

# Causal Data Science

Lecture 12.2: Recap of the course

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# Summary of the course

6/02/2023	Introduction
9/02/2023	Probability recap
13/02/2023	Graphical models, d-separation
16/02/2023	Causal graphs, Interventions, SCMs
20/02/2023	Covariate adjustment: backdoor criterion
23/02/2023	Covariate Frontdoor criterion, Instrumental variables
27/02/2023	Counterfactuals, potential outcomes, estimating causal effects 1
2/03/2023	Estimating causal effects 2 (matching, IPW)
6/03/2023	Constraint based structure learning
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Background on causal graphs

We know the causal graph, how do we estimate causal effects?

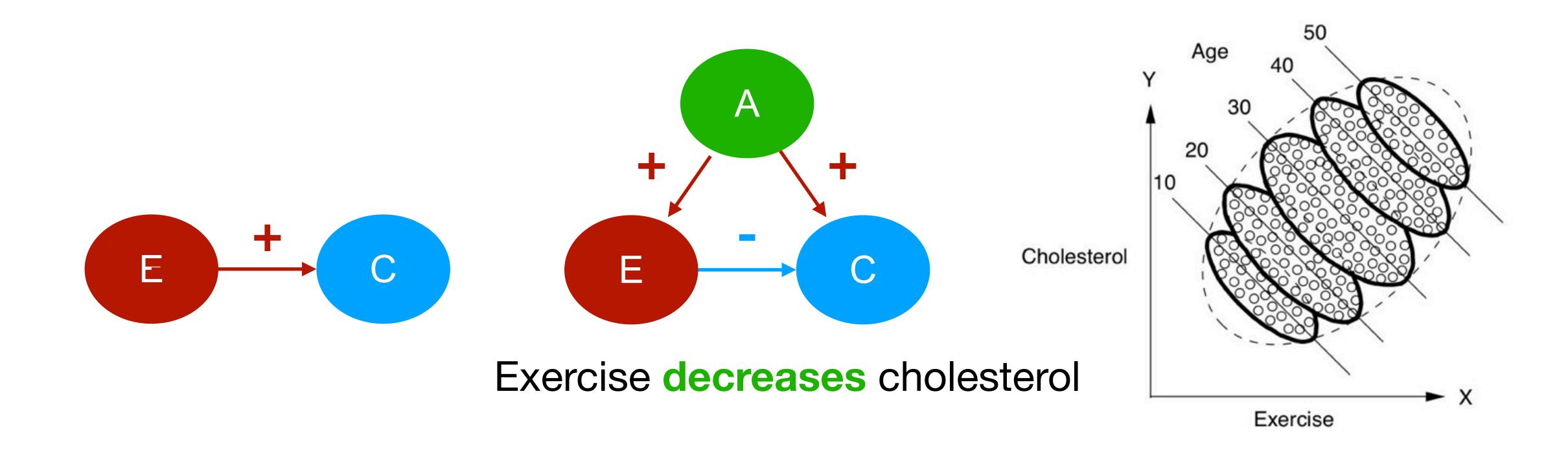
What happens if the graph is unknown?

Cutting edge research



# Motivation: Simpson's paradox - confounding

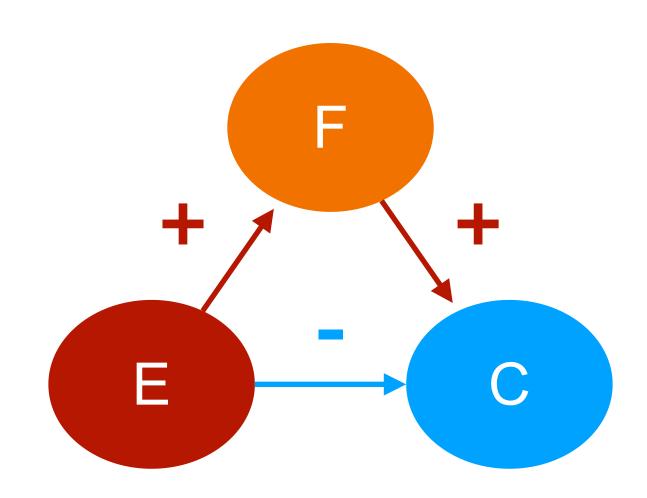
Let's assume we have observational data (e.g. data collected by hospitals)



