In R from library 'astsa', I have imported the quarterly US GNP data spanning from 1947-2023. ## [1] "ts" ## [1] 244.142 247.063 250.716 260.981 267.133 274.046 [1] 23718.26 24530.59 24929.18 25456.41 25885.43 26289.49 The dataset provides an insight into the economic performance of US post-war to modern times in units of billions of dollars. Insight into source and structure confirms quarterly Time-series data. Time-Series [1:304] from 1947 to 2023: 244 247 251 261 267 ...

U.S. Gross National Product Time-Series

The motivation of this analysis is to understand the future behaviour of the time series data 'Quarterly U.S. GNP' (Gross National Product) from 1947-2023. To this end, I will use Meta's Prophet forecasting system to generate a prediction of up to this year's (2025) values. To verify the robustness of the forecast, I will compare the prophet prediction with a non-parametric forecasting method. To further understand the data, I will compare parametric methods to investigate to what extent the dependence of GNP growth is linear

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analysis

Preface:

over time.

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Coursework 1 220306122

Understanding the data:

1960 1980 2000 2020 Time From the apparent exponential growth in initial plot, It is worth theorising a multiplicative model of Time series: $x_t = m_t S_t Y_t$. Decomposition of multiplicative Time-series model: I can continue in identifying historical trends, seasonal patterns and any measurable randomness. Decomposition of multiplicative time series

observed 10000 25000

trend 10000 0 1.0000

random seasonal 10000090 0.98 0.94 1960 1980 2000 2020 Time The decomposition above provides an insight in behaviour of components. Further analysis will refine these insights later on. But from graph, we can see an exponential trend $m_t = \beta_0 exp(\beta_1 t)$, typically

associated with a multiplicative model: $x_t = m_t S_t Y_t$. The seasonal component in the multiplicative

model oscillates around one, acting as a small scaling factor for trend. Whilst the random/residual component is mostly centered around one, the clear downward spike represents an anomaly in the effect of COVID on GNP (circa 2020). Suggesting the data correctly reflects the economic downturn during the COVID period. I will account for this deviation in the prediction. Logarithmic decomposition: From the source: https://fred.stlouisfed.org/series/GNP I had gathered that the data was given in units of billions of dollars. The logarithm of this financial data will reflect the linearity of the apparent exponential growth, making data easier to interpret, as well as reflect the relevant increment of growth rates rather than absolute monetary value. The multiplicative model $x_t = m_t S_t Y_t$ becomes additive with log-transformation: $log(x_t) = log(m_t) + log(S_t) + log(Y_t)$ *Logarithmic* Decomposition of additive time series observed ω trend ω 9 seasonal 0e+00 0.02 - 03

random -0.06 1960 1980 2000 2020 The log-transformation reflects a more linear form in trend, seasonal and residual components centered around zero, confirming the initial assumption of a multiplicative model was appropriate. The continuation of using a log-additive model will be useful moving forward in parametric analysis. More on Logarithmic transformation: Building on the findings of the logarithmic decomposition, I wanted definitive proof that a log transformation of the data would reflect an accurate representation of the linearity. I would continue in

defining a log-data-frame and log-linear model: ## ## $lm(formula = y \sim ds, data = USGNP.df)$ ## Residuals: 1Q Median 3Q ## -0.46330 -0.14035 -0.03449 0.18382 0.29836 ## Coefficients: ## Estimate Std. Error t value Pr(>|t|) ## (Intercept) -1.212e+02 9.795e-01 -123.7 <2e-16 *** 6.512e-02 4.934e-04 132.0 <2e-16 *** ## ---## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 ## Residual standard error: 0.1888 on 302 degrees of freedom ## Multiple R-squared: 0.983, Adjusted R-squared: 0.9829 ## F-statistic: 1.742e+04 on 1 and 302 DF, p-value: < 2.2e-16 From the summary of the log-transformed model, we can see an adjusted \mathbb{R}^2 value of 98.29%. We can compare this to the original model's value by transforming the response back to reflect the exponential growth:

Call: $\#\# \ lm(formula = exp(y) \sim ds, data = USGNP.df)$ ## ## Residuals: 1Q Median 3Q Min ## ## -2860.5 -2438.9 -532.6 1882.5 7923.6

Coefficients:

6

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Estimate Std. Error t value Pr(>|t|) ## (Intercept) -5.922e+05 1.314e+04 -45.08 <2e-16 *** 3.018e+02 6.618e+00 45.61 ## ds ## ## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1 ## ## Residual standard error: 2532 on 302 degrees of freedom ## Multiple R-squared: 0.8732, Adjusted R-squared: 0.8728 ## F-statistic: 2080 on 1 and 302 DF, p-value: < 2.2e-16 The adjusted \mathbb{R}^2 value of the original data is 87.28%. \mathbb{R}^2 , known as the coefficient of determination, is the percentage of total variation in the response explained by the model. A higher value suggests a large proportion of variance is accounted for, indicating a better fit of the model and to what extent the relationship between time and GNP is linear. Therefore, the approximate 11% increase in the logtransformed model reflects a better fit of the the data. Concluding the log transformation provides a stabilisation of variance and linearises exponential growth, making the data easier to interpret and improves the predictive power. Parametric analysis: In continuation of understanding the growth in the series, I will fit parametric models to analyse the optimal number of parameters which best explains the data. Ultimately, to summarise the trend.

I set the seasonal and residual component to zero solely for analysing the trend.

 $log(x_t) = log(m_t)$ where m_t has functional form: $m_t = eta_0 exp(eta_1 t)$

Call: $\#\# lm(formula = y \sim ds, data = USGNP.df)$ ## ## Residuals: Min 1Q Median 3 Q Max ## -0.46330 -0.14035 -0.03449 0.18382 0.29836 ## ## Coefficients: Estimate Std. Error t value Pr(>|t|) ## (Intercept) -1.212e+02 9.795e-01 -123.7 6.512e-02 4.934e-04 132.0 <2e-16 *** ## ds ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 ## Residual standard error: 0.1888 on 302 degrees of freedom ## Multiple R-squared: 0.983, Adjusted R-squared: 0.9829 ## F-statistic: 1.742e+04 on 1 and 302 DF, p-value: < 2.2e-16 Linear-Log-model 9

1960 1980 2000 2020 Time $log(x_t) = log(eta_0) + eta_1 t$ Fitting a log-linear model provides an understanding of growth of GNP over time. The fitted values/estimates are represented through the red line. Whilst it captures the general trend, we can further capture the fluctuations of the data by increasing the number of parameters via polynomial regression: $log(x_t) = log(eta_0) + eta_1 t + eta_2 t^2$ ~ Quadratic-log-model ## ## Call: $\#\# lm(formula = y \sim +ds + I(ds^2), data = USGNP.df)$ ## Residuals: 1Q Median 3Q -0.22928 -0.08675 0.04068 0.09457 ## ## Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) -1.435e+03 6.410e+01 -22.39 <2e-16 *** 1.389e+00 6.460e-02 21.51 ## ds -3.336e-04 1.627e-05 -20.50 <2e-16 *** $## I(ds^2)$ ## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 ## Residual standard error: 0.1221 on 301 degrees of freedom ## Multiple R-squared: 0.9929, Adjusted R-squared: 0.9928 ## F-statistic: 2.101e+04 on 2 and 301 DF, p-value: < 2.2e-16

 $log(x_t) = log(eta_0) + eta_1 t + eta_2 t^2 + eta_3 t^3$ ~ Cubic-log-model

1Q Median -0.11734 -0.06032 -0.01876 0.05347

-1.363e+02

6.905e-02

-1.165e-05

Residuals:

Coefficients:

(Intercept)

accuracy of the forecast.

311 2024-07-01 ## 312 2024-10-01 313 2025-01-01 314 2025-04-01 315 2025-07-01 316 2025-10-01

311 2024-07-01

312 2024-10-01

313 2025-01-01 314 2025-04-01

315 2025-07-01

316 2025-10-01

10

7

6.5

6

5.5

model:

9

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1950

1960

consistent in numerical and graphical representation.

1970

1980

This interactive plot is a variation of the simple representation that allows for better exploration of the observed and fitted trend. For example, I can verify the predictive outcomes of the graph match that with the confidence interval produced earlier on of the last 6 quarters. Of Q4 2025, the prediction from the graph is 10.21 which coincides with the prediction interval of 10.17-10.24. Confirming the analysis is

Non-parametric forecasting via Holt-Winters

The seasonal component of the series refers to the regular, predictable patterns that repeat over specific

1990

Prophet forecast plot:

##

ds

##

I(ds^2)

 $## I(ds^3)$

lm(formula = y ~ +ds + I(ds^2) + I(ds^3), data = USGNP.df)

Estimate Std. Error t value Pr(>|t|)

6.052e+00

3.049e-03

5.121e-07

Residual standard error: 0.0741 on 300 degrees of freedom ## Multiple R-squared: 0.9974, Adjusted R-squared: 0.9974 ## F-statistic: 3.822e+04 on 3 and 300 DF, p-value: < 2.2e-16

