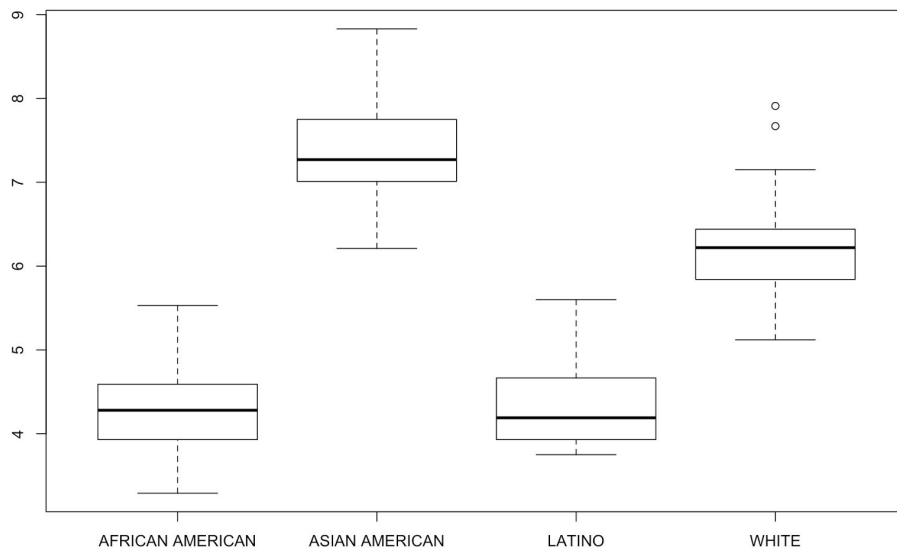


## # Mid Term 2

```
install.packages("readxl")
setwd("/opt/Code/My/R-Tests/regression/mt2")
demoData <- readxl::read_xlsx("MOADData.xlsx")
colnames(demoData) <- c("race", "year", "area", "hdIndex",
"lifeExpAtBirth", "gradDeg", "schEnroll", "medianEarn", "healthIndex", "eduIndex", "incomeIndex")
par(mfrow=c(1,1))
```

### a) Compare HD index for the other races

```
boxplot(demoData$hdIndex~demoData$race)
```



**Comments:** By looking at the plot it seems that mean HD index is over all higher for asian american people. It is relatively lower for african american and latino people. Where mean HD Index for white people is higher than Latino and African american but lower than Asian American people. There are a few outlier in the higher side of HD index for white people.

### b) Predictor - Race & median income, additive regression model for response HD Index

```
demoHdLm1<-lm(hdIndex~factor(race)+medianEarn,demoData)
summary(demoHdLm1)
```

Call:

```
lm(formula = hdIndex ~ factor(race) + medianEarn, data = demoData)
```

Residuals:

```
    Min     1Q  Median     3Q     Max
-0.7308 -0.2324  0.0073  0.1906  0.8993
```

Coefficients:

```
              Estimate Std. Error t value Pr(>|t|)
(Intercept)    1.20537586  0.22447945   5.37 0.00000066 ***
factor(race)ASIAN AMERICAN  2.25357499  0.11743121  19.19 < 2e-16 ***
factor(race)LATINO        0.68021373  0.10995107   6.19 0.00000002 ***
```

```
factor(race)WHITE      0.70494933 0.12802229  5.51 0.00000037 ***
medianEarn             0.00011090 0.00000772 14.36 < 2e-16 ***
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.34 on 86 degrees of freedom

Multiple R-squared: 0.949, Adjusted R-squared: 0.946

F-statistic: 398 on 4 and 86 DF, p-value: <2e-16

#### Comments:

The model is  $Y = B_0 + B_1 X_1 + B_2 X_2 + B_3 X_3 + B_4 X_4$

Here

$B_0 = 1.205$  (Intercept)

$B_1 = 2.254$ ;  $X_1 = 1$  (race = Asian American), 0 (Otherwise)

$B_2 = 0.68$ ;  $X_2 = 1$  (Latino), 0 (Otherwise)

$B_3 = 0.704$ ;  $X_3 = 1$  (White), 0 (Otherwise)

$B_4 = 0.00011$ ;  $X_4 \Rightarrow$  median earn

R-squared value is large enough to validate the significance of the model. Also p-value being very small confirms the validity of the coefficients.

F-statistic is slightly high so we can try including interaction terms.

c) Full model with all possible interaction terms