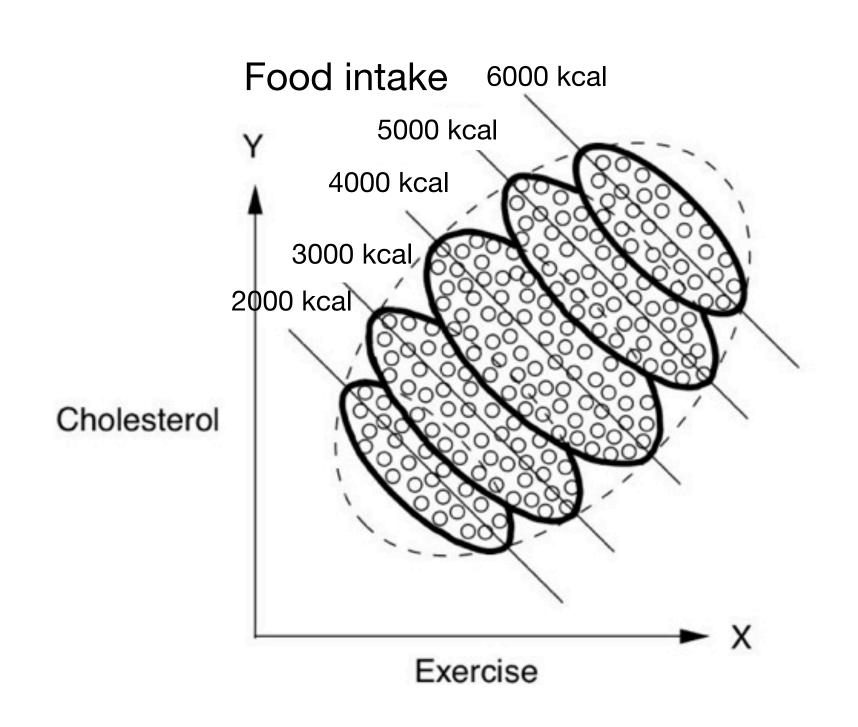


## Motivation: Shall we always just control on everything?

#### In alternative universe:



Exercise increases cholesterol



Takeaway: The covariates to adjust for can be different based on the graph.



# Bayesian networks (BNs)

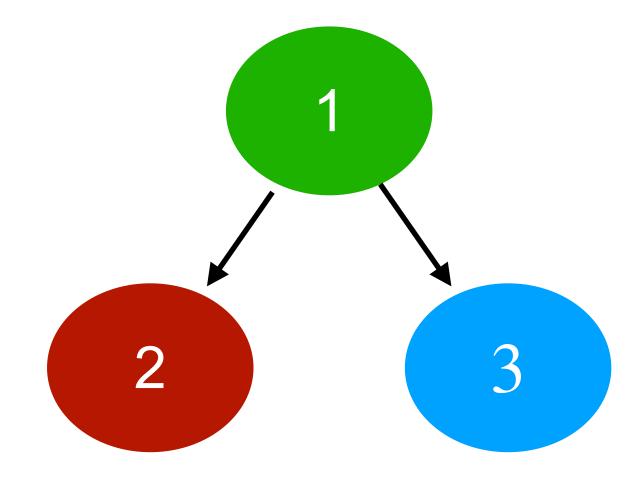
- We have a set of random variables  $X_1, \ldots, X_p$  with joint  $p(X_1, \ldots, X_p)$
- We have a DAG G, s.t. each random variable  $X_i$  is represented by node i
- We then say  $p(X_1, ..., X_p)$  factorizes over G if

$$p(X_1, ..., X_p) = \prod_{i \in V} p(X_i | \mathbf{X}_{pa(i)})$$

factorisation

They can help We can easily simplify the read conditional independences

They can represent causal models





# D-separation (summary)

- A path between i and j is blocked by  $A \subseteq V$  at least one condition holds:
  - There is a *non-collider* on the path that is in A, or
  - There is a collider k on the path, but  $k \notin A$  and  $\operatorname{Desc}(k) \cap A = \emptyset$
- Otherwise it is active
- Nodes i and j is d-separated by A if all paths between i, j are blocked
  - We denote d-separation as  $i \perp j \mid A$

