

Proof for why HPP does not work if samples are normally distributed

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1 HPP against normally distributed samples

In the following, we see what happens to the computations the *Learning a parallelepiped* attack is based on if we replace the uniform distribution by a normal distribution. The key component and assumption of the *Learning a parallelepiped* attack is that the provided samples are distributed uniformly over $\mathcal{P}(V)$. One runs into trouble if the sampled vectors are on the form $\mathbf{v} = \mathbf{x}V$ where \mathbf{x} follows a normal distribution, i.e. $x_i \sim \mathcal{N}(\mu, \sigma^2)$. Although one might be able to approximate the covariance matrix $V^t V$ and transform the hidden parallelepiped to a hidden hypercube, one can not do a gradient descent based on the fourth moment given such samples using the method from the original attack [2]. We will show that if samples follow a normal distribution, the fourth moment of $\mathcal{P}(C)$ over \mathbf{w} on the unit circle is constant, and therefore a gradient descent can not reveal any information about the secret key V .

Adapting the definition of $\mathcal{P}(V)$ Recall that $\mathcal{P}(V)$ is defined as $\{\sum_{i=1}^n x_i \mathbf{v}_i : x_i \in [-1, 1]\}$ where \mathbf{v}_i are rows of V and x_i is uniformly distributed over $[-1, 1]$ (generally one can take another interval than $[-1, 1]$ and do appropriate scaling). Firstly, the normal distribution $\mathcal{N}(\mu, \sigma^2)$ is defined over $(-\infty, \infty)$, so it does not make sense to talk about samples "normally distributed over $\mathcal{P}(V)$ " without tweaking any definitions. Therefore, let $[-\eta, \eta]$ be a finite interval on which to consider a truncated normal distribution $\mathcal{N}_\eta(\mu, \sigma^2)$ such that $\int_{-\eta}^{\eta} f_X(x) dx = 1 - \delta$ for some negligibly small δ where $f_X(x)$ is the probability density function of $\mathcal{N}(\mu, \sigma^2)$, given by $f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$. Now we consider $\mathcal{P}_\eta(V) = \{\sum_{i=1}^n x_i \mathbf{v}_i : x_i \in [-\eta, \eta]\}$ and proceed as in the original HPP with $\mathcal{P}_\eta(V)$ instead of $\mathcal{P}(V)$.

Approximating $V^t V$ Let $V \in \mathcal{GL}_n(\mathbb{R})$. Let \mathbf{v} be chosen from a truncated normal distribution $\mathcal{N}_\eta(0, \sigma^2)$ over $\mathcal{P}_\eta(V)$. Then $\lim_{\eta \rightarrow \infty} \mathbb{E}[\mathbf{v}^t \mathbf{v}] = V^t V \cdot \sigma^2$.

Proof. Let samples be on the form $\mathbf{v} = \mathbf{x} V$, where \mathbf{x} is a row vector where each element $x_i \sim \mathcal{N}_\eta(0, \sigma^2)$. Then $\mathbf{v}^t \mathbf{v} = (\mathbf{x} V)^t (\mathbf{x} V) = (V^t \mathbf{x}^t) (\mathbf{x} V) = V^t \mathbf{x}^t \mathbf{x} V$. Considering $\mathbb{E}[\mathbf{x}^t \mathbf{x}]$, we see that for $i \neq j$, $\mathbb{E}[x_i x_j] = \mathbb{E}[x_i] \mathbb{E}[x_j] = 0 \cdot 0 = 0$ due to independent random variables. For $i = j$, $\lim_{\eta \rightarrow \infty} \mathbb{E}[x_i^2] = \mathbb{V}[x_i] = \sigma^2$ since $\mathbb{V}[x_i] = \mathbb{E}[x_i^2] - \mathbb{E}[x_i]^2 = \mathbb{E}[x_i^2] - 0 = \sigma^2$. Therefore, $\lim_{\eta \rightarrow \infty} \mathbb{E}[\mathbf{x}^t \mathbf{x}] = I_n \cdot \sigma^2$, i.e. the matrix with σ^2 on the diagonal. Consequently, $\lim_{\eta \rightarrow \infty} \mathbf{v}^t \mathbf{v} = V^t \mathbb{E}[\mathbf{x}^t \mathbf{x}] V = V^t (I_n \cdot \sigma^2) V = (V^t V) \cdot \sigma^2$ and conversely $\lim_{\eta \rightarrow \infty} V^t V = (\mathbf{v}^t \mathbf{v}) / \sigma^2$. \square

This means that we can in theory approximate the covariance matrix $V^t V$ by averaging over $\mathbf{v}^t \mathbf{v}$ and dividing by σ^2 . However, it is not immediately clear if one needs more samples for this approximation than in the original attack due to the difference in distributions.