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

8.967e+04 4.004e+03 22.39 <2e-16 ***

-22.52

22.64

-22.75

<2e-16 ***

<2e-16 ***

Varying-degree-Log-model **Original Data** Log Model Quadratic Log Model Cubic Log Model 6 ∞ 9 1980 2000 2020 1960 Time Both quadratic and cubic log models effectively capture the fluctuations of growth in the series graphically. However, from the summaries of models, there is a 0.01% decrease in adjusted ${\it R}^2$ from linear to quadratic, suggesting the quadratic term lacks the explanatory power to justify the increase in model complexity. This in fact leads me to believe that the increase seen in the summary of the cubic variation is not an increase in explanatory power but an instance of having over-fit the model, which could lead to a sub-optimal performance with new data/predictions. Hence I will restrict from further analysis of increasing the number of parameters. Having provided a basic understanding of the data in question, I will now continue in providing a forecast of values for the years 2023-2025 via Meta's Prophet forecasting system. Forecast via Meta's Prophet package: Importing the 'prophet' package from the library will provide the necessary functions to forecast future values of the log-transformed data. I also implemented the zoo package for confirming a quarterly time index.

After loading the necessary packages I created a data-frame with the log-transformation of GNP data. Fitted the prophet function to model the data-frame. Then the last step of initialisation was to create the future data-frame for 12 quarterly periods, Q1 of 2023 to Q4 of 2025, for which values I will forecast. I've also accounted for the impact of COVID so that the prophet function will take quarter 2 of 2020 as an anomaly in an otherwise positively correlated trend. Accounting for such events will also improve the

The predict() function uses the prophet model and future dates to forecast the values for the quarters in

question. I also implemented intervals for the predictions in the last 6 quarters, the numerical representation was purely for informative purposes of what to expect from the upcoming plots:

10.23405

10.24571

1980

ds

Using the tail() function in R I can confirm the last 6 quarters of future dates:

ds yhat_lower yhat_upper

10.14232

10.14738 10.15768

10.16602

10.17388

10.17898

1960

The plot above is a simple representation of the linear relationship between GNP and time. The positively correlated blue line is a continuation/representation of prediction of the black line (observed GNP). I can further analyse the plot components: 10 -9 -1960 1980 2000 2020 ds 0.000 --0.050 **-**-0.075 **-**-0.100 **-**-0.125 **-**1980 1960 2000 2020 ds April 01 January 01 July 01 January 01 October 01 Day of year Displaying the predictive linear trend along with the yearly seasonal component and the random component centered around zero. While accounting for the spike in the forecast initialisation is important, it does not imply that the randomness is fully explained. Instead, it highlights a significant deviation from the usual pattern. Accounting for this improves the accuracy of the trend. Interactive plot: **Actual** — Predicted 10.5 10 9.5 9 8.5 8 7.5

2000

2020

ليا

2020

2000

2010

periods. In either of the multiplicative or log additive models, It's important that the seasonal component effects the trend as minimal as possible, as if it weren't to, it would introduce bias and artificial trends. Here, through Holt-Winters forecasting, I can show that the seasonal component does not significantly alter the predictions. **Holt-Winters filtering** 10 6 ∞ 9 1980 2000 1960 2020 Time The graph above is the instance in which the model applies triple exponential smoothing to the series, considering trend, level and seasonal components.

Holt-Winters with no seasonality

9 2000 1960 1980 2020 Time Here, having set seasonal smoothing to false, we have produced a prediction exclusive of seasonal components, ultimately seeming indifferent to the full method. Triple exponential smoothing prediction with prediction interval: ## fit upr lwr ## [7,] 10.27757 10.36207 10.19308 [8,] 10.28929 10.38223 10.19636 [9,] 10.29999 10.40211 10.19786 ## [10,] 10.31007 10.42066 10.19948 ## [11,] 10.33375 10.45289 10.21462 [12,] 10.34548 10.47324 10.21771 Disabling seasonal component: ## fit upr ## [7,] 10.27239 10.35200 10.19278 [8,] 10.28602 10.37347 10.19858 [9,] 10.29966 10.39490 10.20442 **##** [10,] 10.31330 10.41633 10.21026 [11,] 10.32694 10.43778 10.21609 [12,] 10.34057 10.45926 10.22189

The notion is also satisfied through numerical representation, without seasonal components the predictions are not significantly altered. In comparison to the prophet forecast, there is an explanation to the slight differences in prediction intervals, it being the methods differ in flexibility and approach. Prophet has flexibility in accounting for deviations/impacts as evidently shown, but Holt-Winters offers a straightforward approach in accessing predictions. Concluding that the choice of either method is dependent on the complexity of the model.