We mostly assumed:  $\mathbf{A} \perp \mathbf{B} \mid \mathbf{C} \iff X_{\mathbf{A}} \perp \perp X_{\mathbf{B}} \mid X_{\mathbf{C}}$ 



# BNs vs causal BNs - example

• Fire (F) and Alarm (A) with p(F,A) and  $A \not\perp \!\!\! \perp F$  can be factorized as:



# Causal Bayesian networks

 We introduced a new operator that can represent a hypothetical intervention on the whole population, i.e. a perturbation of the system:

do(X = x) ... force X to the value x for all samples

• If (G = (V, E), p) is a Bayesian network and if for all  $W \subset V$ :

$$p(X_{\mathbf{V}}|\operatorname{do}(X_{\mathbf{W}}=x_{\mathbf{W}})) = \prod_{i \in \tilde{\mathbf{V}} \setminus \tilde{\mathbf{W}}} P(X_{i}|X_{\operatorname{Pa}(i)}) \cdot \#(X_{\tilde{\mathbf{w}}}=x_{\tilde{\mathbf{w}}})$$

then (G, p) is a causal Bayesian network

Parents are now direct causes



# Structural causal models (SCMs)

- Let (G, p) be a causal Bayesian network
- We can write each variable  $X_i$  for  $i \in V$  as a function of its parents in G and a noise term  $e_i$  in a structural equation:

$$X_i \leftarrow h_i(X_{\text{Pa}(i)}, \epsilon_i)$$
 often linear often Gaussiah

• We assume all noises are independent of each other  $\forall i \neq j : \epsilon_i \perp \!\!\! \perp \epsilon_j$ 



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We know the causal graph, how do we estimate causal effects?

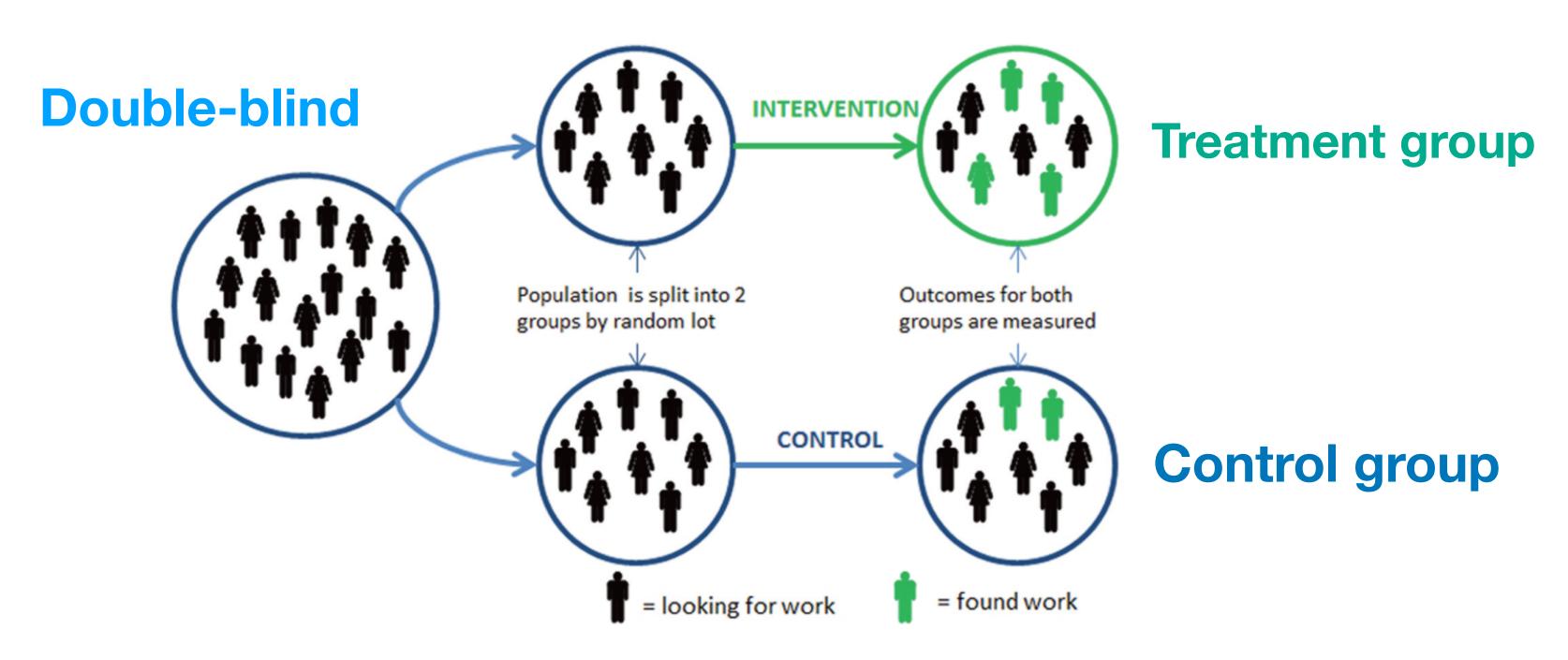
 $p(x_j | \operatorname{do}(x_i))$ ?