```
demoHdLm2<-lm(hdIndex~factor(race)+medianEarn + factor(race)*medianEarn,demoData)
summary(demoHdLm2)
```

Call:

```
lm(formula = hdIndex ~ factor(race) + medianEarn + factor(race) *
    medianEarn, data = demoData)
```

Residuals:

```
    Min      1Q  Median      3Q     Max
-0.7198 -0.2167 -0.0191  0.1733  0.8536
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	1.38075138	0.45994217	3.00	0.0035
factor(race)ASIAN AMERICAN	1.98979746	0.67911040	2.93	0.0044
factor(race)LATINO	-2.53625134	1.27979430	-1.98	0.0508
factor(race)WHITE	0.70284335	0.63238786	1.11	0.2696
medianEarn	0.00010457	0.00001641	6.37	0.0000000099
<b>factor(race)ASIAN AMERICAN:medianEarn</b>	<b>0.00000879</b>	<b>0.00002143</b>	<b>0.41</b>	<b>0.6829</b>
factor(race)LATINO:medianEarn	0.00014448	0.00005659	2.55	0.0125
<b>factor(race)WHITE:medianEarn</b>	<b>0.00000186</b>	<b>0.00001977</b>	<b>0.09</b>	<b>0.9251</b>

```
(Intercept)          **
factor(race)ASIAN AMERICAN      **
factor(race)LATINO          .
factor(race)WHITE
```

```

medianEarn          ***
factor(race)ASIAN AMERICAN:medianEarn
factor(race)LATINO:medianEarn      *
factor(race)WHITE:medianEarn

```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.328 on 83 degrees of freedom  
Multiple R-squared: 0.953, Adjusted R-squared: 0.949  
F-statistic: 239 on 7 and 83 DF, p-value: <0.0000000000000002

### Comments:

After adding the interaction terms, F-Statistic value has decreased. R-squared is high enough to validate the model.

Full model :  $Y = B_0 + B_1 X_1 + B_2 X_2 + B_3 X_3 + B_4 X_4 + B_5 X_1 X_4 + B_6 X_2 X_4 + B_7 X_3 X_4$

$B_0 = 1.381$  (Intercept)

$B_1 = 1.99$ ;  $X_1 = 1$  (race = Asian American), 0 (Otherwise)

$B_2 = -2.54$ ;  $X_2 = 1$  (Latino), 0 (Otherwise);

$B_3 = 0.70$ ;  $X_3 = 1$  (White), 0 (Otherwise);

$B_4 = 0.00010$ ;  $X_4 \Rightarrow$  median earn

$B_5 = 0.00000879$  (Interaction of Asian American race with Median Income); **p-value = 0.68**

$B_6 = 0.00014448$  (Interaction of Latino race with Median Income)

$B_7 = 0.00000186$  (Interaction of White race with Median Income); **p-value= 0.9251**

By looking at p values for  $B_5$  and  $B_7$ , we should remove these terms as p-value is higher than 0.5 but since the coefficients are very small, we can ignore these.

### d) Visual displays for both additive and full model

```
> anova(demoHdLm1, demoHdLm2)
```

Analysis of Variance Table

Model 1:  $hdIndex \sim \text{factor}(\text{race}) + \text{medianEarn}$

Model 2:  $hdIndex \sim \text{factor}(\text{race}) + \text{medianEarn} + \text{factor}(\text{race}) * \text{medianEarn}$

```
Res.Df RSS Df Sum of Sq  F Pr(>F)
```

```
1    86 9.68
```

```
2    83 8.94  3    0.738 2.28 0.085 .
```

