{q_{T_G}} \right\} = \gamma_N \left(\text{Ma}_{T_G:T}^{\mu, \Sigma} \right)^N \sqrt{|\Sigma|}. \quad (4.236)$$

$$q_{T_G} \equiv \text{Ma}_{T_G:T}^{\mu, \Sigma} \quad (4.235)$$



ROBUST ESTIMATORS – HIGH BREAKDOWN

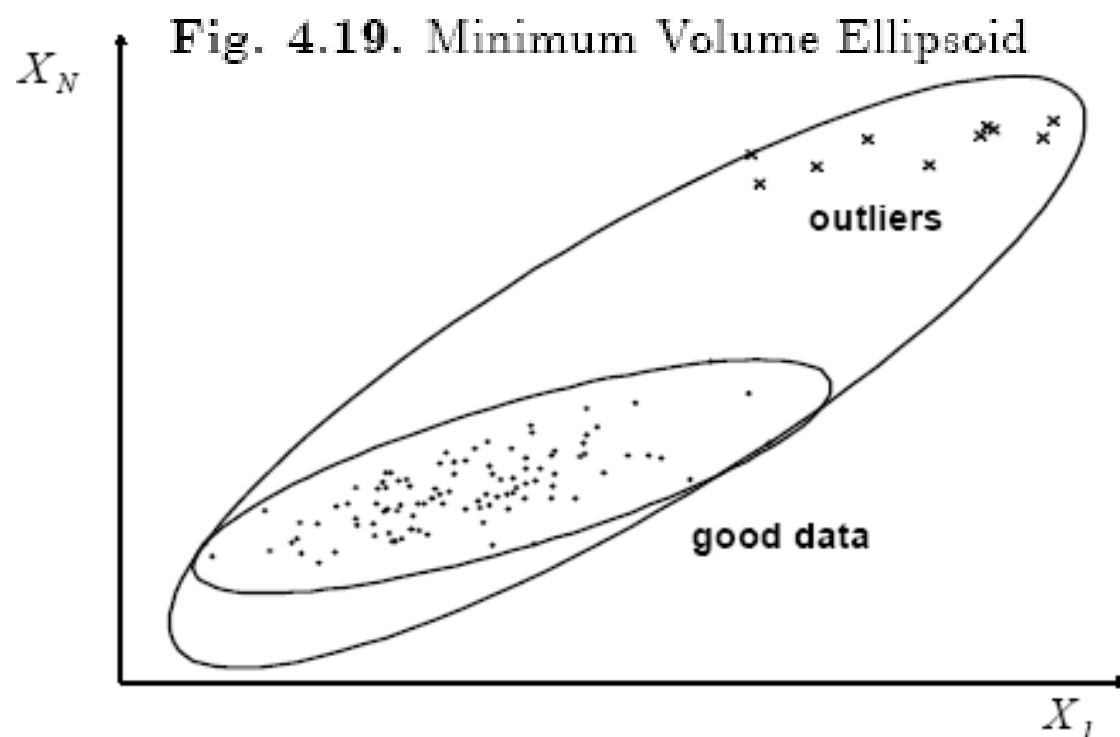
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$$\text{Vol} \left\{ \mathcal{E}_{\mu, \Sigma}^{q_{T_G}} \right\} = \gamma_N \left(\text{Ma}_{T_G:T}^{\mu, \Sigma} \right)^N \sqrt{|\Sigma|}. \quad (4.236)$$

$$q_{T_G} \equiv \text{Ma}_{T_G:T}^{\mu, \Sigma} \quad (4.235)$$



$$\left(\hat{\mu}_{T_G}, \hat{\Sigma}_{T_G} \right) = \underset{\mu, \Sigma \succeq 0, |\Sigma|=1}{\operatorname{argmin}} \left\{ \text{Ma}_{T_G:T}^{\mu, \Sigma} \right\} \quad (4.237)$$

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Formulas and figures in this presentation refer to the book **Risk and Asset Allocation**, Springer.