EFFECT OF TREATHENT ON OUTCOME



# Estimating causal effects: special case RCT

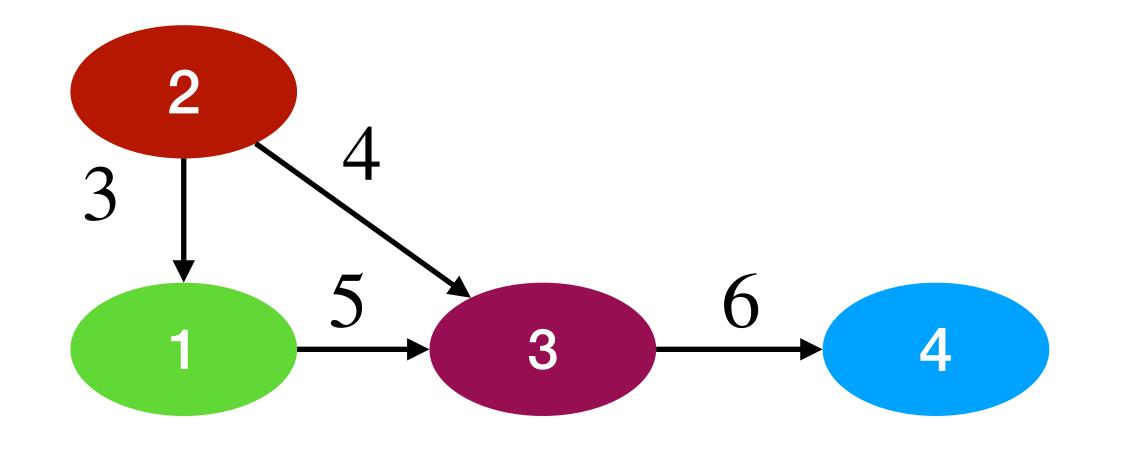
Randomised Controlled Trials (RCT): intervening on the treatment





### Estimating causal effects: special case linear SCMs

- In a linear SCM we estimate the total average causal effect of  $X_i$  on  $X_j$  :
  - For each directed path from  $X_i$  to  $X_j$ , multiply the edge weights
  - Sum the weights from all paths



$$E[X_{4}|X_{2}=1]-E[X_{4}|X_{2}=0]$$

$$= 3.5.6 + 4.6 = 114$$



# Identification strategies for causal effects

- Given a causal graph G, an identification strategy is a formula to estimate an interventional distribution from a combination of observational ones
- Backdoor criterion, Adjustment criterion

$$p(x_j | \operatorname{do}(x_i)) = \int_{x_{\mathbf{Z}}} p(x_j | x_i, x_{\mathbf{Z}}) p(x_{\mathbf{Z}}) dx_{\mathbf{Z}}$$