---

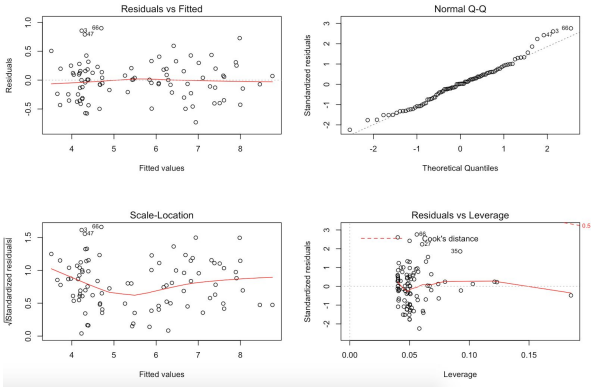
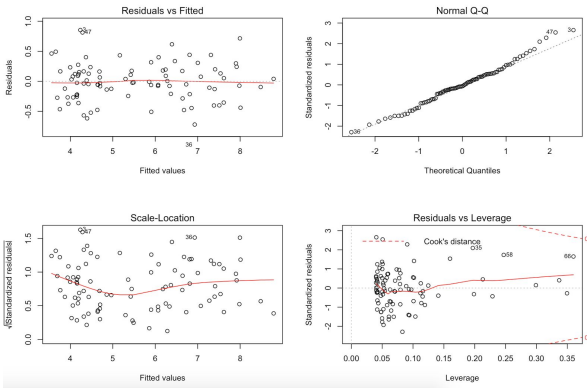
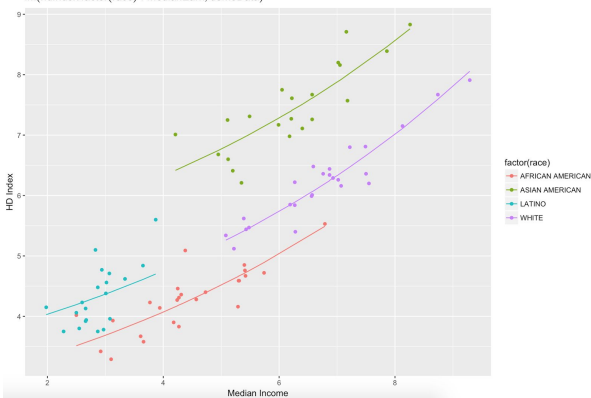
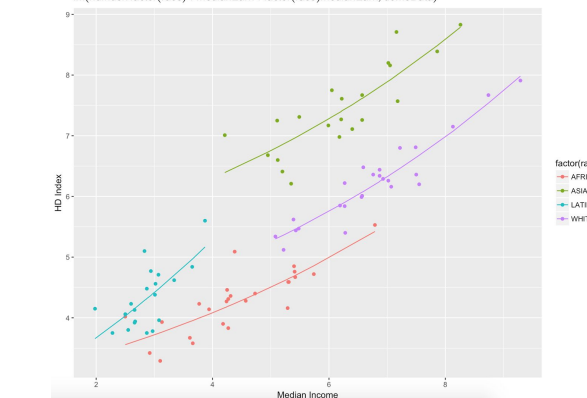
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
par(mfrow=c(2,2))
```

```
install.packages("ggplot2")
```

```
library(ggplot2)
```

$Y = B_0 + B_1 X_1 + B_2 X_2 + B_3 X_3 + B_4 X_4$	$Y = B_0 + B_1 X_1 + B_2 X_2 + B_3 X_3 + B_4 X_4 + B_5 X_1 X_4 + B_6 X_2 X_4 + B_7 X_3 X_4$
---	---

<code>plot(demoHdLm1)</code>	<code>plot(demoHdLm2)</code>
	
	<ul style="list-style-type: none"> <li>• Residual vs fitted plot is better with this model</li> <li>• Slightly better r-squared</li> </ul>
<pre>ggplot(data=demoData, aes(x=incomeIndex, y=hdIndex, colour=factor(race))) + geom_point() + xlab("Median Income") + ylab("HD Index") + geom_line(aes(y=demoHdLm1\$fitted.values)) + ggtitle(summary(demoHdLm1)\$call)</pre>	<pre>ggplot(data=demoData, aes(x=incomeIndex, y=hdIndex, colour=factor(race))) + geom_point() + xlab("Median Income") + ylab("HD Index") + geom_line(aes(y=demoHdLm2\$fitted.values)) + ggtitle(summary(demoHdLm2)\$call)</pre>
	
	<p>As we can see, this model is improved as the observations (for African american and Latino) are closer to regression line.</p>

**e) Include life expectancy as predictor to build additive model**