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$$\mathbf{X} \equiv \begin{pmatrix} \mathbf{X}_A \\ \mathbf{X}_B \end{pmatrix} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (2.160)$$

$$\boldsymbol{\mu} \equiv \begin{pmatrix} \mu_A \\ \mu_B \end{pmatrix}, \quad \boldsymbol{\Sigma} \equiv \begin{pmatrix} \Sigma_{AA} & \Sigma_{AB} \\ \Sigma_{BA} & \Sigma_{BB} \end{pmatrix} \quad (2.161)$$


MISSING DATA – EM ALGORITHM

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$$\mathbf{X} \equiv \begin{pmatrix} \mathbf{X}_A \\ \mathbf{X}_B \end{pmatrix} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \quad (2.160)$$

$$\boldsymbol{\mu} \equiv \begin{pmatrix} \boldsymbol{\mu}_A \\ \boldsymbol{\mu}_B \end{pmatrix}, \quad \boldsymbol{\Sigma} \equiv \begin{pmatrix} \boldsymbol{\Sigma}_{AA} & \boldsymbol{\Sigma}_{AB} \\ \boldsymbol{\Sigma}_{BA} & \boldsymbol{\Sigma}_{BB} \end{pmatrix} \quad (2.161)$$

$$\boldsymbol{\mu}_B | \mathbf{x}_A \equiv \boldsymbol{\mu}_B + \boldsymbol{\Sigma}_{BA} \boldsymbol{\Sigma}_{AA}^{-1} (\mathbf{x}_A - \boldsymbol{\mu}_A) \quad (2.165)$$

$$\mathbf{X}_B | \mathbf{x}_A \sim N(\boldsymbol{\mu}_B | \mathbf{x}_A, \boldsymbol{\Sigma}_B | \mathbf{x}_A) \quad (2.164)$$

$$\boldsymbol{\Sigma}_B | \mathbf{x}_A \equiv \boldsymbol{\Sigma}_{BB} - \boldsymbol{\Sigma}_{BA} \boldsymbol{\Sigma}_{AA}^{-1} \boldsymbol{\Sigma}_{AB} \quad (2.166)$$

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$$\begin{pmatrix} X_{t,\text{mis}(t)} \\ X_{t,\text{obs}(t)} \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_{\text{mis}(t)} \\ \mu_{\text{obs}(t)} \end{pmatrix}, \begin{pmatrix} \Sigma_{\text{mis}(t),\text{mis}(t)} & \Sigma_{\text{mis}(t),\text{obs}(t)} \\ \Sigma_{\text{obs}(t),\text{mis}(t)} & \Sigma_{\text{obs}(t),\text{obs}(t)} \end{pmatrix} \right) \quad (4.257)$$

$$\mu_B|x_A \equiv \mu_B + \Sigma_{BA}\Sigma_{AA}^{-1}(x_A - \mu_A) \quad (2.165)$$

$$\Sigma_B|x_A \equiv \Sigma_{BB} - \Sigma_{BA}\Sigma_{AA}^{-1}\Sigma_{AB} \quad (2.166)$$

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$$\begin{pmatrix} X_{t,\text{mis}(t)} \\ X_{t,\text{obs}(t)} \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_{\text{mis}(t)} \\ \mu_{\text{obs}(t)} \end{pmatrix}, \begin{pmatrix} \Sigma_{\text{mis}(t),\text{mis}(t)} & \Sigma_{\text{mis}(t),\text{obs}(t)} \\ \Sigma_{\text{obs}(t),\text{mis}(t)} & \Sigma_{\text{obs}(t),\text{obs}(t)} \end{pmatrix} \right) \quad (4.257)$$

$$x_{t,\text{obs}(t)}^{(u)} \equiv x_{t,\text{obs}(t)} \quad (4.261)$$

$$x_{t,\text{mis}(t)}^{(u)} \equiv \mu_{\text{mis}(t)}^{(u)} \quad (4.262)$$

$$\mu_B | x_A \equiv \mu_B + \Sigma_{BA} \Sigma_{AA}^{-1} (x_A - \mu_A) \quad (2.165)$$

$$+ \Sigma_{\text{mis}(t),\text{obs}(t)}^{(u)} \left(\Sigma_{\text{obs}(t),\text{obs}(t)}^{(u)} \right)^{-1} \left(x_{t,\text{obs}(t)} - \mu_{\text{obs}(t)}^{(u)} \right)$$

$$\Sigma_B | x_A \equiv \Sigma_{BB} - \Sigma_{BA} \Sigma_{AA}^{-1} \Sigma_{AB} \quad (2.166)$$