Frontdoor criterion

$$p(x_j | do(x_i')) = \int_{x_{\mathbf{M}}} p(x_{\mathbf{M}} | x_i') \int_{x_i} p(x_j | x_{\mathbf{M}}, x_i') p(x_i) dx_i$$

Instrumental variables



# Backdoor criterion [Pearl 2009]

- Given a CBN (G, p) with  $G = (\mathbf{V}, \mathbf{E})$ , a set  $\mathbf{Z} \subseteq \mathbf{V} \setminus \{i, j\}$  satisfies the backdoor criterion for estimating the causal effect of  $X_i$  on  $X_j$  with  $i \neq j$ :
  - Z does not contain any descendant of i,  $Desc(i) \cap Z = \emptyset$ , and
  - $\mathbf{Z}$  blocks all backdoor paths from i to j (all paths that start with an arrow into  $i \leftarrow \ldots j$ )

The backdoor criterion finds some (not necessarily all) valid adjustment sets



#### Complete: Adjustment criterion [Shpitser et al, Perković et al]

- Given a CBN (G, p), a set  $\mathbf{Z} \subseteq \mathbf{V} \setminus \{i, j\}$  satisfies the **adjustment** criterion for estimating the causal effect of  $X_i$  on  $X_j$  with  $i \neq j$ :
  - 1.  ${\bf Z}$  does not contain any descendant of nodes  $r \neq i$  on a directed path from i to j
  - 2.  ${f Z}$  blocks all paths from i to j that are not directed paths from i to j

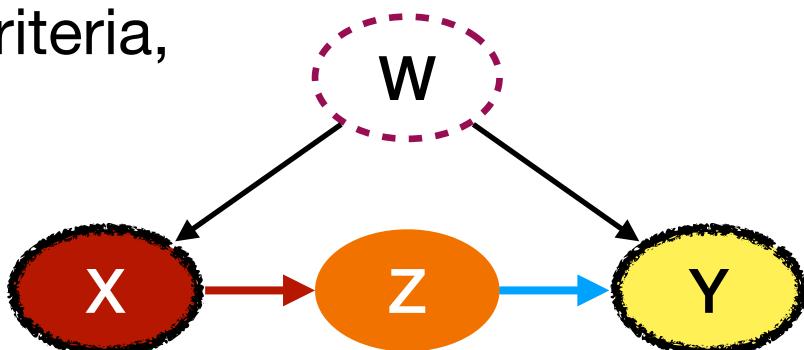
The adjustment criteria finds all valid adjustment sets (but there are other sets that allow identification of total causal effects - e.g. frontdoor criterion)