```
demoHdLm3 <- lm(hdIndex ~ factor(race) + medianEarn + lifeExpAtBirth, demoData)
summary(demoHdLm3)
```

Call:

```
lm(formula = hdIndex ~ factor(race) + medianEarn + lifeExpAtBirth,
    data = demoData)
```

Residuals:

```
    Min     1Q  Median     3Q      Max
-0.2847 -0.0931 -0.0117  0.0912  0.4008
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-10.2242078	0.6101484	-16.76	< 0.00000000000000002
factor(race)ASIAN AMERICAN	0.5251107	0.1046638	5.02	0.000002840516
factor(race)LATINO	-0.6690835	0.0859496	-7.78	0.0000000000015
factor(race)WHITE	0.2202769	0.0617825	3.57	0.0006
medianEarn	0.0000944	0.0000035	26.95	< 0.00000000000000002
lifeExpAtBirth	0.1574541	0.0082948	18.98	< 0.00000000000000002

```
(Intercept)          ***
factor(race)ASIAN AMERICAN ***
factor(race)LATINO    ***
factor(race)WHITE     ***
medianEarn            ***
lifeExpAtBirth        ***
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.147 on 85 degrees of freedom

Multiple R-squared: 0.99, Adjusted R-squared: 0.99

F-statistic: 1.72e+03 on 5 and 85 DF, p-value: <0.00000000000000002

**Comments:**

$Y = B_0 + B_1 X_1 + B_2 X_2 + B_3 X_3 + B_4 X_4 + B_5 X_5$

Where

$B_0 = -10.2242078$  (Intercept)

$B_1 = 0.5251107$ ;  $X_1 = 1$  (race = Asian American), 0 (Otherwise)

$B_2 = -0.6690835$ ;  $X_2 = 1$  (Latino), 0 (Otherwise)

$B_3 = 0.2202769$ ;  $X_3 = 1$  (White), 0 (Otherwise)

$B_4 = 0.0000944$ ;  $X_4 \Rightarrow$  median income

$B_5 = 0.1574541$ ;  $X_5 \Rightarrow$  lifeExpAtBirth

R-squared value is large enough (and has increased to 0.99 in comparison with other models) to validate the significance of the model. Also p-value being very small confirms the significance of the coefficients.

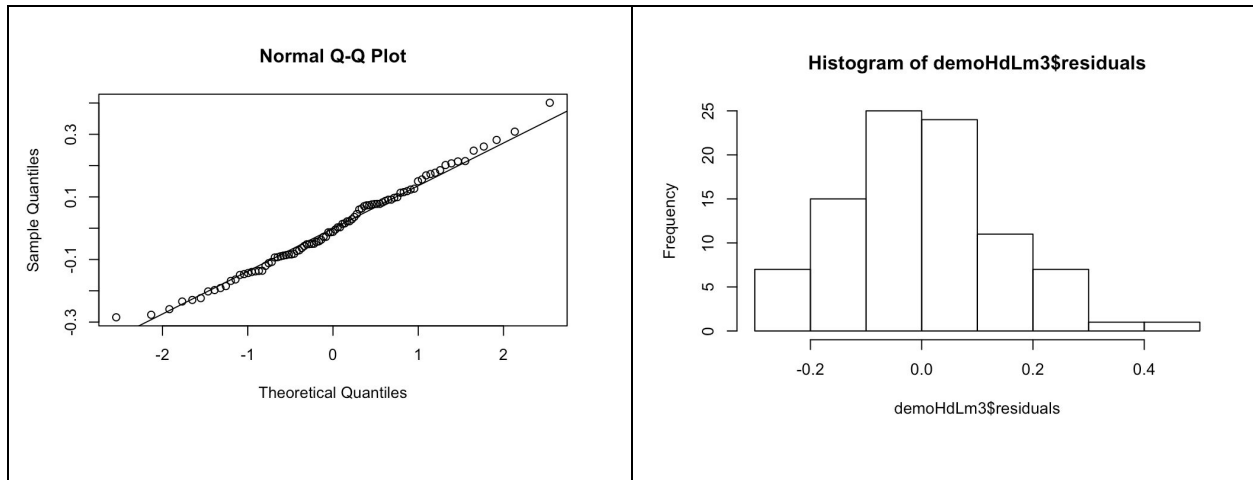
F-statistic is high so we can try including interaction terms.

**Residual plots for normality and formal tests:**

[plot\(demoHdLm3\)](#)

[qqnorm\(demoHdLm3\\$residuals\)](#)