MISSING DATA – EM ALGORITHM

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$$\begin{pmatrix} \mathbf{X}_{t,\text{mis}(t)} \\ \mathbf{X}_{t,\text{obs}(t)} \end{pmatrix} \sim N \left(\begin{pmatrix} \boldsymbol{\mu}_{\text{mis}(t)} \\ \boldsymbol{\mu}_{\text{obs}(t)} \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{\text{mis}(t),\text{mis}(t)} & \boldsymbol{\Sigma}_{\text{mis}(t),\text{obs}(t)} \\ \boldsymbol{\Sigma}_{\text{obs}(t),\text{mis}(t)} & \boldsymbol{\Sigma}_{\text{obs}(t),\text{obs}(t)} \end{pmatrix} \right) \quad (4.257)$$

$$\mathbf{x}_{t,\text{obs}(t)}^{(u)} \equiv \mathbf{x}_{t,\text{obs}(t)} \quad (4.261)$$

$$\boldsymbol{\mu}_B | \mathbf{x}_A \equiv \boldsymbol{\mu}_B + \boldsymbol{\Sigma}_{BA} \boldsymbol{\Sigma}_{AA}^{-1} (\mathbf{x}_A - \boldsymbol{\mu}_A) \quad (2.165)$$

$$\mathbf{x}_{t,\text{mis}(t)}^{(u)} \equiv \boldsymbol{\mu}_{\text{mis}(t)}^{(u)} \quad (4.262)$$

$$+ \boldsymbol{\Sigma}_{\text{mis}(t),\text{obs}(t)}^{(u)} \left(\boldsymbol{\Sigma}_{\text{obs}(t),\text{obs}(t)}^{(u)} \right)^{-1} \left(\mathbf{x}_{t,\text{obs}(t)} - \boldsymbol{\mu}_{\text{obs}(t)}^{(u)} \right)$$

$$\mathbf{C}_{t,\text{obs}(t),\text{mis}(t)}^{(u)} \equiv \mathbf{0}, \quad \mathbf{C}_{t,\text{obs}(t),\text{obs}(t)}^{(u)} \equiv \mathbf{0}, \quad (4.263)$$

$$\mathbf{C}_{t,\text{mis}(t),\text{mis}(t)}^{(u)} \equiv \boldsymbol{\Sigma}_{\text{mis}(t),\text{mis}(t)}^{(u)} \quad (4.264)$$

$$\boldsymbol{\Sigma}_B | \mathbf{x}_A \equiv \boldsymbol{\Sigma}_{BB} - \boldsymbol{\Sigma}_{BA} \boldsymbol{\Sigma}_{AA}^{-1} \boldsymbol{\Sigma}_{AB} \quad (2.166)$$

$$- \boldsymbol{\Sigma}_{\text{mis}(t),\text{obs}(t)}^{(u)} \left(\boldsymbol{\Sigma}_{\text{obs}(t),\text{obs}(t)}^{(u)} \right)^{-1} \boldsymbol{\Sigma}_{\text{obs}(t),\text{mis}(t)}^{(u)}.$$

estimate

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$$\mathbf{x}_{t,\text{obs}(t)}^{(u)} \equiv \mathbf{x}_{t,\text{obs}(t)} \quad (4.261)$$

$$\boldsymbol{\mu}_B | \mathbf{x}_A \equiv \boldsymbol{\mu}_B + \boldsymbol{\Sigma}_{BA} \boldsymbol{\Sigma}_{AA}^{-1} (\mathbf{x}_A - \boldsymbol{\mu}_A) \quad (2.165)$$

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$$+ \boldsymbol{\Sigma}_{\text{mis}(t),\text{obs}(t)}^{(u)} \left(\boldsymbol{\Sigma}_{\text{obs}(t),\text{obs}(t)}^{(u)} \right)^{-1} \left(\mathbf{x}_{t,\text{obs}(t)} - \boldsymbol{\mu}_{\text{obs}(t)}^{(u)} \right)$$

$$\mathbf{C}_{t,\text{obs}(t),\text{mis}(t)}^{(u)} \equiv \mathbf{0}, \quad \mathbf{C}_{t,\text{obs}(t),\text{obs}(t)}^{(u)} \equiv \mathbf{0}, \quad (4.263)$$

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$$- \boldsymbol{\Sigma}_{\text{mis}(t),\text{obs}(t)}^{(u)} \left(\boldsymbol{\Sigma}_{\text{obs}(t),\text{obs}(t)}^{(u)} \right)^{-1} \boldsymbol{\Sigma}_{\text{obs}(t),\text{mis}(t)}^{(u)}.$$

estimate

$$\boldsymbol{\mu}^{(u+1)} \equiv \frac{1}{T} \sum_t \mathbf{x}_t^{(u)} \quad (4.265)$$

$$\boldsymbol{\Sigma}^{(u+1)} \equiv \frac{1}{T} \sum_t \left[\mathbf{C}_t^{(u)} + \left(\mathbf{x}_t^{(u)} - \boldsymbol{\mu}^{(u)} \right) \left(\mathbf{x}_t^{(u)} - \boldsymbol{\mu}^{(u)} \right)' \right] \quad (4.266)$$

update