# Frontdoor criterion example

- We cannot use the backdoor/adjustment criteria,
- because W is unobserved



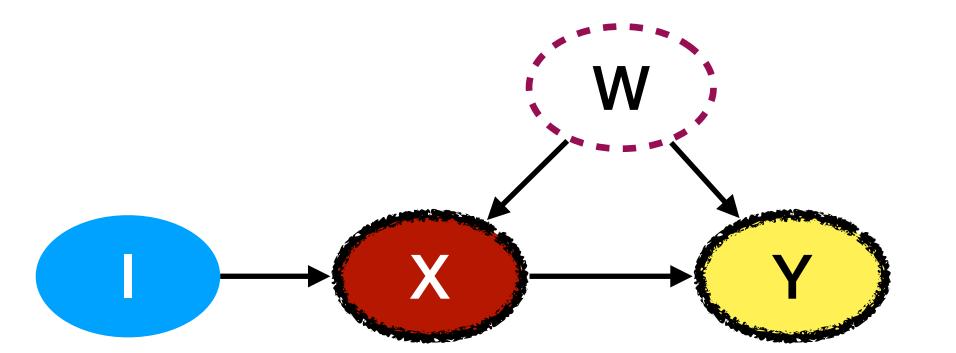


- 1. Find all mediator variables f M on the directed paths between X and Y
- 2. Estimate effect of X on M (no unblocked backdoors from  $i \leftarrow ...M$ )
- 3. Estimate effect of  $\mathbf{M}$  on  $Y(i \text{ blocks all backdoor paths from } \mathbf{M} \leftarrow \dots j)$
- 4. Combine the two effects



#### Instrumental variables

We want to estimate the effect of X on Y



- ullet We cannot use the backdoor/adjustment criteria, because W is unobserved
- We cannot use frontdoor because there is no mediator
- We can exploit the instrumental variable (IV)  $\it I$ 
  - $I \to X$ , but  $I \not\to Y$  directly,  $I \perp \!\!\! \perp W$

$$B = \frac{Cov(I, Y)}{Cov(I, X)}$$
(or 2SLS)



#### Counterfactuals

- Up to now we have been discussing how to estimate the interventional distribution  $P(X_i | do(X_i))$
- But this does not tell us what would have happen to individuals under different interventions that the ones that were actually performed
  - For example: a patient was treated and they recovered, what would have happened if they were not treated? (Retrospectively)
- Assumption: the noise variables stay the same



# Unit-level counterfactuals for linear SCMs

- Linear SCM S with observed variables  $(X_1, ..., X_p)$  and noises  $(\epsilon_{X_1}, ..., \epsilon_{X_p})$
- We can compute counterfactuals for  $do(X_j)$  and unit i with  $(x_1^i, \ldots, x_p^i)$ :
- **1.** Abduction: reconstruct the noise variable values for i using  $S: (\hat{\epsilon}_{X_1}^i, \ldots, \hat{\epsilon}_{X_p}^i)$
- **2. Action:** If  $x_j^i = 0$  in the original data, change the equation for i to  $x_j^i \leftarrow 1$ , else if  $x_j^i = 1$ , change it to  $x_j^i \leftarrow 0$  (the counterfactual assignment)
- 3. Prediction: Recompute  $(\hat{x}_1^i, ..., \hat{x}_p^i)$  using S and  $(\hat{e}_{X_1}^i, ..., \hat{e}_{X_p}^i)$

Issue: We cannot use unit-level counterfactuals to falsify a wrong causal model



### Unit-level causal effects vs average causal effects

- Unit-level causal effect:  $Y_i(t=1)-Y_i(t=0)$
- Fundamental problem of causal inference: we cannot observe a factual and a counterfactual outcome for each unit.
  - In general the treated population and the untreated population are composed by individuals that are not exactly the same.
    - SUTVA, consistency, ignorability, positivity
- But: we can estimate the effect from data at a population level

ATE = 
$$\mathbb{E}[Y(t = 1) - Y(t = 0)]$$



# Estimation method: Matching

We want to estimate the average treatment effect on observational data:

ATE = 
$$\mathbb{E}[Y(t=1) - Y(t=0)] = \mathbb{E}[Y|do(T=1)] - \mathbb{E}[Y|do(T=0)]$$

- Intuition: find the most similar couple of patients in terms of covariates  $\mathbf{X}$ , such that one is in the treatment and the other in the control group
  - Assumption: X satisfy the backdoor criterion
- Goal: discard unmatched units, so we have the same number of units with the same combination of values for X in treatment and control (balancing)



#### Estimation method: Propensity score matching (PSM)

- Assumptions: binary treatment T, X is valid adjustment set
- Propensity score: the probability of getting assigned the treatment

$$\pi := P(T = 1 | \mathbf{X} = x)$$

- We then do matching on propensity scores
- $\pi$  encodes all information of  ${f X}$  that is useful for T, i.e.  $T\perp {f X} \mid \pi$ 
  - If X has a lot of covariates, it might easier to match for since it's a number
  - $\pi$  is estimated from data, e.g. with logistic regression