```
qqline(demoHdLm3$residuals)
hist(demoHdLm3$residuals)
```



Residual distribution looks Normal.

Shapiro-Wilk normality test

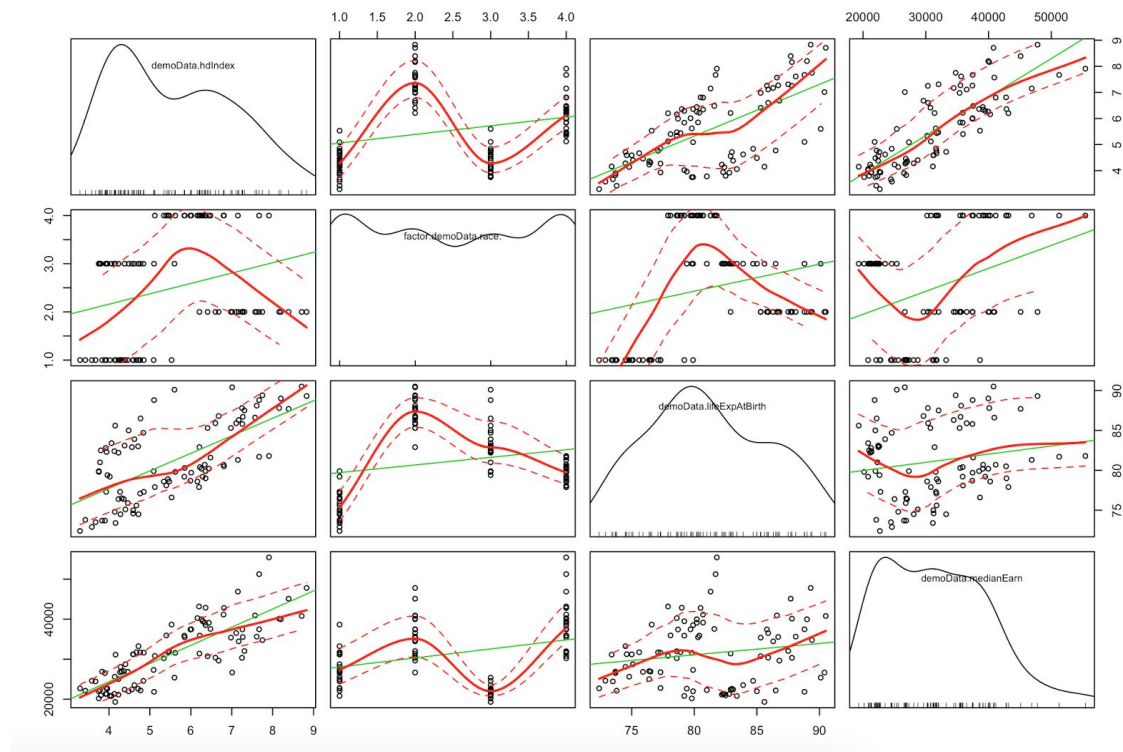
```
data: demoHdLm3$residuals
W = 1, p-value = 0.7
```

### Scatter and marginal model plots