Hypercube transformation Assume now that we know $V^t V$. Consider instead of $\mathcal{P}(V)$, $\mathcal{P}_\eta(V)$. Then by following part 1 of **Lemma 2** and its proof from [2] we can transform our hidden parallelepiped $\mathcal{P}_\eta(V)$ into $\mathcal{P}_\eta(C)$, a hidden hypercube, since this does not depend on the distribution of the samples - it only assumes one knows $V^t V$. For completeness, by adapting the second part of **Lemma 2** to our case:

Proof. Let $\mathbf{v} = \mathbf{x} V$ where \mathbf{x} is normally distributed according to $\mathcal{N}_\eta(0, \sigma^2)$. Then samples \mathbf{v} are distributed according to $\mathcal{N}_\eta(0, \sigma^2)$ over $\mathcal{P}_\eta(V)$. It then follows that $\mathbf{v} L = \mathbf{x} V L = \mathbf{x} C$ has a truncated uniform distribution over $\mathcal{P}_\eta(C)$. \square

Thus, we should be able to map our normally distributed samples from the hidden parallelepiped to the hidden hypercube.

Learning a hypercube It is clear that samples uniformly over $\mathcal{P}_\eta(C)$ centered at the origin form a hypersphere for which any orthogonal rotation leaves the sphere similar in shape. As a consequence, the fourth moment of $\mathcal{P}_\eta(C)$ is constant over the unit circle.

Analogous to [2] we compute the 2nd and 4th moment of $\mathcal{P}_\eta(V)$ over a vector $\mathbf{w} \in \mathbb{R}^n$. The k -th moment of $\mathcal{P}_\eta(V)$ over a vector \mathbf{w} is defined as $mom_{V,k} = \mathbb{E}[\langle \mathbf{u}, \mathbf{w} \rangle^k]$ where $\mathbf{u} = \sum_{i=1}^n x_i \mathbf{v}_i$ and $x_i \sim \mathcal{N}_\eta(0, \sigma^2)$. First we consider $\langle \mathbf{u}, \mathbf{w} \rangle = \langle \sum_{i=1}^n x_i \mathbf{v}_i, \mathbf{w} \rangle = \sum_{i=1}^n x_i \langle \mathbf{v}_i, \mathbf{w} \rangle$. Then for $k = 2$, $\mathbb{E}[(\sum_{i=1}^n x_i \langle \mathbf{v}_i, \mathbf{w} \rangle)^2] = \mathbb{E}[\sum_{i=1}^n \sum_{j=1}^n x_i x_j \langle \mathbf{v}_i, \mathbf{w} \rangle \langle \mathbf{v}_j, \mathbf{w} \rangle]$. Due to independent random variables, $\mathbb{E}[x_i x_j] = 0$ when $i \neq j$, so we have $\sum_{i=1}^n \mathbb{E}[x_i^2] \langle \mathbf{v}_i, \mathbf{w} \rangle^2 = \sigma^2 \sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^2$ for sufficiently large η due to the well known result that $\mathbb{E}[x^2] = \sigma^2$ for $x \sim \mathcal{N}(0, \sigma^2)$. Thus, we end up with:

$$mom_{V,2}(\mathbf{w}) = \sigma^2 \mathbf{w}^t V^t V \mathbf{w} \quad (1)$$

We observe that if $V \in \mathcal{O}(\mathbb{R})$, $\text{mom}_{V,2}(\mathbf{w}) = \sigma^2 \|\mathbf{w}\|^2$

For $k = 4$:

$$\mathbb{E}[(\sum_{i=1}^n x_i \langle \mathbf{v}_i, \mathbf{w} \rangle)^4] = \mathbb{E}[\sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n \sum_{l=1}^n x_i x_j x_k x_l \langle \mathbf{v}_i, \mathbf{w} \rangle \langle \mathbf{v}_j, \mathbf{w} \rangle \langle \mathbf{v}_k, \mathbf{w} \rangle \langle \mathbf{v}_l, \mathbf{w} \rangle]$$

We consider three cases for the indices i, j, k , and l :

1. **None equal**: if i, j, k , and l are different, the expression equals 0 due to independent random variables.
2. **All equal**: if $i = j = k = l$, then we have $\sum_{i=1}^n \mathbb{E}[x_i^4] \langle \mathbf{v}_i, \mathbf{w} \rangle^4$. A well known result for the normal distribution $\mathcal{N}(0, \sigma^2)$ is that $\mathbb{E}[x^4] = 3\sigma^4$.
3. **Pairwise equal**: if either