#### Estimation method: Inverse probability weighting (IPW)

- Inverse probability (of treatment) weighting: weight by inverse of probability of treatment received:
  - For treated T=1: weight by the inverse of  $\pi=P(T=1|\mathbf{X})$
  - For untreated T=0: weight by the inverse of  $1-\pi=P(T=0|\mathbf{X})$

$$\hat{\mathbb{E}}(Y(t=1)) = \frac{1}{n} \sum_{i=1}^{n} Y_i \cdot 1\{T=1\} \cdot \frac{1}{P(T=1|X_i)}$$

$$\hat{\mathbb{E}}(Y(t=0)) = \frac{1}{n} \sum_{i=1}^{n} Y_i \cdot 1\{T=0\} \cdot \frac{1}{P(T=0 \mid X_i)}$$



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What happens if the graph is unknown?



# Causal discovery overview

## Constraint-based causal discovery

- Conditional independence tests
- Observational data
- Output: MEC
- SGS, PC

# Score-based causal discovery

- Penalised likelihood
- Observational data
- Output: MEC
- GES

#### **Restricted models**

- Nonlinear additive noise, Linear Non-Gaussianity
- Observational data
- Output: DAG
- RESIT, LINGAM

# Interventional causal discovery / causal invariance

- Observational and Interventional data
- Output: parents of Y,I-MEC
- ICP, JCI



# SGS algorithm (Spirtes, Glymour, Scheines)

- ullet Assuming P is Markov and faithful to an unknown graph G
- We can estimate a CPDAG from samples of P in three steps:
  - 1. Determine the skeleton
  - 2. Determine the v-structures
  - 3. Direct as many remaining edges as possible
- Note: the directed parts of the CPDAG will agree with G, but some parts might stay undirected

# PC algorithm (Peter Spirtes, Clark Glymour)

- ullet Assuming P is Markov and faithful to an unknown graph G
- We can estimate a CPDAG from samples of P in three steps:
  - 1. Determine the skeleton in an optimised way
  - 2. Determine the v-structures
  - 3. Direct as many remaining edges as possible



# Score-based causal discovery

- Score-based causal discovery: find the graph that maximises a score S(G,D) (fit of graph G on data D)
- Typically we use BIC (Bayesian information criterion)

$$BIC(D, G) := 2 \cdot \log p(D \mid G, \theta^{MLE}) - \log(n) \cdot \#parameters$$

- Score equivalence: all DAGs in a MEC get the same score
- Decomposable, Local consistency



# Greedy Equivalence Search (GES)

- 1. Start with empty CPDAG
- 2. Add edges one by one until local maxima in BIC
- 3. Remove edges one by one until local maxima in BIC

#### Phase 1 neighbours $\varepsilon^+$ :

Given a a starting equivalence class  $\varepsilon$ , another class  $\varepsilon'$  is in the neighbours  $\varepsilon^+$  if there exists a DAG  $G \in \varepsilon$ , such that adding an edge to G results in  $G' \in \varepsilon'$ 

Phase 2 neighbours  $\varepsilon^-$ : same with removing an edge



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# Do-calculus (complete identification strategy)

- For disjoint sets  $A, B, C, W \subseteq V$ :
- Rule 1: insertion/deletion of observations

$$\mathbf{A} \perp_d \mathbf{B} \mid \mathbf{C}, \operatorname{do}(\mathbf{W}) \implies P(X_{\mathbf{A}} \mid X_{\mathbf{B}}, X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}})) = P(X_{\mathbf{A}} \mid X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}}))$$

Rule 2: action/observation exchange

$$\mathbf{A} \perp_d I_{\mathbf{B}} | \mathbf{B}, \mathbf{C}, \operatorname{do}(\mathbf{W}) \implies P(X_{\mathbf{A}} | \operatorname{do}(X_{\mathbf{B}}), X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}})) = P(X_{\mathbf{A}} | X_{\mathbf{B}}, X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}}))$$