```
install.packages("car")
```

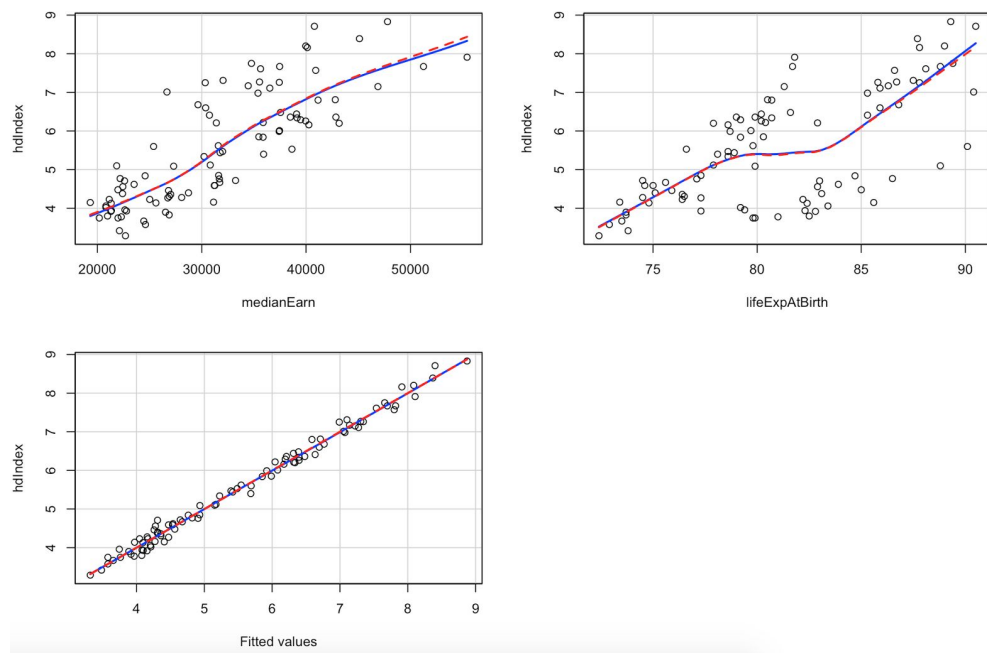
```
library(car)
```