- $i = j \neq k = l$
- $i = l \neq j = k$
- $i = k \neq j = l$

we have the following:

$$\sum_{i \neq j} \mathbb{E}[x_i^2 x_j^2] \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \langle \mathbf{v}_j, \mathbf{w} \rangle^2$$

Since $\mathbb{E}[x_i^2 x_j^2] = \mathbb{E}[x_i^2] \mathbb{E}[x_j^2] = \sigma^4$ due to independent random variables, we have

$$\sigma^4 \sum_{i \neq j} \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \langle \mathbf{v}_j, \mathbf{w} \rangle^2$$

Putting together the expressions above we have

$$\text{mom}_{V,4}(\mathbf{w}) = 3\sigma^4 \sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^4 + 3(\sigma^4 \sum_{i \neq j} \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \langle \mathbf{v}_j, \mathbf{w} \rangle^2)$$

since there are three cases where indices pair up two and two. The final result becomes:

$$\text{mom}_{V,4}(\mathbf{w}) = 3\sigma^4 (\sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^4 + \sum_{i \neq j} \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \langle \mathbf{v}_j, \mathbf{w} \rangle^2) \quad (2)$$

Claim: If $V \in \mathcal{O}(\mathbb{R})$, and \mathbf{w} is on the unit sphere, $\text{mom}_{V,4}(\mathbf{w})$ is constant.

Proof. This can be shown by rewriting (3.2) as

$$\begin{aligned} \text{mom}_{V,4}(\mathbf{w}) &= 3\sigma^4 \left(\sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^4 + \sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \sum_{j=1}^n \langle \mathbf{v}_j, \mathbf{w} \rangle^2 - \sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^4 \right) \\ \text{mom}_{V,4}(\mathbf{w}) &= 3\sigma^4 \left(\sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^2 \sum_{j=1}^n \langle \mathbf{v}_j, \mathbf{w} \rangle^2 \right) = 3\sigma^4 (\sigma^2 \|\mathbf{w}\|^2)^2 = 3\sigma^8 \end{aligned}$$

because $\text{mom}_{V,2}(\mathbf{w}) = \sum_{i=1}^n \langle \mathbf{v}_i, \mathbf{w} \rangle^2 = \sigma^2 \|\mathbf{w}\|^2$ when $V \in \mathcal{O}(\mathbb{R})$ and $\|\mathbf{w}\|^2 = 1$ when \mathbf{w} lies on the unit sphere. \square

In conclusion, if samples over the secret parallelepiped $\mathcal{P}_\eta(V)$ follow a continuous normal distribution, a gradient descent based on the fourth moment described in [2] is impossible because the fourth moment is constant over the unit sphere of \mathbb{R}^n .

The discrete Gaussian distribution We now do the same computations considering the discrete Gaussian distribution as described in [1]. It turns out that for the parameters used in Hawk, this distribution behaves like its corresponding Normal distribution.

Consider the discrete Gaussian distribution as described in [1].

If $X \sim \mathcal{D}_{2\mathbb{Z}+c,\sigma}$, we have $\text{Supp}(X) = 2\mathbb{Z} + c$ where $c \in \{0, 1\}$, and $\Pr[X = x] = \frac{\rho_\sigma(x)}{\sum_{y \in 2\mathbb{Z}+c} \rho_\sigma(y)}$ where $\rho_\sigma(x) = e^{-\frac{x^2}{2\sigma^2}}$. For $c \in \{0, 1\}$, $\mathbb{E}[X] = 0$ and $\mathbb{V}[X] = \sigma^2$ for appropriate choices of σ .

... Show definitions and stuff here ...

Then by computational results, as η grows large enough, $\mathbb{E}[x^2] = \sum_{-\eta}^{\eta} x^2 f_X(x)$ and $\mathbb{E}[x^4] = \sum_{-\eta}^{\eta} x^4 f_X(x)$ approaches σ^2 and $3\sigma^4$ respectively.

... Show table of computations here ...

References

- [1] Joppe W. Bos, Olivier Bronchain, Léo Ducas, Serge Fehr, Yu-Hsuan Huang, Thomas Pornin, Eamonn W. Postlethwaite, Thomas Prest, Ludo N. Pulles, and Wessel van Woerden. Hawk. Technical report, NXP Semiconductors, Centrum Wiskunde & Informatica, Mathematical Institute at Leiden University, NCC Group, PQShield, Institut de Mathématiques de Bordeaux, September 2024.
- [2] Phong Q. Nguyen and Oded Regev. Learning a parallelepiped: Cryptanalysis of ggh and ntru signatures, 2009.