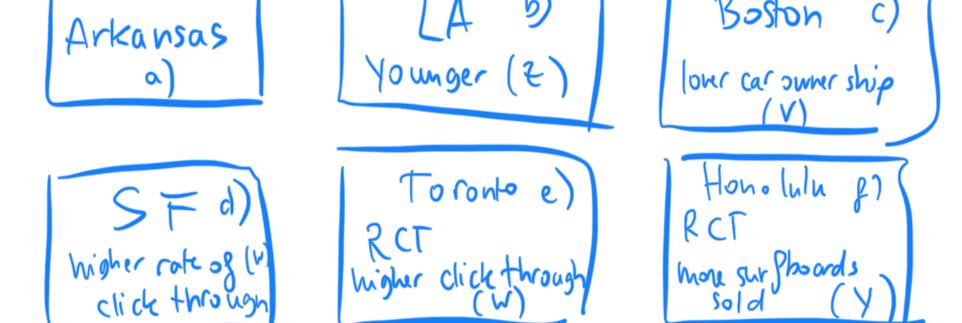
Rule 3: insertion/deletion of actions

$$\mathbf{A} \perp_d I_{\mathbf{B}} | \mathbf{C}, \operatorname{do}(\mathbf{W}) \implies P(X_{\mathbf{A}} | \operatorname{do}(X_{\mathbf{B}}), X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}})) = P(X_{\mathbf{A}} | X_{\mathbf{C}}, \operatorname{do}(X_{\mathbf{W}}))$$



# Transportability [Bareinboim and Pearl 2016]

 How to combine the data from different observational and experimental conditions, each conducted on a different population, to estimate a causal effect on a target population?



 Given the true causal graph in the target setting and the selection diagrams showing the differences in the other settings, one can find an estimated by applying do-calculus



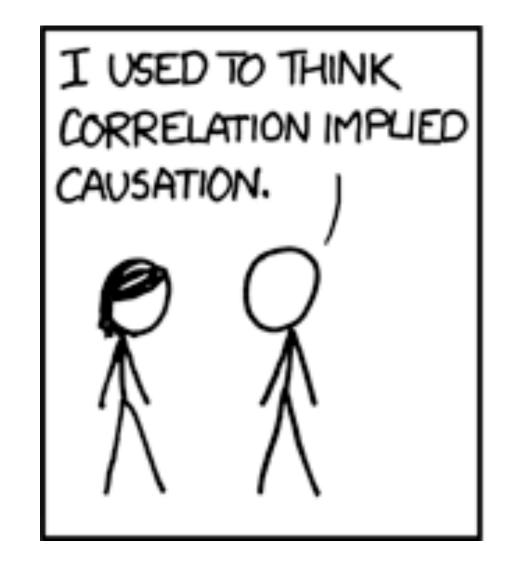
## Causality + machine learning (non-exhaustive list)

- 1. Machine learning (ML) helps causality
  - Causal discovery learning causal graphs from data
  - Causal effect estimation matching, weighting, double ML
  - (Causal) representation learning
- 2. Causality (in the most general definition) helps machine learning
  - Robustness, Transfer learning
  - Reinforcement Learning
  - Bias mitigation, fairness

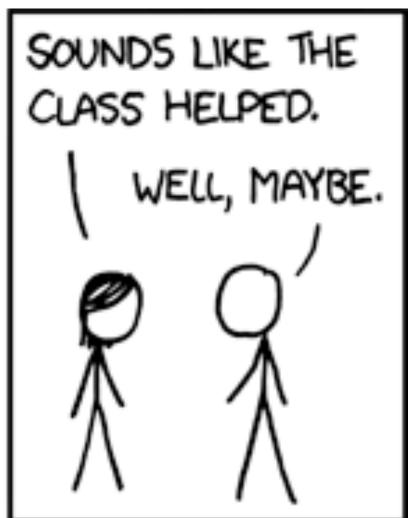


## Questions??

If you have any follow-up question use the Discussions tab in Canvas







https://xkcd.com/552/