```
scatterplotMatrix(~demoData$hdIndex+factor(demoData$race)+demoData$lifeExpAtBirth+demoData$incomeIndex)
```



`marginalModelPlots(demoHdLm3)`

Marginal Model Plots



By looking at the fitted vs observed value plot, it seems that the model fits great.

**Constancy of Variance:**

`> ncvTest(demoHdLm3)`

Non-constant Variance Score Test

Variance formula:  $\sim$  fitted.values

Chisquare = 0.0001 Df = 1 p = 0.992

**High p value confirms the constancy in variance.**

**f) Checking outliers in data**

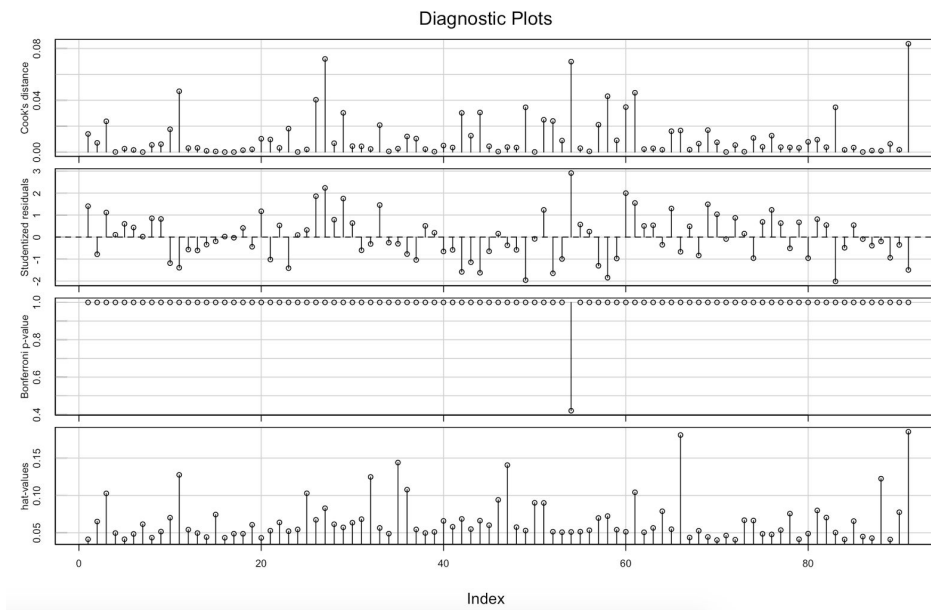
`outlierTest(demoHdLm3)`

No Studentized residuals with Bonferonni  $p < 0.05$

Largest  $|rstudent|$ :

	rstudent	unadjusted p-value	Bonferonni p
54	2.91	0.00461	0.42

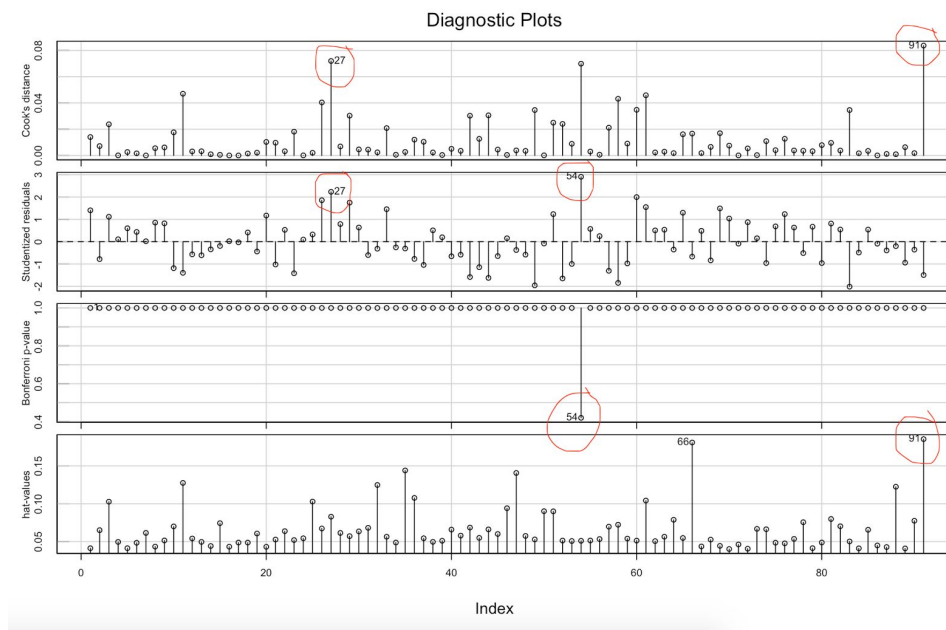
`influenceIndexPlot(demoHdLm3)`



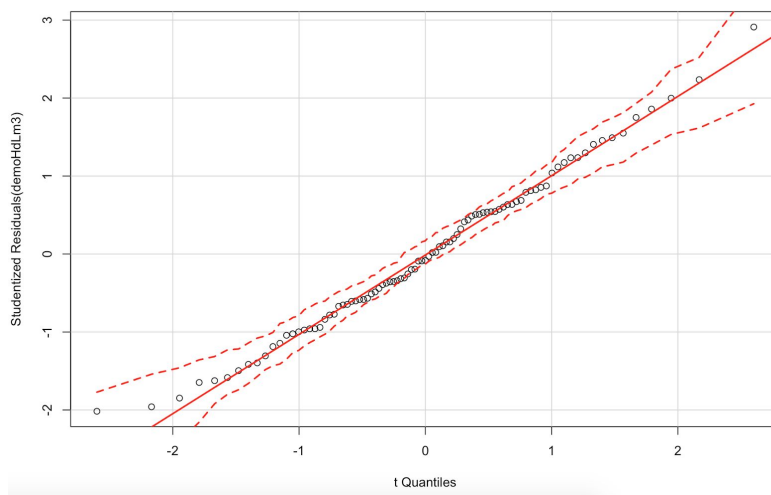
Finding top two in each

`> influenceIndexPlot(demoHdLm3, id.n=2)`



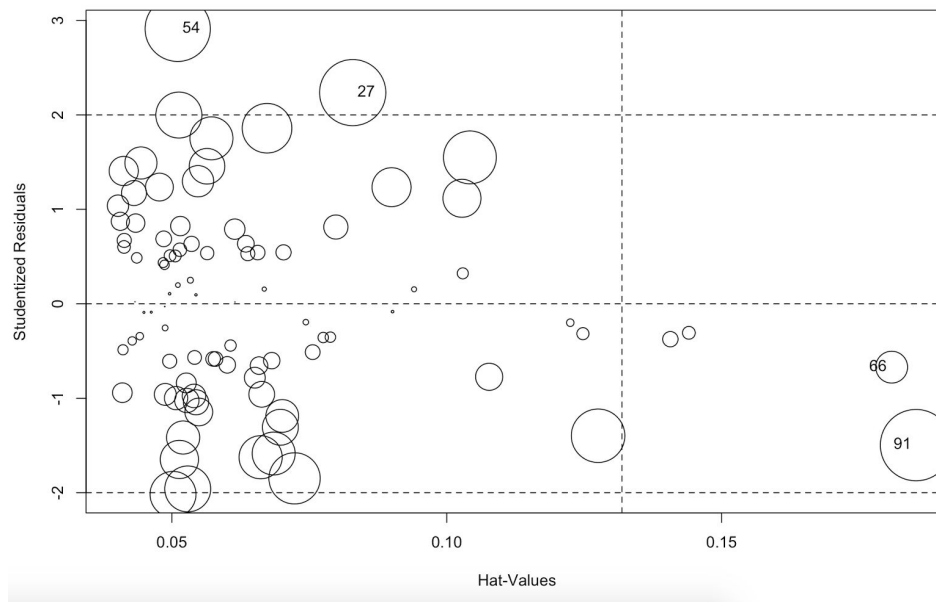


`qqPlot(demoHdLm3)`



QQ-Plot for studentized residuals.

**Influence Plot for Cook's distance confirms the outliers to 27, 54, 66, 91.**  
 Two points with largest influence are 54 and 91.



#### g) Using AIC to justify the choice of model with 3 predictors

# Step function

```
nullModel <- lm(hdIndex~1,demoData)
```

```
summary(nullModel)
```

Call:

```
lm(formula = hdIndex ~ 1, data = demoData)
```

Residuals:

Min	1Q	Median	3Q	Max
-2.262	-1.277	-0.152	1.088	3.278

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	5.552	0.152	36.5	<0.0000000000000002 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.45 on 90 degrees of freedom

```
step(nullModel, scope=list(lower=nullModel, upper=demoHdLm3), direction="forward")
```

Start: AIC=68.5

```
hdIndex ~ 1
```

	Df	Sum of Sq	RSS	AIC
+ factor(race)	3	156.1	32.9	-84.6
+ medianEarn	1	126.7	62.3	-30.5
+ lifeExpAtBirth	1	82.3	106.7	18.4
<none>			189.0	68.5

Step: AIC=-84.6

hdIndex ~ factor(race)

	Df	Sum of Sq	RSS	AIC
+ medianEarn	1	23.2	9.7	-193.9
+ lifeExpAtBirth	1	15.3	17.6	-139.3
<none>		32.9		-84.6

Step: AIC=-194

hdIndex ~ factor(race) + medianEarn

	Df	Sum of Sq	RSS	AIC
+ lifeExpAtBirth	1	7.83	1.85	-343
<none>		9.68		-194

Step: AIC=-343

hdIndex ~ factor(race) + medianEarn + lifeExpAtBirth

Call:

lm(formula = hdIndex ~ factor(race) + medianEarn + lifeExpAtBirth,  
data = demoData)

Coefficients:

(Intercept)	factor(race)ASIAN AMERICAN
-10.2242078	0.5251107
factor(race)LATINO	factor(race)WHITE
-0.6690835	0.2202769
medianEarn	lifeExpAtBirth
0.0000944	0.1574541

hdIndex ~ factor(race) + medianEarn + lifeExpAtBirth

Above model is the most significant and efficient model as this one has lowest AIC value.